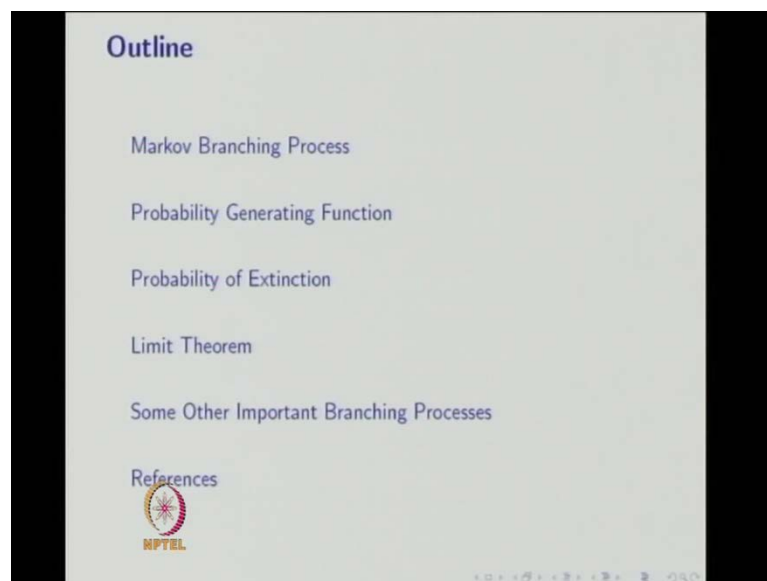


Stochastic Processes
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Module - 9
Branching Processes
Lecture - 2
Markovian Branching Process

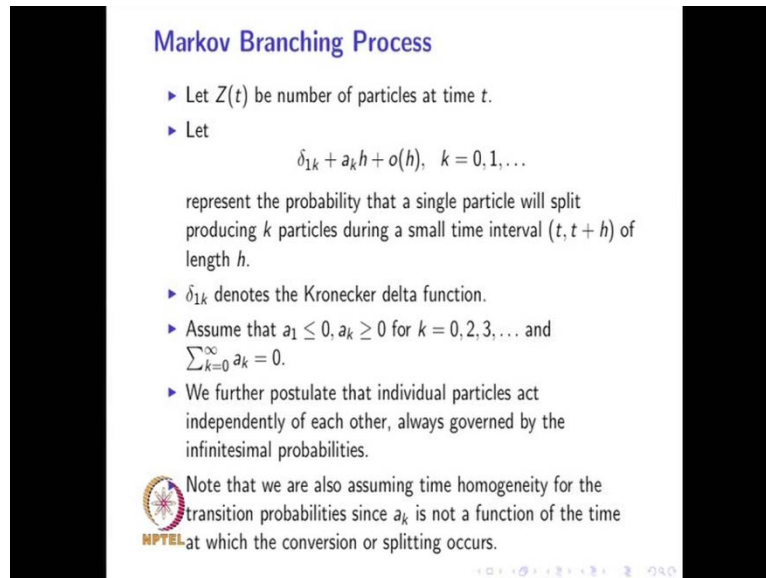
This is a stochastic processes module nine branching processes. In the lecture one, we have discussed the definition and examples of branching processes, important discrete type branching process, Galton-Watson process is discussed in detail. We found mean and variance of Galton-Watson process. Then we have find you will find the probability of extinction for the Galton-Watson branching process. This is a lecture two, in this lecture we are going to discuss Markov branching process.

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This is a very important branching process of a continuous time. This we are going to start with the probability generating function. Then, we are finding the probability of extinction and we discuss the limit theorem. Finally, we are going to discuss some other important branching processes at the end.

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Markov Branching Process

- ▶ Let $Z(t)$ be number of particles at time t .
- ▶ Let $\delta_{1k} + a_k h + o(h)$, $k = 0, 1, \dots$ represent the probability that a single particle will split producing k particles during a small time interval $(t, t + h)$ of length h .
- ▶ δ_{1k} denotes the Kronecker delta function.
- ▶ Assume that $a_1 \leq 0$, $a_k \geq 0$ for $k = 0, 2, 3, \dots$ and $\sum_{k=0}^{\infty} a_k = 0$.
- ▶ We further postulate that individual particles act independently of each other, always governed by the infinitesimal probabilities.

 Note that we are also assuming time homogeneity for the transition probabilities since a_k is not a function of the time at which the conversion or splitting occurs.

What is a Markov branching process? Let $Z(t)$ be the number of particles at time t . The sequence is Z of t , the collection of random variables Z of t for t greater than or equal to 0 form a Markovian branching process with the following assumptions. Let $\delta_{1k} + a_k h + o(h)$ for k is equal to, 1, 0, 1, 2 and so on, represents the probability, that a single particle is split producing k particles during a small time interval t to $t + h$ of length h . δ_{1k} denotes the Kronecker delta function.

Assume, that a_1 is less than or equal to 0 for k is equal to 0, 2, 3 and so on. a_k 's are greater than or equal to 0 and summation of a_k 's starting from 0, 1, 2 and so on, that will be 1, that will be 0. We further postulate, that individual particles act independently of each other, always governed by the infinitesimal probabilities. Note that we are also assuming time homogeneity for the transition probabilities, since a_k is not a function of time at which the conversation conversion or split occurs; since a_k is not a function of the time at which the split occurs. The similar assumptions we have taken care in the discrete type branching process also.

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Markov Branching Process ...

