

Stochastic Processes
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Module - 1
Probability Theory Refresher
Lecture - 4
Problems in Sequences of Random Variables

So, this is stochastic processes module one probability theory refresher lecture four problems in sequence of random variables. As a illustrative examples we are going to discussed four problems in this lecture.

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1). Let Z_1, Z_2, \dots be a sequence of r.v.s each having Poisson distribution with parameter n .

$Z_n \sim P(n), n=1, 2, \dots$

Limiting distribution of the r.v.

$Y_n = \frac{Z_n - n}{\sqrt{n}}$

The first problem, let z_1, z_2 , and so on be a sequence of random variables each having Poisson distribution with parameter n , that is z_n is Poisson distribution with the parameter n , for n is equal to 1, 2, 3 and so on. Our interest is to find the limiting distribution of the random variable that is defined as y suffix n that is z_n minus n divided by square root of n .

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Handwritten derivation of the Moment Generating Function (MGF) of a Poisson distribution with parameter n :

$$\begin{aligned}
 M_{Z_n}(t) &= E(e^{Z_n t}) \\
 &= \sum_{k=0}^{\infty} e^{kt} \frac{e^{-n} n^k}{k!} \\
 &= e^{-n} \sum_{k=0}^{\infty} \frac{(e^t \cdot n)^k}{k!} \\
 &= e^{-n} e^{n e^t} = e^{n(e^t - 1)}
 \end{aligned}$$

Below this, the MGF of the standardized variable is written as:

$$M_{Y_n}(t) = M_{\frac{Z_n - n}{\sqrt{n}}}(t)$$

So given Z_n is Poisson distribution with the parameter n . We can find out the MGF of Z_n . MGF of Z_n is nothing but expectation of $e^{Z_n t}$. That is same as summation k is equal to 0 to infinite e^{kt} , e^{-n} n^k by k factorial, because it is the expectation of $e^{Z_n t}$, where Z_n is Poisson distribution with the parameter λ . Therefore, this is going to be m is k is equal to 0 to infinity this one.

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Handwritten derivation of the MGF of the standardized variable $Y_n = \frac{Z_n - n}{\sqrt{n}}$:

$$\begin{aligned}
 &= E\left(e^{\frac{Z_n - n}{\sqrt{n}} t}\right) \\
 &= e^{-t\sqrt{n}} M_{Z_n}\left(\frac{t}{\sqrt{n}}\right) \\
 &= e^{-t\sqrt{n}} e^{n(e^{t/\sqrt{n}} - 1)} \\
 &= e^{-t\sqrt{n}} e^{n\left(1 + \frac{t}{\sqrt{n}} + \frac{t^2}{2n} + \frac{t^3}{3!n^{3/2}} + \dots - 1\right)}
 \end{aligned}$$

So, you can take $e^{\text{power } n}$ outside. So, the remaining term becomes, k is equal to zero to infinity $e^{\text{power } t}$ multiplied by n to whole thing power k by k factorial. That is same as $e^{\text{power } n}$ times, $e^{\text{power } t}$ that can be rewritten as $e^{\text{power } n}$ times $e^{\text{power } t \text{ minus } 1}$. Now, we will find out the M z f of the random variable y_n where y_n is the z_n minus n divided by square of n .

Therefore, the M z f of the random variable y_n as a function of t that becomes M z f of z_n minus n divided by square root of n function of t , that is same as expectation of $e^{\text{power } z_n \text{ minus } n \text{ divided by root } n \text{ multiplied by } t}$, you know the rules of a moment generating function the constant is out, so you can use that logic. So, it become $e^{\text{power } n}$ times $e^{\text{power } t \text{ times root } n}$ because n t by root n , therefore it becomes a t times root n , then M z f of the random variable z_n use the another will of a moment generating function instead of t if the comes it t divided by square root of n . So, that is same as $e^{\text{power } n}$ times $e^{\text{power } t \text{ times root } n}$ just now we found what is the moment generating function of a z_n so use the same thing, but replace t by t divided by square root of n therefore, this becomes $e^{\text{power } n}$ times wherever the t you replace by t by square root of n so t by square of n minus 1.

