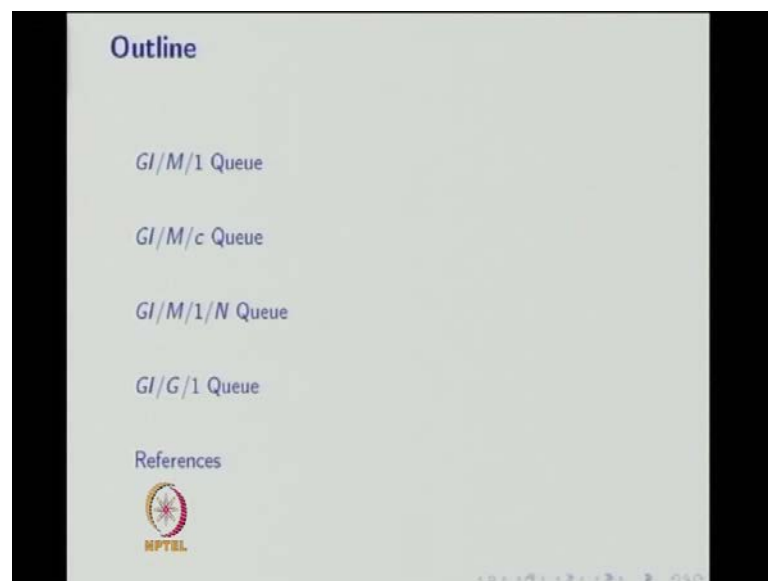


Stochastic Processes
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Module - 8
Renewal Processes
Lecture - 5
Non Markovian Queues

This is a stochastic processes module 8, Renewal process. In the lecture 1, we have discussed the renewal function and renewal equation. In the lecture 2, we have discussed the generalize renewal processes and renewal limit theorems. In lecture 3, we have covered Markov renewal and regenerative processes. In lecture 4, we have discussed non Markovian queues such as $M/G/1$ queue, $M/G/1$ and queue, $M/G/C/C((\infty))$ systems.

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This is a lecture-5, non Markovian queues, in this lecture we are going to discuss $GI/M/1$ Queue, $GI/M/c$ Queue then $GI/M/1/N$ Queue and finally, $GI/G/1$ Queue. What is $GI/M/1$ Queue? It means that the inter arrival time follows non-exponential distribution, which are independent. Therefore, the GI some books, they use only G as a notation; M stands for the service time is a exponential distribution only one server in the system with infinite capacity. So, consider the customers arrive at a time points t_0, t_1, t_2 and so on. Let Z_n is equal to $t_{n+1} - t_n$; be the i.i.d.; random variables with the distribution function with the CDFA with the mean $1/\lambda$.

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GI/M/1 Queue

- ▶ Consider the queueing model GI/M/1.
- ▶ Customers arrive at time points $0 = t_0, t_1, t_2, \dots$
- ▶ Let $Z_n = t_{n+1} - t_n, n = 1, 2, \dots$ be i.i.d. random variables with distribution function $A(\cdot)$ with mean $1/\lambda$.
- ▶ Let the service time distribution be exponential with mean $1/\mu$.
- ▶ Let $Q(t)$ be the number of customers in the system at time t




Therefore, as a special case if you assume that the inter arrival time is exponential distribution with the mean, 1 by λ . Then, the arrival follows a poisson process with the parameter λ . But in this G/M/1 model the inter arrival time is non exponential distribution with the CDF function, A with the mean 1 by λ . Let the service time distribution be exponential with the mean 1 by μ . Let $Q(t)$ be the number of customers in the system at time t .

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GI/M/1 Queue ...

- ▶ Define $Q(t_n - 0) = Q_n, n = 1, 2, \dots$
- ▶ Thus Q_n is the number in the system just before the n th arrival.
- ▶ Then $\{(Q_n, t_n), n = 0, 1, \dots\}$ is a Markov renewal process, where t_n is the instant when the n th customer arrives and $Q_n = Q(t_n - 0)$.
- ▶ Define X_n as the number of potential service completions during the inter-arrival period Z_n .
- ▶ Let $\{b_j, j = 0, 1, \dots\}$ be the distribution of X_n .
- ▶ Then,

$$b_j = P(X_n = j) = \int_0^\infty \frac{e^{-\mu t} (\mu t)^j}{j!} dA(t)$$


So, $Q(t)$ for t greater than or equal to 0 is a stochastic process. Since, the $Q(t)$ is the number of customers in the system at any time t . Therefore, the corresponding stochastic process is a discrete state continuous time stochastic process. So, the underlying stochastic process in the G/M/1 Queue is a $Q(t)$, t greater than or equal to 0.

Now, define $Q(t_n - 0)$ has a Q_n . Thus, Q_n is a number of customers in the system just before the n th arrival. Therefore, the Q_n for n is equal to 1, 2 and so on. This follows a discrete state, discrete time stochastic process. So, this is an embedded stochastic process from $Q(t)$, in the $Q(t)$ is a discrete state continuous time stochastic process. Whereas, $Q(n)$ is a discrete state discrete time stochastic process because the Q_n is the number of customers in the system just before the n th arrival. The bivariate random variables Q_n, t_n for different values of n forms a Markov renewal process; where, t_n is the instant when the n th customer arrives and Q_n is defined $Q(t_n - 0)$.

