

Statistical Inference
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Lecture - 61
Interval Estimation – I

So, far this in this course on Statistical Inference we have concentrated on 2 main topics; one was the point estimation where for estimating a parametric function $g(\theta)$ we propose a point estimator that is we call it $t(x)$. And we have discussed methods of obtaining various types of estimators and how to evaluate that is which estimator should be used so what are the criteria like unbiasedness, consistency in minimum variance and bias estimation minimum risk and so on.

Another area that we have covered so far is the problem of testing of hypothesis; that means, we want to make certain assertion about a probability distribution and then we want to test whether our assertion is tenable in the light of the data that we are available towards or not.

A third and equally important area of statistical inference is the Interval Estimation. In the point estimation we propose a value. So, for example, we say suppose we are considering estimation of the average income of a population or average per capita expenditure of a population, then we may consider sample and based on that we calculate a value. And then we say for example, the statement like this per month per capita expenditure on say health is 200 rupees.

However, there are some statisticians and also from practical point of view sometimes it is more useful if in place of a value we give a range of values. So, rather than saying that the percentage of cure by a certain medicine is 75 percent, it is better to give a statement that it is more than 70 percent, which may give more faith in people to use that medicine. If we say it is more than 70 percent rather than giving exact an value if we say it is more than 70 percent.

Similarly, if we want to say per capita expenditure as I mentioned just now for a point estimation problem in place of saying say 200 rupees we may say it is between 175 rupees to 225. Now when we propose an interval so first of all this is called an interval

estimator. Now secondly, there is a probability associated with this thing or you can say a probability statement associated with this because how often this interval will actually include the true value and how often it will not include the true value, that is known as the confidence level.

So, in the problem of interval estimation we propose confidence intervals. I will start with the general definition of a confidence set and then we will move over to confidence intervals when we deal with the real valued parameters.

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Lecture 39

Interval Estimation :

Let $X = (X_1, \dots, X_n)$ be a random sample from a population with distⁿ $P_\theta, \theta \in \mathcal{H} \subset \mathbb{R}^k$. A family of subsets $S(X)$ of \mathcal{H} is said to be a family of confidence sets at confidence level $(1-\alpha)$ if

$$P(\theta \in S(X)) \geq 1-\alpha \quad \forall \theta \in \mathcal{H}.$$

In case of $k=1$

$S(X) = (a(X), b(X))$ is said to be $(1-\alpha)$ -level confidence interval for θ if

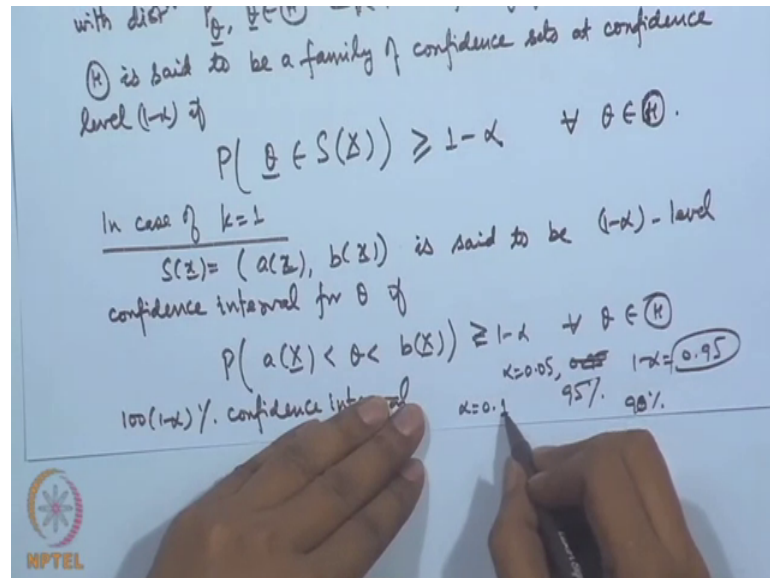
$$P(a(X) < \theta < b(X)) \geq 1-\alpha \quad \forall \theta \in \mathcal{H}$$

So, as before let X_1, X_2, \dots, X_n , so, this is a random sample X_n be a random sample from a population with probability distribution say P_θ ; θ belonging to say a script θ . Now in general θ can be a vector so let me assume in the beginning it is k dimensional and then specific retention will be paid to the case of k is equal to 1.

