

Statistical Inference
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Lecture - 36
Application of NP – Lemma – II

Let me take another example, in which the range of the variable may be dependent upon the range of the parameter and let us see in that case how the Neyman Pearson lemma works.

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4. Let $X_1, \dots, X_n \sim U(0, \theta), \theta > 0$

$H_0: \theta = 1 \in (\theta_0)$
 $H_1: \theta = 2 \in (\theta_1)$

$f(x, \theta) = \begin{cases} \frac{1}{\theta^n}, & 0 \leq x_{(1)} \leq x_{(n)} \leq \theta \\ 0, & \text{ew} \end{cases}$

$f_0(x) = \begin{cases} 1, & 0 \leq x_{(1)} \leq x_{(n)} \leq 1 \\ 0, & \text{ew} \end{cases}, \quad f_1(x) = \begin{cases} \frac{1}{2^n}, & 0 \leq x_{(1)} \leq x_{(n)} < 2 \\ 0, & \text{ew} \end{cases}$

Case		f_0	f_1
1	$x_{(1)} \leq 0$	0	0
2	$0 < x_{(1)} \leq x_{(n)} \leq 1$	1	$\frac{1}{2^n}$
3	$x_{(1)} > 0, 1 \leq x_{(n)} < 2$	0	$\frac{1}{2^n}$
4	$x_{(n)} > 2$	0	0

Let us consider say X_1, X_2, \dots, X_n from uniform 0 to θ distribution and we consider a hypothesis testing problem say θ is equal to 1 against say θ is equal to 2 . We may also right here say θ is equal to 1 and here I may right $\theta = 1$ and then I may consider the case θ is less than 1 or θ is not greater than 1 . So, for convenience we have considered this is special case; $\theta = 1$ and $\theta = 2$.

So, let us consider the most powerful test here. So, the joint distribution is 1 by θ to the power n , and here the range of the variables is from 0 to θ . So, we write it in this particular form. So, when we write for f_0 and f_1 for f_0 this is simply 1 . So, this is simply; see we may if we write here open interval then we need not put equality

here we may put it like this otherwise may put a quality. The probability of those points will be 0 so it does not make any difference.

Similarly if we consider f_1 , then under f_1 theta is equal to 2. So, it will become 1 by 2 to the power n 0 less than x_1 less than or equal to x_n less than 2 it is equal to 0. I was mentioning here that the range of the densities where the two densities are positive is not the same. Here you can see this density is positive for 0 to 1 and this density is positive for 0 to 2. So, let us look at the various cases of f_1 and f_0 ok. So, we will make it in the form of a table, let us consider say case 1 2 3 4 like that we will write 1 2 3 4.

So, I will write all the cases which may be trivial and nontrivial cases. If we observe x_1 to be less than 0; obviously, this is not possible. So, both the densities f_1 and f_0 they are 0 here. If we considered the case 0 is less than, so x_1 and x_n is less than or equal to 1, in this case the first density is 1 and second density is 1 by 2 to the power n. Then we may have say x_1 of course, may be greater than 0, but x_n is say beyond it is beyond 1 in that case what will happen? That this first density becomes 0; however, the second density remains 1 by 2 to the power n.

And then we may have the extreme case that is x_n is greater than 2, then this is 0 and this is 0. So, broadly speaking we have to consider the f_1 greater than $k f_0$ from the Neyman Pearson lemma under these 4 cases.

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Handwritten notes on a blue background. At the top, there are some scribbles and the text $f_1(x) = \begin{cases} \frac{1}{2^n}, & 0 < x_1 \leq x_n < 2 \\ 0, & \text{else} \end{cases}$. Below this is a table with two columns, f_0 and f_1 , and four rows corresponding to different ranges of x_1 and x_n . To the right of the table, there is a handwritten note: "By NP lemma the MP test is rejecting H_0 if $f_1(x) > k f_0(x)$ ".

	f_0	f_1
$x_1 \leq 0$	0	0
$0 < x_1 \leq 1$	1	$\frac{1}{2^n}$
$0 < x_1 \leq 2$	0	$\frac{1}{2^n}$
	0	0

By NP lemma the MP test is rejecting H_0 if $f_1(x) > k f_0(x)$

By NP lemma, the MP test is rejecting H_0 if $f_1(x)$ is greater than k times $f_0(x)$. So, here the values of k we have to choose; so, let us write here.

