

**Probability and Statistics**  
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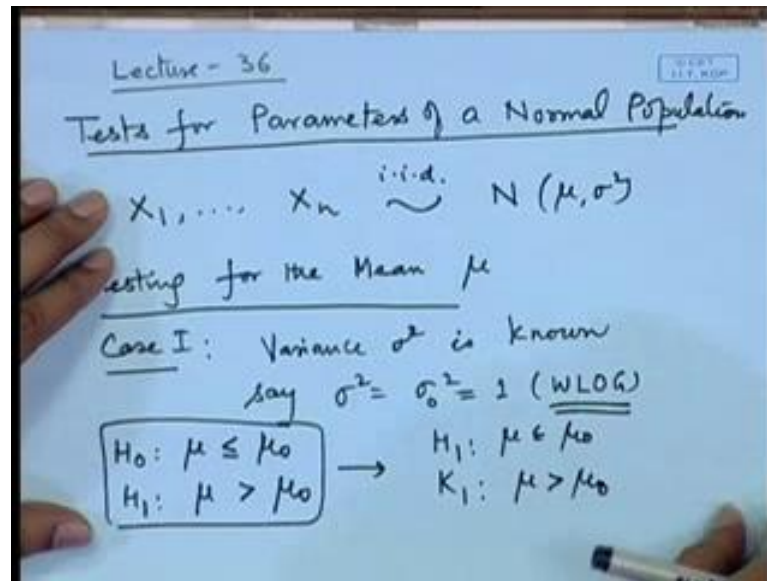
**Lecture - 70**  
**Testing for Normal Mean**

Now, more than the situations will correspond to the test of simple versus composite hypothesis are composite versus composite hypothesis. So, to handle such situations the generalization of Neyman-Pearson Lemma was carried out. And by using concepts of the distribution satisfying a monotone likelihood ratio property, there was uniformly most powerful test for certain hypothesis.

Certain composite versus composite hypotheses are certain simple versus composite hypothesis; and even then there were situations when we have Newson's parameters, and we do not have the uniformly most powerful test. In certain situations the concept of unbiasedness in the test was introduced. We had the concepts of similar test. And so uniformly most powerful and biased test are uniformly most powerful invariant test have been studied. Another approach for testing is through likelihood ratios and various test have been discovered for these situations also.

The theoretical derivations of the test for all these situations will be part of another course called statistical inference. In this particular course we will discuss only the applications of the test for parameters of normal distributions the test for proportions etcetera. So, let me begin with the test for the parameters of the normal distribution where the testing problems may be composite.

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So, let us take the case of say testing for the mean. So, consider the situations say  $X_1, X_2, \dots, X_n$  following normal  $\mu, \sigma^2$ . So, we have a random sample from a normal distribution with parameters  $\mu$  and  $\sigma^2$ . We may be testing for the mean  $\mu$ .

So, there may be two cases: as in the case of confidence intervals we have the case when  $\sigma^2$  that is variance  $\sigma^2$  is known or unknown. So, if the variance  $\sigma^2$  is say known- say  $\sigma^2 = \sigma_0^2$  or we may take without loss of generality is equal to 1, without loss of generality you may take it to be 1 also. In this case, now let us go back to the application of the NP lemma what we have observed here that the test function is based on the value of  $\bar{X}$ , we have considered the testing for normal  $\mu$  when  $\mu$  is equal to say  $\mu_0$  and against  $\mu$  is equal to  $\mu_1$ .

