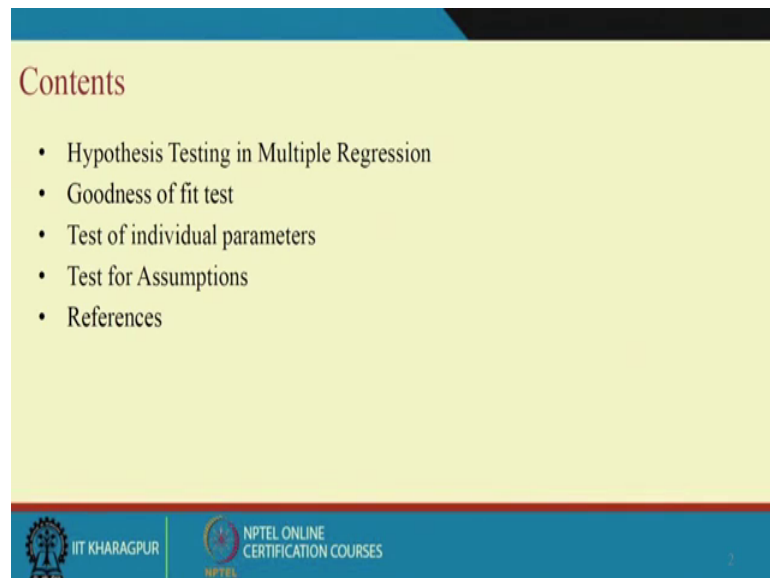


Design and Analysis of Experiments
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Lecture – 22
Multiple Linear Regression: Hypothesis Testing and Model Adequacy Test

Welcome. We will continue Multiple Linear Regression. Today I will talk about model adequacy test primarily whether regression equation is adequate or not from the data or the experimental data you obtain that basically I have given you a proper response surface or not.

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So, first I we will start with the overall hypothesis testing using f test then I will give you some goodness of fit test measure like R square, adjusted R square. Test of individual parameters we will revisit because earlier I have shown you how to do and another important one is the test of assumptions. So, when I talk about test of assumptions I will give you what are the assumptions and how those assumptions must be must be verified that they are really true.

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Estimation of Error

The fitted regression (predicted) model is -


$$\hat{y} = X\hat{\beta}$$

In scalar notation, the fitted model is

$$\hat{y}_i = \hat{\beta}_0 + \sum_{j=1}^k \hat{\beta}_j x_{ij} \quad \text{where } i=1,2,\dots,n$$

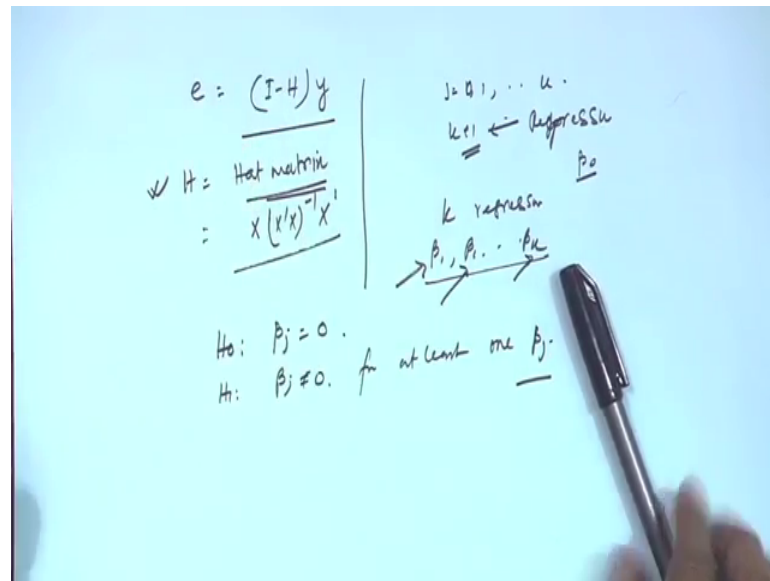
Residual of i-th observation $e_i = y_i - \hat{y}_i$

Overall residual (in matrix form)

$$e_{(n \times 1)} = y - \hat{y}$$
$$e = y - \hat{y}$$
$$e = y - X\hat{\beta}$$
$$e = y - X(X'X)^{-1}X'y$$
$$e = (I - H)y \quad \text{Hat matrix} = H = X(X'X)^{-1}X'$$


So, straightway let us go to that a regression equation here that y equal to X beta plus epsilon now if you when you estimate the y it will be y equal to X beta cap. And all of you know then in case scalar notation we can write like this and ultimately the difference between the actual observation minus the predicted one or otherwise other way I can say fitted one is the residual. So, $e_i = y_i - \hat{y}_i$. So, as there are n number of y values so that means, there will be n number of residuals. So, in matrix form then that e equal to $y - \hat{y}$ cap it is n cross one vector and if we just do certain level of manipulation for e we will find out that e will be $(I - H)y$ where H is $X(X'X)^{-1}X'$.

