

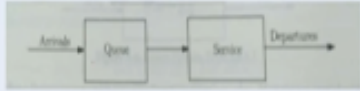
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**Lecture - 60**  
**Application to Queuing Theory and Reliability Theory**

Hello friends. Welcome to my lecture on application of probability theory to queuing theory and reliability theory. Queuing theory is the mathematical study of waiting lines or queues. (Refer Slide Time: 00:46)

**Basic characteristics of queueing models**

Queueing theory is the mathematical study of waiting lines or queues. A queue or waiting line is formed when units (customers or clients) requiring some kind of service arrive at a service counter or service channel. A simple queueing model is shown in the following figure:



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graph LR; Arrivals --> Queue; Queue --> Service; Service --> Departures
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Basic characteristics of a Queueing systems are:

- (a) The input or arrival pattern,
- (b) The service mechanism (or service pattern),
- (c) The queue discipline,
- (d) The system capacity.

A queue or waiting line is found when units that is customers or clients requiring some kind of service arrive at a service counter or service channel. A simple queueing model is shown in the following figure. There you can see arrivals, queue is formed and then service is done and then they depart. So basic characteristics of queueing system are the input or arrival pattern, the service mechanism or service pattern, the queue discipline, the system capacity. (Refer Slide Time: 01:14)

#### Input or arrival pattern

The input describes the manner in which customers (or units) arrive and join the queueing system. It is not possible to observe or detect the actual amount of customers arriving at the queue for service. Hence, we express the arrival pattern of customers by the probability distribution of the number of arrivals per unit of time or of the inter arrival time.

Let us consider those queueing systems in which the number of arrivals per unit of time is a Poisson random variable with mean  $\lambda$ . We know that in this case, the time between consecutive arrivals, i.e. the inter arrival time of the Poisson process, has an exponential distribution with mean  $\frac{1}{\lambda}$ .

Let us discuss all these characteristics one by one. So input or arrival pattern. The input describes the manner in which customers or units arrive and join the queueing system. It is not possible to observe or detect the actual amount of customers arriving at the queue for service. Hence, we express the arrival pattern of customers by the probability distribution of the number of arrivals per unit of time or of the inter arrival time.

Let us consider those queueing systems in which the number of arrivals per unit of time is a Poisson random variable with mean  $\lambda$ . We know that in this case, the time between consecutive arrivals that is the inter arrival time of the Poisson process has an exponential distribution with mean  $1/\lambda$ .

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#### Service mechanism

It can be described by a service rate (i.e. the number of customer serviced in one time period) or by the inter service time (i.e. the time required to complete the service for a customer).

Let us consider those Queueing systems in which the number of customers serviced per unit of time has a Poisson distribution with mean  $\mu$  or, equivalently, the inter service time has an exponential distribution with mean  $\frac{1}{\mu}$ .

Now let us discuss the service mechanism. It can be described by a service rate that is the number of customer serviced in one time period or by the inter service time that is the time required to complete the service for a customer. Let us consider those queueing systems in

which the number of customers serviced per unit of time has a Poisson distribution with mean  $\mu$  or equivalently the inter service time has an exponential distribution with mean  $1/\mu$ .  
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**Queue Discipline**

In this procedure, the customers are selected for service when a queue is formed. The various types of queue disciplines are tabulated in the following table:

No.	Queue Discipline	Description
1	FIFO or FCFS	First In First Out or First Come First Served. This is the most commonly used procedure in servicing customers.
2	LIFO or LCFS	Last In First Out or Last Come First Served. This procedure is used in inventory systems.
3	SIRO	Selection for service in Random Order.
4	PIR	Priority in Selection, i.e. customers are prioritized upon arrival. This procedure is used in manual transmission managing systems.

**Figure : Fig.2**

Let us assume that service is provided by the FCFS or FIFO procedure, i.e. on a first come first served basis.

Now next we consider queue discipline. In this procedure, the customers are selected for service when a queue is formed. The various types of queue disciplines are tabulated in the following table. Number one FIFO or FCFS. FIFO means First in First Out or First Come First Served. This is the most commonly used procedure in servicing customers. Number two LIFO or LCFS, Last in First Out or Last Come First Served. This procedure is used in inventory systems.

SIRO, SIRO Selection for Service in Random Order. This is not very usual. PIR, Priority in Selection that is customers are prioritized upon arrival. This procedure is used in manual transmission managing systems. If a message is very important, it is transmitted; first it is given priority over other messages. So it is priority in selection. Now let us assume that service is provided by the FCFS or FIFO procedure. That is on a first come first served basis.  
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### System Capacity

The maximum number of customers in the queueing system can be either finite or infinite. In some queueing models, only limited customers or units are allowed in the system.

### Transient and steady states

A Queueing system is said to be in transient state when the operating characteristics of the system depend on time. A Queueing system is said to be in steady state when the operating characteristics of the system are independent of time. For example, if  $P_n(t)$  is the probability that there are  $n$  customers in the system at time  $t$  (when  $t > 0$ ), finding  $P_n(t)$  is quite a difficult task even for a simple case. Thus, we are most interested in the steady state analysis of the system, i.e. in determining  $P_n(t)$  in the long run i.e. as  $t \rightarrow \infty$ . Thus, we look for steady state probabilities, i.e.  $P_n(t) \rightarrow P_n(\text{a constant})$  as  $t \rightarrow \infty$ .

So the maximum number of customers, system capacity. The maximum number of customers in the queueing system can be either finite or infinite. In some queueing models, only limited customers or units are allowed in the system. Now let us talk about transient and steady states. A queueing system is said to be in transient state when the operating characteristics of the system depend on time.

A queueing system is said to be in steady state when the operating characteristics of the system are independent of time. For example, if  $p_n(t)$  is the probability that there are  $n$  customers in the system at time  $t$ . When  $t > 0$ , finding  $P_n(t)$  is quite a difficult task even for a simple case. Thus, we are most interested in the steady state analysis of the system that is in determining  $P_n(t)$  in the long run.

That is as  $t$  goes to infinity. Thus, we look for a steady state probabilities that is  $P_n(t)$  goes to  $p_n$ , a constant as  $t$  goes to infinity.

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### Kendall's Notation of a Queueing System

In Kendall's notation (Kendall, 1951) Queueing system has the form

$$(a/b/c) : (d/e)$$

where

*a* = Inter – arrival distribution

*b* = Service time distribution

*c* = Number of channels or servers

*d* = System capacity

*e* = Queue discipline

Now let us discuss Kendall's notation of a queueing system. In Kendall's notation, Kendall gave this notation in 1951. So queueing system has this form  $a/b/c$  such that  $d/e$  where  $a$  is inter arrival distribution,  $b$  is service time distribution,  $c$  is number of channels or servers,  $d$  is system capacity,  $e$  is queue discipline.

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### Kendall's Notation of a Queueing System cont...

$a$  and  $b$  usually take one of the following symbols:

M: for Markovian or Exponential distribution

G: for arbitrary or general distribution

D: for fixed or deterministic distribution.

The four important Queueing systems describe in Kendall's notation are as follows:

(a)  $(M/M/1) : (\infty/\text{FIFO})$

(b)  $(M/M/s) : (\infty/\text{FIFO})$

(c)  $(M/M/1) : (k/\text{FIFO})$

(d)  $(M/M/s) : (k/\text{FIFO})$

Next, we define the line length and queue length for a queue.

Now  $a$  and  $b$  usually take one of the following symbols, M for Markovian or Exponential distribution, G for general distribution and D for fixed or deterministic distribution. The 4 important queueing systems described in Kendall's notation are as follows. MM1 infinity oblique FIFO, MMS infinity oblique FIFO, MM1 k oblique FIFO, MMS k oblique FIFO okay.

In this lecture, we will be discussing the first queueing system MM1 infinity oblique FIFO.

Next, we define the line length and queue length for a queue.

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**Definition:** For a queue, the line length or queue size is defined as the number of customers in the Queueing system. Also, the queue length is defined as the difference between the line length and the number of customers being served, i.e.

$$\text{Queue length} = \text{Line length} - \text{Number of customers being served}$$

#### Terminology

$n$  = Number of customers (units) in the system

$N(t)$  = Number of customers (units) in the system at time  $t$

$P_n(t)$  = The probability that there are exactly  $n$  customers at time  $t$  i.e.  $P[N(t) = n]$

$P_n$  = The steady state probability that exactly  $n$  customers are in the system

$\lambda_n$  = Mean arrival rate when there are  $n$  customers in the system

$\mu_n$  = Mean service rate when there are  $n$  customers in the system

$\lambda$  = Mean arrival rate when  $\lambda_n$  is constant for all  $n$

$\mu$  = Mean service rate when  $\mu_n$  is constant for all  $n$

So, for a queue, the line length or queue size is defined as the number of customers in the queueing system. Also, the queue length is defined as the difference between the line length and the number of customers being served. That is queue length=line length-number of customers being served.

Now we will be using the following terminology,  $n$  is the number of customers units in the system,  $N(t)$  number of customers in the system at time  $t$ ,  $P_n(t)$  the probability that there are exactly  $n$  customers at time  $t$  that is probability that  $N(t)=n$ .  $P_n$ , the steady state probability that exactly  $n$  customers are there in the system. And then,  $\lambda_n$  is the mean arrival rate okay, mean arrival rate when there are  $n$  customers in the system.

