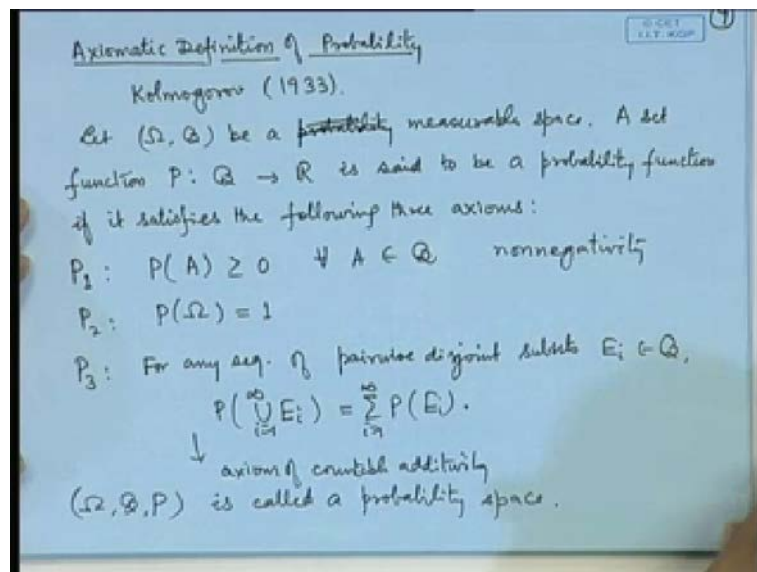


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**Module No. #01**

**Lecture No. #04**  
**Laws of Probability-I**

In the last lecture, I have introduced the axiomatic definition of probability. This takes care of the deficiencies or drawbacks left by the classical definition or the relative frequency definition of probability. So, in this definition we give a general framework under which a probability function is defined. This does not tell you how to calculate a probability. But, a probability function must satisfy these axioms in order to be a proper probability function.

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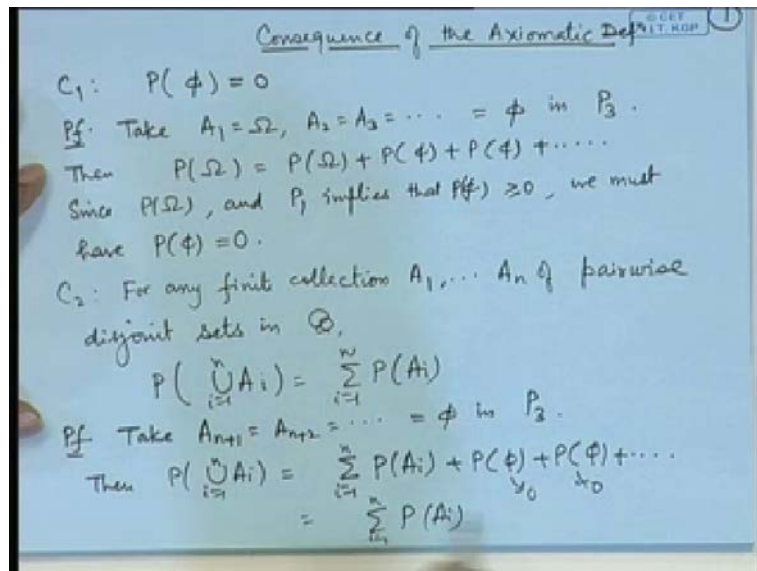


So, in particular if we have a sample space and a sigma field of subsets of that sample space let us call it as script B then a set function p from b to R is said to be a probability function if it satisfies the given 3 axioms which we name P1, P2, P3. The first is the non negativity axiom; that is, the probability of a is greater than or equal to 0 for all A belonging to B.

So, this is the axiom of non negativity; then probability of the full sample space is 1.

Basically, it makes the  $P$  be a finite function and the third axiom is the axiom of countable additivity. That is, for a given pair wise disjoint sequence of sets probability of the union is equal to the sum of the individual probabilities. Thus, this  $\Omega$   $\mathcal{B}$  and  $P$  is called a probability space.

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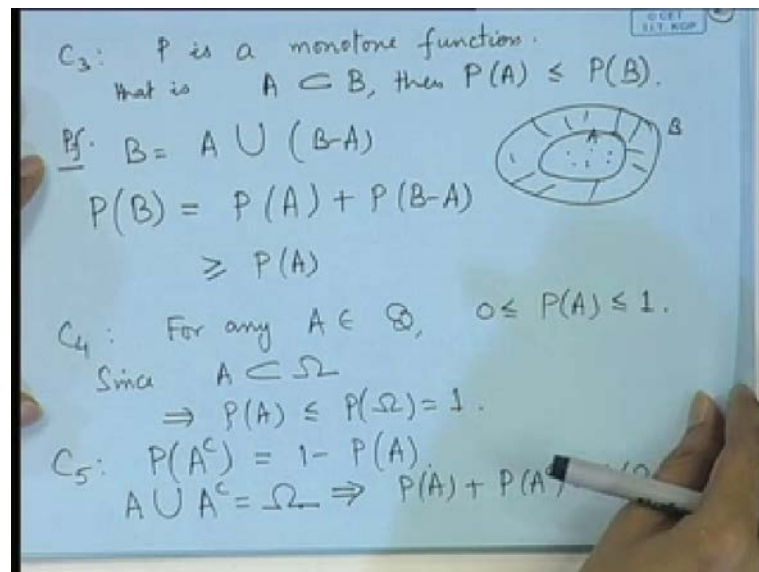
Now, some of the consequences of the axiomatic definition are as follows: the first consequence is, let me call it,  $C_1$ . That probability of the impossible event must be 0. To prove this statement let us take  $A_1$  is equal to  $\Omega$  and  $A_2, A_3, \dots$  to be  $\phi$  in axiom  $P_3$ ; then, we will get probability of  $\Omega$  is equal to probability of  $\Omega$  plus probability of  $\phi$  plus  $P$  of  $\phi$  plus  $P$  of  $\phi$ , etcetera. Since  $\Omega$  is 1 and  $P_1$  implies that  $P$  of  $\phi$  is greater than or equal to 0, we must have  $P$  of  $\phi$  equal to 0.

