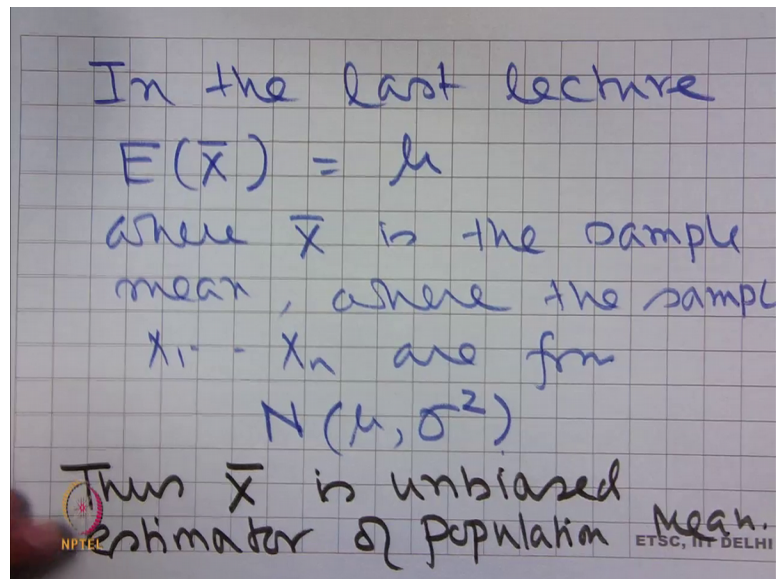


Statistical Inference
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Lecture – 08
Statistical Inference

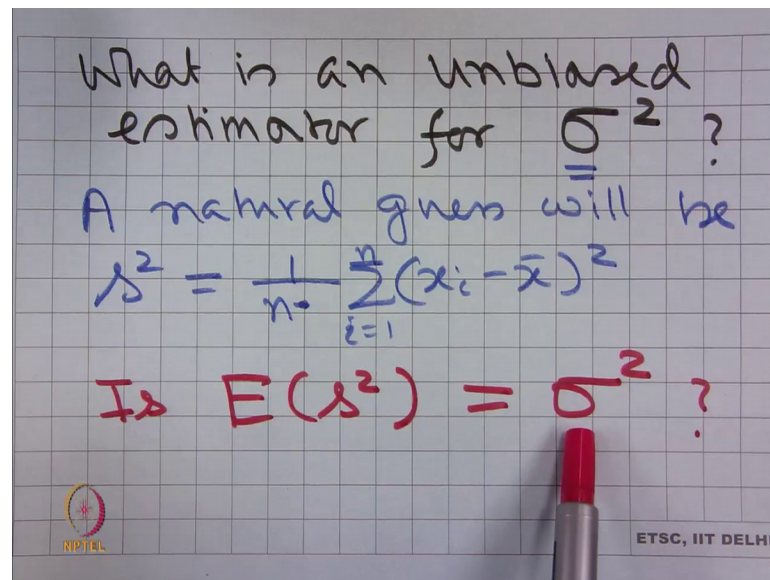
Welcome students to lecture number 8 on the MOOC's course on Statistical Inference.

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In the last lecture, we found that expectation of \bar{X} is equal to μ where \bar{X} is the sample mean, where the sample X_1 up to X_n are from normal μ sigma square. Thus, sample mean is unbiased estimator of population mean.

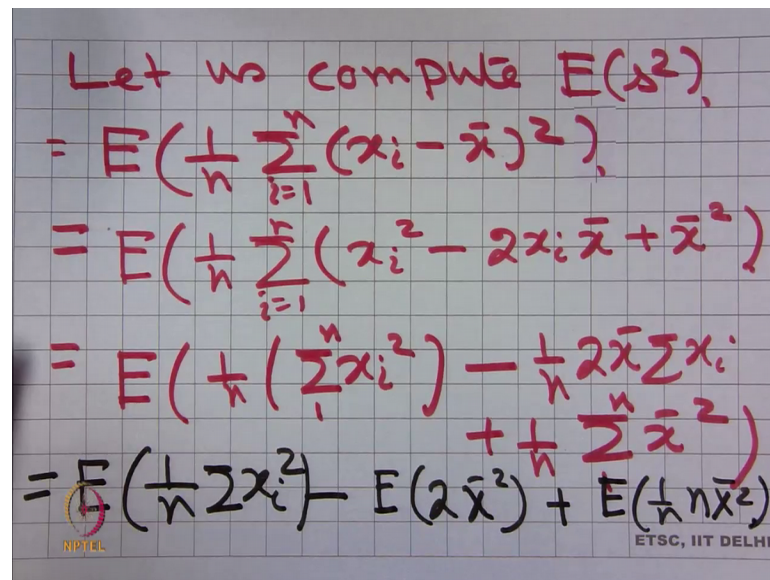
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This we have seen; obviously, the next question is what is an unbiased estimator for sigma square?

That is, we have taken sample x_1, x_2, \dots, x_n and based on that we want to find an unbiased estimate for the population variance sigma square. A natural guess will be sample variance. So, x_1, x_2, \dots, x_n are my sample, \bar{x} is the sample mean, therefore, $\sum (x_i - \bar{x})^2$ upon n that is the sample variance. The question is $E(s^2)$ is the expectation of sample variance is equal to sigma square or in other words whether the sample variance is an unbiased estimator for population variance.

