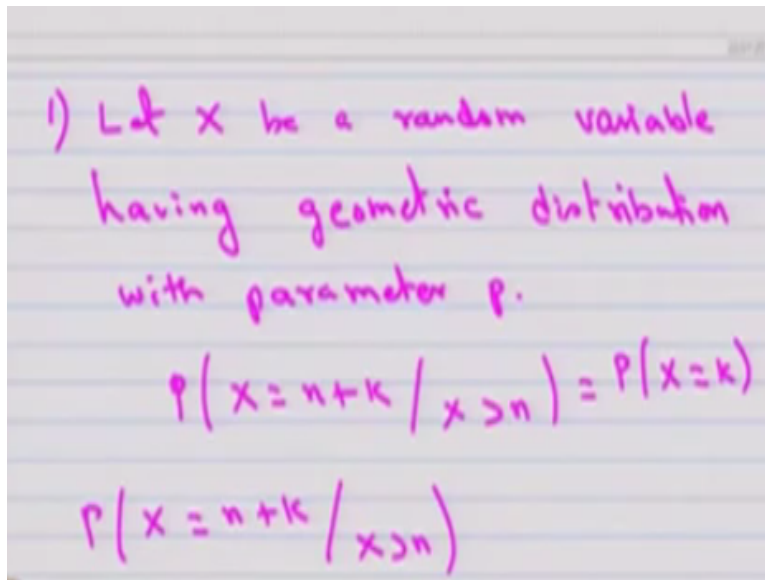


Stochastic Processes - 1
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Lecture – 09
Problems in Random Variables and Distribution

This is the Stochastic processes, model 1; probability theory refresher, lecture 3; problems in random variables and distributions.

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Let as a first problem; let x be a random variable having geometric distribution with the parameter p . Our interest is to find; our interest is to prove that the probability of $x=n+k$, given x takes the value greater than n that is as same as the probability that x takes the value k for every integers, n and k .

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$$\begin{aligned}
 &= \frac{P(x = n+k \cap x > n)}{P(x > n)} \\
 &= \frac{P(x = n+k)}{P(x > n)} \quad \text{---} \\
 &= \frac{(1-p)^{n+k-1} \cdot p}{\sum_{i=n+1}^{\infty} (1-p)^{i-1} \cdot p}
 \end{aligned}$$

You can prove this result by starting from the left hand side that is probability of x takes the value, n + k, given x greater than n by definition this is same as probability of x = n + k intersection, x greater than n, divided by probability of x greater than n, that is same as; that is same as the numerator; x = greater than n means, all possible values, n = n + k that means that the intersection is going to be probability of x takes the value, n + k.

Whereas the denominator is the probability of x is greater than n, that is same as since x is the geometric distribution with the parameter p, the probability of x = n + k, that is nothing but 1-p time p power n+k-1 * p. Whereas the denominator, probability of x is greater than n, that means summation I=n+1 to infinity 1-p power I-1 multiplied by p, that is same as numerator can keep it as it is.

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$$\begin{aligned}
 &= \frac{P(1-p)^{n+k-1}}{P(1-p)^n (1 + (1-p) + (1-p)^2 + \dots)} \\
 &= \frac{(1-p)^{n+k-1}}{(1-p)^n \left(\frac{1}{1-(1-p)} \right)} \\
 &= (1-p)^{k-1} \cdot p = P(x = k)
 \end{aligned}$$

Whereas the denominator since the summation $I = n + k$ to infinity, you can take p times $1 - p$ power, n common outside, the remaining terms are $1 + 1 - p$, the third term will be $1 - p$ whole square and so on. Therefore, you can still simplify you will get $1 - p$ power $n + k - 1$ divided by $1 - p$ power n , keep it as it is, this series you will have the value $1 - p$. therefore if you further simplify you will get a $1 - p$ power $k - 1$ multiply by p .

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2. Let x be a random variable having Gamma distribution with parameters n (positive integer) and λ . Then the CDF of x is given by

$$F_X(x) = 1 - \sum_{i=0}^{n-1} \frac{(\lambda x)^i e^{-\lambda x}}{i!}$$

That is nothing but probability of $x = k$. So this results are the probability of $x = n + k$ given x is greater than n , that is same as probability of $x = k$ for all n at k . This is the important property of geometric distribution and this property is called a memory less property. We will move into the next problem. Let x be a random variable having gamma distribution with the parameter n , you assume that n is a positive integer, the other parameter is lambda.