The second consequence is that for any finite collection  $A_1, A_2, \dots, A_n$  of pair wise disjoint sets in  $\mathcal{B}$  probability of union  $A_i$  is equal to 1 to  $n$  is equal to sigma probability of  $A_i$  is equal to 1 to  $n$ . Let me explain this that why do we need this finite additivity consequence here to be proved we have assumed the countable additivity axiom, but that does not necessarily imply the finite additivity.

A proof of this can be given using the fact that in  $A_3$ , we can take  $A_{n+1}, A_{n+2}, \dots$  to be  $\phi$  in the third axiom. Then we will get probability of union  $A_i$  is equal to 1 to  $n$  is equal to sigma probability of  $A_i$  is equal to 1 to  $n$  plus  $p$  of  $\phi$  plus  $p$  of  $\phi$  etcetera. Now, if you use consequence 1 here then, these terms are 0 and we get sigma

probability of A is equal to 1 to n.

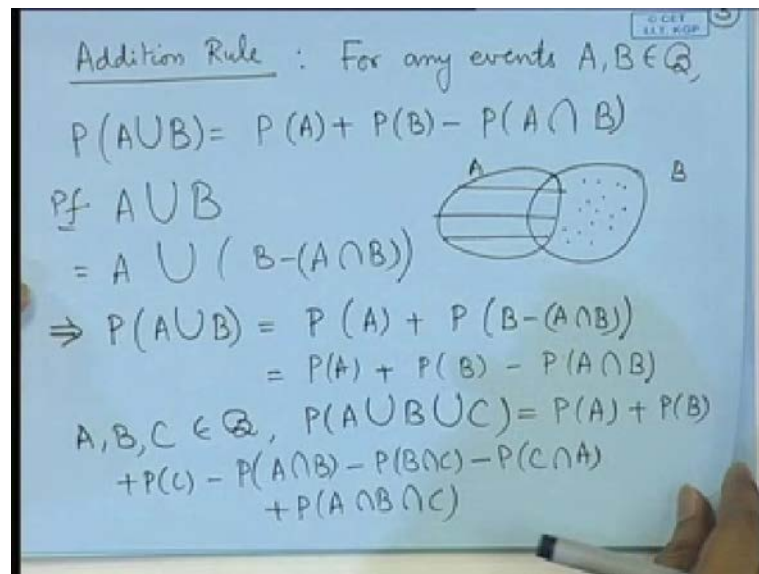
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A third consequence is the  $P$  is a monotone function that is if I take say  $A$  to be a subset of  $B$  then probability of  $A$  will be less than or equal to probability of  $B$  let us look at the proof of this consider say set  $A$  and a set  $B$  then I can write  $B$  as  $A$  union  $B$  minus  $A$  that is this is  $B$  minus  $A$  and this is  $A$  and these 2 are disjoint. So, if I make use of the finite additivity consequence then, probability of  $B$  is equal to probability of  $A$  plus probability of  $B$  minus  $A$ . Naturally, this is greater than or equal to probability of  $A$  since probability of  $B$  minus  $A$  is always greater than or equal to 0.

As a further consequence, we have that for any event a probability of  $A$  lies between 0 and 1. Now, the first part of this is always true because of the  $P_1$  axiom. Now,  $A$  is a subset of  $\Omega$ ; for every  $\Omega$  for every  $A$  this implies that probability of  $A$  is less than or equal to probability of  $\Omega$ , that is equal to 1. If I consider probability of a complement then it is equal to 1 minus probability of  $A$  this follows because, I can write  $A$  union  $A$  complement as  $\Omega$  and therefore, probability of  $A$  plus probability of  $A$  complement is probability of  $\Omega$  that is equal to 1.

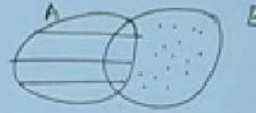
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Addition Rule : For any events  $A, B \in \mathcal{Q}$ ,

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

pf  $A \cup B$

$$= A \cup (B - (A \cap B))$$

$$\Rightarrow P(A \cup B) = P(A) + P(B - (A \cap B))$$
$$= P(A) + P(B) - P(A \cap B)$$

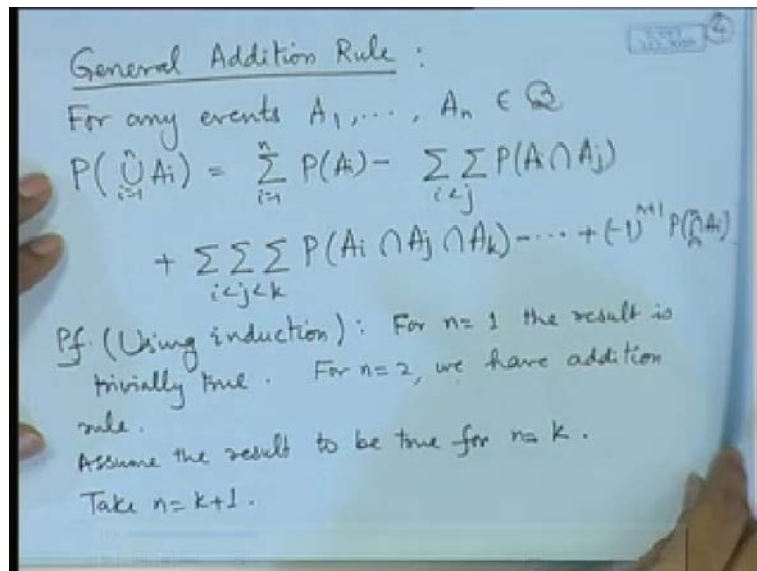
$A, B, C \in \mathcal{Q}$ ,  $P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(A \cap B) - P(B \cap C) - P(C \cap A) + P(A \cap B \cap C)$

