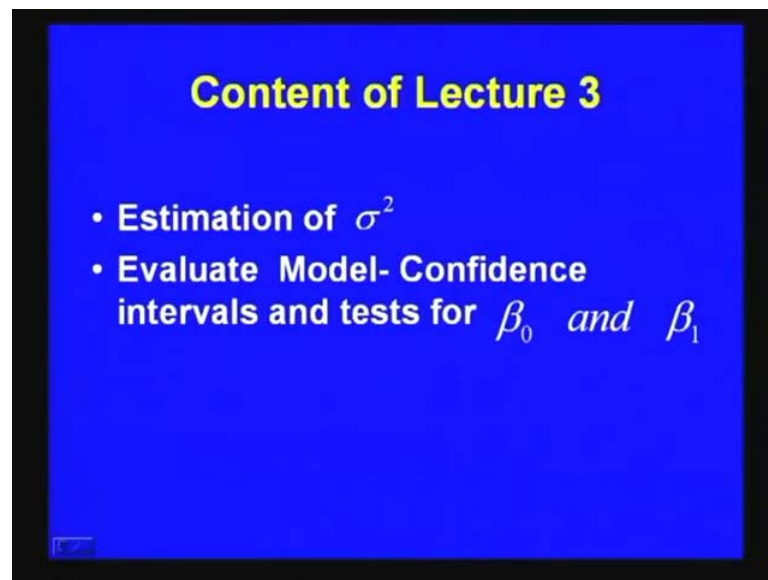


Regression Analysis
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Lecture - 3
Simple Linear Regression (Contd.)

Hi, this is my third lecture in Simple Linear Regression. In the first lecture, we have learned how to fit simple linear regression model to a data set, which fits the data best. That means, we have learned how to fit a simple linear regression model using the least square technique. In the second lecture, we have learned the statistical property of the regression coefficient that is, beta naught hat and beta 1 hat. And we have observed that, both the beta naught hat and beta 1 hat, they are unbiased estimator of beta naught and beta 1 respectively. And we computed the variance of beta naught hat and beta 1 hat and we found that, both the variance of beta naught hat and beta 1 hat, they involve sigma square. So, sigma square is the population variance, which is unknown, so what we need to do is that, we need to estimate the population variance sigma square.

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So, here is the content of today's lecture, we are going to estimate the population variance, sigma square. So, we will give an unbiased estimated of sigma square and next, we evaluate the performance of the fitted model. So, we will talk about the confidence intervals and tests for the regression coefficients, beta naught and beta 1.

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Estimation of σ^2

$$SS_{Res} = S_{yy} - \hat{\beta}_1^2 S_{xy} \quad E\left(\frac{SS_{Res}}{n-2}\right) = \sigma^2$$

$$E(SS_{Res}) = E(S_{yy}) - E(\hat{\beta}_1^2 S_{xy})$$

So, first we talk about the estimation of sigma square, the estimation of sigma square is obtained from S S residual. And in lecture 2, we have proved that, S S residual, this can be written in the form S y y minus beta 1 hat square S x y. Now, our ultimate aim is to prove that, S S residual by n minus 2, this is an unbiased estimator of sigma square. So, what you need to do is that, we will find the expected value of residual sum of square. So, expected value of residual sum of square, S S residual is equal to expectation of S y y minus expectation of beta 1 hat square S x y. So, first let me find the value of expected value of S y y.

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$$E(S_{yy}) = E \sum (y_i - \bar{y})^2$$

$$= E [\sum y_i^2] - n E [\bar{y}^2] = \sum E [y_i^2] - n E [\bar{y}^2]$$

$$= n \sigma^2 + \sum (\beta_0 + \beta_1 x_i)^2 - \sigma^2 - n (\beta_0 + \beta_1 \bar{x})^2$$

$$= (n-1) \sigma^2 + \beta_1^2 (\sum x_i^2 - n \bar{x}^2) \quad E(y_i^2) = V(y_i) + [E(y_i)]^2$$

$$= (n-1) \sigma^2 + \beta_1^2 S_{xx} \quad = \sigma^2 + (\beta_0 + \beta_1 x_i)^2$$

$$E(\bar{y}^2) = V(\bar{y}) + [E(\bar{y})]^2 = \frac{\sigma^2}{n} + (\beta_0 + \beta_1 \bar{x})^2$$

$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$
 $E(y_i) = \beta_0 + \beta_1 x_i$
 $V(y_i) = \sigma^2$

So, expected value of $\sum y_i^2$, which is equal to, expectation of $\sum y_i^2$ is nothing but $\sum (y_i - \bar{y})^2$ and this can be written as, expectation of $\sum y_i^2 - n \bar{y}^2$. Now, again this one is equal to $\sum E(y_i^2) - n \bar{y}^2$. So, what is expectation, let me recall the model $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$.