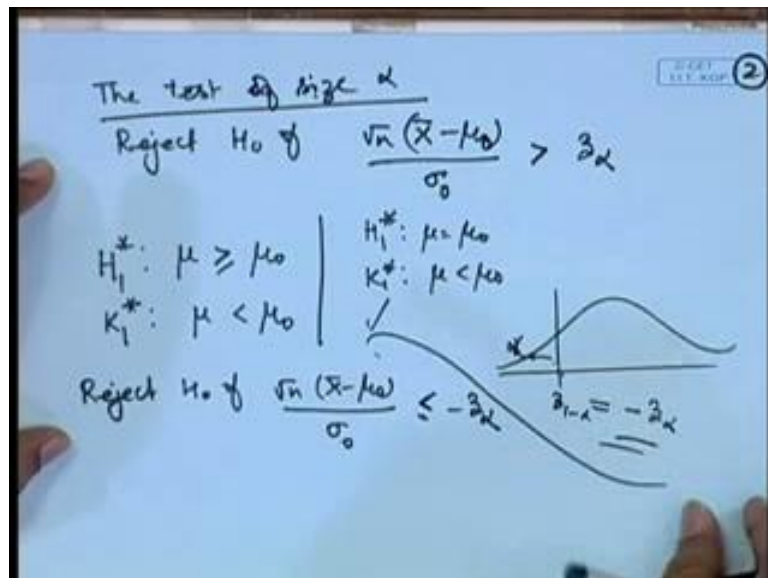
So, we also observed that if  $\mu < \mu_0$  then you reject for larger values of  $\bar{X}$  when  $\mu > \mu_0$  then you reject for by smaller values of  $\bar{X}$ . So, that gives rise to the general situations such as say  $H_0: \mu \leq \mu_0$  against  $H_1: \mu > \mu_0$ . So, the situation may be like this that we are having certain efficiency level, certain measurement regarding a previous procedure, now a new procedure is adopted and we want to see whether the efficiency or the measurement or the effectiveness etcetera has decreased or increased corresponding to the previous one. Or you may be a control kind of variable so you want to test

whether the value or you can say the mean is better than the control or worse than the control.

So, accordingly we may said the null and alternative hypothesis. So, we may interchange the roles also, but let me take up this case. In particular I will be considering four types of hypothesis. So, for convenience let me give some names to this because I will be describing them in detail. So, I will give a new notation to this I will call H 1 as mu is less than or equal to mu naught and K 1 is mu is greater than mu naught, where null hypothesis is denoted by h and the alternative hypothesis is denoted by k.

Now, we have already seen that the test function is dependent upon the value of X bar. So, the test will be; so there will be various situations in this particular case we have uniformly most powerful test. Since we have not introduced the concept of uniformly most powerful or uniformly most powerful unbiased I will not be utilizing this terminology here instead I will just mentioning the kind of the test that you are having.

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So, the test is rejects H naught if root n X bar minus mu naught; so if you have sigma naught then you may put sigma naught here if sigma naught is 1 then you need not put it here. So, this is greater than z alpha. So, the test of size alpha is rejects H naught if this is so. Once again whether you will include equality here or not does not make any difference, because the size does not change, because X bar is a continuous random variable. In fact, this random variable under H naught is having a standard normal

distribution and that is why the probability point of the distribution has turned to be as a  $z_{\alpha}$  point.

Now, from here itself we can look at the other situation also. For example, here if I have  $\mu \geq \mu_0$  and here I will put  $\mu < \mu_0$ . So, accordingly the situations can be altered here. Another point is I may put here say  $H_1$  as  $\mu = \mu_0$  against  $K_1$   $\mu > \mu_0$ . Will the test change? The test will not change, because what we are testing is whether the value of mean is less or more.

In the null hypothesis it is less in the alternative hypothesis it is more. Only the relative position is important, but that is determined by the test statistic or you can say the test function because the value of the control is utilized here. So, whether you say  $\mu \leq \mu_0$  or  $\mu = \mu_0$  does not play much role here in the test function; the test function will remain the same.

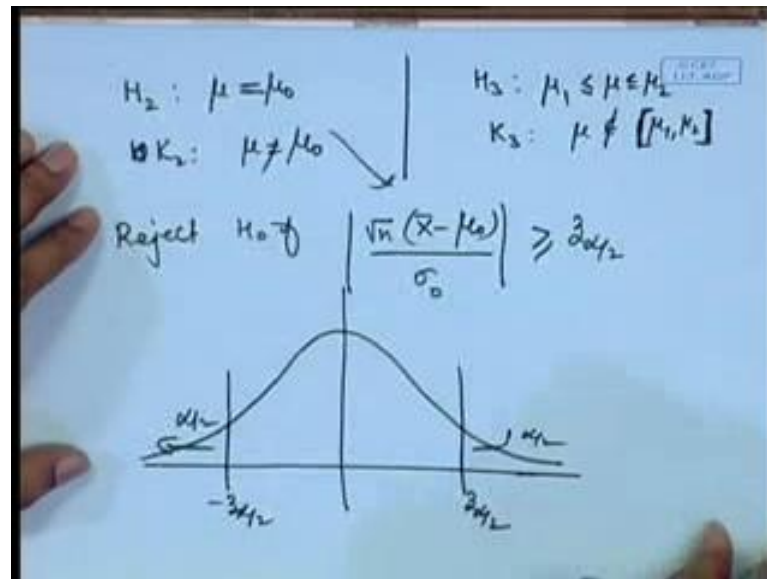
On the other hand if I have considered say  $H_1^*$  say where  $\mu \geq \mu_0$  against say  $K_1^*$  that is  $\mu < \mu_0$ , then the nature of the null and alternative hypothesis has got reversed. So, you will be actually rejecting for a smaller values of  $\bar{X}$  and when the size  $\alpha$  is fixed then the point that you will be getting here this will become this probability will become  $\alpha$ , so this point will be  $z_{1-\alpha}$ .