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The image shows a handwritten derivation in a Notepad window. The text is as follows:

$$= e^{-t\sqrt{n}} e^{t\sqrt{n} + \frac{t^2}{2} + \frac{t^3}{3!n^{3/2}} + \dots}$$

$$M_{Y_n}(t) = e^{\frac{t^2}{2} + \frac{t^3}{3!n^{3/2}} + \dots}$$

As $n \rightarrow \infty$

$$M_{Y_n}(t) \rightarrow e^{t^2/2}$$

Limiting distribution of $\frac{Z_n - n}{\sqrt{n}}$ is a standard normal distribution.

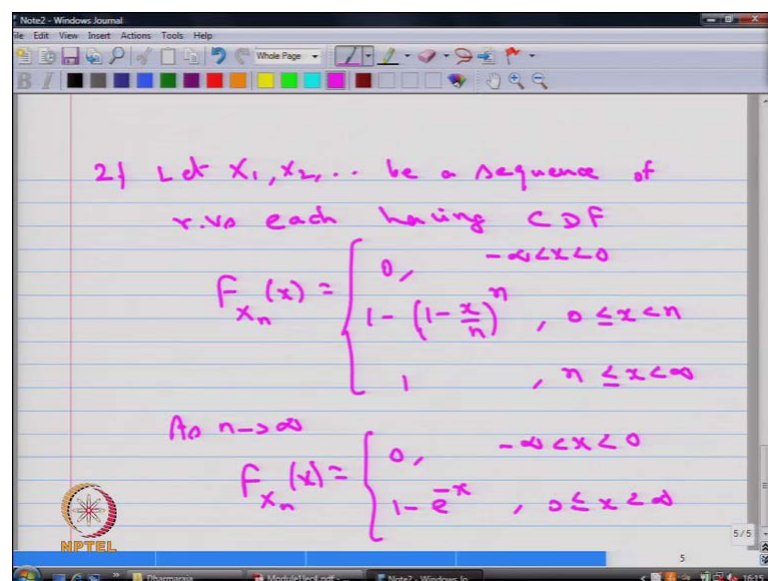
Therefore, we can further simplify by expanding $e^{\text{power } t \text{ by } n}$; that means, you keep this $e^{\text{power } n}$ you expand only $e^{\text{power } t \text{ by square root of } n}$; that is 1 plus t divided by square root of n then the next term will be t square by 2 times n , and the next term will

be t^3 divided by $3!$ $n^{3/2}$ and so on. And the last term is. So, this is an expansion of $e^{t \sqrt{n}}$.

So, close the bracket that is same as $e^{t \sqrt{n}}$ multiplied by. So, this 1 and plus 1 and minus 1 will be cancelled so you will get $e^{t \sqrt{n}}$ that becomes $t \sqrt{n}$, and the next term becomes t^2 by 2 then it becomes t^3 by 3 factorial $n^{1/2}$ and so on. Therefore, this becomes $e^{t^2/2} + t^3/3! \sqrt{n}$ and so on.

Our interest is to find out the limiting distribution of y_n so this is the moment generating function of y_n for n . So, as n tends to infinity, because our interest is to find out the limiting distribution as n tends to infinity the moment generating function of y_n becomes $e^{t^2/2}$. If you recall the moment generating function for standard distributions one can conclude this is the MGF of a standard normal distribution. Therefore we conclude the limiting distribution of y_n is standard normal distribution; that is a the limiting distribution $z_n = (y_n - n)/\sqrt{n}$ is a standard normal distribution. So, this problem is very important in the renewal process therefore we discuss this example has the how to find the limiting distribution of some standard random variables.