Now, we look at certain further consequences of the definition the first of them is the addition rule for any events  $A$  and  $B$ ; probability of  $A$  union  $B$  is equal to probability of  $A$  plus probability of  $B$  minus probability of  $A$  intersection  $B$  in order to prove this statement. Let us consider any 2 sets  $A$  and  $B$  then  $A$  union  $B$  can be expressed as  $A$  union  $B$  minus  $A$  intersection  $B$ . So, we can write  $A$  union  $B$  as  $A$  union  $B$  minus  $A$  intersection  $B$ ; we can observe here that this is a disjoint union. Therefore, if I consider probability of  $A$  union  $B$  it is equal to probability of  $A$  plus probability of  $B$  minus  $A$  intersection  $B$ . Now, at this stage we notice that  $A$  intersection  $B$  is a subset of  $B$  and if we look at the statement that probability of  $B$  is equal to probability of  $A$  plus probability of  $B$  minus  $A$ , this implies that probability of  $B$  minus  $A$  is equal to probability of  $B$  minus probability of  $A$ . That means, if  $A$  is a subset of  $B$  then probability of  $B$  minus  $A$  can be expressed as probability of  $B$  minus probability of  $A$ . Therefore, here we can write this as probability of  $B$  minus probability of  $A$  intersection  $B$ .

Now, naturally I can think of the generalization of this rule for example, if I consider say for 3 events suppose  $A$ ,  $B$  and  $C$  are 3 events then we must have probability of  $A$  union  $B$  union  $C$  is equal to probability of  $A$  plus probability of  $B$  plus probability of  $C$  minus probability of  $A$  intersection  $B$  minus probability of  $B$  intersection  $C$  minus probability of  $C$  intersection  $A$  plus probability of  $A$  intersection  $B$  intersection  $C$ . I can look at this statement from the point of view of set theory or Venn diagram. If I consider 3 events say  $A$ ,  $B$  and  $C$  then the union can be expressed as  $A$  union  $B$  union  $C$ . However, here we

have to remove  $A \cap B$ ,  $B \cap C$  and  $C \cap A$ . If we remove that then the set  $A \cap B \cap C$  has been removed 3 times. So, we have to add it once to get this portion here. So, a  $A \cap B \cap C$  has to be added here.

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So, this gives us a rule for considering a general addition rule and we have the following result: general addition rule. So, if we have events  $A_1, A_2, \dots, A_n$  then, probability of union  $A_i$  is equal to 1 to  $n$  can be expressed as sigma probability of  $A_i$  is equal to 1 to  $n$  minus double summation probability of  $A_i \cap A_j$   $i < j$  plus triple summation probability of  $A_i \cap A_j \cap A_k$   $i < j < k$  minus and so on. Finally, you will have minus 1 to the power  $n+1$  probability of intersection  $A_i$  is equal to 1 to  $n$ .

I can prove this result using induction; for example, if we take  $n$  is equal to 1 the result is trivially true for extension from  $k$  to  $k+1$ . We will need the result for  $n$  is equal to 2 which has already been proved for  $n$  is equal to 2; we have addition rule. Now, assume that the result to be true for  $n$  is equal to  $k$ ; now take  $n$  is equal to  $k+1$ . So, we need to consider probability of union  $A_i$  is equal to 1 to  $k+1$  and we can consider it as probability of union  $A_i$  is equal to 1 to  $k$  union  $A_{k+1}$ .

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The image shows a handwritten derivation on a blue background. It starts with the equation:
$$P\left(\bigcup_{i=1}^{k+1} A_i\right) = P\left(\bigcup_{i=1}^k A_i \cup A_{k+1}\right)$$
Then it uses the distributive property:
$$= P\left(\bigcup_{i=1}^k A_i\right) + P(A_{k+1}) - P\left(\left(\bigcup_{i=1}^k A_i\right) \cap A_{k+1}\right)$$
Next, it applies the inclusion-exclusion principle to the union of k events:
$$= \sum_{i=1}^k P(A_i) - \sum_{i < j} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j \cap A_k) - \dots + (-1)^{k+1} P\left(\bigcap_{i=1}^k A_i\right) + P(A_{k+1})$$
Finally, it subtracts the intersection of the union of k events with the (k+1)th event:
$$- P\left(\bigcup_{i=1}^k (A_i \cap A_{k+1})\right)$$
The final result is:
$$= \sum_{i=1}^{k+1} P(A_i) - \sum_{i < j} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j \cap A_k) - \dots + (-1)^{k+1} P\left(\bigcap_{i=1}^k A_i\right) - \left[ \sum_{i=1}^k P(A_i \cap A_{k+1}) \right]$$

So, now I can apply the result for the union of A and B2 events. So, this becomes probability of union AII is equal to 1 to K plus probability of AK plus 1 minus probability of union AII is equal to 1 to K intersection A K plus 1. Now, the first part of this can be expanded because we have already assumed that this rule is true for n is equal to K. So, this becomes sigma probability of AI Iis equal to 1 to K minus double summation probability of AIntersectionAJI less than Jnow this sums are up to n triple summation probability of A I intersection a j intersection A Kless than Jless than K the sums are up to n and. So, on plus up to minus 1 to the power K plus 1 probability ofintersection A II is equal to 1 to k.

Now, then we have probability of a k plus 1 and here we apply the distributive property of the unions and intersections. So, this becomes minus probability of union A I intersection A K plus 1 I is equal to 1 to k now if you look at this last term it is again union of K events and since we have assumed the probability of union result to be true for n is equal to k we can apply that formula.