A family of subsets $S(X)$ of θ is said to be a family of confidence sets at confidence level $1 - \alpha$ if probability of θ belonging to $S(X)$ that is greater than or equal to $1 - \alpha$ for all θ belonging to script θ . So, you can see here that this interval which is of course a random interval $S(X)$ is a subset of θ . So, a random interval it includes θ the probability of this statement is greater than or equal to $1 - \alpha$, then we say $S(X)$ is a family of confidence sets at confidence level $1 - \alpha$.

Now, in the case of if I consider 1-dimensional k that is real parameter then we may consider this set to be an interval. So, we may say $S(x)$ is equal to an interval of the form $a(x)$ to $b(x)$ this is said to be a $1 - \alpha$ level confidence interval for θ if probability of $a(x) < \theta < b(x)$ is greater than or equal to $1 - \alpha$ for all θ .

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It is also written as sometimes $100(1 - \alpha)$ percent confidence interval. So, in place of saying $1 - \alpha$ we also say $100(1 - \alpha)$ percent. So, for example, 0.95 suppose I take α is equal to 0.05 then $1 - \alpha$ is equal to 0.95. So, this is the level and we may say 95 percent confidence interval or if I choose say α is equal to 0.1 then we will say 90 percent confidence interval because here $1 - \alpha$ will become 0.9.

Now when I say this interval $a(x)$ to $b(x)$ it may also happen that one side is taken to be like $b(x)$ could be plus infinity or $a(x)$ could be minus infinity then it becomes a one-sided confidence interval. So, sometimes like I mentioned the statement that the percentage of cure by a certain medicine is for a certain disease is at least 70 percent.

So, here it is one-sided interval, that is we are saying P is greater than or equal to 0.7 or we may say that then rate of failure is at most 1 percent or at most 10 percent for a certain manufacturing process. So, in that case we are putting an upper bound. So, let me formally will give this one-sided confidence intervals also.

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$(a(x), \infty)$ is left confidence interval of
 $P(\theta > a(x)) \geq 1 - \alpha \quad \forall \theta$

$(-\infty, b(x))$ is right confidence interval of
 $P(\theta < b(x)) \geq 1 - \alpha \quad \forall \theta$

$\inf_{\theta \in \Theta} P(\theta \in S(x)) \rightarrow \text{confidence level}$

A family of $(1 - \alpha)$ level confidence sets $\{S(x)\}$ is said to be uniformly most accurate (UMA) family of confidence sets at level $(1 - \alpha)$ if

$$P_{\theta}(\theta' \in S(x)) \leq P_{\theta'}(\theta' \in S'(x))$$
 for all $\theta \neq \theta'$ & any $(1 - \alpha)$ -level family of confidence sets $S'(x)$.

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So, a x to infinity is you can say left confidence interval if probability of theta greater than $a(x)$ is greater than or equal to $1 - \alpha$. Similarly minus infinity to $b(x)$ is right confidence interval, if probability of theta less than $b(x)$ is greater than or equal to $1 - \alpha$. One point which I will like to repeat here when I say this probability statement this is statement is a probability statement involving the random variable which is in the form of $S(x)$ here because X_1, X_2, \dots, X_n are random variable. So, this $S(x)$ is actually a random variable.

So, one should not understand that we are considering some probability statement for parameter theta; theta is actually fixed. So, in all of these statements the probability statement is valid because of the distribution of X_1, X_2, \dots, X_n is involved here.