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Let us compute $E(s^2)$,
 $= E\left(\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2\right)$,
 $= E\left(\frac{1}{n} \sum_{i=1}^n (x_i^2 - 2x_i\bar{x} + \bar{x}^2)\right)$
 $= E\left(\frac{1}{n} \left(\sum_{i=1}^n x_i^2\right) - \frac{1}{n} 2\bar{x} \sum_{i=1}^n x_i + \frac{1}{n} \sum_{i=1}^n \bar{x}^2\right)$
 $= E\left(\frac{1}{n} \sum_{i=1}^n x_i^2\right) - E\left(2\bar{x}^2\right) + E\left(\frac{1}{n} n \bar{x}^2\right)$

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In order to find the answer, we have to compute the expected value of small square. The expected value of small square is equal to expected value of $\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$; is equal to expected value of $\frac{1}{n} \sum_{i=1}^n x_i^2 - 2\bar{x} \sum_{i=1}^n x_i + \sum_{i=1}^n \bar{x}^2$; is equal to expected value of $\frac{1}{n} \sum_{i=1}^n x_i^2 - 2\bar{x} \sum_{i=1}^n x_i + \sum_{i=1}^n \bar{x}^2$, is equal to expected value of $\frac{1}{n} \sum_{i=1}^n x_i^2 - 2\bar{x} \sum_{i=1}^n x_i + \sum_{i=1}^n \bar{x}^2$. This is equal to therefore, expected value of this whole thing.

So, this is equal to expected value of $\frac{1}{n} \sum_{i=1}^n x_i^2$ minus expected value of $\frac{1}{n} \sum_{i=1}^n 2\bar{x} x_i$ plus expected value of $\frac{1}{n} \sum_{i=1}^n \bar{x}^2$. Therefore, it is $2\bar{x}^2$ plus $\frac{1}{n}$ plus expected value of $\frac{1}{n} \sum_{i=1}^n x_i^2$.

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$$\begin{aligned} &= E\left(\frac{1}{n} \sum x_i^2\right) - 2E(\bar{x}^2) + E(\bar{x}^2) \\ &= E\left(\frac{1}{n} \sum x_i^2\right) - E(\bar{x}^2). \end{aligned}$$

We know that

$$\begin{aligned} \sigma^2 = V(x_i) &= E(x_i - E(x_i))^2 \\ &= E(x_i^2) - (E(x_i))^2 \\ &= E(x_i^2) - \mu^2 \\ \therefore E(x_i^2) &= \sigma^2 + \mu^2 \end{aligned}$$

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Is equal to expected value of $\frac{1}{n} \sum x_i^2$ minus 2 times expected value of \bar{x}^2 plus expected value of \bar{x}^2 ; is equal to expected value of $\frac{1}{n} \sum x_i^2$ minus expected value of \bar{x}^2 .

Now, we know that variance of x_i is equal to expected value of x_i minus expected value of x_i whole square. So, variance of each x_i is equal to expected value of x_i square minus expected value of x_i whole square, is equal to expected value of x_i square minus μ squared. And variance of x_i is equal to σ^2 . Therefore, the expected value of x_i square is equal to $\sigma^2 + \mu^2$.

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By a similar logic

$$V(\bar{x}) = E(\bar{x}^2) - (E(\bar{x}))^2$$
$$\therefore \frac{\sigma^2}{n} = E(\bar{x}^2) - \mu^2$$
$$\therefore E(\bar{x}^2) = \frac{\sigma^2}{n} + \mu^2$$

① $E(\bar{x}^2) = \frac{\sigma^2}{n} + \mu^2$

② $E(\bar{x}^2) = \frac{\sigma^2}{n} + \mu^2$

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By a similar logic variance of \bar{x} is equal to expected value of \bar{x} square minus expected value of \bar{x} whole square. And we know that variance of \bar{x} is equal to sigma square by n . Therefore, sigma square by n is equal to expected value of \bar{x} square minus mu square. Therefore, expected value of \bar{x} square is equal to sigma square by n plus mu square. So, we find that one expected value of \bar{x} square is equal to sigma square plus mu square. 2 the expected value of \bar{x} square is equal to sigma square by n plus mu square.

Now, we have already found that expectation of small square is equal to this.

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$$\begin{aligned}\therefore E(s^2) &= E\left(\frac{1}{n} \sum_{i=1}^n x_i^2\right) - E(\bar{x}^2) \\ &= \frac{1}{n} \sum_{i=1}^n (\sigma^2 + \mu^2) - \left(\frac{\sigma^2}{n} + \mu^2\right) \\ &= \sigma^2 + \mu^2 - \frac{\sigma^2}{n} - \mu^2 \\ &= \sigma^2 \left(1 - \frac{1}{n}\right) = \sigma^2 \left(\frac{n-1}{n}\right) \\ \therefore s^2 &\text{ is NOT unbiased for } \sigma^2\end{aligned}$$

Therefore by putting these values $\frac{1}{n}$ by $\frac{n-1}{n}$ summation over 1 to n sigma square plus mu square minus sigma square by n plus mu square, is equal to sigma square plus mu square minus sigma square by n minus mu square, is equal to sigma square into $1 - \frac{1}{n}$ upon n.

Thus we find that expectation of small square is equal to sigma square into $1 - \frac{1}{n}$ upon n is equal to sigma square into $\frac{n-1}{n}$ upon n. Therefore, small square is not unbiased for sigma square.