Since, inter arrival time is a non exponential distribution, with the mean $1/\lambda$ and the service time is exponential distribution with the mean $1/\mu$. Single server in the system and infinite capacity. Therefore, the Q_n, t_n form a Markov renewal process and the t_n is a time instant in which the arrival occurs. Now, define the random variable X_n as a number of potential service completions during the inter-arrival period Z_n . Z_n is nothing but $t_{n+1} - t_n$; that is an inter arrival time. X_n is a number of potential service completions; during the inter arrival period Z_n and b_j be the distribution of X_n . Obviously, X_n is the discrete type random variable for fixed n and the b_j is a probability mass function for the random variable X_n .

Since, the number of potential service completion could be 0, 1 and so on. So, the probability mass function for different values of j , b_j is nothing but the probability of X_n is equal to j that is nothing but the integration 0 to infinity $e^{-\mu t}; \mu^j t^j$ divided by j factorial and the integration with respect to $A(t)$; where $A(t)$ is the distribution function of inter arrival time with the mean $1/\lambda$.


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GI/M/1 Queue ...

- ▶ The relationship between Q_n and Q_{n+1} is given by

$$Q_{n+1} = \begin{cases} Q_n + 1 - X_{n+1}, & Q_n + 1 - X_{n+1} > 0 \\ 0, & Q_n + 1 - X_{n+1} \leq 0 \end{cases}$$

- ▶ Q_{n+1} is independent of X_{n+1} and Q_{n+1} does not depend on any random variable with an earlier index parameter than n .
- ▶ $\{Q_n, n = 0, 1, \dots\}$ is a time homogeneous discrete time Markov chain.

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And, μ is the parameter for exponential distribution of service time. The way we defined Q_n is nothing but the number of customers in the system just before the n th arrival. We can relate the Q_{n+1} in terms of Q_n . We can find the relationship between Q_n and Q_{n+1} , that is Q_{n+1} will be $Q_n + 1$ minus X_{n+1} ; whenever $Q_n + 1$ minus X_{n+1} is greater than 0. Otherwise, it is 0, for $Q_n + 1$ minus X_{n+1} is less than 0. The reason is the number of customers in the system just before the n th arrival plus the $(n+1)$ th arrival is same as the number of customers in the system just before the $(n+1)$ th arrival plus the $(n+1)$ th customer who arrived that is plus 1 minus. How many customers are served during the inter arrival period? That is a X_{n+1} .

So, if you subtract Q_{n+1} the X_{n+1} ; you will get $Q_n + 1$. Whenever, $Q_n + 1$ minus X_{n+1} is greater than 0; if it is less than or equal to 0; then the number of customers will be again 0. Just before the $(n+1)$ th arrival also will be 0. We know that the X_{n+1} is independent of Q_{n+1} . Hence, the Q_{n+1} depends only on Q_n and Q_{n+1} is independent of X_{n+1} . Therefore, the Q_n forms a time homogeneous discrete time Markov chain. The Q_t is a discrete state continuous time stochastic process and the Q_n is the discrete time discrete state stochastic process.

Since, Q_{n+1} is equal to $Q_n + 1$ minus X_{n+1} or 0; and Q_{n+1} is independent of X_{n+1} ; as well as Q_{n+1} depends only on Q_n . Therefore, Q_n for n is equal to 0, 1, 2 and so on, form a time homogeneous. That means, time invariant and

we also satisfy the Markov property. Therefore, this discrete time discrete state stochastic process is called a discrete time Markov chain satisfying the time homogeneous property. Therefore, it is called a time homogeneous discrete time Markov chain.

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GI/M/1 Queue ...

► The transition probability is given by

$$p_{ij} = P(Q_{n+1} = j | Q_n = i)$$

$$= \begin{cases} P(X_{n+1} = i - j + 1), & j > 0 \\ P(X_{n+1} \geq i + 1), & j = 0. \end{cases}$$

► In matrix form,

$$P = \begin{bmatrix} \sum_{i=1}^{\infty} b_i & b_0 & 0 & 0 & \dots \\ \sum_{i=2}^{\infty} b_i & b_1 & b_0 & 0 & \dots \\ \sum_{i=3}^{\infty} b_i & b_2 & b_1 & b_0 & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$


Once, you know that Q_n is a discrete time Markov chain, you can find the one step transition probability matrix, whose elements are p_{ij} . So, the p_{ij} are nothing but what is the probability that Q_{n+1} will be j given that Q_n was i ; that is same as what is the probability that $i + j - 1$ customers are served during the inter arrival time period, that is probability of X_{n+1} is equal to $i - j + 1$. Whenever, j is greater than 0; if j is equal to 0 then, it is nothing but probability that X_{n+1} is equal to $i + 1$. In matrix form you can write it P as a matrix whose elements are p_{ij} .