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So, we can see that the residual can be written like this i minus H into y , where i is the identity matrix and H is known as hat matrix. H is known as hat matrix which is X, X transpose X inverse X transpose this is the part. It is a very interesting matrix because it has a lot of implications; it can be used for doing a lot of tests. Particularly if we are interested to know what are the contributions of individual observations, then you will find out that the diagonal element of the hat matrix will talk about the contribution of individual observations. Suppose you have n number of observations, then diagonally will talk about all these things.

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Data Example

Observations	Temp (x1)	Catalyst feed rate (x2)	Viscosity (y)	Y PREDICTED	Residuals
1	80	8	2256	2244.46	11.54
2	93	9	2340	2352.12	-12.12
3	100	10	2426	2414.06	11.94
4	82	12	2293	2294.04	-1.04
5	90	11	2330	2346.43	-16.43
6	99	8	2368	2389.26	-21.26
7	81	8	2250	2252.08	-2.08
8	96	10	2409	2383.57	25.43
9	94	12	2364	2385.50	-21.50
10	93	11	2379	2369.29	9.71
11	97	13	2440	2416.95	23.05
12	95	11	2364	2384.53	-20.53
13	100	8	2404	2396.89	7.11
14	85	12	2317	2316.91	0.09
15	86	9	2309	2298.77	10.23
16	87	12	2328	2332.15	-4.15

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So, anyhow let us see one a one data example. We have 16 observations we have 3 2 factors X temperature and catalyst feed rate and our dependent variable is viscosity that is to be predicted let it be and we have fitted the model and which we have already shown you earlier and then from there this model we found out the predicted values or the fitted values here exactly these are fitted values. Then this, this viscosity observed values minus fitted value is giving you the residuals these residuals e values this is the last column.

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Test for Significance of Regression

- **Assumption:** The errors ε_j in the model are normally and independently distributed with mean zero and variance σ^2 , i.e.,

$$\varepsilon \approx NID(0, \sigma^2)$$
- The main **objective** of testing the significance is to determine whether a linear relationship exists between the response variable y and a subset of the regressor variables x_1, x_2, \dots, x_k .
- **Hypotheses:**

$$H_0: \beta_1 = \beta_2 = \dots = \beta_k = 0$$

$$H_1: \beta_j \neq 0 \text{ for at least one } j \rightarrow \text{At least one of the regressor variables contributes significantly to the model}$$

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Now, what we will do here we will we will actually the one of the important assumptions here is that the errors are normally distributed identical and normally distribution iid, independent and identically distributed and that is also normally distributed that is one.

So, anyhow that test we will see later on. But what is the hypothesis overall fit test here overall fit test is that we are we have j equal to 0, 1 to k ; that means, k plus 1 regression, regression. If we do not consider this one that the con intercept that beta 0 then we have k regressions, k regressors and stab; that means, beta 1 beta 2 like this beta k and that is important because beta 1 related to variable X_1, X_2 and X_k like this important which of the variables are contributing or not. So, here what we will do we will over for overall test we say that beta j not equal to 0; that means, none of the factors are contributing towards the response what we observed during the experiment and H_1 is beta j not equal

to 0. For at least one beta j let me at least beta 1 or beta 2 or beta 3 sum is contributing and no not that all beta j are 0.

This is overall test, this overall test is done through f statistics.

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Test for Significance of Regression (Contd.)




$$SS_T = SS_R + SS_E$$

$$SS_E = y'y - \hat{\beta}'X'y$$

$$SS_T = \sum_{i=1}^n y_i^2 - (\sum_{i=1}^n y_i)^2/n = y'y - (\sum_{i=1}^n y_i)^2/n$$

$$SS_R = \hat{\beta}'X'y - \frac{(\sum_{i=1}^n y_i)^2}{n}$$

$$SS_E = y'y - \hat{\beta}'X'y$$

$$SS_T = y'y - \frac{(\sum_{i=1}^n y_i)^2}{n}$$




You have seen in ANOVA that sum square total equal to sum square I think that the treatment plus sum square error in one way ANOVA you seen this one.