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2) Let X_1, X_2, \dots be a sequence of r.v.s each having CDF

$$F_{X_n}(x) = \begin{cases} 0, & -\infty < x < 0 \\ 1 - \left(1 - \frac{x}{n}\right)^n, & 0 \leq x < n \\ 1, & n \leq x < \infty \end{cases}$$

As $n \rightarrow \infty$

$$F_{X_n}(x) = \begin{cases} 0, & -\infty < x < 0 \\ 1 - e^{-x}, & 0 \leq x < \infty \end{cases}$$

Next example; let x_1, x_2 and so on be a sequence of random variables, each having c d F cumulative distribution function F suffix x_n of x_0 from minus infinity to 0 detects the

value $1 - (1 - x/n)^n$. For x lies between 0 to n from n onwards till infinity the value is 1 . So, this is the cumulative distribution function for the random variables exercise. It is the function of n therefore I have made it F suffix x suffix n ; that means, this is the c d F for the random variable n for every n you have this, as n tends to infinity we get F suffix x n of x that becomes 0 from minus infinity to 0 , and it takes the value $1 - e^{-x}$ from 0 to infinity, as n tends to infinity this c d F of the random variables x/n becomes 0 between the intervals minus infinity to 0 , the value becomes $1 - e^{-x}$ where x lies between 0 to infinity.

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Handwritten notes on a digital notepad:

$$F_{x_n}(x) = \begin{cases} 0, & -\infty < x < 0 \\ 1 - e^{-x/n}, & 0 \leq x < \infty \end{cases}$$

Suppose x is a r.v. with CDF

$$F_x(x) = \begin{cases} 0, & -\infty < x < 0 \\ 1 - e^{-x}, & 0 \leq x < \infty \end{cases}$$

Then $x_n \xrightarrow{d} x$

$x \sim \text{Exp}(1)$

Suppose, x is a random variable with the c d F that is F_x of x that is 0 between the interval minus infinity to 0 and $1 - e^{-x}$ where x lies between zero to infinity. Then one can conclude x_n converges to x in distribution, since the sequence of F_{x_n} of x tends to F_x of x , for x is greater or equal to 0 and the value is $1 - e^{-x}$. Hence one can conclude the sequence of random variable of x_n converges to random variable in distribution.

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3) Suppose we choose at random n numbers from the interval $[0, 1]$ with uniform distribution. Let X_i be a random variable describing the i th choice. Then, for $i = 1, 2, \dots$

$$E(X_i) = \int_0^1 x \, dx = \frac{1}{2}.$$

Here, the x is a exponential distribution with the parameter 1. So, this is the one example of a all the sequence of random variable converges to a random variable in distribution. Next I will move in to the third example; suppose, we choose at random n numbers from the interval 0 to 1 with uniform distribution. Let capital x_i be a random variable describing the i th choice. Then for i is equal 1, 2 and so on you can find out what is the expectation of x is?

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$$\begin{aligned} \text{Var}(X_i) &= \int_0^1 x^2 \, dx - \left(\frac{1}{2}\right)^2 \\ &= \frac{1}{3} - \frac{1}{4} \\ &= \frac{1}{12}. \end{aligned}$$

Let $S_n = X_1 + X_2 + \dots + X_n$

$$\begin{aligned} E\left(\frac{S_n}{n}\right) &= \frac{1}{2}. \\ \text{Var}\left(\frac{S_n}{n}\right) &= \frac{1}{12n}. \end{aligned}$$

That is nothing but the integration from 0 to 1 x times the probability density function, the probability density function for uniform distribution with the interval 0 to 1 that is 1. Therefore, x into dx if we compute the expectation x_i is going to be 1 by 2. Similarly, one can evaluate the variance of x_i is that is nothing but 0 to 1 $x^2 dx$ minus the mean square, expectation of x^2 minus expectation of x the whole square. So, the expectation of x^2 is 0 to 1 $x^2 dx$. So, if you evaluate this quantity that is 1 by 3 minus 1 by 4 so if you simplify you will get 1 by 12.

If you remember the formulae of variance of uniformly distributed random variable between the interval a to b then the variance of x_i is nothing but you can get it that by substituting the value of a is equal to 0, and b is equal to 1 you will get 1 by 12. Let $S_n = x_1 + x_2 + \dots + x_n$ and so on till x_n , one can find the variance of S_n , because you know the variance of x_i is using that you can find out what is the mean of S_n ? But our interest is not the finding the mean of S_n oriented is to find out the mean of S_n by n , that is basically suppose x_i is the samples then S_n divided by n is nothing, but the sample mean.