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For each case $f_1 > kf_0$ & the values of k are listed below

Case	$f_1(x) > kf_0(x)$	$f_1(x) = kf_0(x)$	$f_1(x) < kf_0(x)$
I	Never	Always	Never
II	$k < \frac{1}{2^n}$	$k = \frac{1}{2^n}$	$k > \frac{1}{2^n}$
III	Always	Never	Never
IV	Never	Always	Never

Case I: We may take $Q_1(x)$ as any value
 Case II: Set $Q_1(x) = 1$ if $k < \frac{1}{2^n}$, $Q_1(x) = 0$ if $k > \frac{1}{2^n}$ and any value of $Q_1(x)$ if $k = \frac{1}{2^n}$.
 Case III: $Q_1(x) = 1 \forall k$. (Always reject H_0)

For each case f_1 greater than k f_0 and the values of k are listed below; so let us consider each case. In the case 1 if we consider $f_1(x)$ greater than k times $f_0(x)$ now both the values are 0. So, this is never possible whatever be the value of k it is never possible; $f_1(x)$ is equal to k times $f_0(x)$ that is always true and $f_1(x)$ less than k times $f_0(x)$ is never possible. Let us consider second case, in the second case if we look at f_1 that is 1 by 2 to the power n greater than k times f_0 .

So, this condition is true if k is less than 1 by 2 to the power n , this is true if k is equal to 1 by 2 to the power n , this is true if k is greater than 1 by 2 to the power n , let us look at the third case. In the third case f_0 is 0. So, f_1 greater than k times f_0 is always true, and therefore this equality is never possible. Let us consider the case 4 once again both of them are 0. So, inequality is never possible whereas, the equality is always true.

So, now based on this we should tell when to reject H_0 and when to accept H_0 ; that means, dependent upon these 4 cases and the choices of k , we should give what is the test function. And at the same time we should also tell that whether the probability of type 1 error is equal to α can be achieved for a given value of α .

So, what we consider in case 1? Since f_1 is equal to $k f_0$ is always possible always true therefore, whatever be the value of ϕ_1 , it does not make any difference.

So, we may take ϕ_1 as any value. In case 2 if k is less than $1/2$ to the power n , in this case f_1 is greater than $k f_0$ that means this is the corresponding case 2 rejecting H_0 . So, if case is less than $1/2$ to the power n we will say, reject H_0 and in this case we will say except H_0 ; that means, ϕ_1 is equal to 0. However, when k is equal to $1/2$ to the power n can we may again say we may accept or reject depends upon we can assign something of course, the probability of this case will be 0.

So, we can say $\phi_1(x)$ that is a test function is equal to 1, if case less than $1/2$ to the power n it is equal to 0, if k is greater than $1/2$ to the power n and any value of ϕ_1 if k is equal to $1/2$ to the power n . If you look at case 3, in the case 3 this condition is always true. So, we always reject that is $\phi_1(x)$ is equal to 1 whatever v k that is always a reject H_0 . Note here that these are also you restrict because what is happening? If we are getting the third case that is x is between 1 and 2 that mean naturally the observations are from the density uniform 0 to 2, otherwise observation cannot be greater than 1 and therefore, we should definitely reject H_0 and accept H_1 .