So, the first element will be p_{11} plus p_{21} and so on, and the second element in the first row will be b_0 . Since, b_j is nothing but the probability mass function for the random variable X_n that is for all n , for all n , it is identically distributed. Therefore, the probability mass function of X_n is b_j and the running index of j is 0, 1 and so on. Therefore, if you make the row sum $b_0 + b_1 + b_2$ and so on, that will be 1, whereas in the second row the first element will be $b_1 + b_2$ and so on. The second row second element will be b_0 , second row third element will be b_0 and so on.

Substitute i and j in the above equation you substitute i and j accordingly, you will get the these values summation of b is starting from 2 and b 1, b 2, b 0 and so on.

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GI/M/1 Queue ...


- ▶ For the DTMC to be irreducible, $b_0 > 0$ and $b_0 + b_1 < 1$. We can easily determine that the DTMC is aperiodic.
- ▶ Let

$$\phi(\theta) = \int_0^\infty e^{-\theta t} dA(t), \quad \text{Re}(\theta) > 0$$
- ▶ The probability generating function of $\{b_j\}$ is obtained as

$$\beta(z) = \sum_{j=0}^{\infty} b_j z^j = \phi(\mu - \mu z), \quad |z| \leq 1$$
- ▶ Hence,

$$E(Z) = \beta'(1) = -\mu \phi'(0) = \frac{\mu}{\lambda}$$
- ▶ We define the traffic intensity

$$\rho = (\text{arrival rate}) / (\text{service rate})$$

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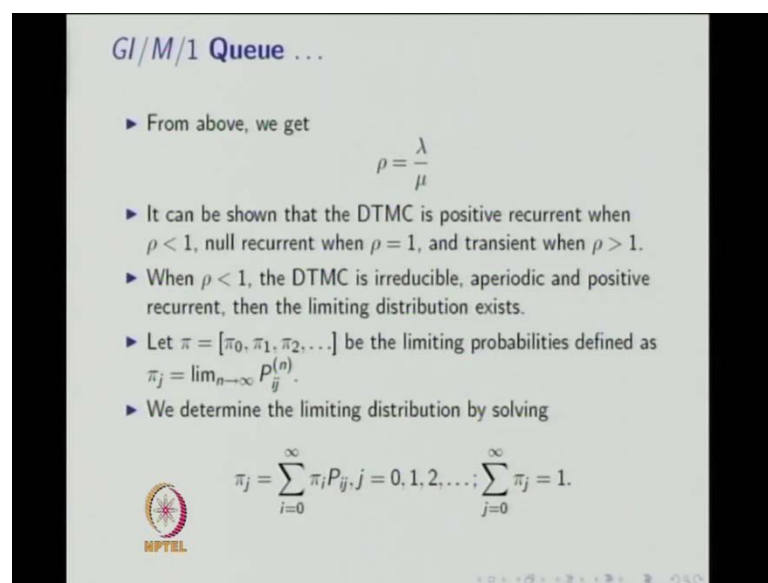
Similarly, you can get the 3rd row you can verify that each element of a b B matrix will be lies between 0 to 1 and the row sum will be 1. So basically stochastic matrix. These are the one step transition probability matrix P. Our interest is to find out the study state or limiting distribution. For that, we need a irreducible and positive recurrent and the aperiodic also. So, for the irreducible you need b 0 has to be greater than 0 as well as b 0 plus b 1 has to be less than 1. If this is satisfied then you will get the conclusion the given time homogeneous discrete time Markov chain will be a irreducible. That means, each state is a communicating with each other states with the condition of b 0 is greater than 0; b 0 plus b 1 is less than 1. We can easily determine that the DTMC is a periodic.

We have discussed the aperiodic in the discrete time Markov chain. So, we can verify that this discrete time Markov chain is a periodic also. That means, if the period is 1; now, we find out the Laplace transform of the CDF of inter arrival time distribution that is a phi naught. So, phi naught is equal to integration 0 to infinity e power minus theta times t d A (t); where A (t) is the CDF; of inter arrival time distribution with the real of theta as to be greater than 0. Now, we are finding the probability generating function for p j is that is a distribution of X n. So, that is beta (z); that the summation j is equal to 0 to infinity b j z power j; that you can write down in terms of Laplace transform. So, this

Laplace strange transform so, Laplace strange transform of $A(t)$ so, it is a psi of mu minus mu times z.

Now, we can find out the expectation of Z that is nothing but, the beta dash of 1; if you differentiate probability generating function then substitute Z is equal to 1; will be the mean number of arrivals that is Z. Mean number of inter-arrival time so, that is nothing but, the mu by lambda. Now we can define the traffic intensity that is nothing but rho, rho is equal to arrival rate divided by the service rate.


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GI/M/1 Queue ...