But in the normal distribution  $z_{1-\alpha} = -z_{\alpha}$ . So, this is reducing to then reject  $H_0$  if  $\sqrt{n}(\bar{X} - \mu_0) / \sigma_0$  is less than or equal to  $-z_{\alpha}$ . Once again whether you include equality or not that does not play any role here. And likewise once again since the relative position is important, therefore  $H_1^*$   $\mu = \mu_0$  against  $K_1^*$   $\mu < \mu_0$  will also have the same test for hypothesis here.

A point about the actual application here when we observe a random sample then the value of  $\bar{X}$  can be calculated, and therefore the value of the test statistics which we call  $\sqrt{n}(\bar{X} - \mu_0) / \sigma_0$  can be found out from the sample. And therefore, and the value of  $z_{\alpha}$  can be seen from the tables of the normal distribution, therefore the test function is a or you can say it is a very precise kind of test here; one can easily find it out here in the given situation.

Now, we may have situations of different type.

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For example, we may like to test; let me give another name say  $H_2$   $\mu$  is equal to  $\mu_0$  against say  $K_2$   $\mu$  is not equal to  $\mu_0$ . Now this kind of situations occurs for example, we are looking at the error in the measurements. So, if there is no error; that means, your measuring device is unbiased then may be  $\mu$  is equal to 0. On the other hand if it is biased then you will have either  $\mu$  to be less than 0 or  $\mu$  is greater than 0.

Suppose, we are completely unaware of whether it is biased or unbiased, so we may not like to test whether  $\mu$  is greater than 0 that is over biased or unbiased, we do not have any interest in under estimation or over estimation. So, we simple want to know whether it is biased or unbiased. In that case a test statistic of this form will be or a test function a null and alternative hypothesis of this nature will be framed. Of course, from the theory of testing of hypothesis when is a generalized name and Pearson lemma etcetera applications of that we get a uniformly most powerful test here. Once again let me not utilize this terminology here.

So, here what happens that you are going to accept, let me write another one which is parallel to this something like saying  $\mu_1$  is less than are equal to  $\mu$  less than or equal to  $\mu_2$  against  $K_3$  when  $\mu$  does not belong to this interval  $\mu_1$  to  $\mu_2$ . If we see actually there is not much difference between the hypothesis  $H_2$  versus  $K_2$  or  $H_3$

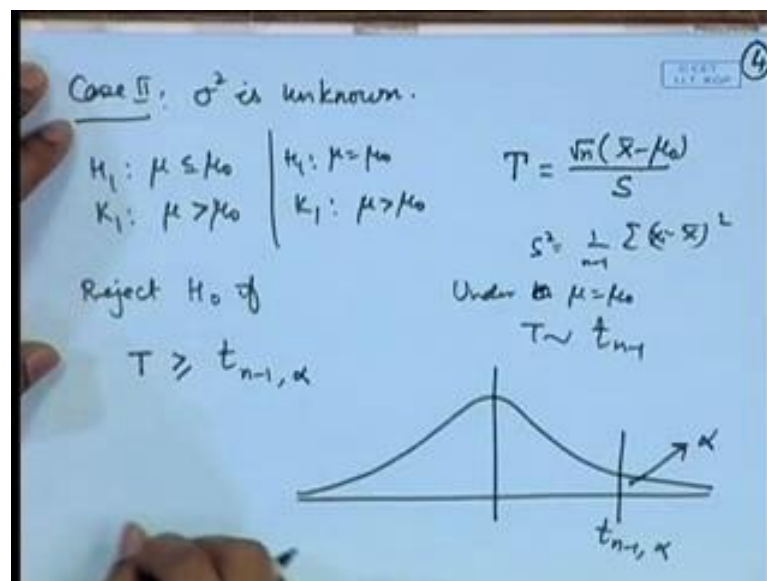
versus  $K_3$  as far as the theory of Neyman-Pearson lemma is concerned, because all that we are concerned about is the nature of the mean here.