So, using that we will get summation of probability AI intersection AK plus 1 for I is equal to 1 to k and that term can be adjusted with this. So, let me write it here. Firstly, sigma probability of A II is equal to 1 to k minus double summation I less than J upto n probability of AI intersection AJ plus triple summation I less than J less than K probability of AI intersection A J intersection A K minus and. So, on plus minus 1 to the power k plus 1 probability of intersection A I Iis equal to 1 to K.



Now, this probability of AK plus 1 can be added to the first term. So, the first term becomes probability of A II is equal to 1 to k plus 1 now let me expand the last union by using the formula for N is equal to K. So, this becomes sigma probability of AI intersection AK plus 1 I is equal to 1 to K minus.

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$$\begin{aligned}
 & \sum_{i < j}^k \sum_{i < j}^k P((A_i \cap A_{k+1}) \cap (A_j \cap A_{k+1})) \\
 & + \sum_{i < j < k}^k \sum_{i < j < k}^k P((A_i \cap A_{k+1}) \cap (A_j \cap A_{k+1}) \cap (A_k \cap A_{k+1})) \\
 & \dots + (-1)^{k+1} P\left(\bigcap_{i=1}^k (A_i \cap A_{k+1})\right) \\
 & = \sum_{i=1}^{k+1} P(A_i) - \sum_{i < j}^{k+1} P(A_i \cap A_j) \\
 & + \sum_{i < j < k}^{k+1} \sum_{i < j < k}^{k+1} P(A_i \cap A_j \cap A_k) - \dots + (-1)^{k+2} P\left(\bigcap_{i=1}^{k+1} A_i\right)
 \end{aligned}$$

Hence the result is true for all positive integral values of  $n$ .

Now, you will have double summation probability of AI intersection AK plus 1 intersection with AJA k plus 1 where I is less than J and the sum goes up to n only goes up to K. So, I think I have made some small mistakes here.

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$$\begin{aligned}
 P\left(\bigcup_{i=1}^{k+1} A_i\right) &= P\left(\bigcup_{i=1}^k A_i \cup A_{k+1}\right) \\
 &= P\left(\bigcup_{i=1}^k A_i\right) + P(A_{k+1}) - P\left(\left(\bigcup_{i=1}^k A_i\right) \cap A_{k+1}\right) \\
 &= \sum_{i=1}^k P(A_i) - \sum_{i < j}^k \sum_{i < j}^k P(A_i \cap A_j) + \sum_{i < j < k}^k \sum_{i < j < k}^k P(A_i \cap A_j \cap A_k) \\
 &\quad - \dots + (-1)^{k+1} P\left(\bigcap_{i=1}^k A_i\right) + P(A_{k+1}) \\
 &\quad - P\left(\bigcup_{i=1}^k (A_i \cap A_{k+1})\right) \\
 &= \sum_{i=1}^{k+1} P(A_i) - \sum_{i < j}^k \sum_{i < j}^k P(A_i \cap A_j) + \sum_{i < j < k}^k \sum_{i < j < k}^k P(A_i \cap A_j \cap A_k) \\
 &\quad - \dots + (-1)^{k+1} P\left(\bigcap_{i=1}^k A_i\right) - \left[\sum_{i=1}^k P(A_i \cap A_{k+1})\right]
 \end{aligned}$$

These sums are up to  $K$ , then you will have triple summation  $I < J < K$  probability of  $A_I \cap A_K + 1$ . So, we may put it  $R \cap A_J \cap A_{K+1} + 1$  and so on minus 1 to the power  $K + 1$  probability of intersection  $A_I \cap A_K + 1$ .

Now, let us look at the terms. The term  $\sum$  probability of  $A_I \cap A_K + 1$  can be combined with this term here where the minus gets adjusted and therefore, if you see here now, we already had all the intersections up to  $K$ . Now, we have  $A_1 \cap A_K + 1$ ,  $A_2 \cap A_K + 1$  and  $A_K \cap A_K + 1$ . So, this gets adjusted here and you will get a term. So, the first term remains as such. Probability of  $A_I$  is equal to 1 to  $k + 1$  in the second term you will get  $I < J$  and now this summation is up to  $K + 1$  probability of  $A_I \cap A_J$ .

Now, let us look at this term this term is  $A_J \cap A_K + 1$  where  $I$  and  $J$ s are varying from 1 to  $k$  and if we look at the third term in the previous expression here we had all the intersections of 3 sets up to  $k$ . So, this term gets adjusted in this 1 and you will get plus triple summation probability of  $A_I \cap A_J \cap A_{K+1}$  less than  $J < r$  up to  $k + 1$ .

In a similar way, if I look at this term here it will be intersection of the 4 terms and the last set is  $A_{k+1}$ ; that means, it is taking care of all the terms of the intersection taken 4 sets at a time. So, in this way all of the terms are combined if you look at this term this is actually intersection of all of the  $A_I$ s from  $I$  is equal to 1 to  $k + 1$  and since there is a minus sign outside of the square bracket this becomes minus 1 to the power  $k + 2$ . So, you will get minus 1 to the power  $k + 2$  probability of intersection  $A_I$  is equal to 1 to  $K + 1$  hence the result is true for all positive integral values of  $n$ .

Let us look at the some applications of this one. Now, before giving the application let me also consider the limit of the probabilities or probability of the limit as I mentioned that we have defined monotonic sequences and for the monotonic sequences of the sets, the limit always exists. So, we have the following result for monotonic sequences of the sets.