- ▶ From above, we get
$$\rho = \frac{\lambda}{\mu}$$
- ▶ It can be shown that the DTMC is positive recurrent when $\rho < 1$, null recurrent when $\rho = 1$, and transient when $\rho > 1$.
- ▶ When $\rho < 1$, the DTMC is irreducible, aperiodic and positive recurrent, then the limiting distribution exists.
- ▶ Let $\pi = [\pi_0, \pi_1, \pi_2, \dots]$ be the limiting probabilities defined as $\pi_j = \lim_{n \rightarrow \infty} P_{ij}^{(n)}$.
- ▶ We determine the limiting distribution by solving

$$\pi_j = \sum_{i=0}^{\infty} \pi_i P_{ij}, j = 0, 1, 2, \dots; \sum_{j=0}^{\infty} \pi_j = 1.$$



From, above you can get rho is nothing but lambda by mu; it can be shown that the DTMC is a positive recurrent when rho is less than 1. We have already made it has a irreducible and aperiodic. Now we are giving the condition for a positive recurrent when rho is less than 1 the given DTMC is a positive recurrent. If rho equal to 1 it is a null recurrent; if rho is greater than 1 then it will be a transient; that means when rho is less than 1; all the states are positive recurrent therefore, the DTMC is called a said to be a positive recurrent. Similarly, when rho is 1 all the states are null recurrent; therefore, the DTMC will be a null recurrent and similarly, for rho is greater than 1. So, our interest is to find out the limiting distribution so, when rho is less than 1 the DTMC is a irreducible aperiodic and positive recurrent. Along with the condition b_0 is greater than 0; and b_0 plus b_1 is less than 1; with this condition it is a irreducible, with rho is less than one; it is a positive recurrent.

We can easily verify it is a periodic. Hence, the limiting distribution exist and it is unique and that will be a independent of a initial state i therefore, P_{ij} will be a limit n tends to infinity; the probability of i to j in n steps. So, define the limiting distribution probability vector has a π_i it consists of π_0, π_1, π_2 and so on, where π_j is are define in this form; limit n tends to infinity P^n_{ij} in n steps. We determine the limiting distribution by solving $\pi_i = \pi_i P$ and the summation of π_i is equal to 1; so, the first one is a homogeneous equation. Including this normalizing condition, we will have a system of non-homogeneous equation. Solve this non-homogeneous, solve this system of non-homogeneous equation you get the limiting probabilities.

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GI/M/1 Queue ...

- We get

$$\pi_j = (1 - \zeta)\zeta^j, \quad j = 0, 1, 2, \dots$$
 where ζ is the unique root of the equation $z = \phi(\mu - \mu z)$ in the interval $(0, 1)$.
- Note that the queue length distribution found just before an arriving customer is geometric with parameter ζ .
- The best computational method for the determination of the limiting distribution seems to be the direct matrix multiplication to get P^n for increasing values of n until the rows can be considered to be reasonably identical.
- It is important to note that the imbedded Markov chain analysis gives the properties of the number in the system at arrival epochs.

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We get the limiting probabilities are $1 - \psi; \psi^j$; where ψ is the unique root of the equation; z is equal to the Laplace strange transform of the inter-arrival distribution with the variable $\mu - \mu z$; in the interval 0 to 1 . That means first you have to solve the equation z is equal to ψ of $\mu - \mu z$. And you know what is the ψ of $\mu - \mu z$ from the beta (z) solve the equation in the interval 0 to 1 . So, the unique root you have to substitute as a ψ ; then substitute ψ in the π_j ; and that will be the since the unique root is in the interval 0 to 1 . Therefore, the π_j $1 - \psi; \psi^j$; that will form a probability mass function for the limiting distribution.

Note that the Queue length distribution found just before the arriving customer is geometric distribution with the parameter ψ . The limiting distribution which we got it

that is just before the arriving customer. The Queue side, the Queue length distribution formed just before an arriving customer which is geometric distribution with the parameter ψ . The best computational method for determination of the limiting distribution since, to be the direct matrix multiplication to get P power n for increasing values of n ; until the rows can be consider to be reasonable identical. Whenever, for larger n , P power n has the identical rows it means a limiting distribution exist.


Therefore, the best computational method is find the P power n for a larger n until the rows can be consider to be a reasonable identical. It is important to note that the embedded Markov chain analysis gives the properties of the number of number in the system at arrival epochs. The $Q(t)$ for t greater than or equal to 0; that is a discrete state continuous time stochastic process whereas, a Q_n for n ; n is equal to 0, 1, 2 and so on.

That is a embedded time homogeneous discrete time Markov chain and that is Q_n is nothing but $Q(t)$ minus 0 that is nothing but number of customers in the system just before the n th arrival. Hence, it is important to note that the embedded Markov chain analysis gives the properties of the number in the system at arrival epochs not the departure epochs or not at the arbitrary time instance. It gives in the embedded Markov chain analysis gives the properties of the number in the system at only at the arrival epochs.

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GI/M/1 Queue ...