So for example, here you are saying mean lies in a range and against  $\mu$  does not lie in that range. So, here we are saying it is equal to a value or not equal to a value. In a sense actually this problem is a generalization because, in place of one value if we say a small range we say we are permitting a variation from say minus 0.5 to plus 0.5 in the measuring device. Then in that case a particular hypothesis will be something like minus half to plus half against whether  $\mu$  is having more variability in the measuring device than that.

So, likewise the test for both of this will be same, and therefore the test will be something like you will be rejecting for both large negative as well as large positive values of  $\bar{X}$ . So, test function will be reject  $H_0$  if  $\sqrt{n}(\bar{X} - \mu_0) / \sigma$ ; specially for this one I am writing  $\mu_0$  by  $\sigma_0$  is greater than or equal to  $z_{\alpha/2}$  why this  $z_{\alpha/2}$  has come because of we are looking at the probability of the type one error here then you are having rejection in both the regions. So, this point has to be then  $\alpha/2$  and this point has to be minus  $z_{\alpha/2}$ .

Now difficulty will arise when  $\sigma$  is unknown, because in that case I will not be able to make use of this  $\sigma_0$  value here. So, in that case what we do?

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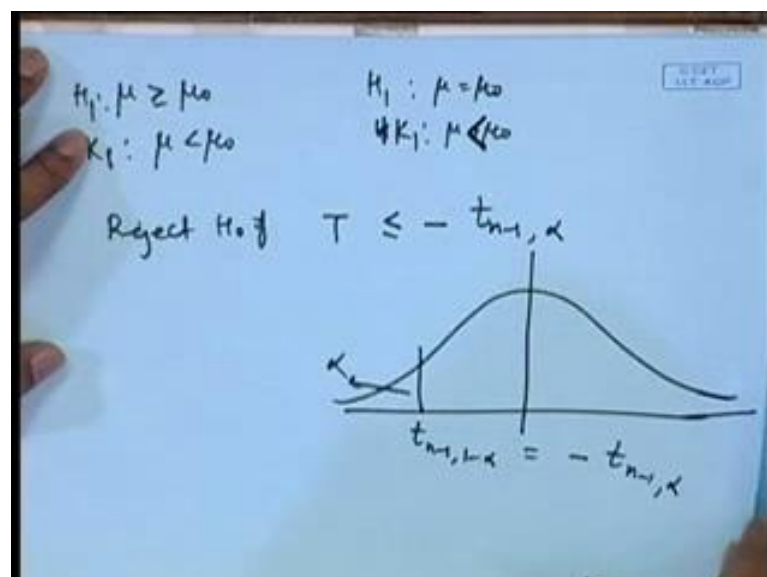
Sigma square is unknown as you remember the case of the confidence interval the sigma square value was replaced by its estimate that is S square. So, if we do that then the test function will be dependent upon a t distribution, because in that case  $\sqrt{n}(\bar{X} - \mu_0)$  by S that will have a t distribution on  $n - 1$  degree of freedom when  $\mu$  is equal to  $\mu_0$  is considered to be true.

So, let me take the hypothesis problem say  $H_1$  that is  $\mu \leq \mu_0$  against say  $K_1$   $\mu > \mu_0$  or a variation of that is  $\mu = \mu_0$  against say  $\mu > \mu_0$ . Then the test will be based on, so let me define the statistic  $T$  that is equal to  $\sqrt{n}(\bar{X} - \mu_0) / S$ ; where  $S^2$  is  $\frac{1}{n-1} \sum (\xi_i - \bar{X})^2$ .

Then under  $H_0$  that is when  $\mu = \mu_0$   $T$  follows a t distribution on  $n - 1$  degree of freedom. So, what happens the test will become reject  $H_0$  if  $T$  is greater than or equal to  $t_{n-1, \alpha}$ . Like the standard normal distribution t distribution is also a symmetric distribution. So, if we keep this probability as  $\alpha$  then  $t_{n-1, \alpha}$  point.