- ▶ Each particle lives a random length of time following an exponential distribution with mean $1/\lambda = a_0 + a_2 + a_3 + \dots$
- ▶ On completion of i -th lifetime, it produces a random number D of descendants of like particles.
- ▶ The probability mass function of D is

$$P(D = k) = \frac{a_k}{a_0 + a_2 + a_3 + \dots}, \quad k = 0, 2, 3, \dots$$

- ▶ The lifetime and progeny distribution of separate individuals are independent and identically distributed.
- ▶ Using the independence assumptions, we get

$$P(Z(t+h) = n+k-1 \mid Z(t) = n) = na_k h + o(h), \quad k = 0, 2, 3, \dots$$

NPTEL $P(Z(t+h) = n \mid Z(t) = n) = 1 + na_1 h + o(h)$

Each particle lives a random length of time following an exponential distribution with the mean $1/\lambda$, that is $(a_0 + a_2 + a_3 + \dots)$. On the completion of i -th lifetime, it produces a random number D of descendants of like particles. The probability mass function of D is probability, that D is equal to k will be a_k divided by $a_0 + a_2 + a_3 + \dots$. The lifetime and progeny distribution of separate individuals are independent and identically distributed. Using the independent assumptions we get the conditional probability of $Z(t+h)$ is equal to $n+k-1$, given $Z(t)$ was n that is same as n times $a_k h$ plus order of h .

We know that small order of h means small order of h divided by h tends to 0, as h tends to infinity, as h tends to 0; order of h divided by h tends to 0 as h tends to 0. And for k equal to 1 the probability of $Z(t+h)$ is equal to k given $Z(t)$ is equal to n . Probability of $Z(t+h)$ is equal to n given $Z(t)$ is equal to n , that will be $1 + na_1 h + o(h)$. So, for $k = 0, 2$ and 3 we have a separate expression, for k equal to 1 we have a different expression.

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Probability Generating Function


► Let

$$P_{ij}(t) = P(Z(t+s) = j \mid Z(s) = i), \quad i, j = 0, 1, \dots, t, s \geq 0$$

and let

$$\phi(t; s) = \sum_{j=0}^{\infty} P_{1j}(t) s^j$$

the probability generating function for $P_{1j}(t)$.



Now, using the condition probability we are going to define the probability generating function. Now, let $P_{ij}(t)$ is nothing but the conditional probability of, probability of $Z(t) + s$ is equal to j , given $Z(s)$ was i . Using these, we define the probability generating function, that is nothing but $\phi(t; s)$, the two variables summation over j $P_{1j}(t) s^j$. So, this will be the probability generating function for $P_{1j}(t)$ where $P_{1j}(t)$ is transition probabilities.

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Theorem 4: PGF of $Z(t)$

► The probability generating function for $P_{ij}(t)$, $\phi(t; s)$, satisfies:

$$\phi(t+v; s) = \phi(t; \phi(v; s))$$

► This is the continuous time analog of Theorem 1, in the case of discrete time branching processes.

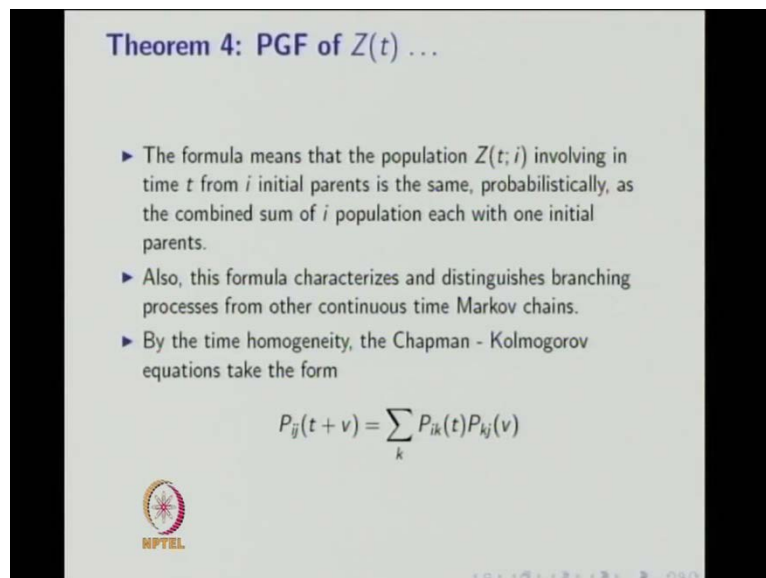
► **Proof:** Since individual particles act independently, we have the fundamental relation

$$\sum_j P_{ij}(t) s^j = \left[\sum_j P_{1j}(t) s^j \right]^i = [\phi(t; s)]^i$$


Now, we are going to discuss the probability generating function of Z of t in the theorem 4. Already first three theorems are discussed are the discrete type branching process. So, here we are going to discuss the fourth theorem, the probability generating function for P_{ij} of t . There is ψ of t, s satisfies ψ of t plus v of ψ of t plus (v, s) , that is same as ψ of t comma ψ of v, s . This is a continuous time analog of theorem 1 in the case of discrete time branching processes.

We discuss the proof. Since individual particles act independently, we have the fundamental relation, the probability generating function for P_{ij} of t , that is nothing but summation over j . Instead of the transition probability, i to j is a transition probability of 1 to j of $t s$ power j the whole power i , that is same as the probability generating function power i .


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Theorem 4: PGF of $Z(t)$...

- ▶ The formula means that the population $Z(t; i)$ involving in time t from i initial parents is the same, probabilistically, as the combined sum of i population each with one initial parents.
- ▶ Also, this formula characterizes and distinguishes branching processes from other continuous time Markov chains.
- ▶ By the time homogeneity, the Chapman - Kolmogorov equations take the form

$$P_{ij}(t + v) = \sum_k P_{ik}(t) P_{kj}(v)$$



The reason is, the formula means, that the population Z of t, i involving in time t from i initial parents is the same, probabilistically, as the combined sum of i population each with one initial parents. Therefore, the left hand side is the probability generating function of, function for P_{ij} of t , that is same as making summation over j , the probability transition, probability of $P(1, j)$ of $t s$ power j the power I , that is nothing but the probability generating function for P_{ij} power i .