We gave a , now definitely when we are having an interval and certainly we would like to have some sort of understanding for example, if I say 95 percent confidence interval. Now certainly if there is an error interval which will give me 99 percent confidence then I will consider it to be better if the length is likely to be the same. On the other hand if I fix a confidence level say I say 90 percent or 95 percent at that level I would like to choose that interval which has the shortest length. Otherwise you can always say like minus infinity to infinity it includes all the parameter values if probability 1.

So, this statement will be a weak statement or you can say it is not going to be useful in the inference point of view. So, we need to make a statement which will be in the valid

as well as which will be more precise or accurate, that brings us to the problem of finding out shortest length confidence intervals or if we fix the width then the maximum confidence level ok. So, we if we put say infimum of say probability ax or say theta belonging to S X.

If I consider this is called the confidence level a family of $1 - \alpha$ level confidence sets S_X is said to be uniformly most accurate. So, we call it UMA; UMA family of confidence sets at level $1 - \alpha$ if probability of theta prime belongs to S_X is less than or equal to $S_{X'}$. For all theta not equal to theta prime and any $1 - \alpha$ level family of confidence sets $S_{X'}$. So, S_X will be said to be uniformly most accurate.

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Examples. 1. Let $X_1, \dots, X_n \sim N(\mu, 1)$.
 We want confidence interval for μ .
 \bar{X} is minimal sufficient statistic. $\bar{X} \sim N(\mu, 1/n)$
 We consider confidence interval of the type $(\bar{X} - c_1, \bar{X} + c_2)$ so that
 $P(\bar{X} - c_1 < \mu < \bar{X} + c_2) \geq 1 - \alpha$
 $\Leftrightarrow P(-c_2 < \bar{X} - \mu < c_1) \geq 1 - \alpha$
 $\Leftrightarrow P(-\sqrt{n}c_2 < \sqrt{n}(\bar{X} - \mu) < \sqrt{n}c_1) \geq 1 - \alpha$
 $\Leftrightarrow P(-\sqrt{n}c_2 < Z < \sqrt{n}c_1) \geq 1 - \alpha$
 We may take equality at $(1 - \alpha)$ as we are

Let me give some example of a confidence interval. Let us consider a normal population let X_1, X_2, \dots, X_n be a random sample from a normal distribution with mean μ and I am considering for convenience variance to be unity. That means, variance is known so, I have converted to unity, we want the confidence interval for say μ . We want confidence interval for μ . Now I am looking at from the practical point of view in this problem \bar{X} is a minimal sufficient statistic; is not it? \bar{X} is a minimal sufficient it is also complete of course.

So, we will actually we have seen when we were doing the estimation. So, \bar{X} was actually given μ mle mini max estimator etcetera. Also when we looked at the test

for μ in this particular problem the test was based on \bar{X} , bar the uniformly most powerful test ump unbiased test or the likelihood ratio test they were all based on \bar{x} bar in fact, the invariant test also. So, when we were setting up a confidence interval it is natural that we will restrict attention to the minimal sufficient a statistic \bar{X} bar. And also what is the distribution of \bar{X} bar? \bar{X} bar follows normal μ 1 by n .

Which is actually a symmetric distribution. So, naturally it suggests that we take an interval around μ which is based on \bar{X} bar. So, we consider confidence interval of the type say \bar{X} bar minus sum C_1 to \bar{x} bar plus sum C_2 and probability of this interval including μ should be at least $1 - \alpha$. Now this statement is equivalent to probability of a \bar{X} bar minus μ is less than C_1 greater than minus C_2 that is greater than or equal to $1 - \alpha$. Now we may also put from here $\sqrt{n} \bar{X}$ bar minus μ that follows normal 0 1.

So, if we consider the probability density function of standard normal distribution then I am choosing two points such that the probability in between is equal to $1 - \alpha$. So, we can write it as minus $\sqrt{n} C_2$ less than $\sqrt{n} \bar{x}$ bar minus μ less than $\sqrt{n} C_1$ greater than or equal to $1 - \alpha$. So, probability of Z lying between these two values minus $\sqrt{n} C_1$ to $\sqrt{n} C_1$ and you can see actually we are dealing with the continuous distribution.