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$$\begin{aligned}
 P\left(\bigcup_{i=1}^{k+1} A_i\right) &= P\left(\bigcup_{i=1}^k A_i \cup A_{k+1}\right) \\
 &= P\left(\bigcup_{i=1}^k A_i\right) + P(A_{k+1}) - P\left(\left(\bigcup_{i=1}^k A_i\right) \cap A_{k+1}\right) \\
 &= \sum_{i=1}^k P(A_i) - \sum_{i < j}^k P(A_i \cap A_j) + \sum_{i < j < k}^k P(A_i \cap A_j \cap A_k) \\
 &\quad - \dots + (-1)^{k+1} P\left(\bigcap_{i=1}^k A_i\right) + P(A_{k+1}) \\
 &\quad - P\left(\bigcup_{i=1}^k (A_i \cap A_{k+1})\right) \\
 &= \sum_{i=1}^{k+1} P(A_i) - \sum_{i < j}^k P(A_i \cap A_j) + \sum_{i < j < k}^k P(A_i \cap A_j \cap A_k) \\
 &\quad - \dots + (-1)^{k+1} P\left(\bigcap_{i=1}^k A_i\right) - \left[ \sum_{i=1}^k P(A_i \cap A_{k+1}) \right]
 \end{aligned}$$

We have the following theorem if  $\{A_n\}$  is a monotonic sequence of sets in  $B$  then, probability of limit of  $A_n$  is equal to limit of probability of  $A_n$ .

To prove this result, let us consider  $A_n$  to be a monotonically increasing sequence. Let  $\{A_n\}$  be a monotonically increasing sequence if that is. So, then limit of the sequence  $A_n$  will become union of  $A_n$ ,  $n$  is equal to 1 to infinity. In order to prove that we have to look at probability of limit - means probability of the union. Now, what we do? We decompose this union by defining a new sequence of sets by saying say  $B_1$  is equal to  $A_1$ ;  $B_2$  is equal to  $A_2$  minus  $A_1$ ;  $B_n$  is equal to  $A_n$  minus  $A_{n-1}$ , for  $n$  greater than or equal to 2. If we look at this one basically, what we have done? This sequence of sets is like this. This is  $A_1$ ; this is  $A_2$ ; this is  $A_3$  and so on. So, if I look at the union of  $A_i$ s, we are decomposing it as a disjoint union; this  $A_1$ ,  $A_2$  minus  $A_1$  will be this portion then  $A_3$  minus  $A_2$  will be this portion. So, we will have that  $B_n$  is a disjoint sequence of sets and  $A_n$  is equal to union of  $B_i$  from 1 to  $n$ . Naturally, this implies that probability of  $A_n$  is equal to probability of  $B_i$  sigma  $i$  is equal to 1 to  $n$ .

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$$\begin{aligned} \lim_{n \rightarrow \infty} A_n &= \lim_{n \rightarrow \infty} \bigcup_{i=1}^n B_i = \bigcup_{n=1}^{\infty} B_n \\ P\left(\lim_{n \rightarrow \infty} A_n\right) &= P\left(\bigcup_{n=1}^{\infty} B_n\right) = \sum_{n=1}^{\infty} P(B_n) \\ &= \lim_{n \rightarrow \infty} \sum_{i=1}^n P(B_i) \\ &= \lim_{n \rightarrow \infty} P\left(\bigcup_{i=1}^n B_i\right) \\ &= \lim_{n \rightarrow \infty} P(A_n) \end{aligned}$$

The case of monotonic decreasing sequences can be proved in a similar way.

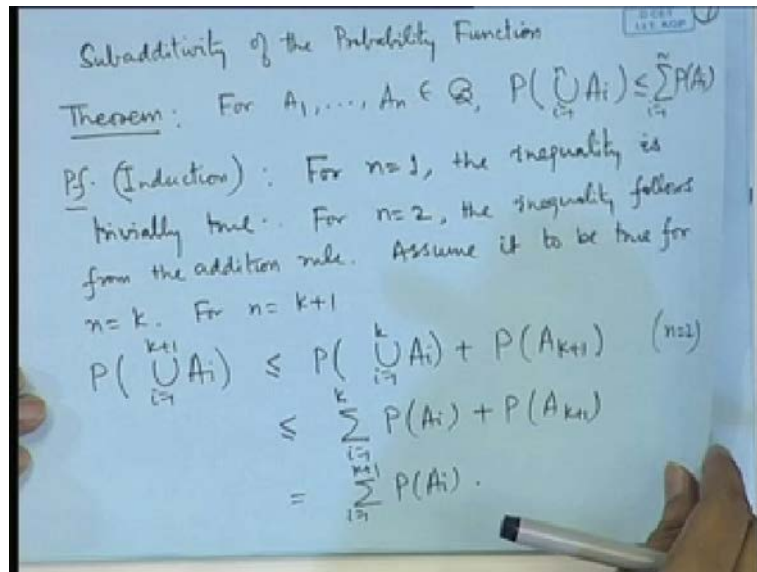
Now, if we look at limit of the sequence  $A_n$  as  $n$  tends to infinity then, it is equal to limit of union  $B_i$  is equal to  $1$  to  $n$ ,  $n$  tending to infinity which is equal to union of  $B_n$ ;  $n$  is equal to  $1$  to infinity because union  $B_i$  is a monotonic increasing sequence and the limit will be the ultimate union of these sets. So, if I look at probability of limit of  $A_n$  as  $n$  tends to infinity then it is equal to probability of union  $B_n$   $n$  is equal to  $1$  to infinity. Now,  $B_n$  is a disjoint sequence of sets; then by the axiom of the countable additivity this becomes probability of sigma probability of  $B_n$   $n$  is equal to  $1$  to infinity.

Now, this we can write as limit as  $n$  tends to infinity sigma  $i$  is equal to  $1$  to  $n$  probability of  $B_i$  which we can write as probability of union of  $B_i$  is equal to  $1$  to  $n$  which is equal to limit as  $n$  tends to infinity probability of  $A_n$ . So, thus we have proved that probability of a limit of a sequence of monotonic sequence of sets is equal to limit of the probability of the sequence of the sets. We may also consider the case of monotonically decreasing; now, that can be obtained by taking the complementations here or you can define in a reverse way.