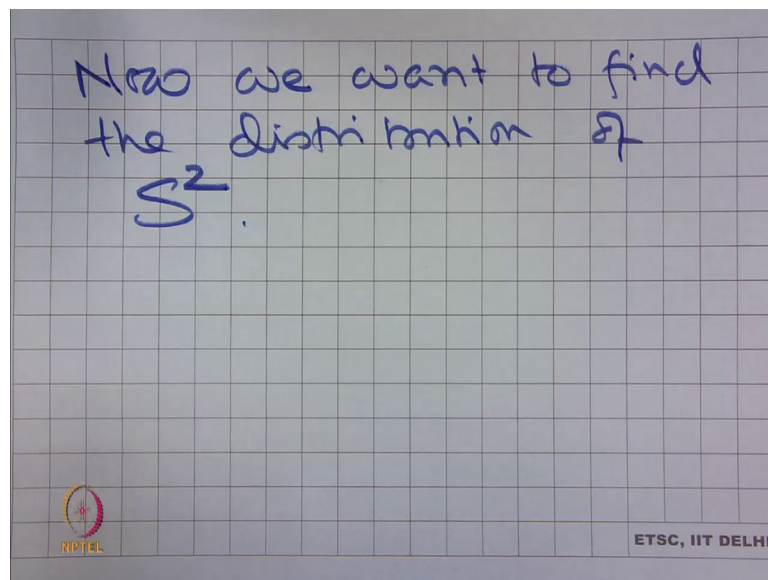
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$$\begin{aligned}\therefore \text{unbiased estimator for } \sigma^2 &= \frac{n}{n-1} \times s^2 \\ &= \frac{n}{n-1} \left(\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 \right) \\ &= \boxed{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}} \\ &= S^2\end{aligned}$$

Therefore, if I ask you what is an unbiased estimator for sigma square, small square is equal to $\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$, it is equal to $\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$ is equal to $\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$ divided by $n - 1$.

So, that shows that our logical guess that sample variance is going to be an unbiased estimator for sigma square is not correct, the unbiased estimator for sigma square is this quantity which is $\sum_{i=1}^n (x_i - \bar{x})^2$ upon $n - 1$ and we often denote it by s^2 .

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Now, we want to find the distribution of or we can say we want to find the sampling distribution of capital S square.

In order to do that we first need some mathematical tricks.

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We have seen that when we transform $(X, Y) \rightarrow (Z, W)$

$$\begin{aligned} Z &= G_1(X, Y) \\ W &= G_2(X, Y) \end{aligned}$$

We can get joint pdf of (Z, W)

$$f_{ZW}(z, w) = f_{XY}(x, y) |J|$$

expressed in z, w

$|J| = \begin{vmatrix} \frac{\partial x}{\partial z} & \frac{\partial x}{\partial w} \\ \frac{\partial y}{\partial z} & \frac{\partial y}{\partial w} \end{vmatrix}$

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We have seen that when we transform from the XY plane to the ZW plane such that Z is equal to G_1 of X Y, and W is equal to G_2 of X Y we can get joint pdf of z w as the joint distribution of XY expressed in zw multiplied by the Jacobian where the Jacobian is the determinant of $\frac{\partial x}{\partial z}$, $\frac{\partial x}{\partial w}$, $\frac{\partial y}{\partial z}$ and $\frac{\partial y}{\partial w}$.

Now, suppose we want to extend it to n dimension. The most important concept was that given z w we can use inverse transformation to get unique solution for x y.

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This we want to extend to n dimension:

$$\begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = \underline{G} \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix}$$

In particular let us consider G to linear transformation.

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Now, we want to extend to n dimension. Suppose we write y_1, y_2, \dots, y_n as a function of x_1, x_2, \dots, x_n . So, we have n random variables x_1, x_2, \dots, x_n , we want to make it transformation to new set of variables y_1, y_2, \dots, y_n using a transformation function g .

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$$\begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = A \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix}$$

where $A = \begin{bmatrix} a_{11} & a_{12} & a_{13} & \dots & a_{1n} \\ a_{21} & a_{22} & \dots & \dots & a_{2n} \end{bmatrix}$

We can get $(x_1, \dots, x_n)^T$ uniquely from $(y_1, \dots, y_n)^T$ if A is invertible.

$$\begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} = A^{-1} \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

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In particular, let us consider g to be linear transformation; that is, y_1, y_2, \dots, y_n is equal to A times x_1, x_2, \dots, x_n where A is equal to $a_{11}, a_{12}, a_{13}, a_{1n}, a_{21}, a_{22}, a_{2n}$.

We can get x_1, x_2, \dots, x_n uniquely from y_1, y_2, \dots, y_n . Let us use transpose notation for column vectors, if A is invertible. Then we get x_1, x_2, \dots, x_n is equal to A^{-1} times y_1, y_2, \dots, y_n . In this case what is going to be the Jacobian?

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What is the Jacobian?

Let us write $A^{-1} = B$

$$= \begin{bmatrix} b_{11} & b_{12} & \dots & b_{1n} \\ \vdots & \vdots & \ddots & \vdots \\ b_{n1} & b_{n2} & \dots & b_{nn} \end{bmatrix}$$

$$\therefore \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} = \begin{bmatrix} b_{11} & \dots & b_{1n} \\ \vdots & \ddots & \vdots \\ b_{n1} & \dots & b_{nn} \end{bmatrix} \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

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Let us write A inverse is equal to B, is equal to $b_{11}, b_{12}, b_{1n}, b_{n1}, b_{n2}, b_{nn}$.
Therefore, x_1 to x_n is equal to $b_{11}y_1 + b_{12}y_2 + \dots + b_{1n}y_n$ up to $b_{n1}y_1 + b_{n2}y_2 + \dots + b_{nn}y_n$.