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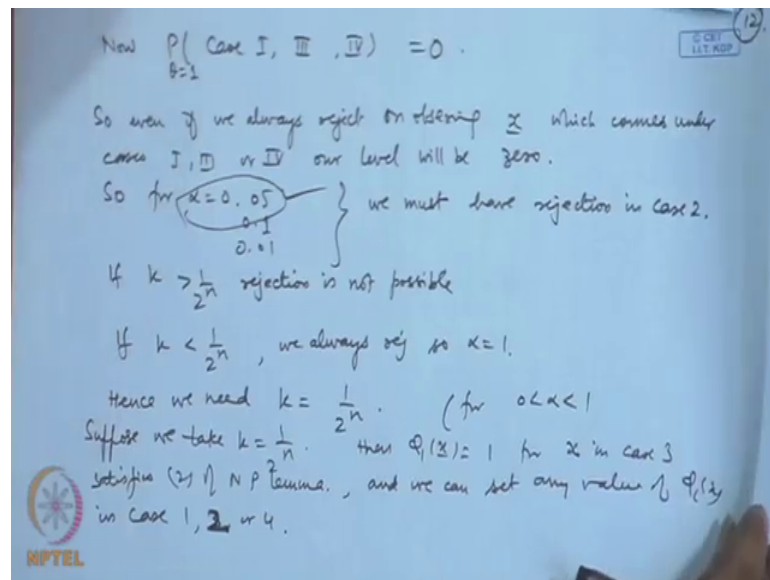
the values of k are listed below

Case	$f_1(x) > k f_0(x)$	$f_1(x) = k f_0(x)$	$f_1(x) < k f_0(x)$
I	Never	Always	Never
II	$k < \frac{1}{2^n}$	$k = \frac{1}{2^n}$	$k > \frac{1}{2^n}$
III	Always	Never	Never
IV	Never	Always	Never

Case I: We may take $\phi_1(x)$ as any value
 Case II: Set $\phi_1(x) = 1$ if $k < \frac{1}{2^n}$, $\phi_1(x) = 0$ if $k > \frac{1}{2^n}$ and any value of $\phi_1(x)$ if $k = \frac{1}{2^n}$.
 Case III: $\phi_1(x) = 1 \forall k$. (Always reject H_0)
 Case IV: Any value of ϕ_1

And in the case 4 once again we may put any value; any value of ϕ_1 .

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Now, probability of case 1 3 and 4 under the null hypothesis this is 0. So, even if we always reject on observing x which comes under cases I, III or IV our level will be 0. So, for alpha is equal to 0.05 at sector some given value of alpha or say 0.01 or 0.1 etcetera then what we should do? We should make the probability of rejection in case 2 to be possible.

So, we must have rejection in case 2. Now again if I take k to be greater than 1 by 2 to the power n rejection is not possible. So, the only rejection is possible for k less than 1 by 2 to the power n here rejection is becoming always true; so alpha will become 1 which is not acceptable. So, therefore, we should have this k equal to 1 by 2 to the power n as a possible value. If k is greater than 1 by 2 to the power n rejection is not possible. If k is less than 1 by 2 to the power n , we always reject.

So, alpha is equal to 1 therefore, we should have k equal to 1 by 2 to the power n for 0 less than alpha less than 1 . So, suppose we take k equal to 1 by 2 to the power n , then $\phi_1(x)$ is equal to 1 for x in case 3 satisfies 2 of NP lemma and we can set any value of $\phi_1(x)$ in case 1 2 1 3 or 4 sorry 1 2 or 4.

Let me give concrete case here when I base our decision on x_n alone. You note here that when I am considering x_1, x_2, \dots, x_n a random sample from you know from 0 to θ , the sufficient statistics actually x_n and x_n is actually playing the role here as you have already noticed here. So, let me explain that part in detail.

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If we base our decision on $Y = X_n$.

In this case, analyzing as before, we write

$$E_{\theta} \phi_1(Y) = \gamma P(\text{Case 2}) + \sum_{\theta=1} P(\text{Case 3, 4 or 5})$$

$$= \gamma \cdot 1 = \gamma = \alpha$$

So a MP test for level α is

$$\phi_1(X) = \begin{cases} \alpha & \text{if } Y \leq 1 \\ \text{otherwise} & \end{cases}$$

$E_{\theta} \phi_1(Y) = \alpha P(Y \leq 1) + P(\text{Case 3, 4 or 5})$

$$= \alpha$$

$$\alpha \left(\frac{1}{\theta}\right)^n + \left(1 - \left(\frac{1}{\theta}\right)^n\right) = \alpha$$

If we base our decision on say Y is equal to X_n . So, in this case what is happening that let me write here.