The probability function we have assuming to be countably additive, but countably additive axiom implies that if we have a disjoint sequence then the probability of union is equal to the sum of the probabilities what if we do not have disjoint sequence for example, if I have 2 sets say  $A$  and  $B$  then we have probability of  $A$  union  $B$  is equal to probability  $A$  plus probability  $B$  minus probability of  $A$  intersection  $B$ . So, that means if I remove probability of  $A$  intersection  $B$  from there then we get probability of  $A$  union  $B$

less than or equal to probability A plus probability of B this is called subadditivity.

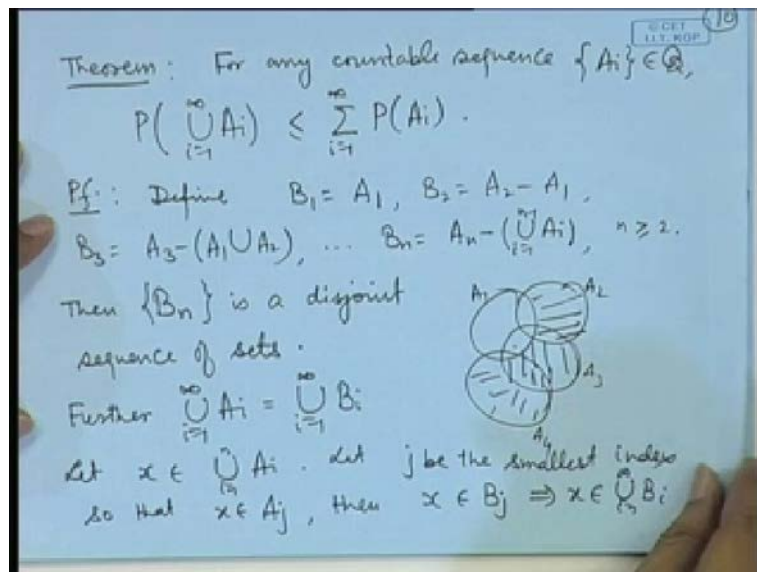
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So, if in general if we consider any sequence of sets then the probability of union will be less than or equal to the sum of the probabilities. So, we have subadditivity of the probability function and we can state it in the form of a theorem for  $A_1, A_2, \dots, A_n$  belonging to  $\mathcal{B}$  probability of union  $A_i$  is equal to 1 to  $n$  is less than or equal to sigma probability of  $A_i$  is equal to 1 to  $n$ . So, I can prove this by induction because for  $n$  is equal to 1 the result is true and if we look at for  $n$  is equal to 2 it is already shown to be true. So, for  $n$  is equal to 1 the inequality is trivially true for  $n$  is equal to 2 which we will require for extension from  $K$  to  $K$  plus 1 case. So, for  $n$  is equal to 2 the inequality follows from the addition rule.

So, assume it to be true for say  $n$  is equal to  $K$  now for  $n$  is equal to  $k$  plus 1 we can write probability of union  $A_i$  is equal to 1 to  $k$  plus 1 as less than or equal to probability of union  $A_i$  is equal to 1 to  $k$  plus probability of  $A_{k+1}$ , by using the result for  $n$  is equal to 2. So, now, on this we can make the use of assumption that upto  $n$  is equal to  $K$  - it is true. So, it becomes less than or equal to probability of  $A_i$  is equal to 1 to  $K$  plus probability of  $A_{k+1}$  which is nothing, but the sum of the probabilities  $i$  is equal to 1 to  $K$  plus 1 therefore, by induction the result is true for all of them.

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Now, if we want to prove the result for a countable number of these then, we can consider the decomposition. So, if we have for any countable sequence say  $A_i$  in probability of union  $A_i$  is equal to 1 to infinity is less than or equal to sigma probability of  $A_i$  is equal to 1 to infinity in order to prove this one. I may consider the decomposition of union  $A_i$  into a disjoint decomposition in the following way. Let us define say  $B_1$  is equal to  $A_1$ ;  $B_2$  is equal to  $A_2$  minus  $A_1$ ;  $B_3$  is equal to  $A_3$  minus a union  $A_2$  and so on. In general  $B_n$  is equal to  $A_n$  minus union of  $A_i$  is equal to 1 to  $n$  minus 1.

If we consider a Venn diagram then, it will be clear that what sets we are defining. Suppose, these sets are  $A_1, A_2, A_3, A_4$ , etcetera, then  $A_1$  and then  $A_2$  minus  $A_1$  is this set; then  $A_3$  minus  $A_1$  union  $A_2$  becomes this set  $A_4$  minus  $A_1$  union  $A_2$  union  $A_3$  becomes this set. So, naturally you can see here that we are considering the union as a disjoint union. So, then  $B_n$  is a disjoint sequence of sets further union of  $A_i$  is equal to 1 to infinity is equal to union of  $B_i$  is equal to 1 to infinity.

To prove this let us observe that union of  $B_i$  is already a subset of union  $A_i$  because each of the  $B_i$  is a subset of the corresponding  $A_i$ . Now, any point of  $A_i$  let us consider say  $x$  belonging to union of  $A_i$  let  $j$  be the smallest index. So, that  $x$  belongs to  $A_j$  then  $x$  will belong to  $B_j$  consequently  $x$  will belong to union of  $B_i$  as a result since union of  $B_i$  is already a subset of union of  $A_i$  we are now getting union of  $A_i$  is a subset of union of  $A_i$ .

(Refer Slide Time: 35:16)

So  $\cup A_i \subset \cup B_i$   
 $\Rightarrow \cup A_i = \cup B_i$   
 $P(\cup_{i=1}^{\infty} A_i) = P(\cup_{i=1}^{\infty} B_i) \leq \sum_{i=1}^{\infty} P(B_i)$   
 $\leq \sum_{i=1}^{\infty} P(A_i)$

Bonferroni Inequality : For any events  
 $A_1, \dots, A_n \in \mathcal{G}$ .