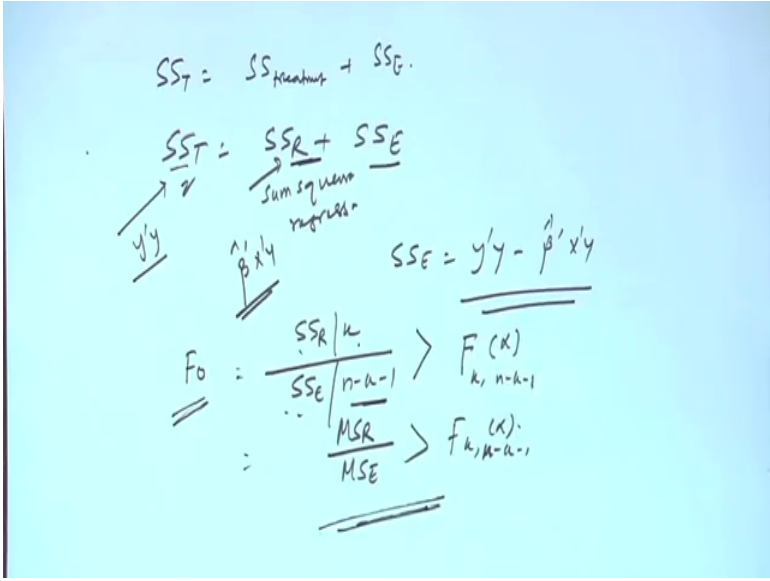
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$$SS_T = SS_{\text{treatment}} + SS_E$$

$$SS_T = \underbrace{SS_R}_{\text{sum square regression}} + SS_E$$

$$SS_E = y'y - \hat{\beta}'X'y$$

$$F_0 = \frac{SS_R/k}{SS_E/(n-k-1)} > F_{k, n-k-1}(\alpha)$$

$$= \frac{MSR}{MSE} > F_{k, n-k-1}(\alpha)$$


Here also in regression we can write sum square total equal to sum square regression plus sum square error this is sum square regression or we can say that it sum square model. And all of you know that $y^T y$ will give you this and sum square regression will be that $\hat{\beta}^T X^T y$ from this and essentially then SSE will be SS_T minus SS_R which is will be $y^T y$ minus $\hat{\beta}^T X^T y$ this will give you the SSE .

So, fine then SS_R also $\hat{\beta}$ and then finally, what happened finally, what happened I have given probably little difficult equation differently. Let us see that SS_T all of we know will be this following this equations. So, here $y^T y$ when is mean subtracted is there that time if you subtract by y is means after that formula, but otherwise what happen is $SS_T - y^T y$ square minus this; that means, this quantity is coming here. Similarly SS_R also that this will be subtracted and SS_E then this $SS_E = y^T y$ minus this and SS_T equal to $y^T y$ minus this by n $y^T y$ minus $y^T y$ square by n that is what we have seen. It is a basically some kind of duplications we have made, but whatever may be the thing you please remember you will be able to compute SS_R you will be able to compare SS_T and SS_E also you will be able to compute from errors residuals, but other way SS_T minus SS_R will give you SSE .

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Test for Significance of Regression (Contd.)


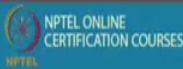
$$F_0 = \frac{SS_R/k}{SS_E/(n-k-1)} = \frac{MS_R}{MS_E}$$

Re ject H_0 if $F_0 > F_{\alpha, k, n-k-1}$

Analysis of Variance for Significance of Regression in Multiple Regression				
Source of Variation	Sum of Squares	Degrees of Freedom	Mean Square	F_0
Regression	SS_R	k	MS_R	MS_R/MS_E
Error or residual	SS_E	$n - k - 1$	MS_E	
Total	SS_T	$n - 1$		

The coefficient of multiple determination: $R^2 = \frac{SS_R}{SS_T} = 1 - \frac{SS_E}{SS_T}$

$$R^2_{adj} = 1 - \frac{SS_E/(n-p)}{SS_T/(n-1)} = 1 - \left(\frac{n-1}{n-p}\right)(1 - R^2)$$

So, then what happened you will create a statistics called f_0 this is nothing, but SS_R divided by k by SS_E divided by its degrees of freedom n minus k minus 1 . So, if this,

this one is greater than $F_{k, n-k-1, \alpha}$. So, then what we will say H_0 rejected, H_0 is rejected.

Now, what is this SS_R by k ? This is a MS_R what is SS_E by degree of freedom MS_E so; that means, this MS_R by MS_E if it is greater than $F_{k, n-k-1, \alpha}$ will this F_0 it is and then we will say that the null hypothesis that none of the factors are contributing is not correct. This is what is in terms of ANOVA table if you see the ANOVA table here it is in terms of a ANOVA table we have shown the same thing.

Now, this is overall f test overall f test will say that whether at least one of the regression coefficient contributing or not if none of the regression coefficient, contribute regression coefficient is significant or none of the factors X variables are contributing then and then is that is what H_0 will be accepted and none will be nothing in no influence. But what happened here you may be interested to know that suppose if test it says that H_0 is rejected then we want also we want to know that what is the variability of y is explained by X that was this is if I say the y variability may be the regression y is X beta plus epsilon then this X beta this portion may be able to explain this much. So, what is this portion?