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$$f_X(x) = \frac{\lambda^n x^{n-1} e^{-\lambda x}}{\Gamma(n)}, \quad x > 0, \lambda > 0$$

Now,

$$F_X(x) = \int_{-\infty}^x f(t) dt$$

$$= \int_0^x \frac{\lambda^n t^{n-1} e^{-\lambda t}}{\Gamma(n)} dt$$

Then the cumulative distribution function CDF of x is given by capital $F(x)$ for the random variable x that is $1 - \sum_{I=0}^{n-1} \frac{\lambda^I x^I}{I!} e^{-\lambda x}$. So, whenever x is a gamma distribution with the parameters n and λ , then the CDF can be written in this way. We know that the probability density function of the gamma distribution is $\frac{\lambda^n x^{n-1} e^{-\lambda x}}{\Gamma(n)}$.

Since n is a positive integer, $\Gamma(n)$ is a $(n-1)!$ factorial. Now you can find out the CDF of this random variable that is nothing but $\int_0^x f(t) dt$, the probability density function that is same as since the $f(x)$ is the; this is valid for x is greater than 0 and λ is greater than 0. So this integration is valid from 0 to x $\frac{\lambda^n t^{n-1} e^{-\lambda t}}{\Gamma(n)} dt$.

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$$\begin{aligned}
 & \text{Put } \lambda t = \mu \\
 & = \int_0^{\lambda x} \frac{\mu^{n-1} e^{-\mu}}{\Gamma(n)} d\mu \\
 & = 1 - \int_{\lambda x}^{\infty} \frac{\mu^{n-1} e^{-\mu}}{\Gamma(n)} d\mu \\
 & = 1 - \frac{1}{(n-1)!} \left[\frac{\mu^{n-1} e^{-\mu}}{-1} \right]_{\lambda x}^{\infty}
 \end{aligned}$$

So now you have to integrate this one and get an expression for this CDF of the random variable x . So what we can do, make a substitution λt , that is same as; we make it as some μ . Therefore, these integration becomes the integration from 0 to λx , $\frac{\mu^{n-1} e^{-\mu}}{\Gamma(n)}$, divided by $\Gamma(n)$. That is same as $1 - \int_{\lambda x}^{\infty} \frac{\mu^{n-1} e^{-\mu}}{\Gamma(n)} d\mu$.

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$$\begin{aligned}
 & - \int_{\lambda x}^{\infty} \frac{(n-1)\mu^{n-2} e^{-\mu}}{-1} d\mu \\
 = & 1 - \frac{1}{(n-1)!} (\lambda x)^{n-1} e^{-\lambda x} \\
 & - \frac{(\lambda x)^{n-2} e^{-\lambda x}}{(n-2)!} \dots \\
 = & 1 - \sum_{i=0}^{n-1} \frac{(\lambda x)^i e^{-\lambda x}}{i!}
 \end{aligned}$$

That is same as 1-; since n is a positive integer gamma of n is n-1 factorial so you can take it outside. You can do this integration by parts so you will get a; Mu power n-1 e power - Mu divided by -1 between the limits lambda x to infinity – integration from lambda x to infinity n-1 times Mu of n-2 e power –Mu divided by -1 d Mu. So the whole thing is multiplied by n-1 factorial.

Now you can integrate the second term again by integration by parts and when you substitute the limits for Mu is infinity and as well as Mu = lambda x and subsequently if you do the integration by parts, you will get a 1- n-1 factorial lambda x power n-1 e power – lambda x. Then the next term will be –lambda x power n-2 e power - lambda x by n-2 factorial. Similarly, the other terms.

The last term will be by doing integration by parts again and again, the last term you will get minus- of lambda x power 0 e power –lambda x by 0 factorial. These we can write it in the form 1-summation I=0 to n-1 lambda x power I e power –lambda x by I factorial. So here we are finding this CDF of the gamma distribution when one of the integer is the positive integer, one of the parameters is a positive integer.

This result will be useful in finding the total time spending the queuing system, that will be discussed in the later models.