And we assume that, expected value of ϵ_i is equal 0, so expected value of y_i equal to $\beta_0 + \beta_1 x_i$, x_i is not a random variable. And we also know, we also assume that, the variance of ϵ_i is equal to σ^2 , so the variance of y_i is also equal to σ^2 . Now, expected value of y_i^2 is equal to variance of y_i plus expectation of y_i whole square, this is from the definition of the variance. Now, the variance of y_i is equal to σ^2 and expectation of y_i is this quantity, so this is equal to $\sigma^2 + (\beta_0 + \beta_1 x_i)^2$.

And similarly, we can find out the expected value of \bar{y}^2 , it can be proved that, expected value of \bar{y}^2 is equal to, of course this is equal to variance of \bar{y} plus expectation of \bar{y} whole square. So, variance of \bar{y} is equal to σ^2/n and expectation of \bar{y} is equal to $\beta_0 + \beta_1 \bar{x}$, this is equal to $\sigma^2/n + (\beta_0 + \beta_1 \bar{x})^2$.

Now, basically what I will do is that, I will plug these values here, so expected value of $\sum y_i^2$ is equal to this thing, $\sum (\sigma^2 + (\beta_0 + \beta_1 x_i)^2) - n(\sigma^2/n + (\beta_0 + \beta_1 \bar{x})^2)$, which is this quantity, so $n \sigma^2 + \sum (\beta_0 + \beta_1 x_i)^2 - n \sigma^2 - n(\beta_0 + \beta_1 \bar{x})^2$. So, this is basically $\sum (\beta_0 + \beta_1 x_i)^2 - n(\beta_0 + \beta_1 \bar{x})^2$. So, little bit algebra, we will prove that, this is nothing but $(n-1) \sigma^2 + \beta_1^2 \sum x_i^2 - n \beta_1^2 \bar{x}^2$. And this is nothing but $(n-1) \sigma^2 + \beta_1^2 S_{xx}$, this is the notation for this term, $\sum (x_i - \bar{x})^2$.

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$$\begin{aligned} E(S_{yy}) &= (n-1)\sigma^2 + \beta_1^2 S_{xx} \\ E(\hat{\beta}_1^2 S_{xx}) &= S_{xx} E(\hat{\beta}_1^2) \\ &= S_{xx} \left[\frac{\sigma^2}{S_{xx}} + \beta_1^2 \right] \\ &= \sigma^2 + \beta_1^2 S_{xx} \end{aligned}$$
$$\begin{aligned} E(\hat{\beta}_1) &= \beta_1 \\ V(\hat{\beta}_1) &= \frac{\sigma^2}{S_{xx}} \\ E(\hat{\beta}_1^2) &= V(\hat{\beta}_1) + [E(\hat{\beta}_1)]^2 \\ &= \frac{\sigma^2}{S_{xx}} + \beta_1^2 \end{aligned}$$

So, what we proved that, expected value of S_{yy} is equal to n minus 1 sigma square plus beta 1 square S_{xx} .

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Estimation of σ^2

$$\begin{aligned} SS_{Res} &= S_{yy} - \hat{\beta}_1^2 S_{xx} \quad E\left(\frac{SS_{Res}}{n-2}\right) = \sigma^2 \\ E(SS_{Res}) &= E(S_{yy}) - E(\hat{\beta}_1^2 S_{xx}) \\ &= (n-1)\sigma^2 + \beta_1^2 S_{xx} - \sigma^2 - \beta_1^2 S_{xx} \\ &= (n-2)\sigma^2 \\ \therefore E\left(\frac{SS_{Res}}{n-2}\right) &= \sigma^2 \end{aligned}$$

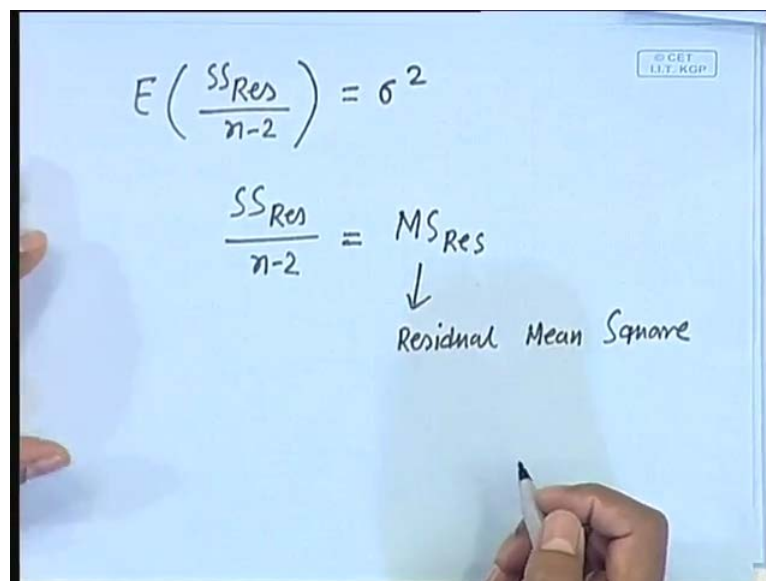
Now, see that, you want to compute the expected value of residual sum of square, so this involves the expected value of S_{yy} and the expected value of beta 1 hat square S_{xx} . Next we will compute the expected value of this one, so this is equal to next expectation of beta 1 hat square S_{xx} , is equal to, I think I did a mistake here, this is not $x y$, this is $x x$, so this is not $x y$, this is $x x$. So, this one is equal to S_{xx} , expected value of beta 1