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$$R^2 = \frac{SSR}{SST} = 1 - \frac{SSE}{SST}$$

$$R^2 = 1 - \frac{\frac{SSE}{n-p}}{\frac{SST}{n-1}}$$

$$R^2 = 1 - \frac{(n-1) s_e^2}{(n-1) s_y^2}$$

$$0 \leq R^2 \leq 1$$

$$R_a^2 \leq R^2$$

So, this portion it is some kind of absolute major that the variability of y explained by the regression model that is if divided by the variability not explained by the regression model or total variability will give you a measure which is known as R square. So, R

square is a major which is SS_R by SS_T that mean variability explained by the model and total variability of y variability of y explained by the model and total variability this is R square this one can be written like this $1 - \frac{SS_E}{SS_T}$ also.

So, what is the problem here? Problem is if I expand this I can write what SS_E we can write that SS_R by SS_T and SS_E by this one. So, we can write like this $\frac{SS_E}{n - p}$ by $\frac{SS_T}{n - 1}$ where p equal to $k + 1$ if you write like this then this quantity will give you a measure which is known as suppose this quantity gives you a measure which is known as R square adjusted.

So, I will explain little further this one. So, R square I can write one minus $\frac{SS_E}{n - k - 1}$ into se square you have seen earlier se square similarly SS_T you have seen $n - 1$ into sy square. So, now, if you divide the thing this, this by degrees of freedom $n - k - 1$ or $m - p$ and this by its degrees of freedom you are getting this one. So, we are creating another coefficient measure are adjusted R square which will become then, then what happened R_a square if I write then this will be $\frac{SS_E}{n - p - 1}$ minus $\frac{SS_T}{n - k - 1}$ that will this se square SS_T means $n - 1$ minus se square by $n - 1$ minus sy square. So, this portion, this value is this R square value is unaffected by n and p that is the sample size as well as the parameters number of parameters to regression coefficient to be estimated.

So, it is a speedo sample size this, this value will get and it is a better measure than R square, R square. R square will be impleted suppose n equal to p then this quantity what happen ultimately this quantity will become for example, here if I read n equal to p this will become 0. So, R square become 1 minus abnormally high R square value you will get. So, and whatever may be the case R square will lie in between 0 to 1 and R_a square also lie in between 0 to 1 and R_a square is greater than equal to R square. This R square, is that is coefficient of multiple determination ok. So, this is the goodness of test.

If R square value is greater than in case of a laboratory experiment whether R square or R_a square if it is greater than equal to 0.90 then it is a fit data, is a good fit to the model regression model.

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Regression statistics

Observations	Temp (x1)	Catalyst feed rate (x2)	Viscosity (y)
1	80	8	2256
2	93	9	2340
3	100	10	2426
4	82	12	2293
5	90	11	2330
6	99	8	2368
7	81	8	2250
8	96	10	2409
9	94	12	2364
10	93	11	2379
11	97	13	2440
12	95	11	2364
13	100	8	2404
14	85	12	2317
15	86	9	2309
16	87	12	2328

Regression Statistics	
Multiple R	0.96279283
R Square	0.926970033
Adjusted R Square	0.915734653
Standard Error	16.35860385
Observations	16

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You just see that multiple R square value that 0.96, I think R square is SS by SS R by SS T, fine. So, R square value 0.92 adjusted R square 0.91 and standard error is 16.3 by observation 16. So, R square and adjusted R square both are more than 0.9. So, we can say that it is good the data is fit to the regression model.

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Goodness of fit test-ANOVA model

Observations	Temp (x1)	Catalyst feed rate (x2)	Viscosity (y)
1	80	8	2256
2	93	9	2340
3	100	10	2426
4	82	12	2293
5	90	11	2330
6	99	8	2368
7	81	8	2250
8	96	10	2409
9	94	12	2364
10	93	11	2379
11	97	13	2440
12	95	11	2364
13	100	8	2404
14	85	12	2317
15	86	9	2309
16	87	12	2328

ANOVA	df	SS	MS	F	Significance F
Regression	2	44157.08654	22078.54	82.5045585	4.09975E-08
Residual	13	3478.85096	267.6039		
Total	15	47635.9375			

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This is the f test what I explained to you. Now f value computed 82 and it is highly significant; that means, what happened model weather predictors of the factors are contributing.

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Individual parameter test

Observations	Temp (x1)	Catalyst feed rate (x2)	Viscosity (y)
1	80	8	2256
2	93	9	2340
3	100	10	2426
4	82	12	2293
5	90	11	2330
6	99	8	2368
7	81	8	2250
8	96	10	2409
9	94	12	2364
10	93	11	2379
11	97	13	2440
12	95	11	2364
13	100	8	2404
14	85	12	2317
15	86	9	2309
16	87	12	2328

	Coefficients	Standard Error	t Stat	P-value	Lower 95%	Upper 95%
Intercept	1566.077771	61.59183583	25.42671	1.8031E-12	1433.0167	1699.138843
X Variable 1	7.621290077	0.618429643	12.32362	1.5177E-08	6.285254059	8.957326094
X Variable 2	8.584845886	2.438683811	3.520278	0.00376481	3.316389819	13.85330195

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Individual parameter test earlier I have shown you and here this example is repeated.