$\mu_n$  is the mean service rate when there are  $n$  customers in the system.  $\lambda$  is mean arrival rate when  $\lambda_n$  is constant for all  $n$ , and  $\mu$  is mean service rate when  $\mu_n$  is constant for all  $n$ .

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### Terminology cont...

$\rho = \frac{\lambda}{\mu}$  = Utilization factor or traffic intensity

$f_s(w)$  = The probability density function of waiting time in the system

$f_q(w)$  = The probability density function of waiting time in the queue

$L_s$  = The expected number of customers in the system or the average line length

$L_q$  = The expected number of customers in the queue or the average queue length

$L_w$  = The expected number of customers in non-empty queues

$W_s$  = The expected waiting time of a customer in the system

$W_q$  = The expected waiting time of a customer in the queue.

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Now some more notations are there. Rho is  $\lambda/\mu$  which is known as utilization factor or traffic intensity,  $f_s(w)$  is the probability density function of waiting time in the system,  $f_q(w)$  is the probability density function of waiting time in the queue,  $L_s$  is the expected number of customers in the system or the average line length. The  $L_q$  is the expected number of customers in the queue or the average queue length.

$L_w$  is the expected number of customers in the non-empty queues.  $W_s$  is the expected waiting time of a customer in the system,  $W_q$  is the expected waiting time of a customer in the queue. (Refer Slide Time: 08:05)

### Transient State Probabilities for Poisson Queue Systems

We now state the differential equations for the transient state probabilities for the Poisson queue systems, which are also known as birth-death process or immigration-emigration process.

Let  $N(t)$  be the number of customers (or units) in the system at time  $t$  and  $P_n(t)$  be the probability that there are  $n$  customers in the system at time  $t$ , where  $n \geq 1$ , i.e.

$$P_n(t) = P[N(t) = n].$$

The differential equations satisfied by  $P_n(t)$ , are given by

$$P'_n(t) = -(\lambda_n + \mu_n)P_n(t) + \lambda_{n-1}P_{n-1}(t) + \mu_{n+1}P_{n+1}(t) \quad (1)$$

where  $n \geq 1$ , and

$$P'_0(t) = -\lambda_0 P_0(t) + \mu_1 P_1(t) \quad (2)$$

Now let us find the differential equations for the transient state probabilities for the Poisson queue systems which are also known as birth-death process or immigration-emigration process. Let  $N(t)$  be the number of customers in the system at time  $t$  and  $p_n(t)$  be the probability that there are  $n$  customers in the system at time  $t$  where  $n \geq 1$  that is  $P_n(t)$  = probability that  $N(t) = n$ .



Then the differential equations satisfied by  $P_n(t)$  are given by  $P_n'(t)$  that is derivative of  $P_n(t)$  with respect to  $t = -\lambda_n + \mu_n P_{n+1}(t) + \lambda_{n-1} P_{n-1}(t) - \mu_n P_n(t)$  where  $n \geq 1$  and when  $n=0$  we have the differential equation  $P_0'(t) = -\lambda_0 P_0(t) + \mu_1 P_1(t)$ .  
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**Transient State Probabilities for Poisson Queue Systems cont...**

The solution of these differential equations yields the transient state probability,  $P_n(t)$ , for  $n \geq 0$  but this is exceedingly difficult even for simple problems.

**Steady State Probabilities for Poisson Queue Systems**

The steady state probabilities for poisson queue systems are derived by assuming that  $P_n(t) \rightarrow P_n$ , independent of  $t$ , as  $t \rightarrow \infty$ . The equations of steady state probabilities,  $P_n$ , can be obtained by putting  $P_n'(t) = 0$  and replacing  $P_n(t)$  by  $P_n$  in equations (1) and (2). Thus we obtain

$$-(\lambda_n + \mu_n)P_n + \lambda_{n-1}P_{n-1} + \mu_{n+1}P_{n+1} = 0 \quad (3)$$

and

$$-\lambda_0 P_0 + \mu_1 P_1 = 0. \quad (4)$$

Transient state probabilities for Poisson queue systems. The solution of these differential equations, the solution of the two differential equations that we have got number 1 and 2. Number 1 is for  $n \geq 1$  and number 2 is for  $n=0$ . So the solution of these differential equations yields the transient state probability,  $P_n(t)$  for  $n \geq 0$  but this is exceedingly difficult even for simple problems.

So we consider steady state probabilities for Poisson queue systems. The steady state probabilities for Poisson queue systems are derived by assuming that  $P_n(t)$  goes to  $P_n$  independent of  $t$  as  $t$  tends to infinity. The equations of steady state probabilities  $P_n$  can be obtained by putting  $P_n'(t) = 0$  because  $P_n(t)$  is now as  $t$  tends to infinity is a constant, so when we take the derivative of  $P_n(t)$  with respect to  $t$ , it will be  $=0$ .

So  $P_n'(t) = 0$  and replacing  $P_n(t)$  by  $P_n$  in equations 1 and 2 and then we obtain, you can see here if you put  $P_n'(t) = 0$  and  $P_n(t) = P_n$  then you get  $-\lambda_n + \mu_n P_{n+1} + \lambda_{n-1} P_{n-1} - \mu_n P_n = 0$  okay. So we have this equation and similarly the other equation, equation number 2 for  $n=0$  becomes  $P_0'(t) = 0$  and this is  $-\lambda_0 P_0 + \mu_1 P_1 = 0$  okay of the right side. So we get  $-\lambda_0 P_0 + \mu_1 P_1 = 0$ , so we get the second equation.  
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### Steady State Probabilities for Poisson Queue Systems cony...

Equations (3) and (4) are called the balance equations or equilibrium equations of Poisson queue systems.

Using the principle of mathematical induction, it follows that

$$P_n = \frac{\lambda_0 \lambda_1 \lambda_2 \dots \lambda_{n-1}}{\mu_1 \mu_2 \mu_3 \dots \mu_n} P_0 \quad (5)$$

and

$$P_0 = \frac{1}{1 + \sum_{n=1}^{\infty} \left( \frac{\lambda_0 \lambda_1 \lambda_2 \dots \lambda_{n-1}}{\mu_1 \mu_2 \mu_3 \dots \mu_n} \right)} \quad (6)$$

From these equations, one can determine the values of  $P_n$  in terms of  $P_0$  and the value of  $P_0$ . Now let us see equations 3 and 4 are called the balance equations this equation 3 and 4 okay this 3 and this 4 are called balance equations or equilibrium equations of the Poisson queue systems. Using the principle of mathematical induction, it can be shown that this  $P_n$  from this equation okay using the principle of mathematical induction one can show that  $P_n = \lambda_0 \lambda_1 \lambda_2 \dots \lambda_{n-1} / \mu_1 \mu_2 \mu_3 \dots \mu_n P_0$ .

And  $P_0 = 1 / (1 + \sum_{n=1}^{\infty} \lambda_0 \lambda_1 \lambda_2 \lambda_{n-1} / \mu_1 \mu_2 \mu_3 \dots \mu_n)$  and so on  $\mu_n$ .

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### Model-I (M/M/1): ( $\infty$ /FIFO) Single Server With Infinity Capacity

For this model, we make the following assumptions:

- (a) The mean arrival rate is constant, i.e.  $\lambda_n = \lambda$  for all  $n$ .
- (b) The mean service rate is constant i.e.  $\mu_n = \mu$  for all  $n$ .
- (c) The mean arrival rate is less than the mean service rate, i.e.  $\lambda < \mu$  or equivalently that  $\rho = \frac{\lambda}{\mu} < 1$ .

The assumption that the traffic intensity  $\rho < 1$  ensures that an infinite queue will not build up. Under assumption (c) and equation (6) simplifies to

$$P_0 = \frac{1}{1 + \sum_{n=1}^{\infty} \left( \frac{\lambda}{\mu} \right)^n} = \frac{1}{1 + \frac{\lambda}{\mu} + \left( \frac{\lambda}{\mu} \right)^2 + \dots} = \frac{1}{\left( 1 - \frac{\lambda}{\mu} \right)^{-1}} = 1 - \frac{\lambda}{\mu} = 1 - \rho \quad (7)$$

Now let us consider the model MM1 infinity oblique FIFO single server with infinite capacity. So here a is m b is also m that means we are considering exponential type distribution number 3, 1 means we are considering single server okay, infinity means the

queue can be infinite okay and FIFO means we are considering the discipline as first in first out okay, first in first out FIFO okay.

That is the queue discipline we are considering,  $a$  is arrival distribution,  $b$  is service time distribution, both we are assuming as that they follow exponential type distribution,  $c$  is the number of channels or servers. So we are considering that the system has one server and  $d$  is the system capacity, system capacity we are taking as infinite. So we have MM1 infinity FIFO model okay.