square and we know that, expected value of beta 1 hat is equal to beta 1, this is an unbiased estimator and the variance of beta 1 hat is equal to sigma square S x x.

So, expected value of beta 1 hat square is equal to variance of beta 1 hat plus expected value of beta 1 hat whole square. We know both the values here, this is equal to sigma square by S x x and this one is equal to beta 1 square. So, this thing is equal to S x x and expected value of beta 1 hat square is equal to sigma square S x x plus beta 1 square, so this is going to be equal to sigma square plus beta 1 square S x x.

Now, just we need to plug these two values here, expected value of residual sum of square is equal to, we proved that, this one is equal to n minus 1 sigma square plus beta 1 square S x x and this one is equal to minus sigma square minus beta 1 square S x x and this one is nothing but n minus 2 sigma square. So, what we proved is that, expected value of S S residual, residual sum of square by n minus 2 is equal to sigma square that means, residual sum of square by n minus 2 is an unbiased estimator of sigma square.

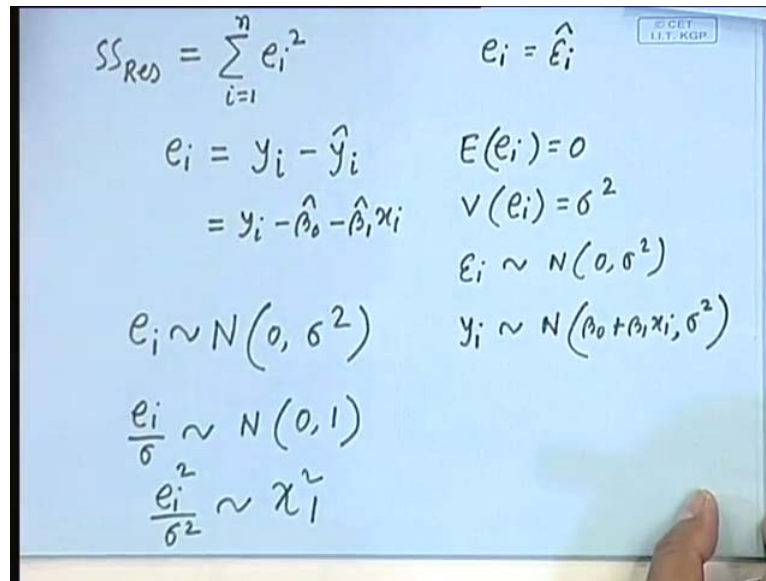
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$$E\left(\frac{SS_{Res}}{n-2}\right) = \sigma^2$$
$$\frac{SS_{Res}}{n-2} = MS_{Res}$$

↓
Residual Mean Square

So, what we proved is that, expected value of S S residual by n minus 2 is equal to sigma square and this S S residual by n minus 2, this is also denoted by M S residual and this is called residual mean square. This is ultimately found an unbiased estimator of sigma square, which is M S residual, next we will talk about the distribution of M S residual.

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The image shows a whiteboard with handwritten mathematical derivations. The equations are as follows:

$$SS_{Res} = \sum_{i=1}^n e_i^2 \quad e_i = \hat{\epsilon}_i$$
$$e_i = y_i - \hat{y}_i \quad E(e_i) = 0$$
$$= y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i \quad V(e_i) = \sigma^2$$
$$e_i \sim N(0, \sigma^2) \quad y_i \sim N(\beta_0 + \beta_1 x_i, \sigma^2)$$
$$\frac{e_i}{\sigma} \sim N(0, 1)$$
$$\frac{e_i^2}{\sigma^2} \sim \chi_1^2$$

So, what is S S residual, residual sum of square is nothing but summation e_i^2 , i is from 1 to n . So, this e_i is nothing but the i th residual, this is the difference between the observed response value and the predicted response value. So, it can be proved that, expected value of e_i is equal to 0, so you can prove it and also it can be proved that, the variance of e_i is equal to σ^2 . See, this e_i is nothing but the estimate of i th error term, the variance of e_i , this can be proved σ^2 .