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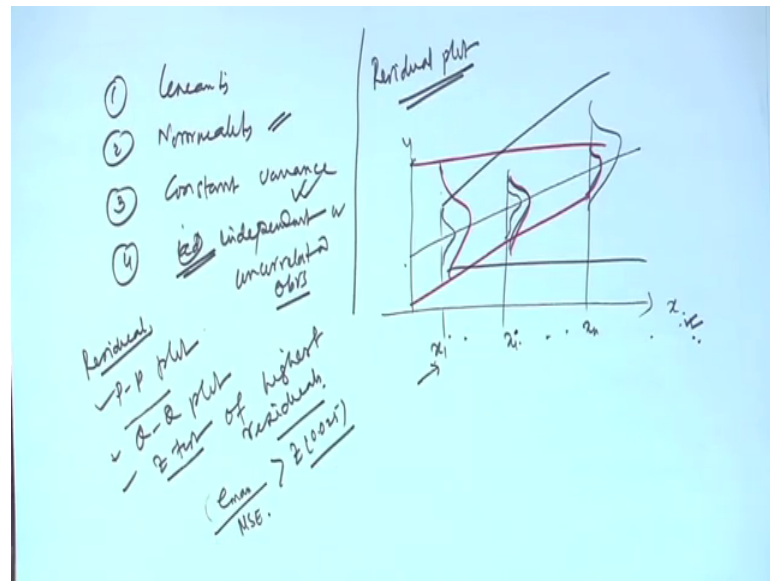
Test of Assumptions: Linearity

Temp (x1)	Residuals	e=x1b1
80	11.54	621.2432
93	-12.12	696.66
100	11.94	774.069
82	-1.04	623.9058
90	-16.43	669.4861
99	-21.26	733.2477
81	-2.08	615.2445
96	25.43	757.0738
94	-21.5	694.9013
93	9.71	718.49
97	23.05	762.3151
95	-20.53	703.4926
100	7.11	769.239
85	0.09	647.8997
86	10.23	665.6609
87	-4.15	658.9022

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Now another, now the test of assumptions; test of assumption is very very important in regression.

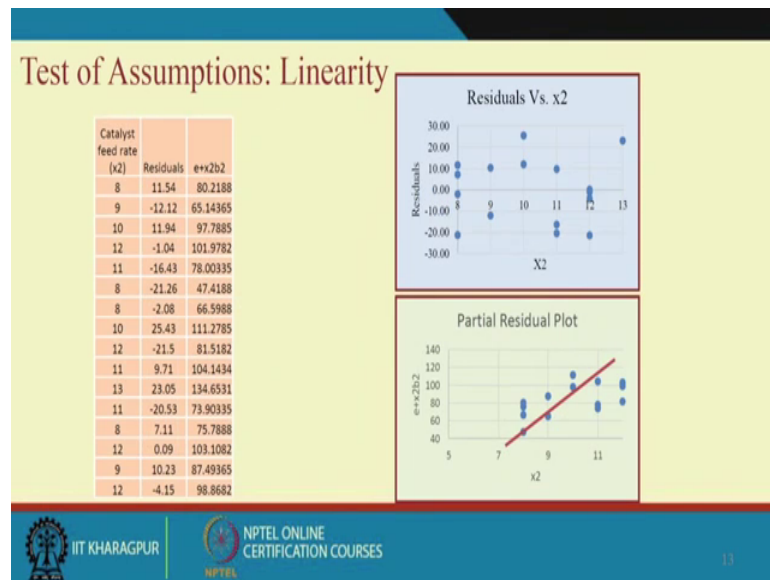
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What are the assumptions, multiple linear regression, one is the linearity normality 3 is your constant variance 4 is iid independent, independent in nature I can say independent or uncorrelated observations independent or uncorrelated observations.