For this model, we make following assumptions. The mean arrival rate is constant that is  $\lambda_n = \lambda$  for all  $n$ . Mean service rate is constant that is  $\mu_1 = \mu$  for all  $n$ . Mean arrival rate is  $<$  the mean service rate. So that means  $\lambda < \mu$  or we can say  $\rho = \lambda/\mu < 1$ ,  $\rho$  is the traffic intensity, so the traffic intensity we are assuming that it is  $< 1$  and which will ensure that an infinite queue will not build up.

Because mean service rate is higher than the mean arrival rate, so the infinite queue will not build up. Now under assumption  $c$  that is mean arrival rate is  $<$  the mean service rate okay, the equation 6 simplifies to let us see the equation 6, this equation 6 becomes  $1/(1 + \lambda/\mu + \lambda^2/\mu^2 + \dots + \lambda^{n-1}/\mu^{n-1})$ , so we have  $\lambda/\mu$  raised to the power  $n$  okay.

So we get the  $1/(1 + \lambda/\mu + \lambda^2/\mu^2 + \dots + \lambda^{n-1}/\mu^{n-1})$  which is  $1/(1 + \lambda/\mu + \lambda^2/\mu^2 + \dots)$  and so on which is geometric series with ratio  $\lambda/\mu$ . So we have the sum of the series as  $1/(1 - \lambda/\mu)$ . So  $1/(1 - \lambda/\mu)$  is  $1/(1 - \rho)$  and  $\lambda/\mu$  we denote by  $\rho$  okay. So  $P_0$  is  $1 - \rho$ .

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Substituting equation (7) into (5), we have

$$P_n = \left(\frac{\lambda}{\mu}\right)^n P_0 = \rho^n (1 - \rho)$$

#### Characteristic of Model-I

**1. Average or expected number of customers in the system ( $L_s$ ):** If  $N$  denotes the number of customers in the Queueing system, then  $N$  is a discrete random variable taking values  $N = 0, 1, 2, \dots$  and we know that

$$P_n = P(N = n) = \rho^n (1 - \rho), \text{ for } n = 0, 1, 2, \dots$$

Now let us consider the case substituting equation 7 into 5 okay. So this  $P_0 = 1 - \rho$ , let us put in equation number 5. Equation number 5 now is  $P_n = (\lambda/\mu)^n P_0$  okay because we are assuming that the mean service rate and the mean arrival rate are constant okay. So we have the  $P_n = (\lambda/\mu)^n P_0$  okay and this is  $\rho$  to the power  $n$  and  $P_0 = 1 - \rho$ , so we have  $P_n = \rho^n (1 - \rho)$  okay.

Now characteristics of model 1 okay so average or expected number of customers in the system. Average or expected numbers or number of customers in the system we denote by  $L_s$ . So if  $N$  denotes the number of customers in the queueing system, then  $N$  is a discrete random variable. It will take values 0, 1, 2, 3 and so on and we know that  $P_n$  is the probability that  $n$  takes the value  $N$  okay. So it is  $\rho^n (1 - \rho)$  as we have seen just now for  $n = 0, 1, 2, 3$  and so on.

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#### Average or expected number of customers in the system ( $L_s$ )

Thus, the expected number of customers in the system is given by

$$L_s = \underline{E(N)} = \sum_{n=0}^{\infty} n P_n = \sum_{n=0}^{\infty} n \rho^n (1 - \rho) = (1 - \rho) \sum_{n=1}^{\infty} n \rho^n$$

We note that

$$L_s = \rho(1 - \rho) \sum_{n=1}^{\infty} n \rho^{n-1} = \rho(1 - \rho)[1 + 2\rho + 3\rho^2 + \dots] = \rho(1 - \rho)(1 - \rho)^{-2} = \rho(1 - \rho)^{-1}$$

$$\Rightarrow L_s = \frac{\rho}{1 - \rho} = \frac{\frac{\lambda}{\mu}}{1 - \frac{\lambda}{\mu}} = \frac{\lambda}{\mu - \lambda}$$

Now thus the expected number of customers in the system  $L_s$  is given by expected value of  $n$  that is  $\sum_{n=0}^{\infty} n P_n$  which is  $\sum_{n=0}^{\infty} n \rho^n$  to the power  $n-1$ . Now when  $n=0$ , first term is 0, so we can write it as  $1-\rho$  is independent of  $n$ , so we can write it outside and then  $\sum_{n=1}^{\infty} n \rho^n$  to the power  $n$ . Now  $L_s$  can also be written as  $\rho \times 1-\rho \times \sum_{n=1}^{\infty} n \rho^{n-1}$ .

So we have  $\rho \times 1-\rho$  and then this series  $\sum_{n=1}^{\infty} n \rho^{n-1}$  and  $\rho$  to the power  $n-1$  becomes  $1+2\rho+3\rho^2$  and so on okay and its sum we know it is  $1-\rho$  to the power  $-2$  okay. So we have  $\rho \times 1-\rho$  to the power  $-1$  and therefore  $L_s = \rho / (1-\rho)$ . Now  $\rho$  is  $\lambda/\mu$ , so we can put the value of  $\rho$  as  $\lambda/\mu$  and we get  $L_s = \lambda/\mu - \lambda$ . So that means that average or expected number of customers in the system is  $\lambda/\mu - \lambda$ .

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Average or expected number of customers in the queue or average length of the queue ( $L_q$ )

The queue length is given by

*Queue length = Line length – Number of customers being served*

Thus, if the number of customers in the system is  $n$ , then the number of customers in the queue or the queue length is  $n - 1$ . Hence the expected length of the queue is

$$L_q = E(N - 1) = \sum_{n=1}^{\infty} (n - 1) P_n = \sum_{n=1}^{\infty} (n - 1) (1 - \rho) \rho^n$$

$$\Rightarrow L_q = E(N - 1) = \sum_{n=2}^{\infty} (n - 1) (1 - \rho) \rho^n = \sum_{m=1}^{\infty} m (1 - \rho) \rho^{m+1}$$

by substituting  $m = n - 1$  in the summation.

Now average or expected number of customers in the queue or average length of the queue. We know that queue length is given by line length–number of customers being served. Thus, if the number of customers in the system is  $n$ , then the number of customers in the queue or the queue length is  $n-1$ . Hence the expected length of the queue is  $L_q = E(N-1)$  and so it is  $\sum_{n=1}^{\infty} (n-1) P_n$  okay

Now when  $n=1$ , the first term becomes 0 okay or before that what we can do is we can put the value of  $P_n$ ,  $P_n$  is  $\rho^n$  to the power  $n-1$ . So this summation becomes  $\sum_{n=1}^{\infty} (n-1) \rho^n$  and therefore using the fact that at  $n=1$  the first term becomes 0 and will start with 2 onwards, so  $n=2$  to infinity  $(n-1) \rho^n$  to the power  $n$  and here what we can do is let us put  $m=n-1$  so that  $m$  runs from 1 to infinity and we have  $m$  times  $1-\rho \times \rho$  to the power  $m+1$ .

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Average or expected number of customers in the queue or average length of the queue ( $L_q$ ) cont...

Hence

$$L_q = \rho^2(1-\rho) \sum_{m=1}^{\infty} m \rho^{m-1} = \rho^2(1-\rho)(1-\rho)^{-2} = \rho^2(1-\rho)^{-1}$$

$$\Rightarrow L_q = \frac{\rho^2}{1-\rho} = \frac{\left(\frac{\lambda}{\mu}\right)^2}{1-\frac{\lambda}{\mu}} = \frac{\lambda^2}{\mu(\mu-\lambda)}$$

And therefore I can write  $L_q = \rho^2 \sum_{m=1}^{\infty} m \rho^{m-1}$  to infinity  $m$  times  $\rho$  to the power  $m-1$  okay and this we know the sum of this is  $1-\rho$  to the power-2, so we have  $\rho^2 \times 1-\rho$  to the power-1 and therefore  $L_q$  which is the average or expected number of customers in the queue is  $\rho^2 / 1-\rho$ .  $\rho$  is  $\lambda/\mu$ , so  $\lambda/\mu$  whole square upon  $1-\lambda/\mu$  which gives us  $L_q$  as  $\lambda^2 / \mu(\mu-\lambda)$ .

(Refer Slide Time: 19:30)

Average or expected number of customers in nonempty queue ( $L_w$ )

$$L_w = E\{N-1 | N-1 > 0\} = \frac{E(N-1)}{P(N-1 > 0)} \quad (8)$$

We know that

$$E(N-1) = L_q = \frac{\rho^2}{1-\rho}$$

and

$$P(N-1 > 0) = 1 - P_0 - P_1 = 1 - (1-\rho) - [\rho(1-\rho)] = \rho - \rho(1-\rho) = \rho^2.$$

Substituting the above in equation (8), we get

$$L_w = \frac{\frac{\rho^2}{1-\rho}}{\rho^2} = \frac{1}{1-\rho} = \frac{1}{1-\frac{\lambda}{\mu}} = \frac{\mu}{\mu-\lambda}$$

Now let us consider the case average or expected number of customers in non-empty queue. So this means that  $L_w$ .  $L_w$  is the expectation of  $N-1$ ,  $N-1$  is the line length okay divided by  $q$   $N-1$  sorry  $L_w = \text{expectation of } N-1 \text{ given that } N-1 > 0$  that is the queue is non-empty okay, at least one customer is there okay, so  $N-1$  is  $> 0$  and therefore expectation of  $N-1$  okay expectation of  $N-1$ ,  $N$  is no,  $L_w = \text{expectation of } N-1 \text{ given that } N-1 \text{ is } > 0$ .  $N$  is the number of customers,  $N$  denotes the number of customers in the queue system.