Also, this formula characterizes and distinguishes branching processes from other continuous time branching continuous time Markov chains. By the time homogeneity, the Chapman-Kolmogorov equations take the form, the one step transition probability of P_{ij} plus

v can be written in the form of summation k P_{ik} to k of t , then P_{kj} of v . Because it satisfies the time homogeneity one can write the Chapman-Kolmogorov equations because this is the Markov branching process.


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Theorem 4: PGF of $Z(t)$...

► Now

$$\begin{aligned}
 [\phi(t+v; s)]^i &= \sum_j P_{ij}(t+v) s^j \\
 &= \sum_j \sum_k P_{ik}(t) P_{kj}(v) s^j \\
 &= \sum_k P_{ik}(t) \sum_j P_{kj}(v) s^j \\
 &= \sum_k P_{ik}(t) [\phi(v; s)]^k \\
 &= [\phi(t; \phi(v; s))]^i
 \end{aligned}$$

► When $i = 1$, we obtain the result.



Now, the probability generating function of the time t plus v the whole power i , that is same as the summation j P_{ij} of t plus v power j . So, using Chapman-Kolmogorov equation you can write the P_{ij} of t plus v is summation over k P_{ik} of t P_{kj} of v , that is same as summation over k P_{ik} of t summation over j P_{kj} of v s power j . We know, that this is nothing but the probability generating function for P_{kj} of v , that is same as the probability generating function of v , s power k . This will be written as the probability generating function of ψ of t comma ψ of v , s the whole power i . When you substitute i is equal to 1 you get the result because ψ of t plus (v, s) is same as ψ of t comma ψ of v , s . When you substitute i is equal to 1 in this equation you will get, ψ of t plus (v, s) is same as ψ of t comma ψ of v , s the whole power i . When i is equal to 1 we will get ψ of t plus (v, s) same as ψ of t comma ψ of v , s .


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Theorem 5

► Let $u(s) = \sum_k a_k s^k$. Then $\phi(t; s)$ satisfies:

1.
$$\frac{\partial \phi(t; s)}{\partial t} = \frac{\partial \phi(t; s)}{\partial s} u(s)$$
2.
$$\frac{\partial \phi(t; s)}{\partial t} = u(\phi(t; s))$$

with the initial condition

$$\phi(0; s) = \sum_j P_{ij}(0) s^j = s$$


Now, we will move into the theorem 5, which discuss the differential equation corresponding to the probability generating function for P_{ij} of t . Let u of s is equal to summation $a_k s^k$ power k summation over k . Then, the probability generating function for P_{ij} of t satisfies partial derivative of ϕ of t, s with respect to t is equal to partial derivative of ϕ of t, s with respect to s multiplied u s . And partial derivative of ϕ of t, s with respect to t is same as u of ϕ of t, s with the initial condition ϕ of $0, s$ is same as summation over j P_{ij} of $0 s^j$, that is nothing but s . So, the theorem 5 gives the partial differential equation and ordinary differential equations satisfied, by the partial, by probability generating function of P_{ij} of t .

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Theorem 5 ...


► **Proof:** We have

$$\begin{aligned} \phi(h; s) &= \sum_j P_{ij}(h) s^j \\ &= \sum_j (\delta_{1j} + a_j h + o(h)) s^j \\ &= s + h \sum_j a_j s^j + o(h) \\ &= s + hu(s) + o(h) \end{aligned}$$

► Now

$$\phi(t + h; s) = \phi(t; \phi(h; s)) = \phi(t; s + hu(s) + o(h))$$

► By Taylor's theorem, we expand the right-hand side with respect to the second variable



$$\phi(t + h; s) = \phi(t; s) + \frac{\partial \phi(t; s)}{\partial s} hu(s) + o(h)$$

Let see the proof. We start with the psi of h of s, psi of h, s; that is nothing but the summation over j P of i comma j of s P of i comma j of h s power j. Substitute P ij of h and simplify, you will get the first term will be s the second term will be h times summation over j a j s j, the second term will be order of h.

You know, that u of s is same as summation over j, a suffix j of s power j. Therefore, the probability generating function for P ij of h, s, that is nothing but s plus hu of s plus order of h. We know that by the theorem 4, psi of t plus (h, s) will be psi of t comma psi of h of s (h, s). So, substitute psi of (h, s) with s plus h of u s plus o, order of h. Therefore, this will be psi of t, s plus h of u s plus order of h. By Taylor's theorem we expand the right hand side with respect to the second derivative. Therefore, the right hand side will be psi of t, s, the second term will be partial derivative of psi with respect to s times h of u s h times u s plus order of h, all the other term vanishes.

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Theorem 5 ...

► Hence


$$\frac{\phi(t+h; s) - \phi(t; s)}{h} = \frac{\partial \phi(t; s)}{\partial s} u(s) + \frac{o(h)}{h}$$

► Taking $h \rightarrow 0^+$, we get

$$\frac{\partial \phi(t; s)}{\partial t} = \frac{\partial \phi(t; s)}{\partial s} u(s)$$

► This is a partial differential equation for the function of two variables $\phi(t; s)$ with the initial condition

$$\phi(0; s) = \sum_j P_{ij}(0) s^j = s$$



(()) divide by h and take psi of t, s in the left side, therefore the left hand side becomes psi of t plus (h, s) minus i of t, s divided by h, whereas in the right hand side will be partial derivative of psi with respect to s times u of s order of h divided by h.

Taking h tends to 0 positive, we get the partial differential equation dou psi divided by dou t is equal to dou psi by dou s times u s. This is the partial differential equation for the function of two variables psi (t, s) with the initial condition psi of 0, s is s. So, we have proved the first part of theorem 5. Similarly, one can prove the second part of theorem 5.