So, when this  $t$  value crosses this value then we reject  $H_0$ . So, you can see that the nature of the test has not changed much, because it is still dependent upon  $\bar{X}$ . However, earlier the scaling factor was known now it is unknown, so we have to replace it by estimate of that that is calculated from the sample here.

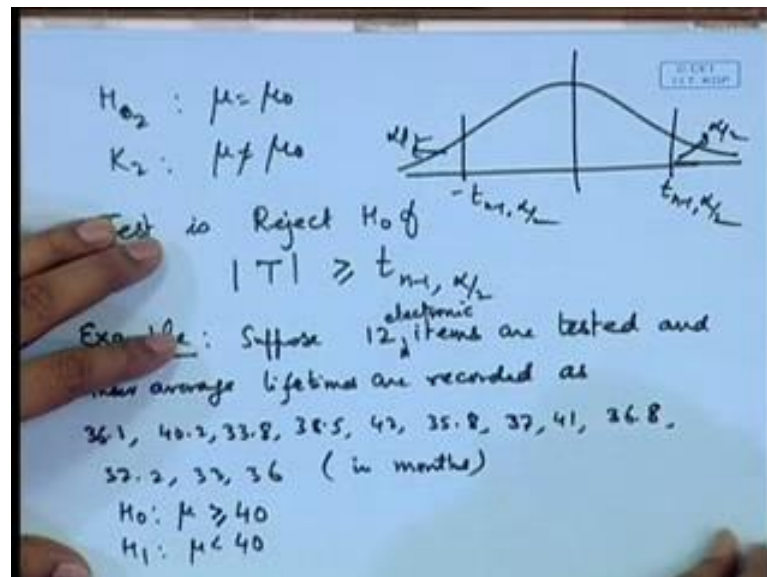
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So, likewise if we get the reverse of this hypothesis say-  $\mu$  is greater than or equal to  $\mu_0$  against say  $\mu$  is less than  $\mu_0$  or a variation of this could be  $H_0: \mu \geq \mu_0$  against  $H_1: \mu < \mu_0$ ; sorry  $\mu$  is less than  $\mu_0$  against  $\mu$  is greater than  $\mu_0$ ; then the test will be reject  $H_0$  if less than or equal to minus  $t_{n-1, 1-\alpha}$ . Because if we are looking at the point on the  $t$  distribution here then this probability is now  $\alpha$ , so point becomes  $t_{n-1, 1-\alpha}$ ; because of the symmetry of the  $t$  distribution this becomes minus  $t_{n-1, \alpha}$  here.

In a similar way we can consider the case of two sided hypothesis.

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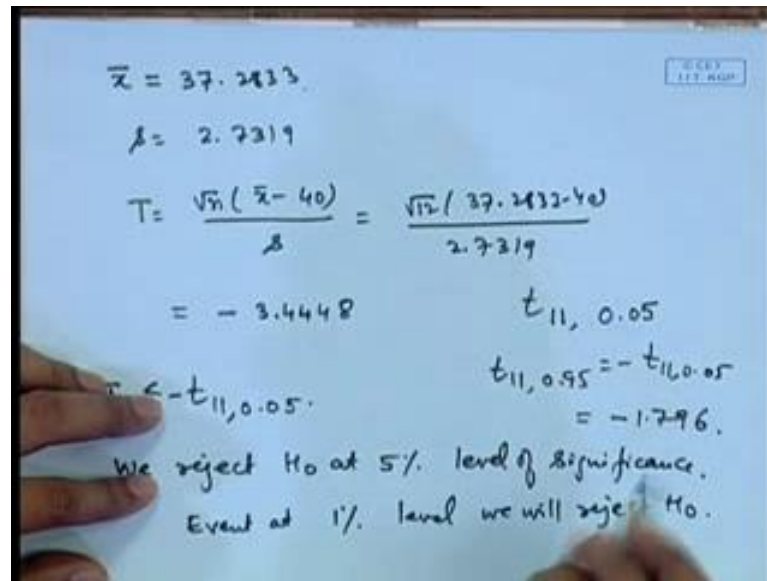


That is say we use the notation say  $H_2: \mu = \mu_0$  against say  $K_2: \mu \neq \mu_0$ , then the test will be reject  $H_0$  if modulus of  $T$  is greater than or equal to  $t_{n-1, \alpha/2}$ . This has happened because now we have the rejection region on both the sides, and therefore this probability will become  $\alpha/2$  now. So, this point becomes  $t_{n-1, \alpha/2}$ , this is minus  $t_{n-1, \alpha/2}$ .