So number of customers is  $>1$  okay, so expectation of  $N-1$ /probability that  $N-1$  is  $>0$  and we know that expectation of  $N-1$  that is  $L_q$  okay  $= \rho^2 / (1-\rho)$  okay.  $L_q = \rho^2 / (1-\rho)$ . So we have  $P(N-1 > 0) = 1 - P_0 - P_1$ .  $P_0$  means probability that  $N=0$  and  $P_1$  is the probability that  $N=1$ , so probability that  $N > 1$  can be obtained from  $1 - P_0 - P_1$  so  $1 - P_0$  is  $1 - \rho$  and  $P_1 = \rho$  times  $1 - \rho$  because  $P_n = \rho^n$  to the power  $n \cdot 1 - \rho$  for all  $n \geq 1$ .

So we have on simplifying this we get  $\rho^2$  okay. So  $P(N-1 > 0)$  is  $\rho^2$ , so  $L_w = \rho^2 / (1-\rho) / \rho^2$  which is  $1 / (1-\rho)$  and that is  $\mu / (\mu - \lambda)$  as  $\rho = \lambda / \mu$ .  
(Refer Slide Time: 21:47)

The probability that the number of customers in the system exceeds k

$$P(N > k) = \sum_{n=k+1}^{\infty} P_n = \sum_{n=k+1}^{\infty} \rho^n (1-\rho) = \rho^{k+1} (1-\rho) \sum_{n=k+1}^{\infty} \rho^{n-k-1} = \rho^{k+1} (1-\rho) \sum_{m=0}^{\infty} \rho^m$$

by substituting  $m = n - k - 1$  in the summation. Thus it follows that

$$P(N > k) = \rho^{k+1} (1-\rho) (1-\rho)^{-1} = \rho^{k+1} = \left( \frac{\lambda}{\mu} \right)^{k+1}.$$

Now let us discuss the probability that number of customers in the system exceeds k okay. So probability that number of customers in the system exceeds k that is  $N > k$ . Now this is  $\sum_{n=k+1}^{\infty} P_n$   $n=k+1$  to infinity, the value of  $P_n$  is  $\rho^n$  to the power  $n \cdot 1 - \rho$ . We can write it as  $\rho^{k+1} \cdot 1 - \rho \cdot \sum_{n=k+1}^{\infty} \rho^{n-k-1}$  and this is  $\rho^{k+1} \cdot 1 - \rho$  and we can put  $m$  as  $n - k - 1$ .

When you take  $m = n - k - 1$  then this summation becomes  $\sum_{m=0}^{\infty} \rho^m$  to infinity  $\rho$  to the power  $m$  and hence probability that  $n > k$  is  $\rho^{k+1} \cdot 1 - \rho$  and this is geometric series with geometric ratio  $\rho$ , its sum is  $1 / (1 - \rho)$ . So we get  $1 - \rho$  to the power  $-1$  and this will give you on simplification the probability that  $N > k = \rho^{k+1}$  that is  $\lambda / \mu$  to the power  $k+1$ .

So this is the probability  $\lambda / \mu$  to the power  $k+1$  is the probability that number of customers in the system exceeds k.  
(Refer Slide Time: 23:09)



Probability density function of the waiting time of a customer in the system ( $f_s(w)$ )

Let  $T_s$  be the random variable that represents the waiting time of a customer in the system,  $f_s(w)$  be the probability density function of  $T_s$  and  $f_s(w|n)$  be the probability density function of  $T_s$  given that there are already  $n$  customers in the system when the customer arrives. Then

$$f_s(w) = (\mu - \lambda)e^{-\mu w} \sum_{n=0}^{\infty} \frac{(\lambda w)^n}{n!} = (\mu - \lambda)e^{-\mu w} e^{\lambda w}, \text{ where } w > 0.$$

Thus, the probability density function of the waiting time of the customer in the system is

$$f_s(w) = (\mu - \lambda)e^{-(\mu - \lambda)w}, \text{ where } w > 0 \quad (9)$$

which is the probability density function of the exponential random variable with parameter  $\mu - \lambda$ .

Now let us discuss probability density function of the waiting time of a customer in the system, we denote it by  $f_s w$ . So let  $T_s$  be the random variable that represents the waiting time of a customer in the system,  $f_s w$  be the probability density function of  $T_s$  and  $f_s w$  given  $n$  be the probability density function of  $T_s$  given that there are already  $n$  customers in the system when the customer arrives.

Then,  $f_s w$  is given by  $\mu - \lambda \cdot e^{-\mu w} \sum_{n=0}^{\infty} \frac{(\lambda w)^n}{n!}$ . Now the sum of the series we are all aware, the sum of the series is  $e^{\lambda w}$ , so we get  $\mu - \lambda \cdot e^{-\mu w} e^{\lambda w}$  where  $w > 0$ . Thus, the probability density function of the waiting time of the customer in the system is  $f_s w = \mu - \lambda \cdot e^{-(\mu - \lambda)w}$  where  $w > 0$ .

This is the probability density function of the exponential random variable with parameter  $\mu - \lambda$ .

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The distribution function of  $T_s$  is given by

$$F_s(w) = P(T_s \leq w) = \begin{cases} 0 & \text{if } w < 0 \\ 1 - e^{-(\mu-\lambda)w} & \text{if } w \geq 0. \checkmark \end{cases}$$

Average or expected waiting time of a customer in the system ( $W_s$ )

The waiting time of a customer in the system is the random variable  $T_s$ , which is exponentially distributed with parameter  $\mu - \lambda$  as given in equation (9). Thus, the average waiting time of the customer in the system is given by

$$W_s = E[T_s] = \frac{1}{\mu - \lambda}.$$

$F_s(w) = \int_0^w f_s(w) dw$   
 $= \int_0^w (\mu - \lambda) e^{-(\mu - \lambda)w} dw$   
 $= \left[ -e^{-(\mu - \lambda)w} \right]_0^w = 1 - e^{-(\mu - \lambda)w}$

Now the distribution function, we know distribution function  $f_s w$  is=probability that  $T_s$  is  $\leq w$ . So when  $w$  is  $< 0$  okay when  $w$  is  $< 0$  because we know that  $f_s w$ ,  $f_s w$  is the probability density function of the random variable, this is  $= \mu - \lambda * e$  to the power  $-\mu - \lambda w$  when  $w$  is  $> 0$ . So when  $w$  is  $< 0$  okay probability is 0, so we get  $f_s w = 0$ . So when  $w$  is  $\geq 0$ , this  $f_s w$  is  $= \int_0^w$  because we want the probability that  $T_s$  is  $\leq w$ .

So  $\int_0^w f_s w dw$  and this will be  $= \int_0^w (\mu - \lambda) e^{-(\mu - \lambda)w} dw$  okay and this is nothing but  $-e$  to the power  $-(\mu - \lambda)w$  from 0 to  $w$ . So this is  $= 1 - e^{-(\mu - \lambda)w}$  okay, so this is the distribution function of  $T_s$ . Now average or expected waiting time of a customer in the system. So the waiting time of a customer in the system is the random variable  $T_s$  which is exponentially distributed with parameter  $\mu - \lambda$  as given in equation 9, here okay  $\mu - \lambda$ .

So thus the average waiting time of a customer in the system is  $1/(\mu - \lambda)$  because it follows exponential distribution okay.

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The probability that the waiting time of a customer in the system exceeds t

$$P(T_s > t) = \int_{w=t}^{\infty} f_s(w) dw = \int_{w=t}^{\infty} (\mu - \lambda) e^{-(\mu - \lambda)w} dw = \int_{w=t}^{\infty} d[-e^{-(\mu - \lambda)w}]$$

$$\Rightarrow P(T_s > t) = \left[ -e^{-(\mu - \lambda)w} \right]_{w=t}^{\infty} = e^{-(\mu - \lambda)t}$$

Probability density function of the waiting time of a customer in the queue ( $f_q(w)$ )

The probability density function of the waiting time of a customer in the queue is

$$f_q(w) = \begin{cases} \frac{\lambda}{\mu} (\mu - \lambda) e^{-(\mu - \lambda)w} & \text{for } w > 0 \\ 1 - \frac{\lambda}{\mu} & \text{for } w = 0 \\ 0 & \text{for } w < 0 \end{cases}$$

Now the probability that the waiting time of a customer in the system exceeds t, so probability that  $T_s > t$  is integral over  $w=t$  to infinity  $f_s w dw$ ,  $f_s w$  is  $\mu - \lambda$  \*  $e$  to the power  $-(\mu - \lambda)w$ , so on integration we get probability that  $T_s > t = e$  to the power  $-(\mu - \lambda)t$ . So this is the probability that waiting time of a customer in the system exceeds t.