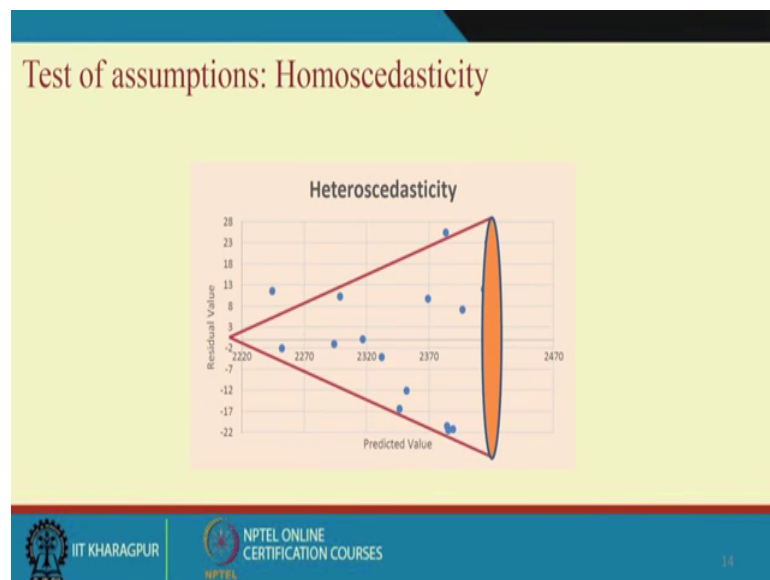
So, from residuals plot we will be able to find all those things residual plot, plot will help you in knowing that whether they are valid or not anyhow. Let us see that for this data set we have found out the residuals this is temperature versus residual. You see the residual plot here now in case of linearity we have shown here partial residual plot this is residual versus X₁; that means, these versus this and we do not find any pattern here. But when you go for partial residual plot; that means, the residual plus the X₁ contribution and versus X₁ then you see the linear plot is obtained so that means, linearly related.

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Now, similarly for X 2, but this linearity is, this linearity is your little weak because there is more spread.

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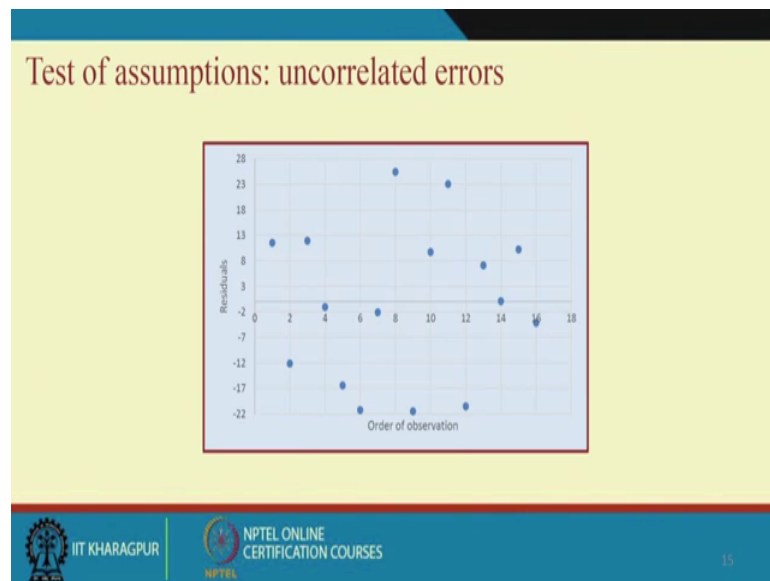


Now second one is that normal here I am showing that heteroscedasticity, it is nothing but that the error variance are not constant over x. If you recall my first I say if it is x and this is y suppose we are fitting a regression line when we are having only one predictor variables or factors then we say that for every fixed observations suppose this is x 1 this is x i and this may be x n every fixed observation, if I keep x at x 1 and do experiment

several times I may get y like this y distribution like this here y distribution like this here. The assumption is that irrespective of your x value the y variability across x, will variability will be same or constant variance across the x this is known as homoscedasticity.

If there is violation like this the when X is low like the value is like this if high like this high like this or other way around. So, this one is big one or other way around, so this is the biggest one this is what is this and this is what is this. So, then you will have either funnel to left or a funnel to funnel to right this kind of situation then this is a called this is a situation for heteroscedasticity for not un constant error variance across y and this is a violation of constant variance and aggression estimates will be inflated or un or it will be wrong that is the issue. So, if you how do you know that it is happening. So, if you plot the residual value versus predicted value and found some kind of funneling effect then it is a heteroscedasticity problem, but if it is random then it is not.

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So, for uncorrelated errors are independent, both or the order in which you have done the experimentation put this side and residual in the vertical y side. So, then you see that whether these observations are showing any, it is a random observations there is no systematic pattern in the plot then it is uncorrelated.