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\bar{X} is minimal sufficient statistic

We consider confidence interval of the type $(\bar{X} - C_1, \bar{X} + C_2)$ so that

$$P(\bar{X} - C_1 < \mu < \bar{X} + C_2) \geq 1 - \alpha$$

$$P(-C_2 < \bar{X} - \mu < C_1) \geq 1 - \alpha$$

$$\Leftrightarrow P(-\sqrt{n} C_2 < \sqrt{n}(\bar{X} - \mu) < \sqrt{n} C_1) \geq 1 - \alpha$$

$$\Leftrightarrow P(-\sqrt{n} C_2 < Z < \sqrt{n} C_1) \geq 1 - \alpha$$

We may take equality at $(1 - \alpha)$ as we are considering continuous distⁿ.

$\sqrt{n}(\bar{X} - \mu) \sim N(0, 1)$

So, we can achieve actually equality here, we may take equality at 1 minus alpha as we are considering continuous distribution, another point you can see when I am putting two points here. So, let me call d 1 and d 2, then there can be many points where the probability can be equal to 1 minus alpha like if I shift on the side or if I shift on the side I can consider so many intervals here.

But once again let us bring down the problem to the finding out shortest length. Now because of symmetry of the standard normal distribution we can consider a symmetric interval and if I choose the symmetric interval it will be the shortest because the probability in the middle is the highest for a normal distribution.

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The image shows a handwritten derivation on a blue background. At the top right, there is a small logo for '© CET I.I.T. KGP'. The text reads: 'We may choose $c_1 = c_2 = c$ ', followed by ' $\Rightarrow c\sqrt{n} = z_{\alpha/2}$ '. Then, 'So $c_1 = \frac{1}{\sqrt{n}} z_{\alpha/2} = c_2$ '. Next, 'So $(\bar{X} - \frac{1}{\sqrt{n}} z_{\alpha/2}, \bar{X} + \frac{1}{\sqrt{n}} z_{\alpha/2})$ ' is identified as a '(1- α)-level confidence interval for μ '. An example is given: 'Let $n = 16, \alpha = 0.05, z_{0.025} = 1.96$ '. The interval is calculated as $(\bar{X} - \frac{1.96}{4}, \bar{X} + \frac{1.96}{4})$, which is simplified to $(\bar{X} - 0.49, \bar{X} + 0.49)$. To the right of the calculations is a hand-drawn normal distribution curve with a central peak labeled '(1- α)'. The area under the curve between two points is shaded, with the tails labeled ' $\alpha/2$ '. The points on the x-axis are labeled ' $-\frac{z_{\alpha/2}}{\sqrt{n}}$ ' and ' $\frac{z_{\alpha/2}}{\sqrt{n}}$ '. At the bottom right, there is a note: 'Suppose X_1, \dots, X_n are observations on measuring time'. The NPTEL logo is visible in the bottom left corner of the slide.

So, we may choose say C 1 is equal to C 2 is equal to say C then this implies C root n that should be equal to z alpha by 2 that is a point on the this is z alpha by 2; that means, this probability is alpha by 2. So, minus z alpha by 2 is the probability where here at this probability is alpha by 2. So, in between probability is 1 minus alpha. So, we are getting here that C 1 is equal to 1 by root n z alpha by 2 that is equal to C 2. So, X bar minus 1 by root n z alpha by 2 to X bar plus 1 by root n z alpha by 2 this is 1 minus alpha level confidence interval for mu.

Let me just take a practical example, suppose I take n is equal to 16 say and alpha is equal to say 0.05. So, then I have to see the value of z 0.025 from the tables of standard

normal distribution this value is 1.96. So, \bar{X} minus 1.96 by 4 to \bar{X} plus 1.96 by 4 that is \bar{X} minus 0.49 to \bar{X} plus 0.49.