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$$\therefore |J| = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} & \dots & \frac{\partial x_1}{\partial y_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial x_n}{\partial y_1} & \frac{\partial x_n}{\partial y_2} & \dots & \frac{\partial x_n}{\partial y_n} \end{vmatrix}$$

$$x_1 = b_{11}y_1 + b_{12}y_2 + \dots + b_{1n}y_n$$

$$\therefore \frac{\partial x_1}{\partial y_1} = b_{11} \quad \frac{\partial x_1}{\partial y_2} = b_{12} \quad \dots \quad \frac{\partial x_1}{\partial y_n} = b_{1n}$$

$$\therefore |J| = |B|$$

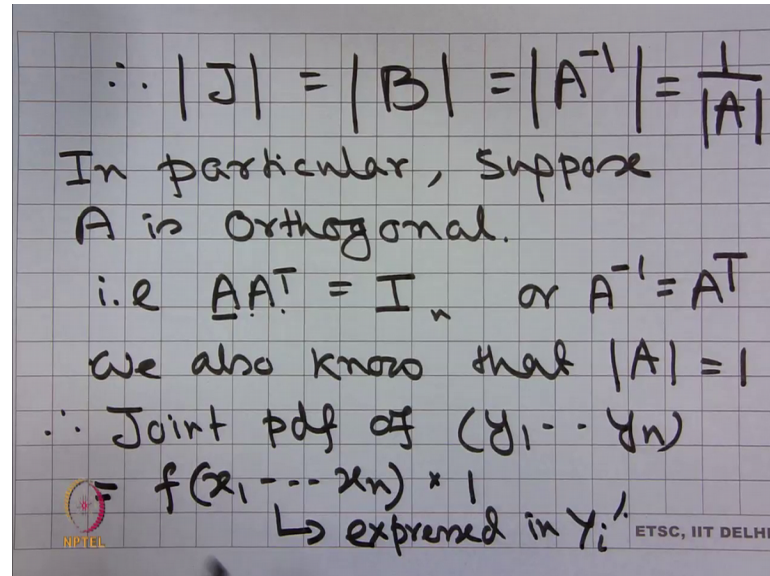
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Therefore, the Jacobian is equal to $\frac{\partial x_1}{\partial y_1}, \frac{\partial x_1}{\partial y_2}, \dots, \frac{\partial x_1}{\partial y_n}$ in the first row, and so on for the other rows.

Now, let us look at one of them. We have x_1 is equal to $b_{11}y_1 + b_{12}y_2 + \dots + b_{1n}y_n$. Therefore, $\frac{\partial x_1}{\partial y_1}$ is equal to b_{11} , $\frac{\partial x_1}{\partial y_2}$ is equal to b_{12} , $\frac{\partial x_1}{\partial y_n}$ is equal to b_{1n} . Or in other words, the first row of the Jacobian matrix is going to be the first row of the matrix B.

to be the first row of b . In a similar way, we can see that this Jacobean matrix J is nothing but the inverse matrix B .

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$$\therefore |J| = |B| = |A^{-1}| = \frac{1}{|A|}$$

In particular, suppose A is orthogonal.
i.e. $AA^T = I_n$ or $A^{-1} = A^T$
we also know that $|A| = 1$
 \therefore Joint pdf of (y_1, \dots, y_n)
 $= f(x_1, \dots, x_n) \times 1$
 \hookrightarrow expressed in y_i 's

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Therefore the Jacobean of the transformation is equal to determinant of B which is equal to determinant of A inverse, which is equal to 1 upon determinant of A .

So, this is a very interesting observation. Now in particular suppose A is orthogonal; that is, AA^T is equal to I . A is an n cross n orthogonal matrix therefore, AA^T is equal to identity matrix, or A inverse is equal to A^T . Therefore, therefore, if we consider that transformation matrix to be orthogonal, we get this and we also know that determinant of A is equal to 1 therefore, joint pdf of y_1, y_2, \dots, y_n is equal to joint pdf of x_1, x_2, \dots, x_n multiplied by 1 this is expressed in y_i 's right?

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Let us now consider x_1, \dots, x_n as independent samples from $N(\mu, \sigma^2)$

$$\therefore f(x_1, \dots, x_n) = \left(\frac{1}{\sqrt{2\pi}\sigma} \right)^n e^{-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}}$$

$$= \frac{1}{(\sqrt{2\pi})^n \sigma^n} e^{-\frac{1}{2} \sum_{i=1}^n \left(\frac{x_i - \mu}{\sigma} \right)^2}$$

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Let us now consider $x_1 \times 2 \times n$ as independent samples from normal μ σ^2 . Therefore, f of $x_1 \times 2 \times n$ which is the joint pdf of $x_1 \times 2 \times n$ is equal to 1 over root over 2π σ whole to the power n , e to the power minus σ^2 x_i minus μ whole square upon $2\sigma^2$, is equal to 1 over root over 2π to the power n σ to the power n e to the power minus half σ^2 x_i minus μ upon σ^2 whole square summation i is equal to 1 to n .

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Now, let us consider the following orthogonal transformation

$$Y_1 = \begin{bmatrix} \frac{1}{\sqrt{n}} & \frac{1}{\sqrt{n}} & \dots & \frac{1}{\sqrt{n}} \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix}$$

$$Y_2 = \begin{bmatrix} \frac{1}{\sqrt{2}} & -\frac{1}{\sqrt{2}} & 0 & \dots & 0 \\ \frac{1}{\sqrt{6}} & \frac{1}{\sqrt{6}} & -\frac{2}{\sqrt{6}} & 0 & 0 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix}$$

In general the i^{th} row of the transformation matrix is

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Now, let us consider the following orthogonal transformation. y_1 is equal to $1/\sqrt{n}$. So, y_1 is considered to be the dot product of $1/\sqrt{n}$ into $x_1 + 1/\sqrt{n}$ into $x_2 + 1/\sqrt{n}$ into x_n .