Probability density function of the waiting time of a customer in the queue  $f_q w$  is given by this the probability density function of the waiting time of the customer in the queue is  $f_s w = \lambda / \mu$  \*  $\mu - \lambda$  \*  $e$  to the power  $-(\mu - \lambda)w$  for  $w > 0$ ,  $1 - \lambda / \mu$  okay for  $w = 0$  and 0 for  $w < 0$ .

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The distribution function of the waiting time of the customer in the queue is

$$F_q(w) = P(T_q \leq w) = \begin{cases} 0 & \text{for } w < 0 \\ 1 - \frac{\lambda}{\mu} & \text{for } w = 0 \\ \frac{\lambda}{\mu} - \frac{\lambda}{\mu} e^{-(\mu - \lambda)w} & \text{for } w > 0 \end{cases}$$

*Handwritten notes:*  $\int_0^w \frac{\lambda}{\mu} (\mu - \lambda) e^{-(\mu - \lambda)w} dw$ ,  $F_q(w) = \int_0^w f_q(w) dw$ ,  $\frac{\lambda}{\mu} [1 - e^{-(\mu - \lambda)w}]$ ,  $\frac{\lambda}{\mu}$

The probability that the waiting time of a customer in the system exceeds t

$$P(T_q > t) = \int_t^{\infty} f_q(w) dw = \int_t^{\infty} \frac{\lambda}{\mu} (\mu - \lambda) e^{-(\mu - \lambda)w} dw = \frac{\lambda}{\mu} \int_t^{\infty} d[-e^{-(\mu - \lambda)w}]$$

$$\Rightarrow P(T_q > t) = \frac{\lambda}{\mu} \left[ -e^{-(\mu - \lambda)w} \right]_t^{\infty} = \frac{\lambda}{\mu} [0 + e^{-(\mu - \lambda)t}] = \frac{\lambda}{\mu} e^{-(\mu - \lambda)t}$$

The distribution function of the waiting time of the customer in the queue is  $f_q w$  which is probability that  $T_q \leq w$  okay probability that  $T_q \leq w$ , so we have defined  $T_q$  earlier,  $f_s w$  is the probability density function of  $T_s$ . where  $T_s$  is the random variable that represents

the waiting time of a customer in the system, and  $T_q$  denotes the random variable that represents the waiting time of a customer in the queue okay.

So this is the probability density function of the waiting time of the customer in the queue okay where  $T_q$  is the random variable which denotes the waiting time of the customer in the queue. So  $f_s(w)$  is given by this and then we have this is not  $f_s(w)$ , this is  $f_q(w)$  okay. So this is  $f_q(w)$ ,  $f_q(w)$  is given by these values okay,  $f_q(w) = \lambda/\mu e^{-(\mu-\lambda)w}$  for  $w > 0$ ,  $1-\lambda/\mu$  for  $w=0$  and 0 for  $w < 0$  okay.

And then the distribution function of the waiting time of the customer in the queue is given by  $F_q(w)$  where probability that  $T_q$  is  $\leq w$ . So this can be, these are given by 0 when  $w < 0$ ,  $1-\lambda/\mu$  when  $w=0$ ,  $1-\lambda/\mu + \lambda/\mu e^{-(\mu-\lambda)w}$  for  $w > 0$ . We can easily obtain these values because  $f_q(w) = \int_{-\infty}^w f_q(w) dw$  okay and  $f_q(w)$  we have here  $f_q(w)$  is 0 when  $w < 0$  okay.

So this will go to 0 to  $w f_q(w) dw$  okay, so from here we can easily find these values. Now the probability that the waiting time of a customer in the system exceeds  $t$ . Here 0 to  $w$  when  $w > 0$  we can put the value of  $f_q(w)$  okay  $f_q(w)$  for  $w > 0$  is given by this expression okay  $\lambda/\mu e^{-(\mu-\lambda)w}$ . So when you integrate this okay we will get  $\lambda/\mu [1 - e^{-(\mu-\lambda)w}]$  okay.

So we will get this, if you integrate this, 0 to  $w f_q(w) dw$  okay, what we will get 0 to  $w \lambda/\mu e^{-(\mu-\lambda)w} dw$  and this will be  $\lambda/\mu$  here we will have  $-e^{-(\mu-\lambda)w}$  0 to  $w$ . So we will get this as  $\lambda/\mu [1 - e^{-(\mu-\lambda)w}]$  okay. So that is the formula for probability that  $T_q$  is  $\leq w$ . Now probability that the waiting time of a customer in the system exceeds  $t$  okay.

Probability that  $T_q > t$ , so this is  $\int_t^{\infty} f_q(w) dw$  and  $f_q(w) dw$  is  $\lambda/\mu e^{-(\mu-\lambda)w} dw$ . So this  $\lambda/\mu$  and this quantity is the derivative of  $-e^{-(\mu-\lambda)w}$  okay. So probability that  $T_q > t$  is  $\lambda/\mu$  times  $-e^{-(\mu-\lambda)w}$  integral over the limits are  $t$  to infinity, so  $\lambda/\mu$  now this goes to 0 when the  $w$  goes to infinity.

So this 0 and when you put the lower limit, you get  $0 + e^{-(\mu-\lambda)t}$ . So we get  $\lambda/\mu e^{-(\mu-\lambda)t}$  okay.

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The average waiting time of a customer in the queue ( $W_q$ )

$$W_q = \frac{\lambda}{\mu(\mu - \lambda)} \checkmark$$

The average waiting time of a customer in the queue if he/she has to wait

It is given by

$$E[T_q | T_q > 0] = \frac{E(T_q) \checkmark}{P(T_q > 0)} = \frac{W_q \checkmark}{1 - P(T_q = 0)} = \frac{W_q}{1 - (1 - \frac{\lambda}{\mu})} = \frac{W_q}{\frac{\lambda}{\mu}} = \frac{W_q \mu}{\lambda} = \frac{\frac{\lambda}{\mu(\mu - \lambda)} \mu}{\lambda} = \frac{1}{\mu - \lambda} \checkmark$$

Now the average waiting time of a customer in the queue  $W_q$  is given by  $\lambda/\mu * \mu - \lambda$ . The average waiting time of a customer in the queue if he, she has to wait is given by  $E$  expectation of  $T_q$  where  $T_q$  is  $> 0$ . Expectation of  $T_q$  is nothing but  $W_q$  and  $T_q > 0$  is  $1 - P(T_q = 0)$ . Now expectation of  $T_q$  given  $T_q > 0$  is then  $W_q = \lambda/\mu$  times  $\mu - \lambda$  and probability that  $T_q = 0$  okay.

Probability that  $T_q = 0$  we have seen here probability that  $T_q = 0$  okay  $= 1 - \lambda/\mu$  okay, probability that  $T_q = 0 = 1 - \lambda/\mu$ , so  $1 - \text{this}$  is  $W_q / (1 - \lambda/\mu)$  and this gives you  $W_q / (\lambda/\mu)$  but  $W_q = \lambda/\mu * \mu - \lambda / (\lambda/\mu)$ . So this is  $1/\mu - \lambda$ , so we get this.

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Theorem (Little's formulae)

For the M/M/1:∞/FIFO model, the relations between  $L_s$ ,  $L_q$ ,  $W_s$  and  $W_q$  are given as follows:

(a)  $L_s = \lambda W_s$

(b)  $L_q = \lambda W_q$

(c)  $W_s = W_q + \frac{1}{\mu}$

(d)  $L_s = L_q + \frac{\lambda}{\mu}$

Now let us discuss Little's formulae for the MM1 infinity FIFO model. The relation between  $L_s$ ,  $L_q$ ,  $W_s$  and  $W_q$  are given as follows.  $L_s = \lambda W_s$ ,  $L_q = \lambda W_q$ ,  $W_s = W_q + 1/\mu$ ,

$L_s = L_q + \lambda/\mu$ . So first let us see  $L_s = \lambda W_s$  okay. We know that  $W_s$  is  $1/(\mu - \lambda)$ .  $L_s$  is what?  $L_s$  is  $\lambda/(\mu - \lambda)$  okay.  
(Refer Slide Time: 34:54)

The distribution function of  $T_s$  is given by

$$F_s(w) = P(T_s \leq w) = \begin{cases} 0 & \text{if } w < 0 \\ 1 - e^{-(\mu - \lambda)w} & \text{if } w \geq 0. \end{cases}$$

Average or expected waiting time of a customer in the system ( $W_s$ )

The waiting time of a customer in the system is the random variable  $T_s$ , which is exponentially distributed with parameter  $\mu - \lambda$  as given in equation (9). Thus, the average waiting time of the customer in the system is given by

$$W_s = E[T_s] = \frac{1}{\mu - \lambda}$$

Handwritten notes:  $L_s = \frac{\lambda}{\mu - \lambda}$ ,  $W_s = \frac{1}{\mu - \lambda} \Rightarrow L_s = \lambda W_s$ ,  $F_s(w) = \int_0^w f_s(w) dw = \int_0^w (\mu - \lambda) e^{-(\mu - \lambda)w} dw = [-e^{-(\mu - \lambda)w}]_0^w = 1 - e^{-(\mu - \lambda)w}$