And also we have assumed that, ϵ_i follows normal zero σ^2 and they are independent, which implies that, the observation y_i , they are also normal with mean $\beta_0 + \beta_1 x_i$ and variance σ^2 . Now see, this e_i is, it is linear combination of y_i , this \hat{y}_i is nothing but $\hat{\beta}_0 + \hat{\beta}_1 x_i$. And both we proved that, $\hat{\beta}_1$ is linear combination of the observations and also $\hat{\beta}_0$ is also a linear combination of the observation.

So, the whole thing this is e_i , is a linear combination of the observations, which implies that, e_i follows normal. Because, e_i is a linear combination of normal variables and the linear combination of normal variables is also normal. So, that is why, e_i follows normal with mean 0 and the variance is equal to σ^2 and from here, we can say that, e_i/σ follows standard normal (0,1). And also since e_i/σ follows standard normal, we can say, e_i^2/σ^2 follows χ_1^2 . And S S residual is, basically it is a sum of e_i^2 , i equal to 1 to n , but the distribution of S S

residual is not chi square n, because all the e_i are not independent, this e_i satisfies the some constant.

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$\hat{\beta}_0$ & $\hat{\beta}_1$ are LSE of β_0 & β_1
 res.

$e_i = y_i - \hat{y}_i$ satisfy

$e_1 + e_2 + \dots + e_n = 0$ — (A)

$e_1 x_1 + e_2 x_2 + \dots + e_n x_n = 0$ — (B)

There are $(n-2)$ degree of freedom for residuals.

$\frac{SS_{Res}}{\sigma^2} = \sum_1^n \frac{e_i^2}{\sigma^2} \sim \chi_{n-2}^2$

We know that, beta naught hat and beta 1 hat are least square estimator of beta naught and beta 1 respectively. And this e_i , which is equal to y_i minus y_i hat, they satisfy the constant, that e_1 plus e_2 , e_n is equal to 0. That is, this is what we proved before also, the sum of the residuals is equal to 0 and also it satisfies, this basically the first normal equation, which residuals satisfy. And the second normal equation is summation $e_i x_i$ is equal to 0 that is, $e_1 x_1$ plus $e_2 x_2$ plus $e_n x_n$ is equal to 0.

So, what I want to prove here is that, this SS_{Res} by σ^2 , which is equal to summation e_i^2 by σ^2 . This does not follow chi square n, it follows chi square with degree of freedom $n - 2$, because there are $n - 2$ degree of freedom for the residuals. All the e_i are not independent, the first I mean, you can choose $n - 2$ residuals independently and then the remaining two residuals have to be chosen in such a way that, they satisfy the condition that, summation e_i is equal to 0 and summation $e_i x_i$ is equal to 0.

So, you have the freedom of choosing $n - 2$ residuals or $n - 2$ e_i independently and the remaining two residuals must be chosen in such a way that, they satisfy these two conditions. That is why, the distribution of SS_{Res} by σ^2 follows chi square with degree of freedom $n - 2$, it is not chi square n. If all the e_i are, if there

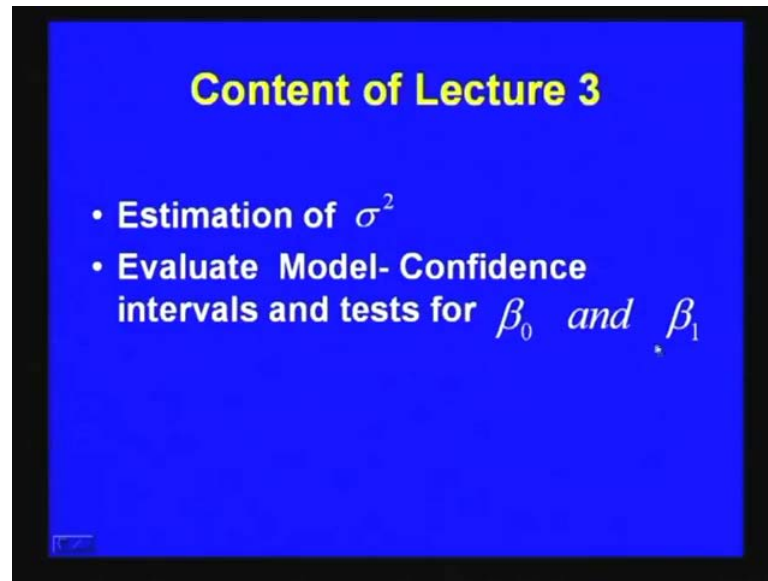
is no constrain on e_i then this could follow chi square n , but since we have two constrain on the residuals, it does not follow chi square n , it follows chi square n minus 2. So, I repeat just that, we have the freedom of choosing n minus 2 residuals independently and the remaining two residuals have to be chosen in such a way that, they satisfy these two conditions like summation e_i equal to 0 and summation $e_i x_i$ equal to 0.