Let us consider y_2 to be the second row is $1/\sqrt{2}$ minus $1/\sqrt{2}$ 0 0. Therefore, first thing you note that the sum of squares of these values is equal to 1. Sum of squares of these values is equal to half plus half is equal to 1. Not only that if we look at the dot product of this vector with this we get 0. Therefore, these 2 vectors are mutually orthogonal.

Similarly, we can have the third row is equal to $1/\sqrt{6}$, $1/\sqrt{6}$ minus $2/\sqrt{6}$ rest are 0. Again if you look at it is sum of square is equal to $1/6 + 1/6 + 4/6$ is equal to 1. And not only that it is orthogonal to this and it is orthogonal to this.

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$$\frac{1}{\sqrt{i(i-1)}} \left[\underbrace{1 \ 1 \ \dots \ 1}_{(i-1) \text{ times}} \ - (i-1) \ 0 \ \dots \ 0 \right]$$

$$\therefore \underbrace{\frac{1}{i(i-1)} + \frac{1}{i(i-1)} + \dots + \frac{1}{i(i-1)}}_{(i-1) \text{ times}} + \frac{(i-1)^2}{i(i-1)}$$

$$= \frac{1}{i} + \frac{i-1}{i} = 1$$

Also its dot product with each previous row is zero.

In general, the i th row of the transformation matrix is $1/\sqrt{i}$ into i minus $1/\sqrt{i}$ into 1 1 1 minus 1 times minus $1/\sqrt{i}$ into 1 and rest are all 0's.

Therefore, the sum square of the element is $1/\sqrt{i}$ into i minus $1/\sqrt{i}$ plus $1/\sqrt{i}$ into i minus 1 plus $1/\sqrt{i}$ into i minus 1. This is i minus 1 times plus $1/\sqrt{i}$ squared upon i into i minus 1, is equal to $1/\sqrt{i}$ plus i minus $1/\sqrt{i}$ is equal to 1. Thus the norm the

length of each of the rows of the matrix is 1. You can easily verify that it is dot product with all the previous rules going to be 0.

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Thus the matrix is an orthogonal matrix

Let us use this matrix for our calculation

Let

$$\begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = A \begin{bmatrix} \frac{x_1 - \mu}{\sigma} \\ \vdots \\ \frac{x_n - \mu}{\sigma} \end{bmatrix}$$

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Thus the matrix is an orthogonal matrix. Let us use this matrix for our calculation. Let y_1, y_2, \dots, y_n be A into $x_1 - \mu, x_2 - \mu, \dots, x_n - \mu$ upon σ . So, this A is the particular A that I have just constructed.

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$$\therefore \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = \begin{bmatrix} \frac{a_{11}}{\sigma} & \frac{a_{12}}{\sigma} & \dots & \frac{a_{1n}}{\sigma} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{a_{n1}}{\sigma} & \frac{a_{n2}}{\sigma} & \dots & \frac{a_{nn}}{\sigma} \end{bmatrix} \begin{bmatrix} x_1 - \mu \\ x_2 - \mu \\ \vdots \\ x_n - \mu \end{bmatrix}$$

$$\therefore |J| = \left| \frac{1}{\sigma^n} \right| \times \begin{bmatrix} x_1 - \mu \\ x_2 - \mu \\ \vdots \\ x_n - \mu \end{bmatrix}$$

If we consider the inverse of this then $|J| = \sigma^n$

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Therefore, y_1, y_2, \dots, y_n I can write it as $a_{11} \frac{x_1 - \mu}{\sigma}, a_{12} \frac{x_2 - \mu}{\sigma}, \dots, a_{1n} \frac{x_n - \mu}{\sigma}$ up to $a_{n1} \frac{x_1 - \mu}{\sigma}, a_{n2} \frac{x_2 - \mu}{\sigma}, \dots, a_{nn} \frac{x_n - \mu}{\sigma}$, multiplied by x_1, x_2, \dots, x_n .

minus μ x 2 minus μ up to x_n minus μ . This a 1 is the original, this a ones are the matrix of the A transformation that we have considered.

Therefore, needs Jacobean is going to be 1 upon sigma to the power n, if I take the modulus if I take 1 upon sigma to be out from each of the n columns the determinant is equal to 1. Therefore, that mod of the Jacobean, if I consider this transformation that is going to be 1 upon sigma to the power n. Therefore, if we consider the inverse of it, the Jacobean is going to be sigma power n.

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$$\begin{aligned} \therefore \text{Joint pdf of } (y_1, \dots, y_n) &= \frac{1}{(\sqrt{2\pi})^n \sigma^n} e^{-\frac{1}{2} \sum \left(\frac{x_i - \mu}{\sigma}\right)^2} \times \sigma^n \\ &= \frac{1}{(\sqrt{2\pi})^n} e^{-\frac{1}{2} \sum_{i=1}^n \left(\frac{x_i - \mu}{\sigma}\right)^2} \end{aligned}$$

Now $\begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = A \begin{pmatrix} \frac{x_1 - \mu}{\sigma} \\ \vdots \\ \frac{x_n - \mu}{\sigma} \end{pmatrix}$

Therefore joint pdf of y_1, y_2, \dots, y_n is equal to 1 over $\sqrt{2\pi}$ sigma to the power n into e to the power minus half sigma x i minus mu upon sigma whole square multiplied by sigma power n, is equal to 1 over $\sqrt{2\pi}$ whole to the power n e to the power minus half and it is sigma x i minus mu upon sigma whole square, i is equal to 1 to n.

Now, y_1, y_2, \dots, y_n is equal to A times x_1 minus mu upon sigma up to x_n minus mu upon sigma.