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We may choose $c_1 = c_2 = c$
 $\Rightarrow c\sqrt{n} = z_{\alpha/2}$
 So $c_1 = \frac{1}{\sqrt{n}} z_{\alpha/2} = c_2$
 So $(\bar{X} - \frac{1}{\sqrt{n}} z_{\alpha/2}, \bar{X} + \frac{1}{\sqrt{n}} z_{\alpha/2})$
 is $(1-\alpha)$ -level confidence interval for μ .
 Let $n=16$, $\alpha=0.05$, $z_{0.025} = 1.96$
 $(\bar{X} - \frac{1.96}{4}, \bar{X} + \frac{1.96}{4})$
 $\equiv (\bar{X} - 0.49, \bar{X} + 0.49)$

Suppose X_1, \dots, X_n are observations on measuring heights in (cm) of girls in a class (V). Suppose they

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So, if I in a practical problem suppose we are considering suppose X_1, X_2, \dots, X_n are observations on measuring say heights in say centimeters of girls in a class suppose class is 5 ok.

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are normally dist. with mean μ & variance 1.
 A r. s. of size 16 gives $\bar{X} = 130$ cm.
 Then 95% confidence interval is $(130 - 0.49, 130 + 0.49)$
 $\equiv (129.51, 130.49)$

2. Let X_1, \dots, X_n be a random sample from $N(\mu, \sigma^2)$ where both μ & σ^2 are unknown. We want the confidence interval for μ . We consider interval $(\bar{X} - c_1, \bar{X} + c_2)$
 $P(\bar{X} - c_1 < \mu < \bar{X} + c_2)$

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And the heights in centimeters are suppose they are normally distributed with mean μ and variance 1 ok. A random sample of size 16 gives say \bar{X} is equal to the observed

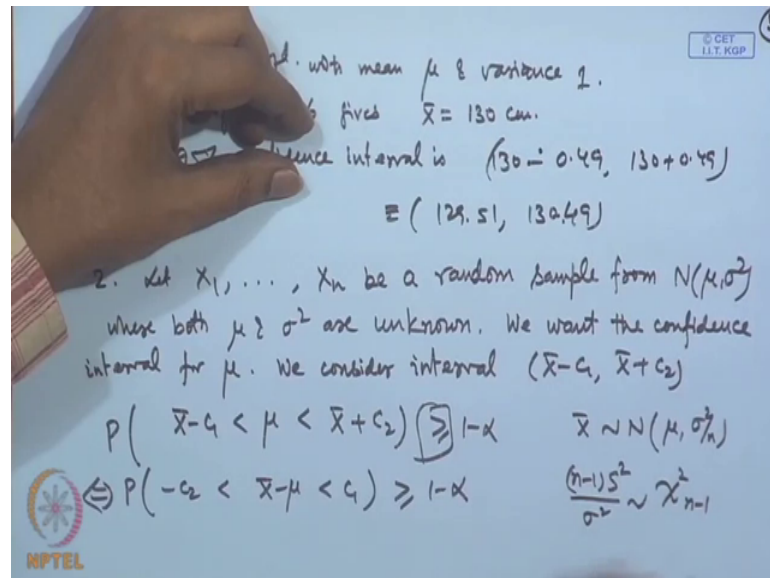
value of the \bar{X} suppose it turns out to be. So, in class 5 we may assume something like say 48.

So, something like say 130 centimeters say then 95 percent confidence interval is 130 point 130 minus 0.49 to 130 plus 0.49, that is 129.51 to 130.49. That means, we are saying with 95 percent confidence that the average height of the girls in a class 5 will be between 129.51 centimeter to 130.49 centimeters.