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The image shows a whiteboard with handwritten mathematical equations. At the top, it states $\frac{SS_{Res}}{\sigma^2} \sim \chi^2_{n-2}$. Below this, a box contains the equation $\frac{(n-2) MS_{Res}}{\sigma^2} \sim \chi^2_{n-2}$. To the right of the box, it concludes with $\therefore MS_{Res} = \frac{SS_{Res}}{n-2}$. A small logo in the top right corner of the whiteboard reads '© CET I.I.T. RGP'.

What we proved is that, summation $SS_{Residual}$ by σ^2 follows chi square n minus 2, which is equivalent to say that, n minus 2 $MS_{Residual}$ by σ^2 follows chi square n minus 2, because $MS_{Residual}$ is nothing but $SS_{Residual}$ by n minus 2. So, this is the result we proved and this is very much useful in testing of hypothesis.

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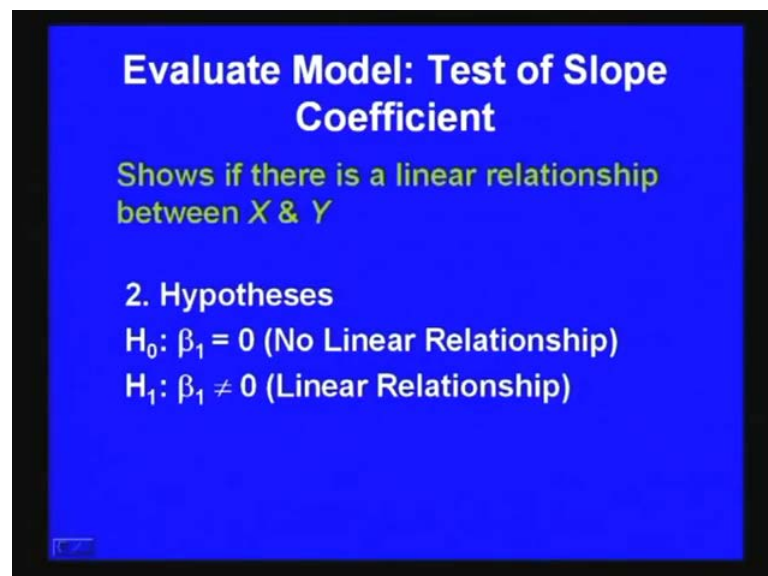


Content of Lecture 3

- Estimation of σ^2
- Evaluate Model- Confidence intervals and tests for β_0 and β_1

So, next we move to the evaluation of model, so what we learnt in the 1 st lecture is that, given the state of data or given the state of observations, we learned how to estimate the regression coefficient. That means, we have learned, how to fit regression model to the data, so once the linear model has been fitted, the next job is to confirm the goodness of the fit. So, what we will do is that, we are going to test the significance of the regression coefficients, beta naught and beta 1.

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Evaluate Model: Test of Slope Coefficient

Shows if there is a linear relationship between X & Y

2. Hypotheses

$H_0: \beta_1 = 0$ (No Linear Relationship)

$H_1: \beta_1 \neq 0$ (Linear Relationship)

First we will test the hypothesis that, beta 1 is equal to 0, this is null hypotheses against the alternate hypothesis that, beta 1 not equal to 0. So, what is the significance of this null hypothesis, beta 1 is equal to 0, if beta 1 is equal to 0 then the model will become Y equal to beta naught plus epsilon that means, there is no linear relationship between the variables X and Y. So, if H naught is accepted then we conclude that, there is no linear relationship between the regressor variable and response variable.

And we say that, X is of little value in explaining the variation in Y, whatever be the value of X, we can estimate the regressor variable y i by y bar. If H naught is rejected that means, if the alternative hypothesis, the two sided alternate hypothesis beta 1 is accepted. That means, it says that, there is a linear relationship between the regressor variable X and the response variable Y and X is of value in explaining the variation in Y, this is the significance of the alternative hypothesis, so how to perform this test, we have to compute the value of the test statistics.

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$$H_0: \beta_1 = 0$$

$$H_1: \beta_1 \neq 0$$

$$\hat{\beta}_1 = \frac{\sum (x_i - \bar{x}) y_i}{\sum (x_i - \bar{x})^2} = \sum c_i y_i$$

$$y_i \sim N(\beta_0 + \beta_1 x_i, \sigma^2)$$

$$\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma^2}{S_{xx}}\right)$$

$$\text{Thus } Z = \frac{\hat{\beta}_1 - \beta_1}{\sqrt{\frac{\sigma^2}{S_{xx}}}} \sim N(0, 1)$$

And here, to test the hypothesis, H naught which is equal to beta 1, is equal to 0, against the alternative hypotheses H 1, beta 1 not equal to 0. The test statistics here is beta 1 hat, which is an estimator of beta 1 and before we have prove that, beta 1 hat which is equal to summation x i minus x bar into y i by summation x i minus x bar whole square, this can be written as summation C i y i. That means, beta 1 hat is linear combination of the

observation y_i and you know that, y_i follows normal distribution with some mean and variance σ^2 and the mean is $\beta_0 + \beta_1 x_i$.