- ▶ As pointed out the under the discussion of the $M/G/1$ queue, the limiting distributions of the number of customers in the system at arrival epochs, at departure epochs, and at arbitrary points in time are the same only when the arrivals occur as a Poisson process.
- ▶ Prabhu [1] and Bhat [2] derived the results of limiting distribution at an arbitrary time t .
- ▶ Let $p_j = \lim_{t \rightarrow \infty} P(Q(t) = j)$ where $Q(t)$ is the number at an arbitrary time t .
- ▶ The limiting distribution $\{p_j, j = 0, 1, 2, \dots\}$ when $\rho < 1$

$$p_0 = 1 - \rho, \quad p_j = \rho(1 - \zeta)\zeta^{j-1}, j \geq 1$$



As pointed out under the discussion of M/G/1 Queue, the limiting distribution of the number of customers in the system at arrival epochs at departure epochs and at arbitrary points in time are the same only if arrivals occur as a Poisson process. So, in the M/G/1 Queue the limiting distribution of the number of customers in the system at arrival epochs at departure epochs, and at arbitrary time points are the same, because the arrival occurs in the Poisson process and inter-arrival follows independent exponential distribution with the same parameter. Whereas, in the GI/M/1 Queue model the embedded Markov chain results give the limiting distribution of the number of customers in the system at the arrival epochs only that is not the same as the limiting distribution at the departure epochs and that is also not the same as the arbitrary time points.

Now, we are finding, now we are going to discuss the limiting distribution at the arrival at the arbitrary time points. Prabhu and Bhat derived the results of limiting distributions. at arbitrary time t . Let p_j is a probability that limit t tends to infinity the probability Q_t is equal to j . So, this is nothing to do with the embedded Markov chain Q_n . We are finding limit t tends to infinity probability that Q_t is equal to j ; where Q_t is the number of customers in the system at arbitrary time t . The limiting distribution exists whenever, $\rho < 1$ and the probability that no customer in the system in the long run or the limiting in a long run that p_0 is equal to $1 - \rho$; and the p_j is equal to ρ times $1 - \rho$ for $j = 1, 2$ and so on. So, this is the limiting distribution at arbitrary time.

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GI/M/1 Queue ...

- ▶ To determine the distribution of the waiting time of a customer, we need the distribution of the number of customers in the system at the time of that customer's arrival.
- ▶ Replace ρ by ζ in M/M/1 results, since the queue length distribution is same.
- ▶ Waiting time distribution is given by

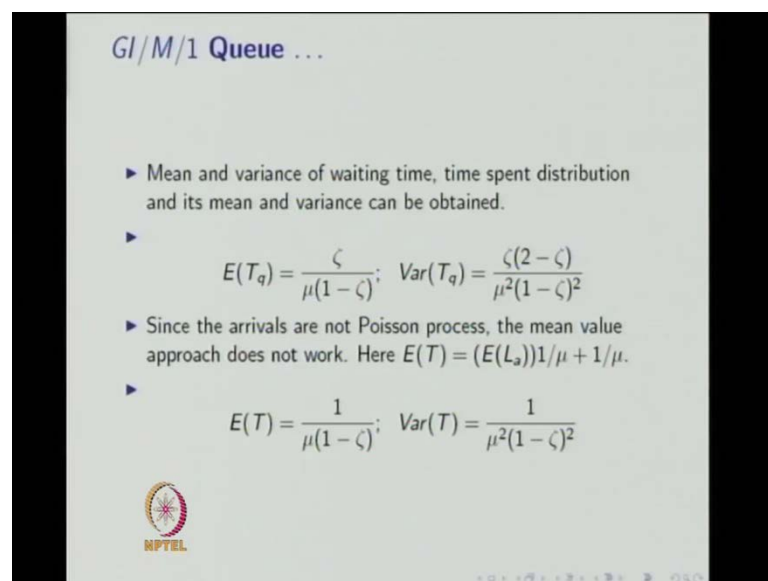
$$F_q(t) = P(T_q \leq t) = 1 - \zeta e^{-\mu(1-\zeta)t}$$


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To determine the distribution of waiting time of a customer we need a distribution of a number of customers in the system at the time of that customer arrives. If you want to find out the distribution of waiting time, replace rho by psi in the M/M/1 results, since the Queue length distribution is same. Our interest is to find out the waiting time distribution. Since, the limiting time distribution at arrival epochs is the same as the limiting distribution of M/M/1 Queue model therefore, you can replace rho in the M/M/1 results by psi to get Queue length distribution if to get the waiting time distribution.

Therefore, by replacing rho by psi in the M/M/1 results of waiting time distribution you get the waiting time distribution for GI/M/1 queue as a the CDF of the waiting time distribution is a $1 - \psi e^{-\mu(1-\psi)t}$ for $t \geq 0$; for $t < 0$ it will be 0. So, this is the waiting time distribution for GI/M/1 queueing system.

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GI/M/1 Queue ...

- ▶ Mean and variance of waiting time, time spent distribution and its mean and variance can be obtained.
- ▶
$$E(T_q) = \frac{\zeta}{\mu(1-\zeta)}; \quad \text{Var}(T_q) = \frac{\zeta(2-\zeta)}{\mu^2(1-\zeta)^2}$$
- ▶ Since the arrivals are not Poisson process, the mean value approach does not work. Here $E(T) = (E(L_a))1/\mu + 1/\mu$.
- ▶
$$E(T) = \frac{1}{\mu(1-\zeta)}; \quad \text{Var}(T) = \frac{1}{\mu^2(1-\zeta)^2}$$

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First we found the limiting distribution at the arrival epochs, next we find the limiting distribution, next we discuss the limiting distribution at the arbitrary time points from the Prabhu and Bhat results, then we have discussed the waiting time distribution, now we are discussing the moments of Queue size. The mean and the variance of waiting time, time spent is to distribution and it is a mean and variance can be obtained. Once, we know the limiting distribution at the arrival epochs as well as that we know the waiting time distribution, you can find the mean and variance of waiting time by adding mean

you can find the mean of time spent also. You cannot use the mean value approach because the arrivals are not Poisson process. So, you can find out the average time spent in the system that is $E(T)$ will be average number of customers seen by at the arrival epochs that is $E(L_a)$ multiplied by $1/\mu$ plus the average service time that is $1/\mu$ will give the average time spent in the system.