So, expectation of S_n divided by n that becomes the 1 by 2. Similarly, if you calculate variance of S_n by n that becomes 1 divided by 12 times n , because the variance of x_i is 1 by 12 so the variance of S_n is the summation of x_i is from 1 to n therefore variance of S_n by n becomes 1 divided by 12 time here.

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For any $\epsilon > 0$, using Chebyshev's inequality

$$P\left\{\left|\frac{S_n}{n} - \frac{1}{2}\right| \geq \epsilon\right\} \leq \frac{1}{12n\epsilon^2}$$

As $n \rightarrow \infty$

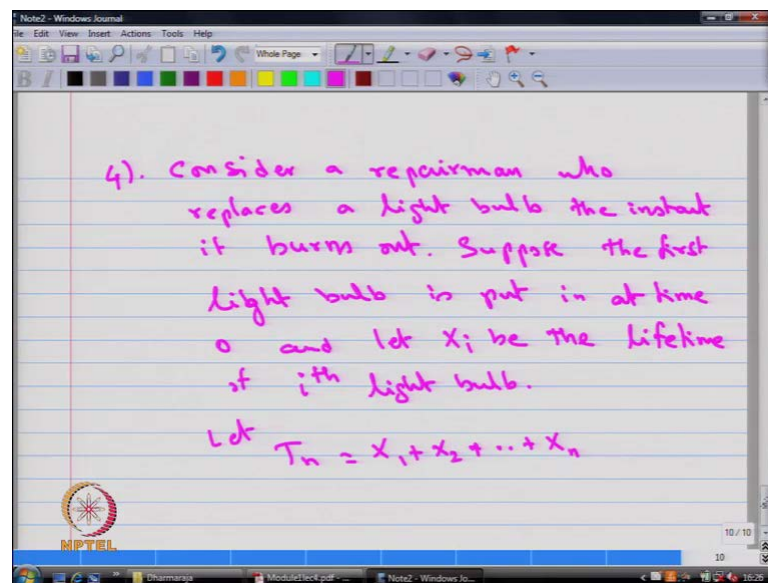
$$P\left\{\left|\frac{S_n}{n} - \frac{1}{2}\right| \geq \epsilon\right\} \rightarrow 0$$

$$\frac{S_n}{n} \xrightarrow{P} \frac{1}{2}$$

For any epsilon greater than 0 using chebyshev's inequality one can conclude the probability of absolute of S_n by n minus $1/2$ greater than or equal to epsilon that is less than or equal to 1 divided by 12 times n epsilon square. I am using the chebyshev's inequality by knowing mean of S_n by n is $1/2$ and variance of S_n by n is 1 divided by $12n$, I get this inequality. Now, as n tends to infinity the probability of absolute of S_n by n minus $1/2$ which is greater than or equal to epsilon will tends to 0, because epsilon is in the n is the denominator, because n is in the denominator as n tends to infinity this probability tends to 0, that is nothing but S_n by n tends to the value $1/2$, and this convergence takes place in probability the sequence of random variables S_n by n converges to $1/2$ in probability.