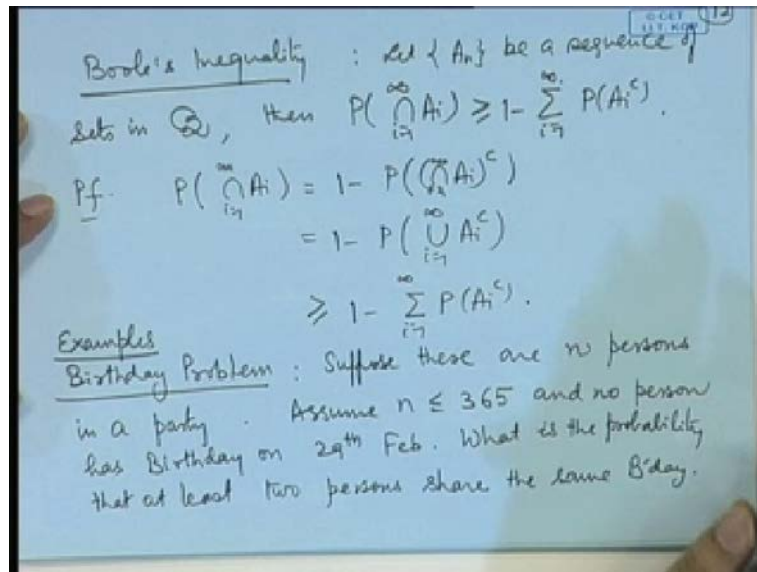
$$\sum_{i=1}^n P(A_i) - \sum_{i < j} P(A_i \cap A_j) \leq P(\cup_{i=1}^n A_i) \leq \sum_{i=1}^n P(A_i)$$

Pf. (Induction)

Therefore, we must have union of  $A_i$  is equal to union of  $B_i$ s. So, now if we consider probability of union of  $A_i$  is equal to 1 to infinity it is probability of union of  $B_i$  is equal to 1 to infinity which is less than or equal to, which is actually equal to sum of the probability of  $B_i$ s because  $B_i$ s are now disjoint and we can use the axiom of countable additivity. Now, each  $B_i$  is a subset of  $A_i$ . Therefore, probability of each  $B_i$  is less than or equal to probability of  $A_i$ . Therefore, this becomes less than or equal to sigma probability of  $A_i$  is equal to 1 to infinity; this proves the countable additivity of the probability function.

We also have something called Bonferroni inequalities which basically give that the probability of the unions are bounded between 2 bounds. So, for any events  $A_1, A_2, \dots, A_n$  in  $\mathcal{B}$  probability of the union which is already less than or equal to sum of the probabilities it is however, greater than or equal to probability of minus. Let me not prove it here. The proof will be by induction we can see the right hand side has already been proved. To prove the left hand side, if we take  $n$  is equal to 1 then it is trivially true for  $n$  is equal to 2 there is equality by the addition rule. So, assuming for  $n$  is equal to  $K$  if we write for  $n$  is equal to  $K + 1$  then we can split it into 2 terms that is union of  $A_i$  is equal to 1 to  $K$  union  $A_{K+1}$ . On that we apply the addition rule and then apply the assumption for  $K$  that will prove the general Bonferroni inequality.

(Refer Slide Time: 38:01)



In a similar way we have what is known as Boole's inequality. The Boole's inequality gives a relation between the intersection. Likewise for example, if I have a sequence of sets in  $\mathcal{B}$  then probability of intersection  $\bigcap_{i=1}^{\infty} A_i$  is equal to 1 to infinity is greater than or equal to 1 minus sigma probability of  $A_i^c$  for  $i=1$  to infinity.

To prove this we simply use the subadditivity because we can write probability of intersection  $\bigcap_{i=1}^{\infty} A_i$  as 1 minus probability of intersection  $\bigcap_{i=1}^{\infty} A_i^c$ . Now, this can be written as 1 minus probability of union  $\bigcup_{i=1}^{\infty} A_i^c$  by using De Morgan's laws. At this stage you can use the countable subadditivity. So, this will become greater than or equal to 1 minus sigma probability of  $A_i^c$ .

Let me give some examples of applications of basic rules of probability. Let me start from a birthday problem. Suppose there are  $n$  persons in a party assuming that the number of persons is less than or equal to 365 and no person has birthday on 29 February. What is the probability that at least 2 persons share the same birthday? Now, in order to analyze this problem, let us consider the set theoretic description.



(Refer Slide Time: 41:22)

Let  $A \rightarrow$  at least two persons share the same B day.

$A^c \rightarrow$  no two persons share the same B day.

$$P(A^c) = \frac{{}^{365}P_n}{(365)^n} = \frac{365 \cdot 364 \cdot \dots \cdot (365 - n + 1)}{(365)^n}$$
$$= 1 \cdot \left(1 - \frac{1}{365}\right) \left(1 - \frac{2}{365}\right) \dots \left(1 - \frac{n-1}{365}\right).$$

So  $P(A) = 1 - \left(1 - \frac{1}{365}\right) \left(1 - \frac{2}{365}\right) \dots \left(1 - \frac{n-1}{365}\right).$

Let us consider  $A$  to be the event that at least 2 persons share the same birthday. Then, if you look at this event it is slightly complicated event in the sense that, 2 persons may share 3 persons may share and so on and finding out the probabilities of each of them may be a little bit complicated because, if we say 2 persons share then which of the day it is and all others must be on some other dates and they should not be the same. Suppose, we say 3 persons share then which one of the 365 days and all other persons must be on distinct days which distinct days. So, this is a complicated way to analyze.

However, if we use these theoretic representations, we can look at the complementary event. A complement - this means no 2 persons have the same birthday. Now, this becomes somewhat simpler because, if we look at the probability of  $A$  complement assuming all the birth dates to be equally likely this number will be simply  ${}^{365}P_n$  divided by  $365$  to the power  $n$ . Here, the denominator denotes the total number of possibilities for  $n$  persons to have birthdays because, each person can have any of the 365 days as a possible birthday and therefore,  $n$  persons can have possible number of birthdays as  $365$  to the power  $n$ .