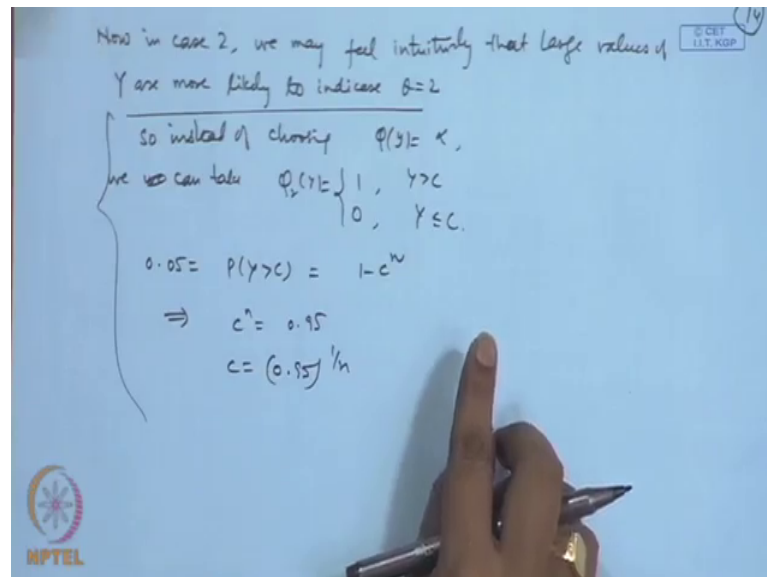
In this case analyzing as before we write expectation of $\phi_1(y)$ for θ is equal to 1 as some γ times probability of case 2, under θ is equal to 1 plus probability of θ is equal to 1 under case 3 or 4 now these are all 0. So, that is equal to γ into 1 that is equal to γ . So, we should choose γ is equal to α . So, a most powerful test for level α is $\phi_1(x)$ or $\phi_1(y)$ is equal to point 0 that is α if Y is less than or equal to 1 it is equal to α otherwise.

That means what we are saying? Reject all the time except for the case when y is less than or equal to 1. If y is less than or equal to 1 then you are rejecting with probability 0.05 and otherwise you are accepting we may also consider the power function here.

So, for example, power function here that is equal to point α into probability of say Y less than or equal to 1 plus probability. So, here I am taking θ is equal to 1 into. So, those cases will not occur this will have probability 0. So, this is actually equal to α for this is $\theta < 1$ and if I take $\theta > 1$, then it will become equal to α into 1 by θ to the power n plus 1 minus 1 by θ to the power n . So, that is equal to well you can simplify this value here.

So, what we are able to do is that, we are able to provide an exact test here for testing parameters in the uniform distribution. We may also consider in a slightly different fashion let me just explain it here.

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Case 2 we may feel intuitively that large values of Y are more likely to indicate $\theta = 2$.

So, instead of choosing $\phi(y) = \alpha$, we can take say $\phi_c(y) = 1$ if y is greater than c it is equal to 0 if y is less than or equal to c and if you consider the probability of this. So, we are getting here say 0.05 is equal to probability of y greater than c that is equal to $1 - c^n$ this means we can take $c^n = 0.95$ or $c = (0.95)^{1/n}$.

So, this is an alternative solution here of course, this is based on the heuristic consideration that large values of y are more likely to indicate $\theta = 2$. So, this part is not coming from NP lemma, in the NP lemma if we write exactly will take that part and the test function is of this nature that $\phi(y) = \alpha$ if Y is less than or equal to c and it is 1 otherwise. So, these are the two forms that I am considering here.

Friends today we have considered in detail various applications of the Neyman Pearson fundamental lemma, how it gives exact tests for testing simple hypothesis works as a simple hypothesis. The important point that you should note here is that we need the

distribution of the criteria. That means our criteria is based on certain function of the random variable which we call test a statistic, we should be able to say something about the distribution of that under the null hypothesis then only the constant k can be determined. If we are unable to determine that then we will not be able to provide the exact form of the test function.

So, in the next lecturer as I mention we will consider extension of the Neyman Pearson lemma to consider the composite hypothesis also. So, in particular we will consider the one sided composite hypothesis testing problems so, that I will be taking up in the following lecture.