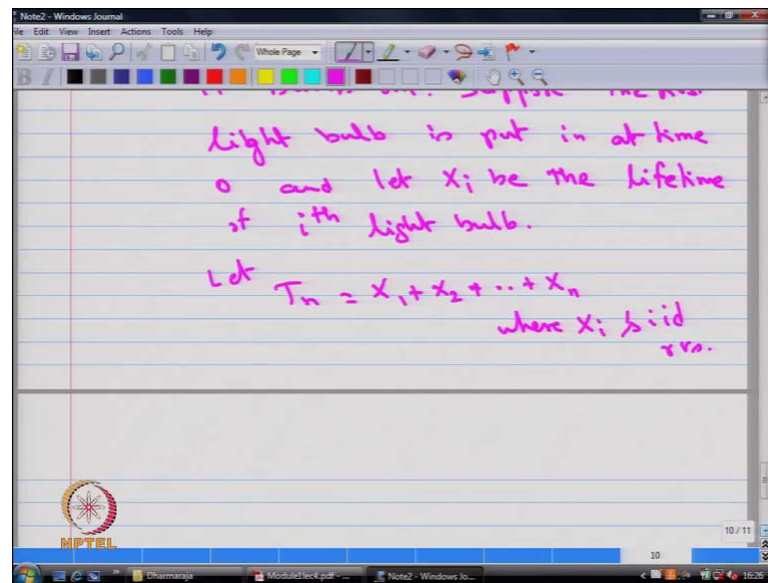
Therefore, we say the sequence of random variable x_n for n is equal to 1, 2 and so on, obeys the weak law of large numbers, because the S_n by n converges to $1/2$ in probability therefore, we say the sequence of random variables x_n is obeys the weak law of large numbers. So, that is the intention of giving this example.

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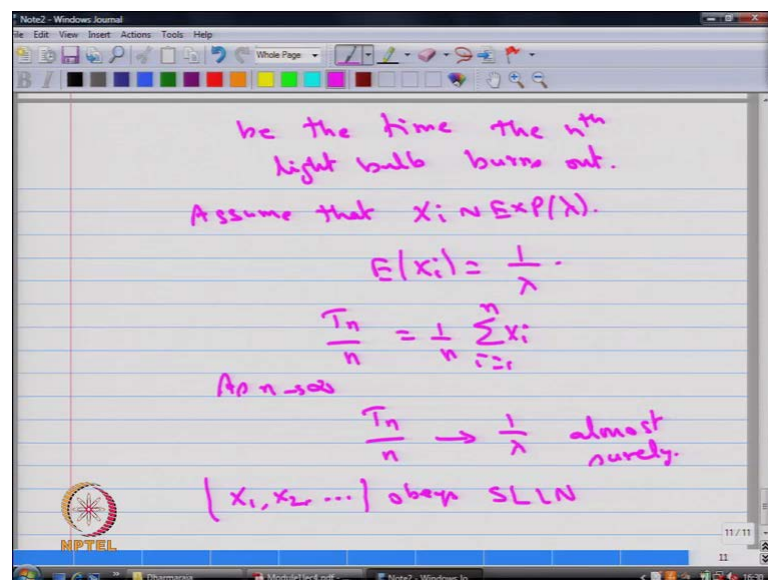
Now, I move in to the fourth example, consider a repair man who replaces a light bulb the instant it burns out. Suppose, the first light bulb is put in at time 0 and let x suffix i be the life time of i th light bulb. You defined the random variables T_n is the sum of n x_i 's.

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where x_i 's are i.i.d random variables, x_i be the life time of i th light bulb, and when x_i are i.i.d random variables you are defining T_n is the x_1, x_2 and x_n and so on.

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So, the T_n be the time of time the n th light bulb burns out, because the T_n is the x_1 plus x_2 and so on, till x_n therefore, T_n be the time the n th light bulb burns out. Assume that x_i is exponential distribution with the parameter λ , you know that already x_i are i.i.d λ from variable now, I am making the further assumption x_i is follows exponential distribution with the parameter λ ; that means, you know what