If we make the assumption that  $n-1$  of them have the same birthday then it becomes a problem of choosing  $n$  numbers out of 365 which are distinct. So, it is nothing but the number of permutations taking  $n$  at a time from 365 that is equal to  $365$  into  $364$  upto  $365$  minus  $n$  plus 1 divided by  $365$  to the power  $n$ , which we may write as a way of representation as  $1; 1 - \frac{1}{365}; 1 - \frac{2}{365}$  and so on upto  $1 - \frac{n-1}{365}$

1 by 365 and so on.. So, probability of A becomes 1 minus the product given by these terms. An interesting thing; what we do? Look at that how many people are required; so, that at least 2 will share a birthday. If we think from a layman point of view then, we may think that the numbers since the number of possible birthdays is 365 to the power n. So, n should be somewhat large in order that this probability is significant. So, let us look at the table of probabilities.

(Refer Slide Time: 45:03)

$n$	$P(A^c)$	$P(A)$
10	0.871	0.129
20	0.589	0.411
23	0.493	0.507
30	0.294	0.706
50	0.030	0.970
60	0.006	0.994

2. Suppose a die is tossed three times independently and the outcomes are recorded as numbers  $a, b, c$ . What is the probability that the roots of equation  $ax^2 + bx + c = 0$  are real?  
 $a, b, c \rightarrow 1, 2, \dots, 6$ .

let us consider sayn probability of A complement and probability of A. So, A simple calculation table can be prepared if I have n is equal to 10; the probability of A complement is point 871 and consequently, probability of A becomes point 129. Ifwe take n is equal to 20, probability of A complement is point 589 and probability of A becomes point 411. Ifwe take n to be 23 then, probability of a complement is point 493 and probability of A becomes point 507. That means, with as less as only 23 persons the probability that at least 2 share a common birthday is more than 50 percent. So, it is from a layman's thinking; this is counter intuitive.

We need very few persons to at least 2 of them to share a common birthday. If I take n is equal to 30 then, this probability becomes point 706. Ifwe take 50, the probability is point 97 and for n is equal to 60 the probability is point 994; it is nearly 1. That means, in a set of 60 people, the probability is nearly 1 that at least 2 of them will share a common birthday.

So, here you can see that the elementary rules of probability have been used for calculation. For example, we have used the property of the complementation to evaluate the actual probability we have used the method of classical probability by assuming all the birth dates to be equally likely for all the persons.

Let us look at some other applications of the basic rules of the probability. Suppose, a die is tossed 3 times independently and the outcomes are recorded as numbers A, B, and C. What is the probability that the roots of equation  $AX^2 + BX + C = 0$  are real.

Now, if we want to calculate this probability here the outcomes A, B, and C are random each of the values of A, B, and C can be numbers 1 to 6. Therefore, the quadratic equation  $AX^2 + BX + C = 0$  will have the real roots if  $B^2 - 4AC$  is positive. So, we have to look at the number of cases where  $b^2 - 4AC$  is greater than or equal to 0.

(Refer Slide Time: 49:35)

b	b <sup>2</sup>	a, c	no. of cases
1	1	--	0
2	4	(1, 1)	1
3	9	(1, 1), (1, 2), (2, 2)	3
4	16	(1, 1), (1, 2), (2, 1), (2, 2), (1, 4), (4, 1), (1, 3), (3, 1)	8
5	25	8 + (1, 5), (5, 1), (1, 6), (6, 1), (3, 3), (3, 4)	14
6	36	14 + (2, 4), (4, 2), (3, 3)	17
			43

A → the roots are real

$$P(A) = \frac{43}{6^3} = \frac{43}{216}$$

So, this has to be done through an enumeration and we can prepare the table that what are the possibilities of B and therefore, the corresponding values of B square. What are the possible values of A and C which lead to  $4AC$  being less than or equal to B square? So, let us take say B is equal to 1 then B square is equal to 1; that means, there is no case which will give me  $4AC$  to be less than or equal to B square. So, there is no possibility here.

So, if we look at the number of cases this is 0 if we take  $b$  is equal to 2 then  $b^2$  is equal to 4 and if I consider  $a$  and  $c$  to be 1 then  $4ac$  will become 4. So, there is 1 case which will give me  $B^2$  greater than or equal to  $4ac$ . If we consider  $B$  is equal to 3 then  $B^2$  is equal to 9 now 1 1 2 and 2 1. There are 3 cases which will give me  $B^2$  greater than or equal to  $4ac$ . If we have  $B$  is equal to 4 then  $B^2$  is equal to 16, we will have the cases 1 1 2 2 1 2 2 which will correspond to 4. So, 14 4 1 13 3 1 1 23 4 5 6 7 8 cases are there which will give me  $B^2$  greater than or equal to  $4ac$ .

If we have  $B$  is equal to 5 then  $B^2$  is equal to 25 then all the above cases that is 8 cases plus we will also have 1 5 1 and possibly 1 4 4 1. So, we will also have 16 6 1 2 3 3 2, basically 1 3 3 4 5 6 more cases. So, 14 cases are there. If I have  $B$  is equal to 6 then,  $B^2$  becomes 36 and all the 14 cases plus we will also have 2 4 2 then 25 is not possible. I think 33 must have come here itself because, no it will not come here 33 will come here because this will give me 9. So, there are 17 cases. So, if we look at the total number of cases it is 39, 42, 43 cases are there. Total number of cases is 43 and the total number of possibilities if I define  $a$  to be the event that the roots are real then, the probability of that will be given by the favorable number of cases divided by the total number of cases which is 6 cube here because, 3 dice - each of them have 6 possibilities. So, the total number of possibilities are 6 cube that is 43 by 216.

Likewise, in this problem we may also find out the probability of the quadratic equation to have complex roots or the real roots, but equal, etcetera. We may consider all types of possibilities. So, I will end today's lecture by this; thank you.