One interpret it in a frequency sense that if we conduct or we collect the sample 100 times then 95 percent of the times the sample will include the true parametric value and 5 percent of the time it is not like it may not include the 2 parameter value. Let me consider some more examples here in this previous case I have considered variance to be known. So, I took it to be 1, but suppose the variance is unknown, if the variance is unknown then the distribution of \bar{X} will involve σ^2 and in that case we need to do little bit of modification.

So, let me take another example let $X_1 X_2 \dots X_n$ be a random sample from normal μ σ^2 where both μ and σ^2 are unknown and as before we want the confidence interval for μ . Now as I argued in the previous case \bar{X} is some sort of you can say good estimator for this problem for μ . So, once again we will write the same thing that we consider interval of the type interval $\bar{X} - C_1$ to $\bar{X} + C_2$ that is probability of $\bar{X} - C_1 < \mu < \bar{X} + C_2$ should be greater than or equal to $1 - \alpha$.

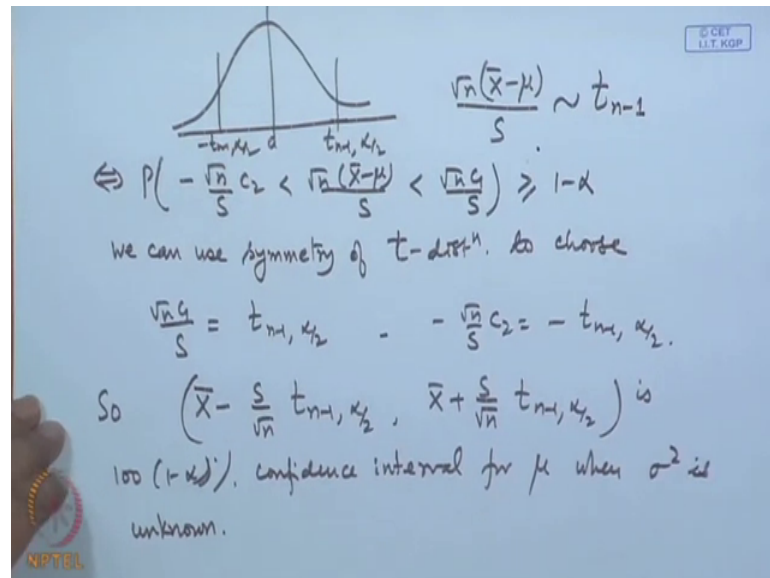
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And of course and since we will be dealing with the continuous distributions I will actually find out the solution corresponding to equal to 1 minus alpha. So, probability of now this is statement you can modify minus C 2 less than X bar minus mu less than C 1 is greater than or equal to 1 minus alpha.

Now in this case if I am again considering X bar the distribution of X bar that is normal mu sigma a square by n. So, if, I write in the state forward fashion this will depend upon; the interval will depend upon sigma here. So, we make use of the t statistic here if you remember n minus 1 S square by sigma a square follows xi square distribution on n minus 1 degrees of freedom.

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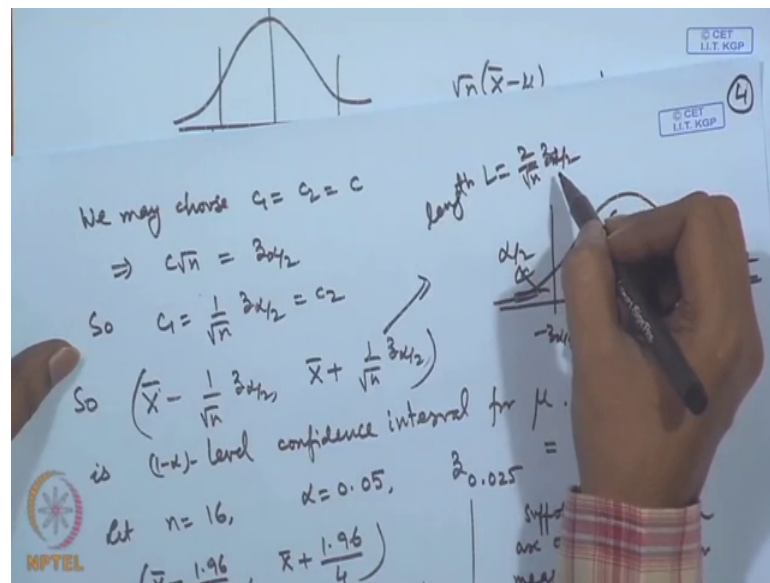