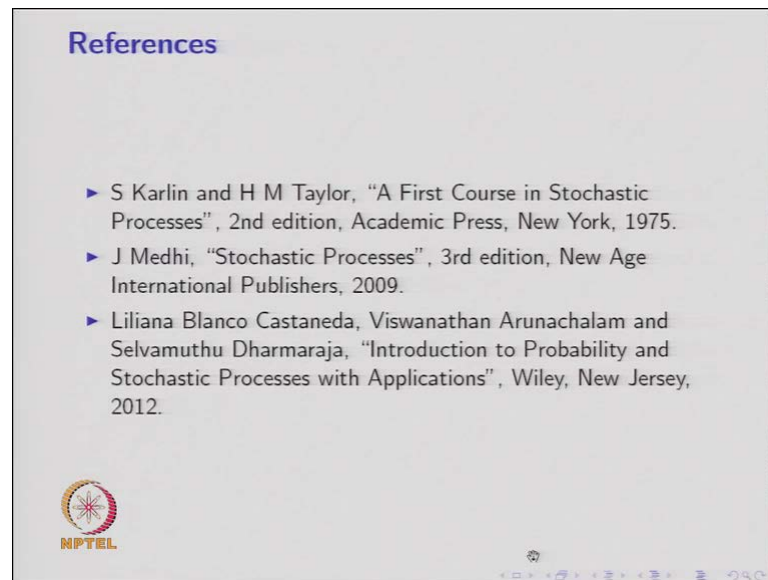
is the mean of this random variable. Since it is exponential distribution with the parameter λ this becomes $1/\lambda$, also one can use the result T_n/n that is nothing but $1/n$ summation of x_i is, where i is running from 1 to n .

As n tends to infinity as n tends to infinity one can prove T_n/n tends to $1/\lambda$; that is the mean of random variable x_i almost surely, I am not proving here the way do the sequence of random variable converges to another random variable converges takes place in probability or in distribution or in (()) mean or almost surely one can prove this the T_n/n converges to $1/\lambda$ almost surely. That means, you can conclude the random variable x_1, x_2 and so on, obeys strong law of large numbers, because that T_n/n is nothing but the $1/n$ summation of x_i 's that converges to the value $1/\lambda$ almost surely we can conclude the sequence of random variable x_i is obeys the strong law of large numbers. Even though, in this problem I made the assumption x_i is follows the exponential distribution with the parameter λ in general the lifetime can be any distribution. So, this problem will be discussed in detail in renewal process.

So, as such here, we are making the assumption of distribution of x_i is exponential distribution, therefore I made it converges takes place almost surely to the value $1/\lambda$ this can be generalized. There are many more problems of the similar kind, but we are discussing only few problems therefore we can use the similar logic of a finding the moment generating function then concluding the distribution and finding the limiting distribution or you verify whether the sequence of random variable converges takes place in mean converges takes place in probability or converges takes place in distribution or converges in the orth mean or converges almost surely this can be used in any problem of the same way what I have done it here.

And I have not discussed any problem in the central limit theorem but that will be used many times, therefore I have not given any problems for the center limit theorem.

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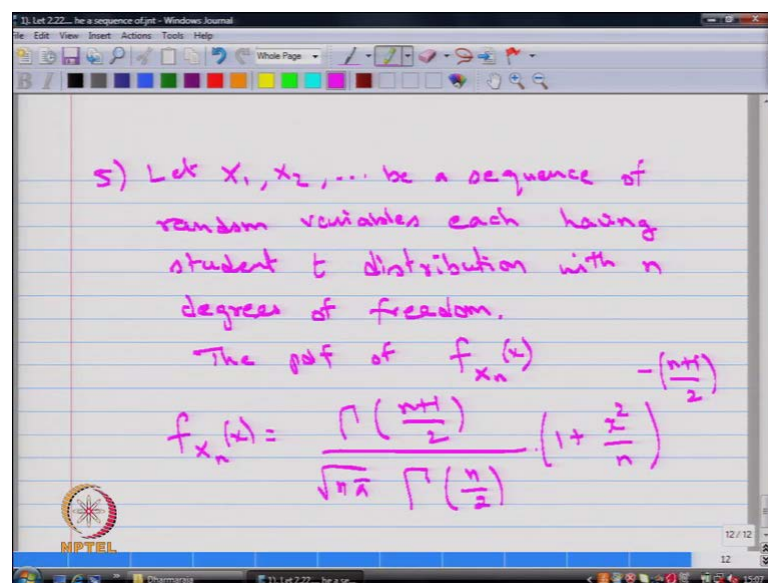
So, if these are all the references reference books we have used for the lecture 3 as well as the lecture 4. It is not end.

Student: (())

So, what to do then we may go to some more problems than...