So here, $\hat{\beta}_1$ is linear combination of normal variables, which implies that, $\hat{\beta}_1$ is also normal with mean. We know the mean of $\hat{\beta}_1$, $\hat{\beta}_1$ is an unbiased estimator, the mean of $\hat{\beta}_1$ is β_1 and the variance of $\hat{\beta}_1$ is we proved before that is, σ^2 / S_{xx} . So, we need the sampling, this is called the sampling distribution of $\hat{\beta}_1$, to find the critical value for the testing of hypothesis. Now, from here, we can say that, the z equal to say, $\hat{\beta}_1 - \beta_1$ by σ^2 / S_{xx} , we know that, this follows normal $(0,1)$ standard normal.

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Test Statistic

$$Z = \frac{\hat{\beta}_1}{\sqrt{\frac{\sigma^2}{S_{xx}}}} \quad \text{under } H_0: \beta_1 = 0$$

if σ^2 is known, we can use Z to test the hypothesis $H_0: \beta_1 = 0$

Reject H_0 if $|Z| > z_{\alpha/2}$

The diagram shows a standard normal distribution curve with the area under the curve in the tails beyond $-z_{\alpha/2}$ and $z_{\alpha/2}$ shaded, representing the rejection region.

And the test statistic is equal to $\hat{\beta}_1$ by σ^2 / S_{xx} , this is z , z equal to this under H_0 , because under H_0 , β_1 is equal to 0. And this is the test statistic to test the hypothesis that, $H_0: \beta_1 = 0$ against the alternative hypothesis that, $\beta_1 \neq 0$. Now, see usually these variance I mean, σ^2 is not known, the population variance is not known. If σ^2 is known, we can use z to test the hypothesis $H_0: \beta_1 = 0$.

And we reject the critical region here, we reject H_0 if z is greater than $z_{\alpha/2}$, so we reject the null hypothesis at α level of significance, if z is greater than $z_{\alpha/2}$. So, this $z_{\alpha/2}$ is nothing but the upper $\alpha/2$ percentage point of standard normal distribution. So, this is the point, $z_{\alpha/2}$ and this is the point minus

z alpha by 2, this is the period of standard normal distribution. So, intuitively I mean, we reject the null hypothesis, if beta 1 is not close to 0.

And formally I mean, we found the test statistic value z and we checked, if the z is greater than z alpha by 2 then we reject the null hypothesis, otherwise we accept the null hypotheses.

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Usually σ^2 is not known

Test Statistic $\sigma^2 = E(MS_{Res}) = E\left(\frac{SS_{Res}}{n-2}\right)$

$$t = \frac{\hat{\beta}_1 - \beta_1}{\sqrt{\frac{MS_{Res}}{S_{xx}}}} \sim t_{n-2}$$

$$= \frac{\hat{\beta}_1}{\sqrt{\frac{MS_{Res}}{S_{xx}}}}$$

under $H_0: \beta_1 = 0$
we reject $H_0: \beta_1 = 0$
if $|t| > t_{\frac{\alpha}{2}, n-2}$

But, I mean, usually sigma square is not known, this is the practical case I mean, we cannot assume that, sigma square is known. And we know that, unbiased estimate of sigma square is M S residual, because we proved that, expectation of M S residual, which is equal to expectation of S S residual by n minus 2, this is equal to sigma square. So, what you do is that, that is, the statistic to test the hypothesis beta 1 is equal to 0, this was of the form, beta 1 hat by root over of sigma square by S x x.

But, sigma square is not known here I mean, this is the more practical case that sigma square is not known. So, what we do is that, we just replace this sigma square by its unbiased estimator, M S residual and we call it t and this is the test statistic say, beta 1 here. This is the test statistic to test a given hypothesis, now this does not follow normal, we need to find the distribution of this one.

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$\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma^2}{S_{xx}}\right)$

$\frac{\hat{\beta}_1 - \beta_1}{\sqrt{\frac{\sigma^2}{S_{xx}}}} \sim N(0,1)$

$\frac{(n-2) MS_{Res}}{\sigma^2} \sim \chi^2_{n-2}$

Let $X \sim N(0,1)$
 $Y \sim \chi^2_n$
then $\frac{X}{\sqrt{Y/n}} \sim t_n$

indep.