Since, the arrivals are not Poisson processes not Poisson process. You cannot use the mean value approach. By simplification you can get already we know what is the distribution of number of customers seen by seen at the arbitrary already we know the limiting distribution at the arrival epochs. So, we can find the mean from those results then multiplied by $1/\mu$ plus $1/\mu$ will give the average time spent in the system. Then, you can find the variance of time spent in the system also. The 1st this one the mean and variance of waiting time since you know the waiting time distribution you can find the mean and variance.

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D/M/1 Queue

- ▶ For D/M/1 queueing model,

$$\phi(\theta) = \int_0^\infty e^{-\theta t} dA(t) = e^{-\theta T^*}$$

for some T^* denoting the inter-arrival time.

- ▶ In this model, the limiting distribution is given by

$$\pi_j = (1 - \zeta)\zeta^j, \quad j = 0, 1, 2, \dots$$

where ζ is the root of the equation $z = e^{-\mu(1-z)}$.

Suppose the inter-arrival time is constant. Then the Laplace strange transform will be $e^{-\theta T}$; you can find the limiting distribution for the D/M/1 model also with the inter-arrival time is a constant then the corresponding queueing model is denoted by D/M/1 Queue. So, we can find the limiting distribution at the arrival epochs as $1 - \psi$; ψ^j ; where ψ is the root of the equation, ψ is equal to $e^{-\mu(1-\psi)}$.

minus μ times $1 - z$; because $\psi(\theta)$ is equal to $e^{-\theta}$. So, to get the root of the equation you have to solve z is equal to $\psi(\mu - \mu z)$.

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D/M/1 Queue

- ▶ For D/M/1 queueing model,

$$\phi(\theta) = \int_0^\infty e^{-\theta t} dA(t) = e^{-\theta}$$
- ▶ In this model, the limiting distribution is given by

$$\pi_j = (1 - \zeta)\zeta^j, \quad j = 0, 1, 2, \dots$$

where ζ is the root of the equation $z = e^{-\mu(1-z)}$.



You have to get the root of the equation z is equal to $\psi(\mu - \mu z)$; so, in the D/M/1 model $\psi(\theta)$ is equal to $e^{-\theta}$. Therefore, it will be z is equal to $e^{-\mu(1-z)}$. So, if you solve this equation you will get ψ ; that is a unique root between the interval 0 to 1; substitute it is here, so that is the limiting distribution of the arrival epochs of D/M/1 queueing model.

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GI/M/c Queue

- ▶ The imbedded Markov chain analysis of the queue GI/M/1 can be easily extended to the multiserver queue GI/M/c. Here, the transition probabilities P_{ij} are different.
- ▶ Consider three cases for the initial value i and the final value j :

$$i+1 \geq j \geq c; \quad i+1 \leq c \text{ and } j \leq c \quad \text{and} \quad i+1 > c \text{ but } j < c$$
- ▶ Case 1: $i+1 \geq j \geq c$

$$P_{ij} = \int_0^\infty e^{-c\mu t} \frac{(c\mu t)^{i+1-j}}{(i+1-j)!} dA(t)$$

This represents $i+1-j$ service completions during an inter-arrival period, when all c servers are busy.



Now, we move into the multiserver queues the embedded Markov chain analysis of the queue GI/M/1 can be easily extended to the multiserver queue GI/M/c, where c is the number of servers in the system. Since the number of servers are greater than or equal to 1. The one step transition probability matrix will be different now that means, each element P_{ij} will be different from the GI/M/1 corresponding one step transition probability, probabilities p_{ij} is so, here you can consider 3 cases for the initial value i and the final value j. So, corresponding to the each cases your p_{ij} can be obtained for the case 1: i plus 1 is greater than or equal to j greater than or equal to c; you will get p_{ij} in this form.

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GI/M/c Queue ...

► Case 2: $i + 1 \leq c$ and $j \leq c$

$$P_{ij} = \binom{i+1}{i+1-j} \int_0^\infty e^{-j\mu t} (1 - e^{-\mu t})^{i+1-j} dA(t)$$

This represents $i + 1 - j$ out of $i + 1$ customers complete service.

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This represents i plus 1 minus j; service completions during inter-arrival period when all c servers are busy. So, the case1 is related to all c servers are busy means, the number of customers in the system is greater than or equal to c that is a situation in which you are getting the one step transition probability for the embedded Markov chain. For the case 2 the i plus 1 is less than or equal to c whereas, j is lesser than or equal to j c also. These cases related to i plus 1 minus j out of i plus 1 customers complete the service.