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Theorem 5 ...


► **Proof of Part 2** Consider

$$\phi(v + h; s) = \phi(v; \phi(h; s))$$

► By Taylor's theorem, we get

$$\phi(v + h; s) = \phi(v; s) + hu(\phi(v; s)) + o(h)$$

► Hence

$$\frac{\phi(v + h; s) - \phi(v; s)}{h} = u(\phi(v; s)) + \frac{o(h)}{h}$$


The proof of part two, you start with partial differential, partial, we start with the probability generating function ψ of v plus (h, s) is same as ψ of v comma ψ of h, s . By Taylor's theorem, the right hand side becomes ψ of v, s plus h of $u \psi$ of v, s plus order of h , then take ψ of v, s in the left hand side, divide throughout by h you will get this equation.


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Theorem 5 ...

► Taking $h \rightarrow 0^+$, then substituting $v = t$, we get

$$\frac{\partial \phi(t; s)}{\partial t} = u(\phi(t; s))$$

► This is an ordinary differential equation with initial condition

$$\phi(0; s) = \sum_j P_{ij}(0) s^j = s$$


Now, limit h tends to 0 plus and then substitutes v is equal to t . In this equation limit h tends to 0 plus and substitute v is equal to h v is equal to t , we get partial derivative of ψ with

respect to t is equal to u of ψ of t, s . This is the ordinary differential equation with the initial condition ψ of $0, s$ equal to s .

So, in the theorem 5 we conclude, the partial, the probability generating function satisfies the partial differential equation and the initial and ordinary differential equations with the initial conditions ψ of $0, s$ is equal to s .

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Mean of $Z(t)$

- ▶ Consider
$$\frac{\partial \phi(t; s)}{\partial t} = \frac{\partial \phi(t; s)}{\partial s} u(s)$$
- ▶ By differentiating with respect to s and interchanging the order of differentiation on the left side, we get
$$\frac{\partial^2 \phi(t; s)}{\partial t \partial s} = u(s) \frac{\partial^2 \phi(t; s)}{\partial s^2} + u'(s) \frac{\partial \phi(t; s)}{\partial s}$$

NPTEL

Now, you will find out the mean of Z of t . You start with the partial differential equation satisfied by probability generating function. By differentiating with respect to s and interchanging the order of differentiation on the left hand side we get, the left hand side is the second order partial derivative of ψ with respect to t . And with respect to s the right hand side, u of s second order partial derivative of ψ with respect to s u dash of s partial derivative of ψ with respect to s .

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Mean of $Z(t)$...


► If $s = 1$, $u(1) = \sum_k a_k = 0$, and then

$$\frac{\partial m(t)}{\partial t} = u'(1)m(t)$$

where

$$m(t) = E[Z(t)] = \left. \frac{\partial \phi(t; s)}{\partial s} \right|_{s=1}$$

► But, since $Z(0) = 1$, then $m(0) = 1$ and the solution is


$$m(t) = e^{u'(1)t}$$


If s is equal to 1, you know, that $u(1)$ will be 0. Suppose the $m(t)$ will be the mean of $Z(t)$ that is nothing but the partial derivative of ϕ with respect to s . Then, substitute s is equal to 1, therefore this equation becomes partial derivative of $m(t)$ with respect to t is equal to $u'(1)m(t)$. Since m is with the single variable, so this is the ordinary differential equation. So, $\frac{dm(t)}{dt}$ is equal to $u'(1)m(t)$, where $m(t)$ is a mean of $Z(t)$. But since $Z(0)$ is equal to 1, $m(0)$ also 1, therefore you can solve this ordinary differential equation with the initial condition $m(0)$ is equal to 1. Hence, the solution will be $m(t)$ is equal to $e^{u'(1)t}$.

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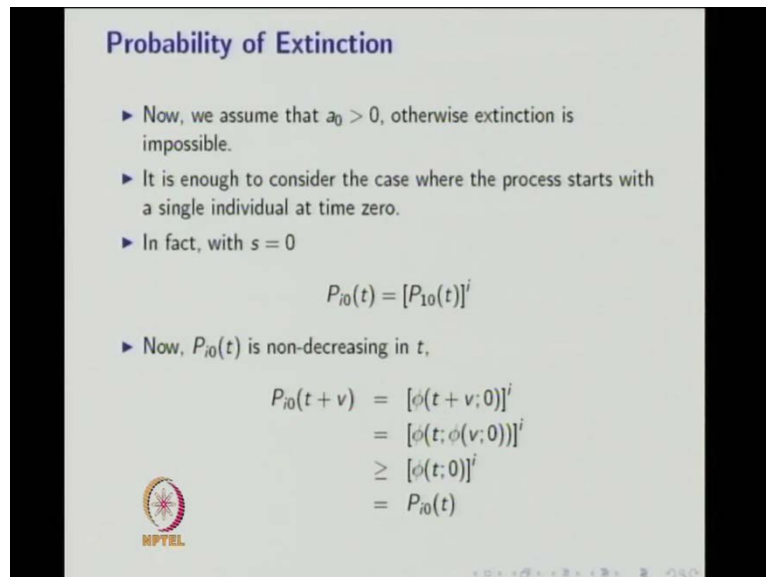
Definition

► The extinction probability, q , is defined by

$$q = \lim_{t \rightarrow \infty} P_{10}(t)$$


Now, we can discuss the mean of Z of t based on the value of u dash of 1. Before that we discuss the probability of extinction that is defined by q , that is nothing but limit t tends to infinity probability of 1 comma 0 of t . This is called a probability of extinction that is denoted by the letter q .