What we need to find is that, $\hat{\beta}_1$ follows normal with mean β_1 and variance σ^2 / S_{xx} . And so from here, we can say that, $\hat{\beta}_1 - \beta_1$ by σ^2 / S_{xx} root, this follows standard normal, normal (0,1) this is one. And also we know that, $(n-2) MS_{Res} / \sigma^2$ follows chi square $n-2$ and it can be proved that, these two are independent. Now, there is one very standard result in sampling distribution, let X follows normal (0,1) and Y follows chi square n . And they are independent then $X / \sqrt{Y/n}$, this follows t distribution with the degree of freedom n . This is a very standard result, I am expecting that, you know sampling distribution well and now, we are going to make use of this result here. So, this one follows standard normal and this one follows chi square $n-2$.

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The image shows a whiteboard with handwritten mathematical derivations. At the top right, there is a small logo for '© CET I.I.T. RGP'. The main derivation consists of two parts. The first part shows the ratio of the estimated coefficient $\hat{\beta}_1 - \beta_1$ to the square root of the variance-covariance matrix $\frac{\hat{\beta}_1 - \beta_1}{\sqrt{\frac{\sigma^2}{S_{xx}}}}$ is approximately equal to a t-distribution with $n-2$ degrees of freedom. The second part shows the same ratio as $\frac{\hat{\beta}_1 - \beta_1}{\sqrt{\frac{MS_{Res}}{S_{xx}}}}$ is approximately equal to a t-distribution with $n-2$ degrees of freedom. The σ^2 in the first part is replaced by $\frac{(n-2)MS_{Res}}{(n-2)\sigma^2}$ in the second part.

And from here, I can say that, $\hat{\beta}_1 - \beta_1$ by $\sigma^2 S_{xx}$ by $n - 2$ MS residual, this is my X, this is my Y. So, Y by $n - 2$, root of this thing, this follows t with $n - 2$ degree of freedom. I am just making use of this result, this is my X, this is my Y and Y follows chi square $n - 2$, instead of n and x follows standard normal. So, X by root of Y by $n - 2$, this follows $t_{n - 2}$ and from here, we get that, $\hat{\beta}_1 - \beta_1$ by, this will cancel out, by MS_{Res} by S_{xx} , this follows $t_{n - 2}$, we obtain this from here only.

So, when we are talking about testing the hypothesis, $\beta_1 = 0$ against the alternative hypothesis, $\beta_1 \neq 0$. And we are considering the case that, σ^2 is not known, so basically what I did here is that, I replaced the σ^2 by MS_{Res} . And we proved that, this follows $t_{n - 2}$ and this t is equal to $\hat{\beta}_1$ by MS_{Res} by S_{xx} under H_0 . Under H_0 , β_1 is equal to 0, I am expecting that, you know testing of hypothesis well and now, we know the distribution of test statistics.

And this is a two sided test, so we reject the null hypothesis, so we reject H_0 , $\beta_1 = 0$. If this t value is greater than $t_{\alpha/2, n - 2}$ with degree of freedom $n - 2$, this $t_{\alpha/2, n - 2}$ is the upper $\alpha/2$ percentage point of t distribution with $n - 2$ degree of freedom, so this is the critical region. And if the t value is greater than this one, the mode of t value is greater than this one and then, we are going to reject


the null hypothesis, β_1 equal to 0. And we conclude that, there is a linear that, rejecting the null hypothesis means, accepting the alternative hypothesis that, β_1 not equal to 0. And we conclude that, there is a linear relationship between the regressor variable and the response variable.

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Test of Slope Coefficient Example

• You're a marketing analyst for Hasbro Toys.

<u>Ad \$</u>	<u>Sales (Units)</u>
1	1
2	1
3	2
4	2
5	4



Is the relationship **significant** at the **.05** level?

Now, let us recall the toy example and these are the cost, the money spent on advertizing and this is the sales amount and what you want to check, before we already have fitted relationship between X and Y. And now, you want to check, whether the relationship is significant at 0.05 level of significance that is, α equal to 0.05 that is, type one error, so let me do this one.

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Ad (x_i)	Y_i (Sales)	$\hat{Y}_i = -0.1 + 0.7X_i$	$e_i = Y_i - \hat{Y}_i$
1	1	0.6	0.4
2	1	1.3	-0.3
3	2	2	0
4	2	2.7	-0.7
5	4	3.4	0.6

$SS_{Res} = \sum e_i^2 = 1.1$ $MS_{Res} = \frac{SS_{Res}}{n-2}$
 $t = \frac{\hat{\beta}_1}{\sqrt{\frac{MS_{Res}}{S_{xx}}}} = \frac{0.7}{\sqrt{\frac{0.3666}{10}}} = 3.3$ $= \frac{1.1}{3} = 0.3666$
 $S_{xx} = \sum x_i^2 - n\bar{x}^2 = 55$

Here X is equal to the cost, the money spent on advertisement that is, the regressor variable. The values for this one is 1 2 3 4 5 and Y is the sales amount and the values are 1 1 2 2 4. And using the least square technique, we already have estimated the regression coefficient and the fitted model is \hat{Y} equal to minus 0.1 plus 0.7 X, so beta 1 is 0.7 and beta naught equal to minus 0.1, so this is the fitted model. Now, corresponds to X equal to 1, the predicted value of the response variable is 0.6, you just put X equal to 1 here, that gives that \hat{Y}_1 for the predicted value is 0.6, corresponds to X equal to 1.