Student: (())

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I did not realize that I may lambda for. Let x_1, x_2 so on, be a sequence of random variables each having student t distribution with n degrees of freedom. Our interest is to find out the limiting distribution of the student t distribution. We know that the probability density function of f of x for the random variable x_n is given by gamma of n plus 1 by 2 divided by square root of n times phi multiplied by gamma of n by 2 multiplied by 1 plus x square by n power minus n plus 1 by 2. So, this is the probability density function of a random variable x_n .

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For large n ,

$$\lim_{n \rightarrow \infty} \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{n\pi} \Gamma\left(\frac{n}{2}\right)} = \frac{1}{\sqrt{2\pi}}$$

Also

$$\lim_{n \rightarrow \infty} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}} = e^{-\frac{x^2}{2}}$$

$$\lim_{n \rightarrow \infty} f_{x_n}(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

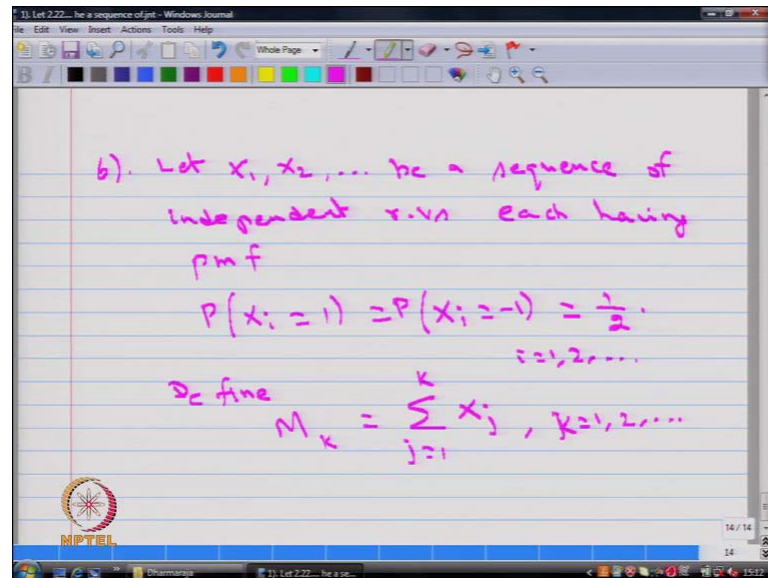
$x_n \xrightarrow{d} Z$ where $Z \sim N(0,1)$

Our interest is to find out the limiting distribution of random variable x_n . For larger for large n , we have the results limit n tends to infinity of gamma of n plus 1 by 2 divided by square root of n phi of gamma of n by 2 is 1 divided by square root of 2 phi using stirling approximation, and also limit n tends to infinity of 1 plus x square by 2 the whole power minus n plus 1 by 2, that we know that is e power minus x square by 2. Hence a limit n tends to infinity of the probability density function of the random variable x_n becomes 1 divided by square root of 2 phi e power minus x square by 2.

Since, right hand side is the probability density function of a standard normal distribution. We conclude for a larger n the sequence of random variable x_1, x_2, x_n and so on, that tends to the random variable z this convergence takes place in distribution, where z is standard normal distribution.

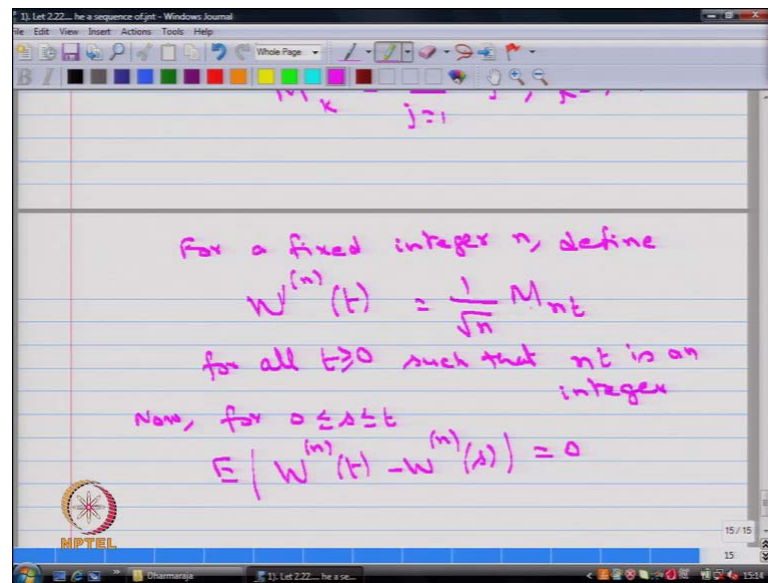
So, this is a simple example of the sequence of random variables, each having a student t distribution. The limiting distribution converges to standard normal in they converge in distribution.