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Probability of Extinction

- ▶ Now, we assume that $a_0 > 0$, otherwise extinction is impossible.
- ▶ It is enough to consider the case where the process starts with a single individual at time zero.
- ▶ In fact, with $s = 0$

$$P_{i0}(t) = [P_{10}(t)]^i$$

- ▶ Now, $P_{i0}(t)$ is non-decreasing in t ,

$$\begin{aligned} P_{i0}(t + v) &= [\phi(t + v; 0)]^i \\ &= [\phi(t; \phi(v; 0))]^i \\ &\geq [\phi(t; 0)]^i \\ &= P_{i0}(t) \end{aligned}$$



Now, we will try to find out the probability of extinction small q . Assume, that a naught is strictly greater than 0, otherwise extinction is impossible. It is enough to consider the case where the process starts with the single individual at time 0 that means Z of 0 is equal to 1. With s equal to 0 you will get P i of 0 of t , that is nothing but P 10 of t power i in the probability generating function of P ij of t .

Now, we will prove, that p i 0 of t is a non-decreasing in t . You start with P i 0 of t plus v , that is nothing but ψ of t plus $(v, 0)$ of power i , that is same as a ψ of t comma ψ of $v, 0$ power i , that will be greater than or equal to ψ of 0, t power i , but that is same as P i 0 of t . Hence, we proved P i 0 of t is non-decreasing in t .

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Derivation of Probability of Extinction

- ▶ Let t_0 be a fixed positive number.
- ▶ Consider the discrete time process

$$Z(0), Z(t_0), Z(2t_0), \dots, Z(nt_0), \dots$$

where $Z(t)$ is the population size at time t .

- ▶ Assume that, the population size at $t = 0$ is one.
- ▶ Since $Z(t)$ is assumed to be a Markov process, the discrete process $Y_n = Z(nt_0)$ will be a Markov chain which is also a discrete time branching process.
- ▶ By the hypothesis of homogeneity of the probability function of $Z(t)$ and



$$\sum_j P_{ij}(t) s^j = \left[\sum_j P_{1j}(t) s^j \right]^i = [\phi(t; s)]^i$$

Let t be the fixed positive number. Consider a discrete time branching process Z of 0, Z of t naught, Z of 2 times t naught and so on, Z of n times t naught, where Z of t is a population size at time t . Assume, that the population size at time 0 is 1 , Z of 0 is equal to 1 . Since Z of t is assumed to be Markov process, the discrete process Y_n , Y suffix n , that is nothing but Z of n of t naught will be a discrete time Markov chain, which is also a discrete time branching process because Z of t is a continuous time branching process. Therefore, Y of n will form a discrete time branching process, which is also a discrete time Markov chain.

By the hypothesis of homogeneity of a probability function of Z of t and the probability generating function of P_{ij} of t , that is nothing but probability generating function of P_{ij} of t power P_1 of $1j$ of t power i . This we have proved it in the earlier; we have proved it in earlier.


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Derivation of Probability of Extinction ...

► We obtain

$$\begin{aligned}
 \sum_k P(Y_{n+1} = k \mid Y_n = i) s^k &= E \left[S^{Y_{n+1}} \mid Y_n = i \right] \\
 &= E \left[S^{Z((n+1)t_0)} \mid Z(nt_0) = i \right] \\
 &= E \left[S^{Z(t_0)} \mid Z(0) = i \right] \\
 &= [\phi(t_0; s)]^i \\
 &= \left[E \left[S^{Z(t_0)} \mid Z(0) = 1 \right] \right]^i \\
 &= \left[E \left[S^{Y_1} \mid Y_0 = 1 \right] \right]^i
 \end{aligned}$$

► This shows that $\{Y_n, n = 0, 1, \dots\}$ is a branching process.



Therefore, using these two we are finding summation over k. The conditional probability of Y_{n+1} is equal to k given Y_n is equal to i multiplied by s power k, that is nothing but expectation of $S^{Y_{n+1}}$ given Y_n is equal to i. That is same as, because Y_n is nothing but Z of n times t_0 , therefore Y_{n+1} is nothing but Z of n plus 1 times t_0 .

We replace Y_n by Z of n times t_0 and Y_{n+1} by Z of n plus 1 times t_0 . This is true for all n, therefore that is same as expectation of $S^{Z(t_0)}$ given $Z(0)$ is equal to I, but that is nothing but $\phi(t_0; s)$ power i, that can be written as expectation of $S^{Z(t_0)}$ given $Z(0)$ is equal to 1 whole power i. That is same as expectation of S^{Y_1} given Y_0 is equal to 1 the whole power i. This shows that Y_n is the branching process; Y_n is a discrete time branching process. So, using these we have proved the Y_n is a discrete time branching processes.


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Derivation of Probability of Extinction . . .

- ▶ The probability generating function of the number of offspring of a single individual in this process is $\phi(t_0; s)$.
- ▶ By Theorem 3, we know that the probability of extinction for the Y_n process is the smallest non-negative root of the equation

$$\phi(t_0; s) = s$$
- ▶ But

$$\begin{aligned}
 P(Y_n = 0 \text{ for some } n) &= \lim_{n \rightarrow \infty} P(Y_n = 0) \\
 &= \lim_{n \rightarrow \infty} P(Z(nt_0) = 0) \\
 &= \lim_{t \rightarrow \infty} P(Z(t) = 0) \\
 &= q
 \end{aligned}$$



The probability generating function for the number of upstream of a single individual in this process is $\psi(t_0; s)$. By theorem 3 we know, that the probability of extinction for Y_n , that is, a discrete time branching process is the smallest non-negative root of the equation $\psi(t_0; s) = s$.