$L_s$  is  $\lambda/(\mu - \lambda)$ ,  $W_s$  is  $1/(\mu - \lambda)$  okay, so these 2 relations imply that  $L_s = \lambda W_s$  okay. So this is how we get the first formulae  $L_s = \lambda W_s$ . Now let us look  $L_q$  and  $W_q$ . Let us look at  $L_q$ .  $L_q = \lambda^2/(\mu(\mu - \lambda))$  okay.  
(Refer Slide Time: 35:28)

The average waiting time of a customer in the queue ( $W_q$ )

$$W_q = \frac{\lambda}{\mu(\mu - \lambda)}$$

Handwritten notes:  $L_q = \frac{\lambda^2}{\mu(\mu - \lambda)}$ ,  $L_q = \lambda W_q$

The average waiting time of a customer in the queue if he/she has to wait

It is given by

$$E[T_q | T_q > 0] = \frac{E(T_q)}{P(T_q > 0)} = \frac{W_q}{1 - P(T_q = 0)} = \frac{W_q}{1 - \frac{\lambda}{\mu}} = \frac{W_q}{\frac{\mu - \lambda}{\mu}} = \frac{\lambda}{\mu - \lambda}$$

Handwritten notes:  $E[T_q | T_q > 0] = \frac{\frac{\lambda^2}{\mu(\mu - \lambda)}}{\frac{\mu - \lambda}{\mu}} = \frac{\lambda}{\mu - \lambda}$

So  $L_q = \lambda^2/(\mu(\mu - \lambda))$  okay and  $W_q$  is  $\lambda/(\mu(\mu - \lambda))$ . So,  $L_q = \lambda W_q$  okay, so  $L_q = \lambda W_q$ .  
(Refer Slide Time: 35:51)



### Theorem (Little's formulae)

For the M/M/1:∞/FIFO model, the relations between  $L_s$ ,  $L_q$ ,  $W_s$  and  $W_q$  are given as follows:

(a)  $L_s = \lambda W_s$  ✓

(b)  $L_q = \lambda W_q$  ✓

(c)  $W_s = W_q + \frac{1}{\mu}$

(d)  $L_s = L_q + \frac{\lambda}{\mu}$

$$W_q + \frac{1}{\mu} = \frac{\lambda}{\mu(\mu-\lambda)} + \frac{1}{\mu}$$

$$= \frac{1}{\mu} \left[ \frac{\lambda + \mu - \lambda}{\mu - \lambda} \right] = \frac{1}{\mu - \lambda} = W_s$$

$$L_q + \frac{\lambda}{\mu} = \frac{\lambda^2}{\mu(\mu-\lambda)} + \frac{\lambda}{\mu}$$

$$= \frac{\lambda}{\mu} \left[ \frac{\lambda}{\mu-\lambda} + 1 \right] = \frac{\lambda}{\mu} \frac{\mu}{\mu-\lambda} = \frac{\lambda}{\mu-\lambda} = L_s$$

Now let us show that  $W_q + 1/\mu$ ,  $W_q + 1/\mu = W_s$  is  $= \lambda/\mu$  times  $\mu - \lambda$ . So this  $= 1/\mu \lambda + \mu - \lambda/\mu - \lambda$ . So this is  $1/\mu - \lambda$  and  $1/\mu - \lambda$  is  $W_s$ .  $L_s$  is this,  $W_s$  is here,  $W_s$  is  $1/\mu - \lambda$ . So this is  $W_s$  okay and then  $L_q + \lambda/\mu$ .  $L_q$  is  $= \lambda^2/\mu$  times  $\mu - \lambda$ . That is the value of  $L_q + \lambda/\mu$ . So this is  $\lambda/\mu$  times  $\lambda/\mu - \lambda + 1$ .

And we get here  $\lambda/\mu \lambda + \mu - \lambda$  that means  $\mu/\mu - \lambda$ . So we can say we get  $\lambda/\mu - \lambda$  and  $\lambda/\mu - \lambda$  is  $L_s$  okay,  $L_s$   $\lambda/\mu - \lambda$ . So if one parameter if  $L_s$ ,  $W_s$  or  $L_q$  or  $W_q$  one is known, the other three can be found from Little's formulae, so they are very useful.

(Refer Slide Time: 37:37)

### Proof

From the characteristics of M/M/1:∞/FIFO model, we know that

$$L_s = \frac{\lambda}{\mu - \lambda}, L_q = \frac{\lambda^2}{\mu(\mu - \lambda)}, W_s = \frac{1}{\mu - \lambda} \text{ and } W_q = \frac{\lambda}{\mu(\mu - \lambda)} \quad (10)$$

(a) and (b) are immediate from equation (10).

(c) We note that

$$W_q + \frac{1}{\mu} = \frac{\lambda}{\mu(\mu - \lambda)} + \frac{1}{\mu} = \frac{1}{\mu - \lambda} = W_s$$

(d) We note that

$$L_q + \frac{\lambda}{\mu} = \frac{\lambda^2}{\mu(\mu - \lambda)} + \frac{\lambda}{\mu} = \frac{\lambda}{\mu - \lambda} = L_s$$

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### Remark

Little's formulae are very useful in the study of the  $M/M/1:\infty/FIFO$  model because if any of the quantities  $L_s$ ,  $L_q$ ,  $W_s$  and  $W_q$  are available, then the other three quantities can be readily obtained by Little's formulae given in the above theorem.

This is the proof. Now so the Little's formulae are very useful in the study of  $MM1$  infinity  $FIFO$  model because if any of the quantities  $L_s$ ,  $L_q$ ,  $W_s$  and  $W_q$  are available, then the other three can be readily obtained from Little's formulae which are given in the above theorem.  
(Refer Slide Time: 37:56)

### Example 1

What is the probability that a customer has to wait more than 15 minutes to get his service completed in  $M/M/1:\infty/FIFO$  queue system if  $\lambda = 6$  per hour and  $\mu = 10$  per hour?

Ans:  $e^{-1}$ .

$$P(T_p > t) = e^{-(\mu - \lambda)t} = e^{-(10-6)t} = e^{-4t}$$

$$\text{Here } t = 15 = \frac{1}{4} \text{ Hours}$$

$$P(T_p > \frac{1}{4}) = e^{-4 \cdot \frac{1}{4}} = \underline{\underline{e^{-1}}}$$

Now let us look at the first question. What is the probability that a customer has to wait more than 15 minutes to get his service completed in  $MM1$   $FIFO$  queue system? If  $\lambda=6$  per hour and  $\mu=10$  per hour okay. So we are given the mean service rate that is  $\mu=10$  per hour and mean arrival rate  $\lambda=6$  okay. Now let us see which formula gives us the probability that a customer has to get wait more than 15 minutes to get the service okay.

So let us go to that formula. So the probability that the waiting time of a customer in the system exceeds  $t$  okay. So if the customer has to wait more than 15 minutes then probability that  $T_s$  is  $\geq 15$  we have to find, so we have probability that  $T_s > t$  okay. So we have

probability that  $T_s \geq t$  is  $e^{-\mu t}$ . So this equal to  $e^{-\mu t}$ . So  $\mu$  is 10,  $\lambda$  is 6\*t so we get  $e^{-4t}$ .

So if the customer has to wait more than  $t$  minutes then probability is  $T_s > t$ . Now here we want the probability that the system that the customer has to wait more than 15 minutes okay. So  $t$  here  $t=15$ . okay so probability that  $T_s \geq 15$ , now 15 we have to convert into hours, so  $1/4$  hours, so probability that  $T_s \geq 1/4 = e^{-4 \cdot 1/4}$  which is  $e^{-1}$  okay. So the probability that the customer has to wait more than 15 minutes to get the service is  $e^{-1}$ .

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**Example 2**

Consider an M/M/1 Queueing system. Find the probability that there are at least  $n$  customers in the system.

Ans:  $\left(\frac{\lambda}{\mu}\right)^n$

$$P(N \geq n) = \sum_{k=n}^{\infty} P_k = \sum_{k=n}^{\infty} e^{-\rho} \frac{\rho^k}{k!} = e^{-\rho} \frac{\rho^n}{n!} \sum_{k=n}^{\infty} \frac{\rho^{k-n}}{(k-n)!} = e^{-\rho} \frac{\rho^n}{n!} \sum_{i=0}^{\infty} \frac{\rho^i}{i!}$$

$$= e^{-\rho} \frac{\rho^n}{n!} \cdot \frac{1}{1-\rho} = \rho^n = \left(\frac{\lambda}{\mu}\right)^n$$

Now find consider MM1 queue system find the probability that there are at least  $N$  customers in the system, so number of customers are  $\geq n$ . let us go to that formula. Probability that number of the customers in the system exceeds  $k$  okay. So probability that the number of customers in the system exceeds  $k$  is given by this formula probability that  $N > k = \sum_{n=k+1}^{\infty} P_n$  okay.

And  $P_n$  is  $\rho^n$ . So let us go to that formula. So we want the probability that there are at least  $n$  customers. So probability that  $N \geq n$ . So probability that  $\sum_{k=n}^{\infty} P_k$  okay so this is  $\sum_{k=n}^{\infty} \rho^k$  okay so we will have  $\rho^n$  we can write outside and here we will have  $\sum_{k=n}^{\infty} \rho^{k-n}$ .