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Test of assumptions: Uncorrelated errors

Residual (e _{i-1})	e _i	e _{i-1} *e _i	e _i ²
11.54	-12.12	-139.865	133.1716
-12.12	11.94	-144.713	146.8944
11.94	-1.04	-12.4176	142.5636
-1.04	-16.43	17.0872	1.0816
-16.43	-21.26	349.3018	269.9449
-21.26	-2.08	44.2208	451.9876
-2.08	25.43	-52.8944	4.3264
25.43	-21.5	-546.745	646.6849
-21.5	9.71	-208.765	462.25
9.71	23.05	223.8155	94.2841
23.05	-20.53	-473.217	531.3025
-20.53	7.11	-145.968	421.4809
7.11	0.09	0.6399	50.5521
0.09	10.23	0.9207	0.0081
10.23	-4.15	-42.4545	104.6529
-4.15			17.2225
		-1131.05	3478.408
	r	-0.32516	
	DW	2.650328	

Durbin Watson Test

$$DW = \frac{\sum_{i=2}^n (\hat{\epsilon}_i - \hat{\epsilon}_{i-1})^2}{\sum_{i=1}^n (\hat{\epsilon}_i^2)}$$

$$r = \frac{\sum_{i=2}^n \hat{\epsilon}_i \hat{\epsilon}_{i-1}}{\sum_{i=1}^n \hat{\epsilon}_i^2}$$

DW=2(1-r)

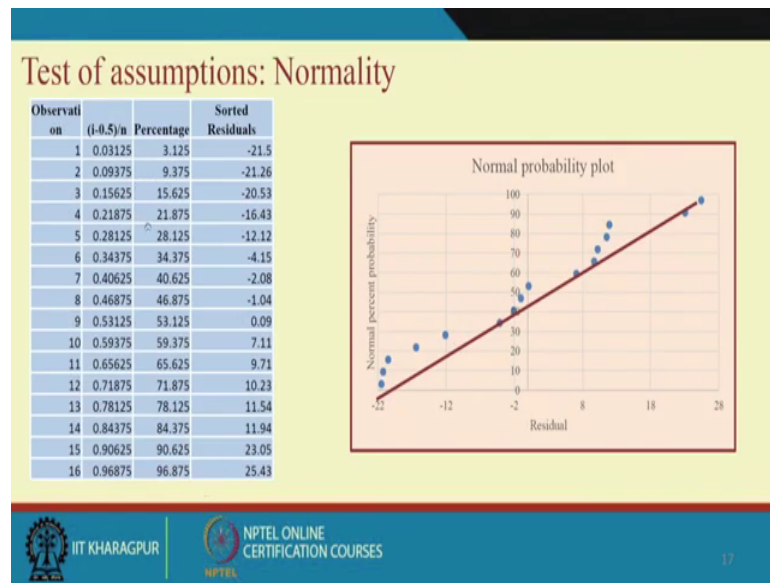
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So, what about the linearity a normality part I think I have told you p-p, p-p plot are told earlier. So, for all residuals you do this, this is used for p-p plot, q-q plot, q-q plot or also you can say individual z test, suppose z test of highest error highest residual that test of highest residual maximum residual. So, you just find out the maximum residual e max from this data and then you divide it by your error variance that is MS E and if this one this is greater than suppose z 0.025 then we can say that there is the violation of normality.

So, for uncorrelated errors there is another test called Durbin Watson test and here what happened just to see that whether the errors are correlated or not, you create the you take the error in or in this order. And you keep create some lag maybe for 1 lag or some lag that 1 lag, 2 lag, k lag maybe and then you find out the correlation between the 2 and here what happened this is basically the residuals and then we created the i minus 1 and e i this then we have found used this formula and found out the R value and it is shown that the R value is minus 3.33 and Durbin Watson test value is 2.65 it is correlated. But 0.3 is significant I think because it is not less.

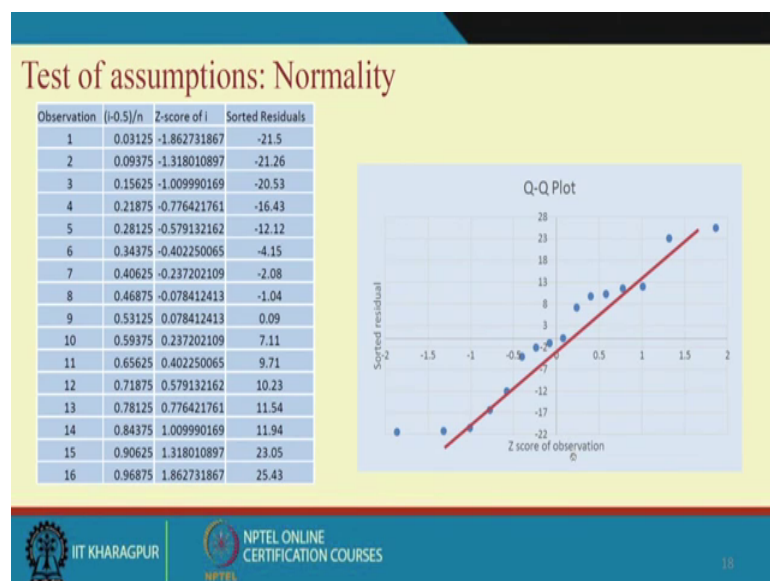
Again R square Durbin Watson value is also more than 2, but it can be considered other way also whether it is also R square and other things giving you better, but this is one assumption which is supposed to be violated here because the constant error sorry independent part.