So, by using the theorem 3 we conclude, the probability of extinction for the Y_n process is the smallest non-negative root of the equation $\psi(Z_0; s) = s$. But we know that probability of Y_n is equal to 0 for some n , that is same as limit n tends to infinity of probability of Y_n is equal to 0; that is same as limit n tends to infinity of probability of Z_n times z_0 is equal to 0, but that is same as limit t tends to infinity of probability $Z(t)$ is equal to 0. By definition this is nothing, but small q , that is a probability of extinction.

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Derivation of Probability of Extinction ...

- ▶ The extinction probability q of the continuous time branching process $Z(t)$ is the smallest non-negative root of the equation

$$\phi(t_0; s) = s$$
 where t_0 is any positive number.
- ▶ Hence, we expect that we should also be able to calculate q from an equation that does not depend on time.



Hence, the probability of extinction q of a continuous time branching process Z of t is the smallest non-negative root of the equation $\phi(t_0; s) = s$. Here, we have to conclude by theorem 3 the probability of extinction for the discrete time branching process. Y_n is the smallest non-negative root of the equation $\psi(t_0; s) = s$. Because of this we conclude the probability of extinction of the continuous time branching process. Z of t is the smallest non-negative root of the equation $\phi(t_0; s) = s$ where t_0 is any positive number.

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Theorem 6: Probability of Extinction


- ▶ The probability of extinction q is the smallest non-negative root of the equation

$$u(s) = 0$$
 Hence, $q = 1$ if and only if $u'(1) \leq 0$.
- ▶ **Proof:** Since q satisfies

$$\phi(t_0; s) = s$$
 for any $t_0 > 0$, we have

$$\frac{\phi(v+h; s) - \phi(v; s)}{h} = u(\phi(v; s)) + \frac{o(h)}{h}$$
- ▶ If $s = q$, then

$$\frac{q - q}{h} = u(q) + \frac{o(h)}{h}$$



Hence, we expect, that we should be, we should also be able to calculate q from equation, that does not depends on time. From this equation can able to calculate q from the equation, does not depend on time.

Now, we are moving into theorem 6, how to find the probability of extinction. We conclude, q is the probability of extinction for the continuous time branching process. So, here in this theorem we are giving the probability of extinction q is the smallest non-negative root of the equation u of s equal to 0. Hence, q is equal to 1 depend only if q dash is lesser than or equal to 0. So, whenever q dash of 1 is less than or equal to 0, then the probability of extinction will be sure, probability will be 1; extinction event will be sure, the probability of extinction will be 1.

Now, we give the proof of probability of extinction. In the earlier theorem we have concluded, q satisfies ψ of t naught comma s is equal to s for any t naught greater than 0. We have, this relation we have in the theorem 4. The theorem 4 discusses, theorem 5 discusses the partial differential equations and ordinary differential equation satisfies by ψ of t comma s .

So, we are using these equations to find the probability of extinction. So, here by using theorem 5, ψ of v plus (h, s) minus ψ of v , s divided by h will be u of ψ of v , s plus order of h by h . We know, that this will be tend to 0 has a h tends to 0 if s equal to q . Thus, q is the probability of extinction in the smallest non-negative root of the equation ψ of t naught comma s equal to s .

Therefore, if you substitute s is equal to q here, then above equation becomes the, left hand side become 0 when you put s is equal to q here, then the ψ of v , q will be q . Therefore, this will be u of q plus order of h divided by h . By substituting s equal to q in the above equation the left hand side becomes 0, the right hand side, first term, ψ of v , q will be q . Therefore, it will be u of q plus order of h divided by h .

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Theorem 6: Probability of Extinction ...

- Hence, for any $h > 0$,

$$0 = u(q) + \frac{o(h)}{h}$$


- Taking $h \rightarrow 0^+$, we obtain

$$u(q) = 0$$

- Since

$$u''(s) = \sum_{k=1}^{\infty} a_k k(k-1)s^{k-2} \geq 0$$

- $u(s)$ is convex in the interval $[0, 1]$.




Hence, for any h greater than 0 it will be, 0 is equal to u of q plus order of h divided by h . As h tends to 0 plus you will get, u of q will be 0. Therefore, the earlier theorem we have concluded the probability of extinction will be ψ of t naught comma s equal to 0 where q will be the probability of the extinction is the smallest non-negative root of the equation. But here, by using this we concluded u of q is equal to 0. Hence, the probability of extinction q is the smallest non-negative root of the equation, u of s is equal to 0 because we concluded, u of q is equal to 0.

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Theorem 6: Probability of Extinction ...

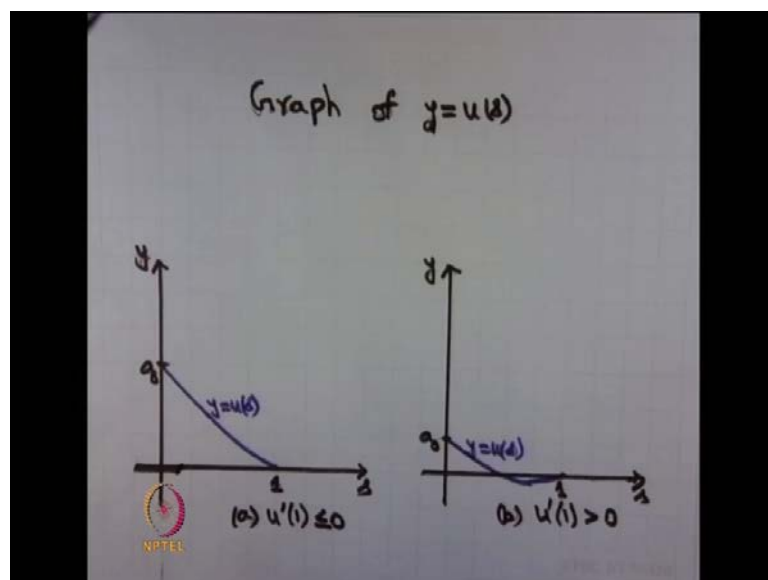
- As $u(1) = 0$ and $u(0) = a_0 > 0$, $u(s)$ may have at most one zero in the interval $(0, 1)$.
- According to whether $u'(1) \leq 0$ or $u'(1) > 0$ holds, we have the cases $q = 1$ or $q < 1$ respectively.
- Note that $E[Z(t_0)] = E[Y]$ if and only if $u'(1) > 0$.
- This means that for the discrete time branching process $Z(nt_0)$, $n = 0, 1, 2, \dots$ ($t_0 > 0$ fixed), extinction occurs with probability less than one, and therefore the same is true for the process $Z(t)$.
- The probability of extinction q is in this case necessarily the smallest zero of $u(s)$ in $[0, 1]$.