Let me take one example here: suppose 12 items are tested and their average life times are recorded as say 36.1, 40.2, 33.8, 38.5, 42, 35.8, 37, 41, 36.8, 37.2, 33, 36; suppose certain electronic items are tested and this is in months. Now the claim here is that the average life is at least 40 against say  $H_1: \mu < 40$ . Now if we want to do the test of hypothesis at a certain level of significance here then we will be making use of the  $t$  variable here.



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$$\begin{aligned}\bar{x} &= 37.2833 \\ s &= 2.7319 \\ T &= \frac{\sqrt{n}(\bar{x} - 40)}{s} = \frac{\sqrt{12}(37.2833 - 40)}{2.7319} \\ &= -3.4448\end{aligned}$$
$$t_{11, 0.05}$$
$$t_{11, 0.95} = -t_{11, 0.05} = -1.796$$

$T < -t_{11, 0.05}$

We reject  $H_0$  at 5% level of significance.  
Even at 1% level we will reject  $H_0$ .

So for example here we can calculate;  $\bar{X}$  turns out to be 37.2833, the  $s$  value turns out to be 2.7319, say the  $t$  value that is  $\sqrt{n}(\bar{X} - 40)$  divided by  $s$  that is  $\sqrt{12}$  into 37.2833 minus 40 divided by 2.7319; this value turns out to be minus 3.4448.

So, now the test function will be to reject  $H_0$  if this value of  $t$  is less than or equal to  $t_{n-1}$  that is 11 at  $\alpha$ . So, suppose here I take  $\alpha$  is equal to 0.05. So, we may consider  $t_{11, 0.95}$  that is equal to minus  $t_{11, 0.05}$ ; that is equal to minus 1.796. This we can see from the tables of the  $t$  distribution. Now you see here  $t$  is less than or equal to  $t_{11, 0.05}$ .

So, we reject  $H_0$  at 5 percent level of significance. Suppose we change the level of significance to another value we may take say 10 percent. Let us see the values of  $t$  distribution from; so in case we decide to modify the level of significance here as say 0.1 or 0.01 etcetera. So at the  $t$  value you can see here, suppose in place of this we make 0.01 then you can see the value of  $t$  is 2.718, but this value minus 3.4448 is a even a smaller than that.

So, even at say one percent level we will reject  $H_0$ . So, now you can see here the manufacturer of the items claims that the average life is more than or equal to 40 months, but his sample that does not support the hypothesis, because you can see from here the values are much smaller. Another point is that  $\bar{X}$  is 37 which as of course less than

40. But is it really significantly smaller? So the answer is yes, because the standard deviation also does not help too much it is 2.73. So, even with the 12 observations you are getting this value to be pretty high that is pretty negative value.

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$H_0: \mu = 40$   
 $H_1: \mu \neq 37$

$$\frac{\sqrt{n}(\bar{x} - 37)}{s} = \frac{\sqrt{12}(.2832)}{2.7319} = ( )$$

at 1%, 5% we will not reject  $H_0$ .

On the other hand if we had tested something like; suppose we want to test here  $H_0$  naught say  $\mu$  is equal to 40 against say  $H_1$  in place of 40 suppose I put 37 against say  $\mu$  is not equal to 37. Then if you look at the value of  $\sqrt{n}(\bar{X} - 37) / S$  that is  $\sqrt{12}$  that becomes 0.2833 divided by 2.7319; this value is much smaller and in fact at say 1 percent or 5 percent etcetera we will not reject  $H_0$ .

For example if we are looking at say 5 percent level say- so 5 percent means we have to see the value of 0.0 to 5 here at 11 that is 2.201 which is pretty high. And this value will be much smaller, because if we are looking at  $\sqrt{12}$  this value will be say three something and this is 1.5. So, this value will become 0.3 something this is much smaller. So, even if we take say 0.1 so the 0.05 you have to see that is 1.796 etcetera.

So, all these values you will not able to reject  $H_0$  naught. Why, because the value  $\mu$  is equal to 37 pretty close to the sample mean here and the variance also supports that; in the sense that the value of this is not too small. If the variability was extremely small then even this difference would have become large here.