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GI/M/c Queue ...

► Case 3: $i + 1 > c$ but $j < c$

$$P_{ij} = \int_{t=0}^{\infty} \int_{\tau=0}^t e^{-c\mu t} \frac{(c\mu t)^{i-c}}{(i-c)!} c\mu \binom{c}{c-j} e^{-j\mu(t-\tau)} (1 - e^{-\mu(t-\tau)})^{c-j} d\tau dA(t)$$

This represents $i + 1 - c$ customers complete service with rate $c\mu$, and then $c - j$ out of c customers complete service.

► Refer Gross and Harris [3] for the limiting distribution.



The third case is related to the i plus 1 greater than c as well as j is less than c ; this represents i plus 1 minus c customers complete service with the rates $c\mu$ then c minus j out of c ; customers complete the service. To get the complete limiting distribution at the arrival time the arrival epochs you can refer gross sceneries to get the complete result.

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GI/M/1/N Queue

► The queueing model $GI/M/1/N$ is similar to $GI/M/1$, but now the system capacity is restricted to N .

► The relation between Q_n and Q_{n+1} is given by

$$Q_{n+1} = \begin{cases} \min(Q_n + 1 - X_{n+1}, N) & Q_n + 1 - X_{n+1} > 0 \\ 0, & Q_n + 1 - X_{n+1} \leq 0 \end{cases}$$

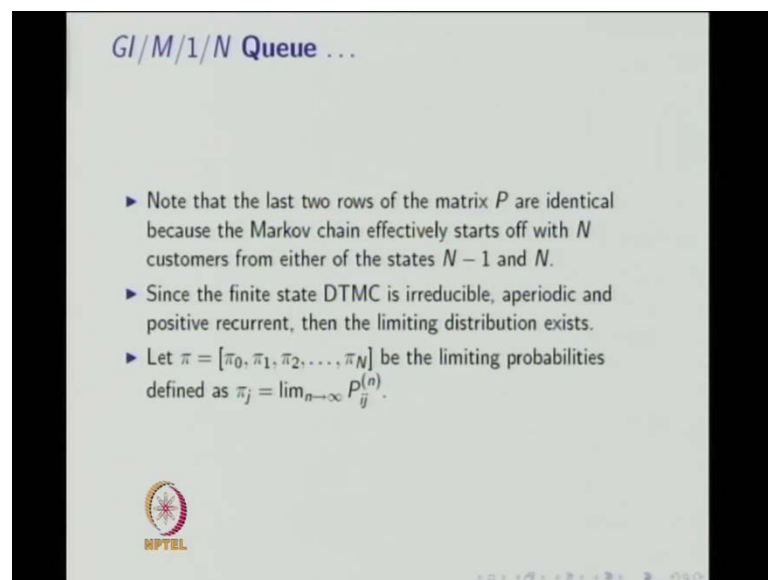
► The transition probability matrix P is given by,

$$P = \begin{bmatrix} \sum_{i=1}^{\infty} b_i & b_0 & 0 & \dots & 0 & 0 \\ \sum_{i=2}^{\infty} b_i & b_1 & b_0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ \sum_{i=N}^{\infty} b_i & b_{N-1} & b_{N-2} & \dots & b_1 & b_0 \\ \sum_{i=N}^{\infty} b_i & b_{N-1} & b_{N-2} & \dots & b_1 & b_0 \end{bmatrix}$$


So, here we have discussed only the one step transition probabilities how that is a different from one step transition probabilities of $GI/M/1$ model. Now, we move into $GI/M/1/N$ Queue; so, this queue is similar to $G/M/1$, but now the system capacity is


restricted to some finite number capital N . Therefore, the relations between Q_n and the Q_{n+1} will be a smaller change in it that is a Q_{n+1} will be minimum of $Q_n + 1$ minus X_{n+1} ; where X_{n+1} is the number of potential service completion during the inter-arrival time period. So, once you know Q_{n+1} in terms of Q_n . You can conclude since Q_{n+1} is in terms of only Q_n and Q_{n+1} is independent of X_{n+1} . Therefore, Q_n form a time homogeneous discrete time Markov chain and I can find the one step transition probability matrix P with the entities p_{ij} . Here also we can verify the row sum will be 1 and each b_i is nothing but the probability mass function of each b_i is are nothing but the distribution of X_n .

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GI/M/1/N Queue ...

- ▶ Note that the last two rows of the matrix P are identical because the Markov chain effectively starts off with N customers from either of the states $N - 1$ and N .
- ▶ Since the finite state DTMC is irreducible, aperiodic and positive recurrent, then the limiting distribution exists.
- ▶ Let $\pi = [\pi_0, \pi_1, \pi_2, \dots, \pi_N]$ be the limiting probabilities defined as $\pi_j = \lim_{n \rightarrow \infty} P_{ij}^{(n)}$.


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Once, you know the one step transition probability matrix we can go for finding the limiting probabilities. Note that the last 2 rows of the matrix P are identical because of the Markov chain effectively starts off with N customers from either of the states N minus 1 and N . Therefore, if you see the last two rows of the transition probability matrix both the rows will be identical because of the Markov chain effectively starts of with N customers from either of the states N minus 1 and N .