And  $k-n$  we can write as  $i$  then this is  $\rho^n \sum_{i=0}^{\infty} \rho^i$ . So we will have the  $\rho^n$  and this is  $1/(1-\rho)$ . So this will be  $\rho^n$  and  $\rho$  is  $\lambda/\mu$  okay. So  $\lambda/\mu$  raised to the power  $n$ . So

probability that there are at least  $n$  customers in the system is given by  $\lambda/\mu$  raised to the power  $n$ .

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### Example 3

Consider an M/M/1 Queueing system. If  $\lambda = 6$  and  $\mu = 8$ , find the probability of at least 10 customers in the system.

Ans:  $(0.75)^{10} = 0.0563$ .

$$P(N \geq 10) = \left(\frac{\lambda}{\mu}\right)^{10} = \left(\frac{6}{8}\right)^{10} = (0.75)^{10} = 0.0563$$

Now consider an MM1 queueing system. If  $\lambda=6$ ,  $\mu=8$  find the probability of at least 10 customers in the system. So we have just now found the probability that there are at least  $n$  customers in the system. So let us take  $n=10$  here okay. So we have probability that is  $N \geq 10$  is  $\lambda/\mu$  raised to the power 10 and  $\lambda$  is 6,  $\mu$  is 8 raised to the power 10, so this is  $3/4$ , so this is 0.75 raised to the power 10 and this comes out to be 0.0563 okay.

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### Example 4

In a given M/M/1 Queueing system, if  $\lambda = 12$  per hour and  $\mu = 24$  per hour, find the average number of customers in the system.

Ans: 1.

$$L_s = \frac{\lambda}{\mu - \lambda} = \frac{12}{24 - 12} = \frac{12}{12} = 1$$

Now let us consider another problem. In a given MM1 queueing system if  $\lambda=12$  per hour,  $\mu=24$  per hour. So mean service rate is 24 per hour and mean arrival rate is 12 per hour. Find the average number of customers in the system. So let us find the formula for the average number of customers in the system. So average or expected number of customers in the system is given by  $L_s$  and  $L_s = \lambda/\mu - \lambda$ , so we will apply this formula.

So average number of customers in the system is given by  $L_s$  which is  $\lambda/\mu - \lambda$ . So  $L_s$  is  $\lambda/\mu - \lambda$  and  $\lambda$  is 12 okay 12 per hour,  $\mu=24$  per hour okay. So we have  $24-12$ , so this is  $12/12$  and we have 1 here okay because this is service rate per hour.

In 1 hour, the arrival rate is 12 and in 1 hour the service rate is 24 okay

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**Example 5**  
 Suppose that the customers arrive at a Poisson rate of one per every 12 minutes and that the service time is exponential at a rate of one service per 8 minutes. What is  
 (a) The average number of customers in the system.  
 (b) The average time a customer spends in the system.  
 Ans: (a) 2 (b) 24 minutes.

$$\lambda = \frac{1}{12}, \mu = \frac{1}{8}$$

$$L_s = \frac{\lambda}{\mu - \lambda} = \frac{\frac{1}{12}}{\frac{1}{8} - \frac{1}{12}} = \frac{\frac{1}{12}}{\frac{3-2}{24}} = \frac{1}{12} \times 24 = 2$$

$$W_s = \frac{1}{\mu - \lambda} = \frac{1}{\frac{1}{8} - \frac{1}{12}} = \frac{1}{\frac{3-2}{24}} = 24 \text{ minutes}$$

Now suppose that the customers arrive at a Poisson rate of one per every 12 minutes okay and that the service time is exponential at a rate of one service per 8 minutes. So  $\lambda$  here is  $1/12$  and  $\mu=1/8$  okay. The average number of customers in the system just now we have seen the average number of customers in the system is given by  $\lambda/\mu - \lambda$ . So  $L_s = \lambda/\mu - \lambda$  and this is  $1/12/1/8 - 1/12$ .

So we have  $1/12$  and we have here 24 and we get  $3-2$ , so we get  $1/12 \times 24$ , so we get 2. So average number of customers in the system is 2. The average time a customer spends in the system. Let us find the average time a customer spends in the system. So average or expected time of a customer in the system  $W_s$ .  $W_s$  is  $1/\mu - \lambda$  okay. So  $W_s = 1/\mu - \lambda$ . Here the average time of customer so  $W_s = 1/\mu - \lambda$  and this is  $1/1/8 - 1/12$  okay, so  $1/24$  and we have  $3-2$ , so this is 24 minutes. So 24 minutes is the answer.

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### Example 6

A super market has a single cashier. During peak hours, customers arrive at a rate of 20 per hour. The average number of customers that can be processed by the cashier is 24 per hour. Calculate:

- The probability that the cashier is idle.  $-\frac{1}{6}$
- The average number of customers in the queueing system.
- The average time a customer spends in the system.
- The average number of customers in the queue.
- The average time a customer spends in the queue waiting for service.

Ans: (i)  $\frac{1}{6}$  or 0.1667 (ii) 5 (iii) 15 minutes (iv) 4.1667 (v)  $\frac{5}{24}$  hours or 12.5 minutes.

$$W_q = W_s - \frac{1}{\mu} = \frac{1}{4} - \frac{1}{24} = \frac{6-1}{24} = \frac{5}{24} \text{ hours} = \frac{5}{24} \times 60 \text{ minutes} = 12.5 \text{ minutes}$$

Now we go to this last question on queueing theory. A super market has a single cashier. During peak hours, customers arrive at a rate of 20 per hour okay. So arrival rate is  $\lambda=20$  and the average number of customers that can be processed by the cashier is 24 per hour. So  $\mu=24$  okay. Calculate the probability that the cashier is idle. When the cashier is idle, we have to find the  $P_0$  okay, the probability that there is no body in the queue okay in the system.

$P_0=1-\rho$ ,  $P_0$ , we found the formula for  $P_0$ ,  $P_0=1-\rho$  okay and  $P_n$  is the probability that there are  $n$  customers in the system. So if the cashier is idle, there will be nobody in the system, so  $P_0=1-\rho$  we have to find. So  $P_0$  okay the first part is  $P_0=1-\rho$  which is  $1-\lambda/\mu$  okay and this is  $1-20/24$  okay, so  $4/24$  okay, so not like this. Here  $\lambda=20$  we have to consider per hour okay.

So  $\lambda=20$  and  $\mu=24$  okay. So  $P_0=1-\rho$  and therefore it is  $1-\lambda/\mu$  which is  $1-20/24$  okay, so  $4/24$  okay, so  $1/6$  hours okay. So probability is  $1/6$ . So  $P_0$  is this is  $1/6$  and then the average number of customers in the queueing system. Average number of customers in the system we know, average number of customers in the system is given by  $L_s$ .

$L_s = \lambda/\mu - \lambda$  okay. So  $L_s = \lambda/\mu - \lambda$  okay. So  $\lambda=20$  and  $\mu=24$  okay. So we get  $1/4$  okay. So average number of customers in the queueing system is given by  $L_q$  and  $L_q = L_s + \frac{1}{\mu}$  okay, so this is  $5$  okay. So average number of customers in the queueing system is  $5$ , so we have got the answer. Now average time a customer spends in the system okay. Average time a customer spends in the system is given by  $W_s$  which is  $1/\mu - \lambda$  okay.



So  $W_s = 1/\mu - \lambda$  which is  $1/24 - 20$ , so  $1/4$  that is 15 minutes. This  $1/4$  hours which is 15 minutes. Then, we have average time a customer spends in the queue waiting for service okay. So average time a customer spends we have got, now we have to first do this, average time of customers in the queue, average number of customers in the queue. We can get from this formula.

So we have got  $W_s = W_q + 1/\mu$  okay. Average number of customers in the queue, so  $L_q$  we have to find  $L_q$  okay. What is the formula for  $L_q$ ?  $L_q = L_s$ ,  $L_s = L_q + \lambda/\mu$  okay.  $L_s = L_q + \lambda/\mu$  okay.  $L_s$  we have found,  $L_s =$  average number of customers in the system okay, average number of customers in the system we have found,  $L_s = 5$  okay. So we have  $5 = L_q + \lambda/\mu$ .

$\lambda = 20$ ,  $20/24$  okay so we have  $5/6$ , so  $L_q = 5 - 5/6$ , so  $25/6$ ,  $25/6$  means 4.1667 okay, so that is  $L_q$ .  $L_q$  is the average number of customers in the queue. Then, average time a customer spends in the queue waiting for service okay. So that means  $W_q$ .  $W_q$  is the average time a customer spends in the queue waiting for service okay. So we have  $W_q$ .  $W_q = W_s - 1/\mu$ ,  $W_s - 1/\mu$ .

So we found to be  $1/4$  okay  $1/4$  hours and  $1/\mu$ ,  $1/\mu$  is  $1/24$  okay. So this is 24 and we get here  $6-1$ , so  $5/24$  okay, so  $5/24$  hours that is to say 12.5 minutes. So that is the average time a customer spends in the queue waiting for service. So this finishes the discussion on queueing theory. Now let us take an example on reliability theory.