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$$\begin{aligned}\therefore y^T y &= \sum_{i=1}^n y_i^2 \\ &= \left(\frac{x_1 - \mu}{\sigma}, \dots, \frac{x_n - \mu}{\sigma} \right) \frac{A^T A}{I} \begin{bmatrix} \frac{x_1 - \mu}{\sigma} \\ \vdots \\ \frac{x_n - \mu}{\sigma} \end{bmatrix} \\ &= \sum_{i=1}^n \left(\frac{x_i - \mu}{\sigma} \right)^2 \\ \therefore \text{Joint pdf of } (y_1, \dots, y_n) \\ &= \boxed{\frac{1}{(\sqrt{2\pi})^n} e^{-\frac{1}{2} \sum_{i=1}^n y_i^2}}\end{aligned}$$

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Therefore $y^T y$ is equal to $\sum_{i=1}^n y_i^2$, is equal to $\sum_{i=1}^n \left(\frac{x_i - \mu}{\sigma} \right)^2$, is equal to $\sum_{i=1}^n \frac{(x_i - \mu)^2}{\sigma^2}$, is equal to $\frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$, is equal to $\frac{1}{\sigma^2} \sum_{i=1}^n x_i^2 - \frac{2\mu}{\sigma^2} \sum_{i=1}^n x_i + \frac{n\mu^2}{\sigma^2}$, is equal to $\frac{1}{\sigma^2} \sum_{i=1}^n x_i^2 - \frac{2\mu}{\sigma^2} \sum_{i=1}^n x_i + \frac{n\mu^2}{\sigma^2}$. Therefore, joint pdf of y_1, y_2, \dots, y_n is equal to $\frac{1}{(\sqrt{2\pi})^n} e^{-\frac{1}{2} \sum_{i=1}^n y_i^2}$ to 1 to n. What does it tell us?

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$$\begin{aligned}\therefore (y_1, \dots, y_n) &\text{ are i.i.d. r.v.s} \\ \Rightarrow \text{each } y_i &\sim N(0, 1)\end{aligned}$$

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It tells that y_1, y_2, \dots, y_n are iid independent identically distributed random variables; such that each y_i follow normal $0, 1$. So, the advantage of the particular transformation that I have made is that, it converts from x_1, x_2, \dots, x_n each of which is normal μ, σ^2 to y_1, y_2, \dots, y_n which are independent, and each y_i is normal $0, 1$.

(Refer Slide Time: 37:46)

$$\begin{aligned}
 \text{Now } y_1 &= \frac{x_1 - \mu}{\sqrt{n}\sigma} + \dots + \frac{x_n - \mu}{\sqrt{n}\sigma} \\
 &= \frac{1}{\sqrt{n}} \sum \left(\frac{x_i - \mu}{\sigma} \right) \\
 &= \frac{1}{\sigma\sqrt{n}} (n\bar{x} - n\mu) \\
 &= \frac{\sqrt{n}}{\sigma} (\bar{x} - \mu) \\
 &= \boxed{\frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim N(0, 1)}
 \end{aligned}$$

Handwritten notes on the slide include a red 'Now' at the start, a red box around the final result, and a red circle around the initial distribution $N(\mu, \sigma^2/n)$ on the left. Logos for NPTEL and ETSC, IIT DELHI are visible at the bottom.

Now, what is y_1 ? y_1 is equal to x_1 minus μ upon root n sigma plus up to x_n minus μ upon root n sigma, is equal to $\frac{1}{\sqrt{n}\sigma} (x_1 - \mu + \dots + x_n - \mu)$, is equal to $\frac{1}{\sqrt{n}\sigma} (n\bar{x} - n\mu)$, is equal to $\frac{\sqrt{n}}{\sigma} (\bar{x} - \mu)$, is equal to $\frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}}$.

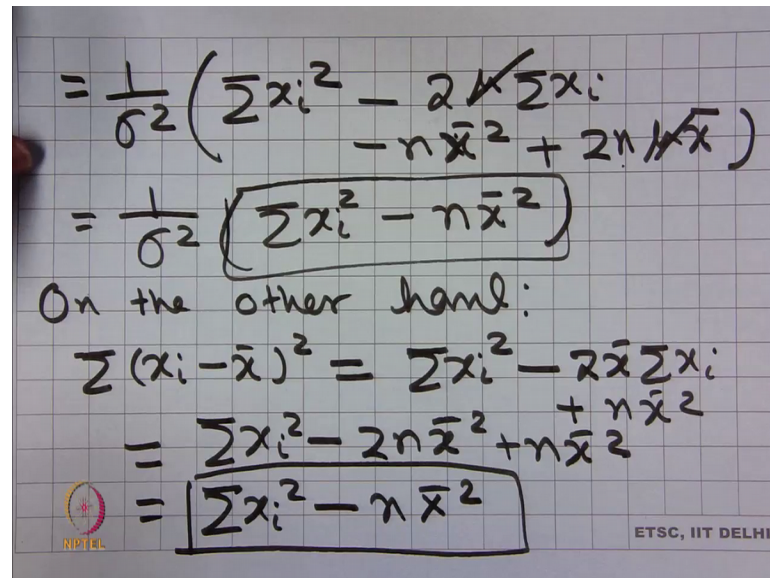
And since y_1 is normal $0, 1$ this is normal $0, 1$ as well. So, this also suggests that \bar{x} is normal with mean is equal to μ , and variance is equal to σ^2/n is equal to σ^2 by n . Although this result I have proved before we can find it here also, but we get something extra. What is that?