Suppose you find the double derivative of u that will be greater than or equal to 0. Hence, we conclude u of s is a convex function in the interval $(0, 1)$. As u of 1 is equal to 0 and u of 0 is equal to a naught, which is greater than 0, u of s may have at most one 0 in the interval $(0, 1)$. The way we defined u of s , u of s is the summation $a_k s^k$, therefore u of 1 will be 0 and u naught will be a naught, which is greater than 0. With that assumption only the probability of extinction is possible.

According to whether u double dash is less than or equal to 0 or greater than 0, we have the case q is equal to 0 or q is less than 1 respectively. That means, when u dash is less than or equal to u dash of 1 is less than or equal to 0, the probability of extinction will be 1. The u dash of 1 is the greater than 0, then the probability of extinction will be less than 1.

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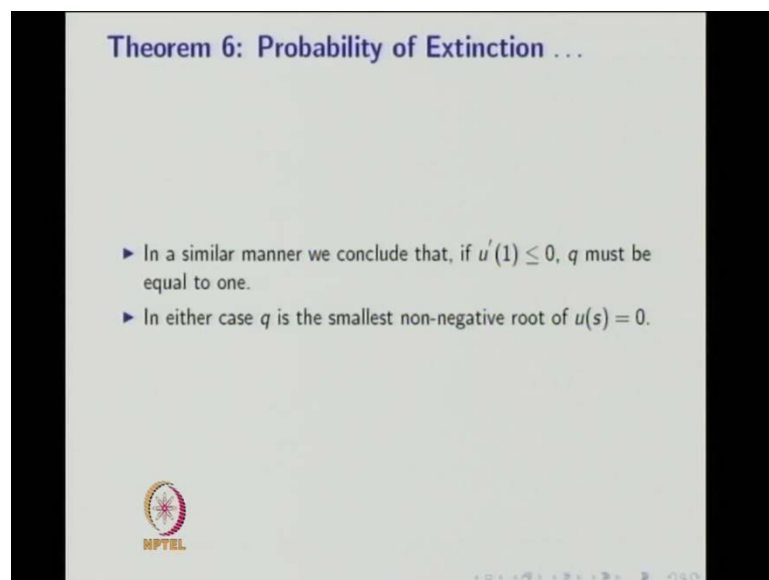
So, graphically one can show, this is a graph of y is equal to u of s . So, here we have two graphs, the graph A related to u dash of 1 is less than or equal to 0. Since u of s is the convex function in the interval 0 to 1 and u naught is a naught u of 1 will be 0. So, this is the graphical representation of y is equal to u of s .

Then, the case u dash of 1 is less than or equal to 0. In the case two, then u dash of 1 is greater than 0, the y is equal to the u of s will cut the x -axis at some point, which is less than 1 because u naught is a naught u of 1 is equal to 0 and u of s is the convex function. u dash of 1 is greater than 0, the u of s will cut in the x -axis before 1. Hence, the probability of extinction

when $u'(1)$ is less than or equal to 0, that will be 1 and the probability of extinction, then $u'(1)$ is greater than 0, it will be less than 1.

Note, that expectation of z naught, expectation of z of t naught is equal to the expectation of y , depend only $u'(1)$ is strictly greater than 0. So, whenever $u'(1)$ is strictly greater than 0, the probability of extinction is less than 1. This means, that for discrete time branching process Z of n times t naught extinction occurs with the probability less than 1 and therefore, the same is true for the process Z of t . The probability of extinction q is in the case necessarily the smallest 0 of u of s in 0 to 1.

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In a similar manner we conclude, that if $u'(1)$ is less than or equal to 0, q must be equal to 1. In either case q is the smallest non-negative root of u of s equal to 0. So, hence the probability of extinction q is the smallest non-negative root of the equation u of s equal to 0. When q is equal to, when $u'(1)$ is less than or equal to 0 the probability of extinction will be 1.

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Limit Theorem

1. If $u'(1) = 0$ and $u''(1) < \infty$, then

$$P(Z(t) > 0 \mid Z(0) = 1) \sim \frac{2}{tu''(1)}, \quad t \rightarrow \infty$$


and

$$\lim_{t \rightarrow \infty} P\left(\frac{2Z(t)}{tu''(1)} > \lambda \mid Z(t) \neq 0\right) = e^{-\lambda}, \quad \lambda > 0$$

2. If $u'(1) > 0$ and $u''(1) < \infty$, then

$$\frac{Z(t)}{e^{tu'(1)}}$$

has a limit distribution as $t \rightarrow \infty$.




Now, we will consider the limit theorem, if u' of 1 is equal to 0 and u'' of 1 is finite. Then, we can show this conditional probability will be approximately 2 divided by t times u'' of t as t tends to infinity. And also, we can conclude the limit t tends to infinity, probability of this event is $e^{-\lambda}$ where λ is strictly greater than 0. When u' of 1 is strictly greater than 0 and u'' of 1 is finite, then the Z of t divided by $e^{tu'}$ of 1 has a limit distribution as t tends to infinity. Without proof we are stating this limit theorem.