Similarly, for X equal to 2 the predicted value is 1.3, for X equal 3 the predicted response value is equal to 2, for X equal to 4 \hat{Y} is equal to 2.7, for X equal to 5 \hat{Y} is equal to 3.4. So, this is \hat{Y}_1 \hat{Y}_2 \hat{Y}_3 \hat{Y}_4 \hat{Y}_5 , now let me compute e_i , the residuals e_i , we will call them e_i , so e_i equal to Y_i minus \hat{Y}_i that is, the observed response and the predicted response, so the values are 0.4, minus 0.3, 0, minus 0.7, 0.6. So, from here, we can compute S S residual, which is equal to summation e_i square, this can be proved that, from here you can compute this, this is equal to 1.1.

And here, M S residual is equal to, you know M S residual is S S residual by n minus 2, here n is equal to 5, we have 5 observations. So, n minus 2 equal to 3, so this gives, this equal to 1.1 by 3, which is equal to 0.3666. Let me compute the test statistics value t, t equal to beta 1 hat by root over of M S residual by S x x. So, we know what is the value of M S residual, beta 1 hat equal to 0.7 and the value of M S residual is 0.3666. What is

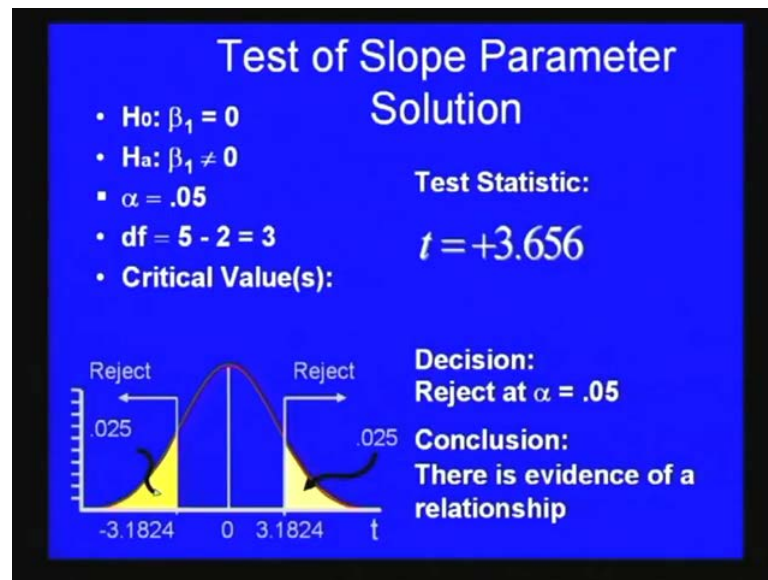
S_{xx} , S_{xy} is equal to summation x_i^2 minus $n \bar{X}^2$, so it is not difficult to compute from here, you are given the x_i 's values. So, this is equal to 55 minus n equal to 5 and \bar{X} equal to 3 square, which is going to be 10 and the t value, ultimate t value is going to be 3.356.

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$t = 3.356$
 $t_{\frac{0.05}{2}, 3} = 3.18$
 $H_0: \beta_1 = 0$ is rejected

So, the t value is equal to, the absorbed t value is 3.356 and from the table, you find the value of $t_{\alpha/2}$. Here, α is equal to 0.05 by 2 and degree of freedom is $n - 2$, n is 5, so this is equal to 3, this is going to be 3.18. So, since t is greater than this quantity, we conclude that, H_0 is that is, $\beta_1 = 0$, is rejected. That means, we say that, there is a linear relationship between the regressor variable and the response variable.

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Here is the summary, we are trying to test this hypothesis, we have the given observation and the level of significance is 0.05, degree of freedom is 5 minus 2 equal to 3. We already computed the test statistic value, which is equal to 3.656 and the critical value is 3.1824, which is nothing but t 0.025 with degree of freedom 2. And since this 3.6 is in the critical region, it lies in the critical observed value, is in the critical region, we are going to reject the null hypothesis at the alpha level of significance. And conclude that, there is evidence of linear relationship between the response variable and the regressor variable and now, we can stop for today.

Thank you.