And as a consequence we had seen that \bar{X} minus μ root n by S this follows t distribution on n minus 1 degrees of freedom. So, from this statement then we write probability of minus root n by S c_2 less than root n \bar{X} minus μ by S less than root n c_1 by S that is greater than or equal to 1 minus α . Since this is again a symmetric distribution this t distribution as you know it is a symmetric distribution around 0 . So, we may choose the point as before as $t_{n-1, \alpha/2}$ and minus t and minus 1 alpha by 2 .

So, we can use symmetry of t distribution to choose root n c_1 by s is equal to $t_{n-1, \alpha/2}$ and minus root n by S c_2 is equal to minus $t_{n-1, \alpha/2}$.

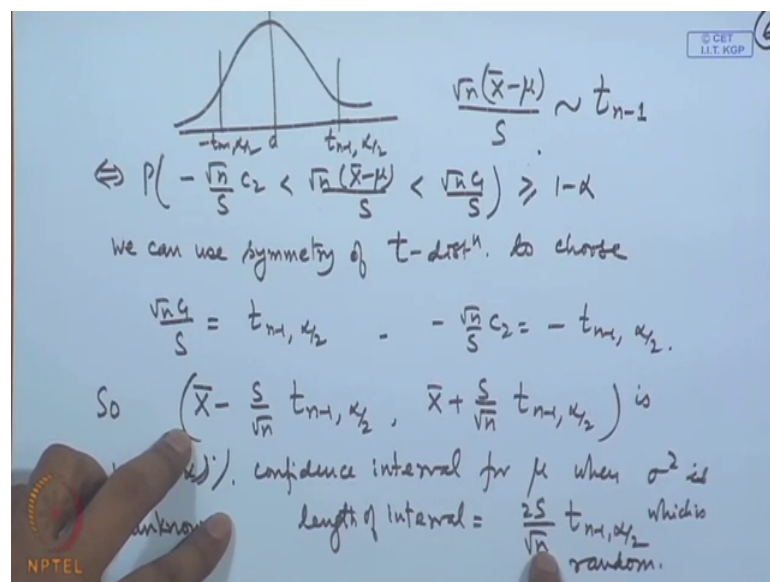
So, from here the confidence interval comes out to be \bar{X} minus S by root n $t_{n-1, \alpha/2}$ to \bar{X} plus s y root n $t_{n-1, \alpha/2}$. So, this is 100 1 minus α percent confidence interval for μ when σ^2 is unknown. When σ^2 was known then here σ would have come here as I solve the problem with σ is equal to 1 . Also let us note down some feature here.

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For example what is the length of the interval, in this case the length of the interval that is equal to 2 by root in z alpha by 2 which is actually free from the parameters and it is free from the random variable.

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In this case if you see the length is length of interval that is equal to 2 S by root n t minus 1 alpha by 2 which is random, but of course, you can see that as n becomes large this will become a small the length of the interval will become a small actually it will go to 0. Another point you can see that as n becomes largest value will converge to z value,

that is $z_{\alpha/2}$ and S will be some number which will actually convert to σ by using the strong law of large numbers or consistency of s for σ .

So, this will converge to; so if n is large actually this is approximately the same as the one for the known σ case. And in both the cases that I have given here the length of the intervals as you can see as n becomes large the length of the interval becomes small. So, the interval becomes sharper for larger sample sizes. So, which again suggests that the more data or you can say more observations are there then the better inference will be drawn because the interval will become more crisp.