(Refer Slide Time: 39:50)

$$\begin{aligned} \text{Let us consider} \\ \sum_{i=2}^n y_i^2 &= \sum_{i=1}^n y_i^2 - y_1^2 \\ &= \sum_{i=1}^n \left(\frac{x_i - \mu}{\sigma} \right)^2 - \frac{1}{\sigma^2} (\bar{x} - \mu)^2 \\ &= \frac{1}{\sigma^2} \left(\sum_{i=1}^n (x_i - \mu)^2 - n(\bar{x} - \mu)^2 \right) \\ &= \frac{1}{\sigma^2} \left(\sum x_i^2 - 2\mu \sum x_i + n\mu^2 - n\bar{x}^2 + 2n\mu\bar{x} - n\mu^2 \right) \end{aligned}$$

Let us consider $\sum_{i=2}^n y_i^2$. This is equal to $\sum_{i=1}^n y_i^2 - y_1^2$. Because we have seen some time back that $\sum y_i^2$ is equal to $\sum (x_i - \mu)^2 / \sigma^2$, which is equal to $\frac{1}{\sigma^2} \left(\sum x_i^2 - 2\mu \sum x_i + n\mu^2 \right)$. So, it is n times \bar{x}^2 plus $2\mu \bar{x}$ minus n times μ^2 .

(Refer Slide Time: 42:16)



The image shows a handwritten derivation on a grid background. It starts with the expression
$$= \frac{1}{\sigma^2} \left(\sum x_i^2 - 2 \cancel{\bar{x} \sum x_i} - n \bar{x}^2 + 2n \cancel{\bar{x}} \right)$$
 where the terms $2 \bar{x} \sum x_i$ and $2n \bar{x}$ are crossed out. This simplifies to
$$= \frac{1}{\sigma^2} \left(\sum x_i^2 - n \bar{x}^2 \right)$$
 which is boxed. Below this, it says "On the other hand:" followed by
$$\sum (x_i - \bar{x})^2 = \sum x_i^2 - 2\bar{x} \sum x_i + n\bar{x}^2$$

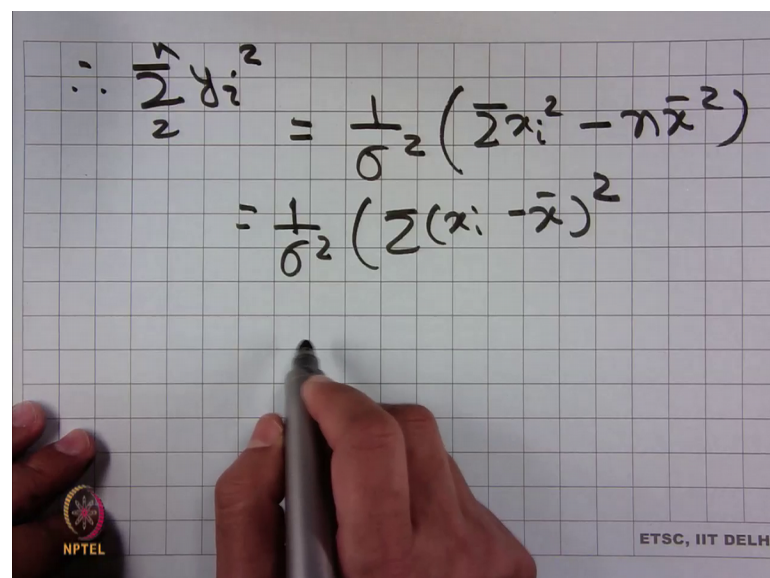
$$= \sum x_i^2 - 2n\bar{x}^2 + n\bar{x}^2$$

$$= \sum x_i^2 - n\bar{x}^2$$
 which is also boxed. The NPTEL logo is in the bottom left and "ETSC, IIT DELHI" is in the bottom right.

So, this cancels is equal to $\frac{1}{\sigma^2}$ by sigma square into sigma x_i square minus 2 mu sigma x_i minus $n \bar{x}^2$ plus 2 $n \mu \bar{x}$. Now sigma x_i is equal to n times \bar{x} . So, this also cancels out, is equal to $\frac{1}{\sigma^2}$ by sigma square into sigma x_i square minus n times \bar{x} square.

On the other hand,, sigma x_i minus \bar{x} whole square is equal to sigma x_i square minus 2 \bar{x} sigma x_i plus $n \bar{x}^2$, is equal to sigma x_i square minus 2 $n \bar{x}$ square plus $n \bar{x}^2$ is equal to sigma x_i square minus $n \bar{x}^2$.

(Refer Slide Time: 44:16)



The image shows a handwritten derivation on a grid background. It starts with
$$\therefore \frac{\sum y_i^2}{2} = \frac{1}{\sigma^2} \left(\sum x_i^2 - n \bar{x}^2 \right)$$

$$= \frac{1}{\sigma^2} \left(\sum (x_i - \bar{x})^2 \right)$$
 A hand holding a pen is visible at the bottom. The NPTEL logo is in the bottom left and "ETSC, IIT DELHI" is in the bottom right.

So, these 2 terms are same therefore we can write as $\sum_{i=2}^n y_i^2$ which we came out to be $\sum_{i=2}^n (x_i - \bar{x})^2$ is equal to $\sum_{i=2}^n (x_i^2 - 2x_i\bar{x} + \bar{x}^2)$.

(Refer Slide Time: 44:44)

Now $\frac{1}{n} \sum (x_i - \bar{x})^2 = s^2$
 $\therefore \sum (x_i - \bar{x})^2 = ns^2$
 $\therefore \sum_{i=2}^n y_i^2 = \frac{ns^2}{\sigma^2}$

Two important consequences.

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Now, $\frac{1}{n} \sum (x_i - \bar{x})^2$ is equal to s^2 . Therefore, $\sum (x_i - \bar{x})^2$ is equal to n times s^2 . Therefore, $\sum y_i^2$ is equal to $\frac{ns^2}{\sigma^2}$. This gives us 2 important consequences.