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Bellman-Harris Processes

- ▶ Consider a classical branching process in which progeny are born at the moment of the parents death.
- ▶ Let $Z(t)$ be the number of particles alive at time t .
- ▶ The distribution of particle lifetime τ is an arbitrary non-negative random variable, the resulting process is called an "age-dependent" or Bellman - Harris process.
- ▶ Assume that all particles reproduce and die independently of each other.
- ▶ This model generalizes the birth and death process in two respects: first, the lifespan of individual particles need not have the exponential distribution, and second, more than one particle can be born.

The process $Z(t)$ is not a Markov process and its analysis is usually done by using renewal theory.



Till now we have discussed the two important branching processes: the first one is Galton-Watson discrete time branching process, the second one is a Markov branching process, which is a continuous type branching process. Now, we are going to discuss some more or some other important branching processes. The first one is Bellman-Harris processes. Consider classical branching processes in which the progeny are born at the moment of parent's death.


Let $Z(t)$ be the number of particles alive at time t . The distribution of a particle lifetime τ is an arbitrary non-negative random variable, the resulting process is called age-dependent or Bellman-Harris processes. So, in the Markov branching process the random variable τ , which is exponential distribution, that here is an arbitrary non-negative random variable, then the resulting process is the age dependent or Bellman-Harris process. So, when τ becomes exponential distribution, then age dependent Bellman-Harris process becomes Markov branching process.

Assume that all particles reproduce and die independently of each other. In the similar assumption, we have taken care in the discrete type, as well as, continuous type branching processes. This model generalizes the birth, death process in two aspects: the first, the lifespan of individual particles need not have the exponential distribution and second, more than one particle can born. Because of these two aspects this model generalizes the birth, death process. The process Z of t is not a Markov process and it is analyzes, usually done by using renewal theory, we have discuss renewal processes in module 8.

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Bellman-Harris Processes with Disasters

- ▶ Consider the population model which follows a Bellman - Harris process.
- ▶ At random times, disasters beset the population and each particle alive at the time of the disaster survives with probability p .
- ▶ The survival of any particle is assumed independent of the survival of any other particle.
- ▶ Measure of interest is limiting behavior when extinction does not occur.




Now, we discuss the Bellman-Harris processes with disasters. Consider the population model, which follows a Bellman-Harris process. At random times, disasters beset the population and each particle are alive at the time of disasters survives with the probability p . The survival of any particle is assumed independent of survival of any other particle. In this model, the measure of interests is limiting behavior when extinction does not occur.

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Bellman - Harris Processes with Immigration

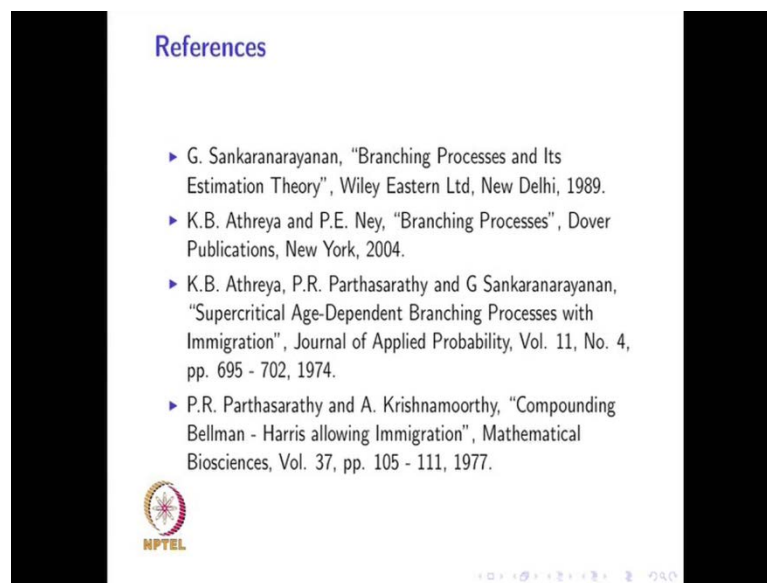
- ▶ Consider a Bellman - Harris process $\{Z(t), t \geq 0\}$ where $Z(t)$ denotes the population size at time t .
- ▶ In addition, we allow the sudden appearance into the system of newly born particles called immigrants.
- ▶ Immigrants are assumed to arrive in groups of various sizes with the probability of n immigrants in a group immigrating at time t given by $p_n(t)$, $n = 1, 2, \dots$
- ▶ Once these particles arrive, they reproduce and die according to the Bellman-Harris process.
- ▶ Measures of interest are the mean, the limiting distribution and the asymptotic behavior.
- ▶ This branching process is widely used to describe growth and of biological populations.



We move into the other important branching process, that is, Bellman-Harris processes with immigrants. In addition to the Bellman-Harris process we allow a sudden appearance into the

system of newly born particles called immigrants. Immigrants are assumed to arrive in group of various sizes. The probability of n immigrants in a group immigrating at time t given by p_n of t . Once these particles arrive, they reproduce and die according to the Bellman-Harris process. In this model measures of interest are the mean of Z of t , the limiting distribution and asymptotic behavior of Z of t . This process is widely used to describe growth and decay of biological populations.

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We are not discussing in detail of Bellman-Harris process in this lecture. In this module we have discussed in detail two important branching processes, Galton Watson process and Markov branching process. We have briefed Bellman-Harris process with disasters and with immigration. In the first two branching processes we have discussed mean and variance of Z of t , limiting distribution probability of extinction in both branching processes. Here are the references for this module 9 branching processes.