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GI/M/1/N Queue ...


► We determine the limiting distribution by solving

$$\pi_j = \sum_{i=0}^N \pi_i P_{ij}, j = 0, 1, 2, \dots; \sum_{j=0}^N \pi_j = 1.$$


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GI/G/1 Queue

- Consider a single-server queue with arbitrary arrival process and arbitrary service time distribution, with infinite buffer.
- Let S be a random variable representing the service time and let $E[S] = 1/\mu$, i.e., μ denotes the service rate.
- Let λ be the mean arrival rate.
- Assume that $\lambda < \mu$ so that the queue is stable.
- One can show that, for a stable GI/G/1 queue, utilization is given by λ/μ .

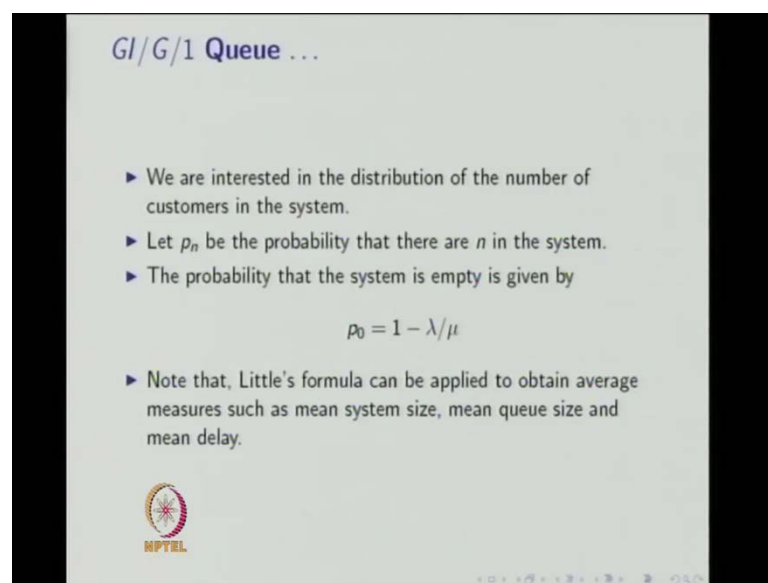


Since the finite state DTMC is irreducible a periodic and positive recurrent that the limiting probabilities limiting distribution exist. So, let π_i be the limiting probabilities defined as π_{ij} is equal to $\lim_{n \rightarrow \infty} p_{ij}^{(n)}$; we can determine the limiting distribution by solving $\pi_i = \sum_j \pi_j P_{ji}$ and the summation of π_i is equal to 1. Here there is a restriction on row because it is a finite state Markov chain. At the end we move into GI/G/1 Queue, in this queueing model a single server queue with the arbitrary arrival process that means the inter-arrival time or non exponential distribution and

arbitrary service time distribution that means a service time also non exponential distribution with the infinite buffer.

Let S be the random variable representing the service time and expectation of S will be $1/\mu$; where μ denotes the service rate. Let λ be the mean arrival rate; assume that λ is less than μ so, that the q is stable, for a stable system you need λ is less than μ ; where λ is related to the mean arrival rate and the μ is related to the service rate.

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GI/G/1 Queue ...

- ▶ We are interested in the distribution of the number of customers in the system.
- ▶ Let p_n be the probability that there are n in the system.
- ▶ The probability that the system is empty is given by

$$p_0 = 1 - \lambda/\mu$$

- ▶ Note that, Little's formula can be applied to obtain average measures such as mean system size, mean queue size and mean delay.

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One can show that for the stable GI/G/1 Queue the utilization is given by λ/μ . Utilization is nothing but what is the probability that the server is busy. So, that is given by λ/μ ; we are interested in the distribution of number of customers in the system in the limiting distribution we are interested in the limiting distribution of a number of customers in the system. Let p_n be the probability that there are n customers in the system in a long run. The probability that the system is empty is given by p_0 that is nothing but $1 - \text{utilization}$; the utilization is given by λ/μ that is probability, that the server is busy in a long run. Therefore, the probability that the system is empty in a longer run is same as $1 - \text{utilization}$; that is $1 - \lambda/\mu$. Note that Little's formula can be applied to obtain average measure such as a mean system size, mean queue size, and mean delay.

Since, it is stable system with the λ is less than μ have the probabilities and once you know the limiting distribution you can use the Little's formula to find out the all other average measures, that means using the limiting distribution you can find out the average number of customers in the system in a long run. Once you know the average number of customers in the long run with the mean arrival rate λ you can find the average number, average time spent in the system. Once you know the average time spent in the system you can find the average time spent in the queue also.

And, once you know the average time spent in the queue, using the little is formula you can find the average queue size also. So, you using the little formula with the help of limiting distribution we can find the average number of customers in the system, then average time spent in the system, average delay, and average queue size also. So, in this lecture we have covered the non Markovian queues namely GI/M/1, GI/M/1/N, GI/M/c, GI/G/1 Queues with these 5 lectures the renewal process is over. Here is the reference of renewal processes.