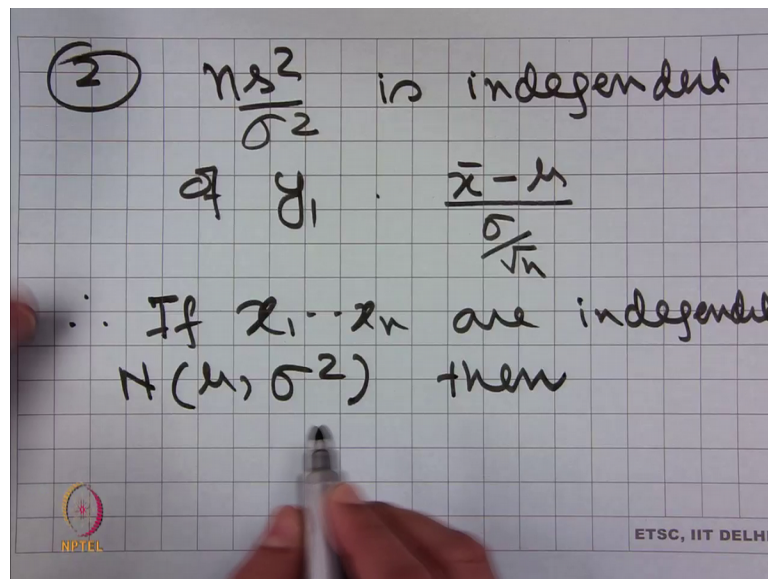
(Refer Slide Time: 45:50)

(i) $\frac{ns^2}{\sigma^2} = \frac{(n-1)s^2}{\sigma^2}$
 is sum of square
 of $n-1$ independent
 $(N(0,1))^2$
 i.e. $\frac{ns^2}{\sigma^2} \sim \chi^2_{(n-1)}$

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One $n s^2$ square sigma square which is same as n minus 1 into s^2 upon sigma square is sum of square of n minus 1 independent normal $0, 1$ square; that is, $n s^2$ square upon sigma square is distributed as chi square with n minus 1 degrees of freedom. Because $n s^2$ square upon sigma square is sum of sigma y_i square from 2 to n ; that means, n minus 1 of independent normal $0, 1$ square. Therefore, $n s^2$ square upon sigma square is distributed at chi square with n minus 1 degrees of freedom.

(Refer Slide Time: 47:10)



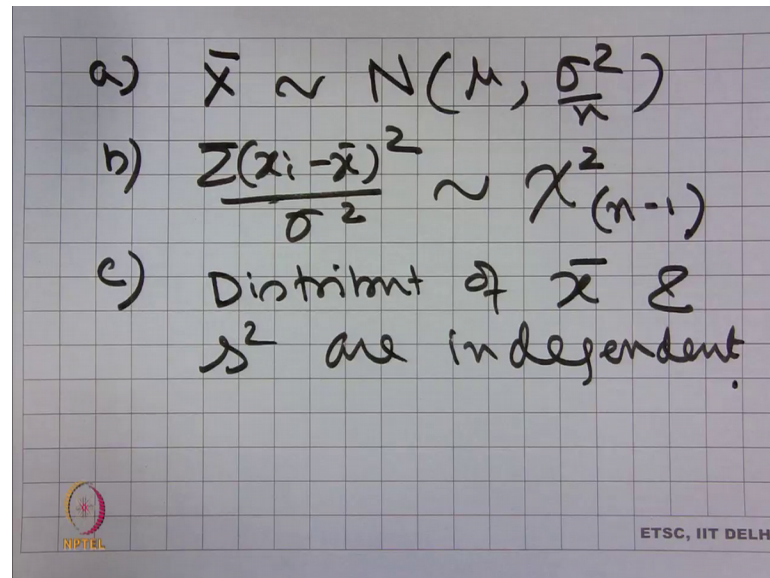
(2) $\frac{n s^2}{\sigma^2}$ is independent of $y_1, \frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}}$

\therefore If x_1, \dots, x_n are independent $N(\mu, \sigma^2)$ then

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And the second point is that $n s^2$ square upon sigma square is independent of y_1 . Because we have found that y_1, y_2, \dots, y_n are n independent normal $0, 1$ variate of which $n s^2$ square sigma square depends only on from y_2 to y_n . Therefore, it is independent of y_1 which is nothing but \bar{x} minus μ upon sigma by root n .

(Refer Slide Time: 48:13)



The image shows a grid background with handwritten mathematical results in black ink. The results are listed as follows:

- a) $\bar{X} \sim N(\mu, \frac{\sigma^2}{n})$
- b) $\frac{\sum (x_i - \bar{x})^2}{\sigma^2} \sim \chi^2_{(n-1)}$
- c) Distribution of \bar{x} & s^2 are independent.

In the bottom left corner, there is a small logo for RIPTIL. In the bottom right corner, the text "ETSC, IIT DELHI" is visible.

Therefore, if x_1, x_2, \dots, x_n are independent normal μ, σ^2 , then \bar{x} is normal with mean μ variance σ^2/n , $\sum (x_i - \bar{x})^2 / \sigma^2$ is distributed as chi square with $n-1$ degrees of freedom. And c distribution of \bar{x} and s^2 are independent.

So, these are 3 important findings that we get by making a particular orthogonal transformation of x_1, x_2, \dots, x_n to y_1, y_2, \dots, y_n . So, these are some important results that we obtain when x_1, x_2, \dots, x_n are independent samples from normal μ, σ^2 . With that I stop for today.

Thank you.