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### Reliability

Consider an equipment or a system which is subjected to failure. Let  $F(t)$  denote the probability of the equipment or system failure within  $t$  units of time. The system reliability  $R(t)$  is the probability that failure will not occur in  $t$  units of time. Thus

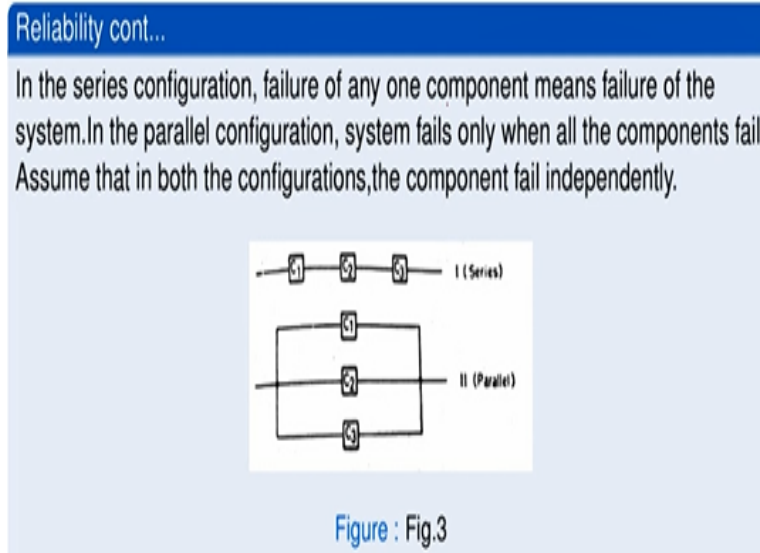
$$R(t) = 1 - F(t)$$

Assume that a number of components are involved in the system. Let us consider two configurations: Series and Parallel configuration.

Consider an equipment or a system which is subjected to failure. Let  $F(t)$  denotes the probability of the equipment or system failure within  $t$  units of time. The system reliability  $R$

$t$  is the probability that failure will not occur in  $t$  units of time and thus we can say that  $R = 1 - F(t)$ . Assume that a number of components are involved in the system. Let us consider 2 configurations, series and parallel configuration.

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In the series configuration, failure of any one component means failure of the system. In the parallel configuration, system fails only when all the components fail. Assume that in both the configurations the component fail independently. Now this is the system in series configuration, this is the system in parallel configuration. So there are 3 components;  $C_1$ ,  $C_2$ ,  $C_3$  which are connected in series and here there are 3 components which are connected in parallel configuration.

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Reliability cont...

Let  $F_1(t)$ ,  $F_2(t)$  and  $F_3(t)$  denotes the failure probabilities of the three components  $C_1$ ,  $C_2$ ,  $C_3$  respectively.  $R_1(t)$ ,  $R_2(t)$  and  $R_3(t)$  are the corresponding reliabilities. Then

$$R_i(t) = 1 - F_i(t), i = 1, 2, 3.$$

The system failure probability is denoted by  $F_s(t)$  and the corresponding system reliability  $R_s(t) = 1 - F_s(t)$

For the series configuration

$$R_s(t) = (1 - F_s(t))$$

$$= (1 - F_1(t))(1 - F_2(t))(1 - F_3(t))$$

So let  $F_1(t)$ ,  $F_2(t)$  and  $F_3(t)$  denote the failure probabilities of the 3 components  $C_1$ ,  $C_2$ ,  $C_3$  respectively.  $R_1(t)$ ,  $R_2(t)$ ,  $R_3(t)$  are the corresponding reliabilities. Then,  $R_i(t) = 1 - F_i(t)$ ,  $i = 1, 2, 3$ . The system failure probability will be denoted by  $F_s(t)$  and the corresponding system

reliability  $R_s(t)$  will be then  $= 1 - F_s(t)$  okay. For the series configuration,  $R_s(t) = 1 - F_s(t) = 1 - F_1(t) * 1 - F_2(t) * 1 - F_3(t)$ . Why?  
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Reliability cont...

This is because, the probability that the system does not fail before  $t$  units of time is equal to the probability that none of the three components  $c_1, c_2, c_3$  fails before  $t$  units of time. For the parallel configuration

$$F_s(t) = (1 - R_1(t))(1 - R_2(t))(1 - R_3(t))$$

This is because in the parallel configuration, the system fails only if all the components  $c_1, c_2, c_3$  fails. Then

$$R_s(t) = 1 - F_s(t).$$

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Because the probability that the system does not fail before  $t$  units of time okay. We are looking at the probability that the system does not fail. We are finding the probability of the reliability okay. That is the system does not fail before  $t$  units of time and we are considering the series configuration. So none of the 3 components must fail and therefore  $R_s(t) = 1 - F_1(t) * 1 - F_2(t) * 1 - F_3(t)$ .

Because  $1 - F_1(t)$  is the probability that the system component  $C_1$  does not fail before  $t$  units of time and similarly  $1 - F_2(t)$  is the probability that the component  $C_2$  does not fail before  $t$  units of time and similarly  $1 - F_3(t)$  is the probability that the component  $C_3$  does not fail before  $t$  units of time. So their product then gives us the probability that the system does not fail before  $t$  units of time in the case of series configuration okay.

So for the parallel configuration what will be the probability?  $F_s(t) = 1 - R_1(t) * 1 - R_2(t) * 1 - R_3(t)$  because in the parallel configuration the system fails only if all the 3 components fail okay. So  $1 - R_1(t)$  is the probability that the component  $C_1$  fails and  $1 - R_2(t)$  is the probability that the component  $C_2$  fails and  $1 - R_3(t)$  is the probability that the component  $C_3$  fails. So when all of them fail, the probability will be given by  $F_s(t)$  and then  $F_s(t)$  gives you the system reliability that is  $R_s(t)$ ,  $R_s(t) = 1 - F_s(t)$  okay.

So in the case of parallel configuration,  $R_s(t)$  will be given by  $1 - F_s(t)$  where  $F_s(t)$  is given by this product okay.

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## Reliability cont...

Suppose

$$F_i(t) = 1 - e^{-\alpha_i t}, \alpha_i > 0, i = 1, 2, 3.$$

Then time T up to failure then follows the exponential distribution. For the series configuration,

$$R_s(t) = e^{-(\alpha_1 + \alpha_2 + \alpha_3)t}$$

and for the parallel configuration

$$R_p(t) = 1 - (1 - e^{-\alpha_1 t})(1 - e^{-\alpha_2 t})(1 - e^{-\alpha_3 t})$$

$$R_p(t) = (1 - (1 - e^{-\alpha_1 t})) (1 - (1 - e^{-\alpha_2 t})) (1 - (1 - e^{-\alpha_3 t}))$$

So suppose let us consider an example suppose  $F_i(t)$  is given by  $1 - e^{-\alpha_i t}$ ,  $\alpha_i > 0$ ,  $i = 1, 2, 3$ . So  $F_i(t)$  as we have seen,  $F_i(t)$  is the probability that the component  $C_i$  fails okay. This  $F_i(t)$  is the probability of failure of the component okay so  $F_i(t)$  is given by  $1 - e^{-\alpha_i t}$  that means  $F_1(t)$  is  $1 - e^{-\alpha_1 t}$ . This is the probability of the failure of component  $C_1$ .

And when you take  $i=2$ , you get the probability of the failure of component  $C_2$  and when you take  $i=3$  you get the probability of the failure of component  $C_3$ . So  $F_i(t)$ 's are given and then the time  $t$  up to the failure then follows the exponential distribution for the series configuration. We have seen in the case of series configuration  $R_s(t) = R_{st}$  is  $1 - F_1(t) * 1 - F_2(t) * 1 - F_3(t)$ , so we have  $R_s(t) = 1 - F_1(t)$ .

$F_1(t)$  is  $1 - e^{-\alpha_1 t}$  okay  $* 1 - e^{-\alpha_2 t} * 1 - e^{-\alpha_3 t}$  okay and you get this expression okay and for the parallel configuration  $R_s(t)$  is = parallel configuration  $R_s(t) = 1 - F_s(t)$  okay,  $F_s(t)$  is given by  $1 - R_1(t) * 1 - R_2(t) * 1 - R_3(t)$  okay. So yeah so we get in the case of parallel configuration  $R_s(t) = 1 - F_s(t)$ ,  $F_s(t) = 1 - R_1(t)$ ,  $1 - R_1(t)$  is the probability that the first component fails that is  $F_1(t)$ ,  $F_1(t) * F_2(t) * F_3(t)$  okay.

So we get  $1 - F_1(t) * F_2(t) * F_3(t)$  okay so that is the system reliability in the case of parallel configuration. So we have discussed both the cases okay. The system reliability in the case of parallel configuration and the system reliability in the case of series configuration okay and we have taken one example where we have taken particular values of the failure probabilities of the 3 components given by  $F_i(t) = 1 - e^{-\alpha_i t}$  where  $\alpha_i > 0$  and  $i = 1, 2, 3$ .

And we have found the values of  $R_s$   $t$  for both the cases the series configuration as well as parallel configuration. So with that I would like to end this lecture. This is our last lecture on advanced engineering mathematics. Thank you very much for your attention.