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Now, move into next example, example 6, let x_1, x_2 so on, be a sequence of independent random variables each having probability mass function, probability of x_i is equal to 1, that is same as probability of x_i takes the value minus 1 probability is 1 by 2. This is valid for that means, it is a sequence of i i d random variable, and they are discrete type. Define M suffix k , thus the sum of first k x_i random variables. So, this running index k is equal to 1, 2 and so on. So, we are defining k sequence of a random variable m_k by summing first k x_i random variables.

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For a fixed integer n , we define another sequence of random variable that is denoted by W superscript n of t , that is nothing but 1 divided by square root of n M suffix n times t . This is for all t greater than equal to 0, such that n time's t is an integer. So, we are defining another sequence of random variable W superscript n of t , that is 1 divided by square root of n times m n of t , where n of t is a integer, so this is valid for all t greater than or equal to 0. To find out the mean and variance for the different of the random variable of a n of t minus W n of S for 0 less than or equal to s , less than or equal to t this quantity will be 0; that means W n of t is the 1 divided by square root of n , M n of t and the way we define the M n of t that is the summation of x_i .

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b). Let X_1, X_2, \dots be a sequence of independent r.v.s each having pmf

$$P(X_i = 1) = P(X_i = -1) = \frac{1}{2}, \quad i = 1, 2, \dots$$

Define

$$M_n = \sum_{j=1}^n X_j, \quad n = 1, 2, \dots$$

And the probability of x_i is equal to 1, and probability of x_i is equal to minus 1 minus 1 is 1 by 2, therefore, the mean of x_i are going to be 0.

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For a fixed integer n , define

$$W_n^{(n)}(t) = \frac{1}{\sqrt{n}} M_{nt}$$

for all $t \geq 0$ such that nt is an integer

Now, for $0 \leq s \leq t$

$$E[W_n^{(n)}(t) - W_n^{(n)}(s)] = 0$$

$$\text{Var}(W_n^{(n)}(t) - W_n^{(n)}(s)) = t - s$$

Because, of that the expectation of or mean of W_n of t minus W_n of s , that is equal to 0. Also, if you evaluate the variance of W_n of t minus W_n of S by finding first variance of x_i 's using that you are find out the variance of M_n of t , then find out the W_n of t minus W_n of s , that is going to be t minus s .

Its need calculation of expectation of x_i square then using the expectation x_i square and expectation of x_i 's you can find out the variance of x_i 's, using variance of x_i 's you can find out the variance of W_n of t , then you find out the variance of W_n of t minus W_n of s .

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$$\text{Var} (W_n^{(m)}(t) - W_n^{(m)}(s)) = t - s$$
 For $t \geq s$,

as $n \rightarrow \infty$

$$W_n^{(m)}(t) \xrightarrow{d} X$$
 where $X \sim N(0, t)$

By using mean and variance, for fixing t greater than or equal to 0, as n tends to infinity you can conclude W_n of t tends to a random variable x , and this convergence takes place in distribution using central limit theorem one can control W_n of t converges to the random variable x , the convergence in distribution, where x is normal distribution with the mean 0 and variance t .

Using a central limit theorem one can prove W_n of t converges to x in distribution, where x is a normal distribution if the mean is 0 and the variance t . This result is very useful in Brownian motion, and this same problem will be discussed in detail, when we are discussing the module of a Brownian motion.