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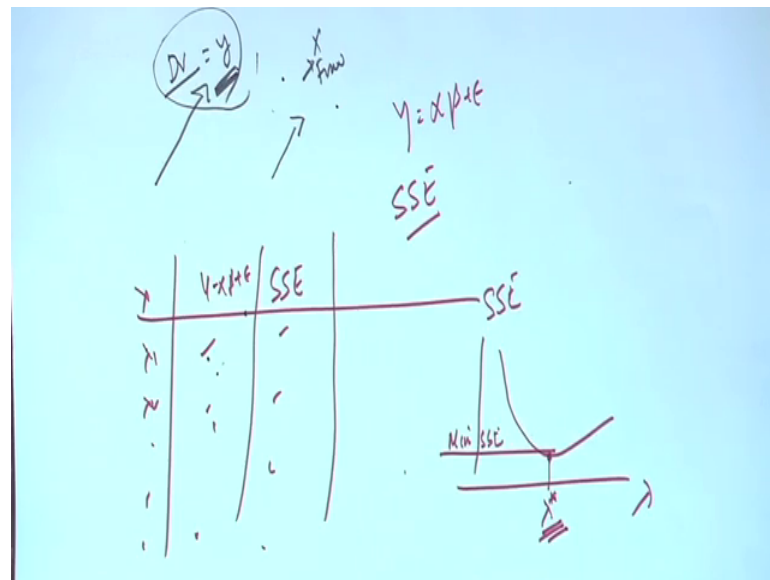
Now, I told, I talk to you about normality and this you know this normality earlier I shown that the residual that you use the normal probability plot and if there is a straight line kind of things and this i minus 0.5 by n . If you use this formula you will get a straight line that is normal there is another one is that residual versus z score means quantile quantile plot. This quantile quantile plot is also showing some kind of normality that is not deviating from normality a much.

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Last, but very important one also that means, what happened when the constant error variance as well as normality assumption is violated you have to transform the data. So, there are many methods of data transformation, and please remember this homoscedasticity it is a feature for the dependent variable y , dependent variable y .

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We are talking about fixed effect model and X are fixed ok. So, when you talk about transformation so it is primarily related to y , but sometimes many times we also do X transformation so that the relationship can be change some of the assumptions can be satisfied for example, linearity; for example, may be normality also.

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Remedy against violation of assumptions

- Heteroskedascity: Transform y, Box-Cox method
- Linearity: Transform y, x or both
- Normality: Box-Cox Method

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So, here what I will show you I will show you very quickly some of the methods for transformation one is Box-Cox method and another one is I think some more methods are there but primarily we will be considering on Box-Cox.

If there is heteroscedasticity then you transform y using Box-Cox method if there is linearity you can transform y, x or both it will help you. If it is normality again go for Box-Cox method.

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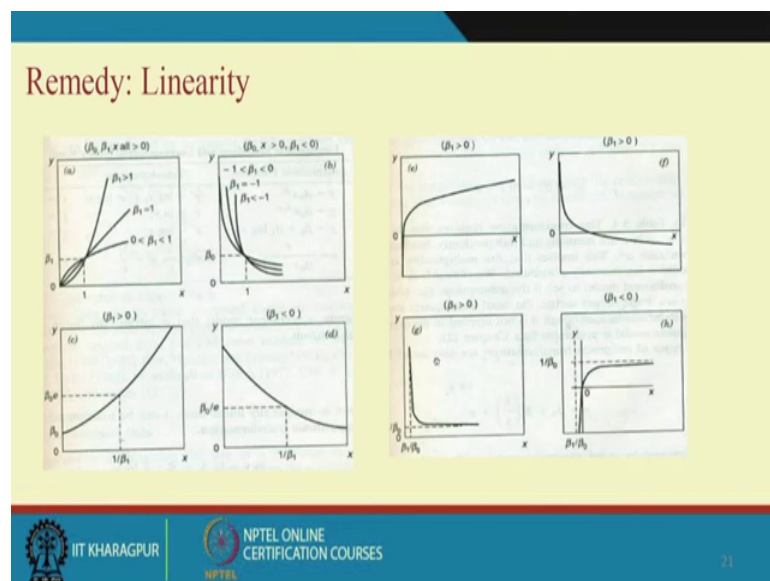
Remedy: Heteroskedasticity

Relationship of σ^2 to $E(y)$	Transformation
$\sigma^2 \propto \text{constant}$	$y' = y$
$\sigma^2 \propto E(y)$	$y' = \sqrt{y}$
$\sigma^2 \propto E(y)[1-E(y)]$	$y' = \sin^{-1}(\sqrt{y}) \quad (0 \leq y_i \leq 1)$
$\sigma^2 \propto [E(y)]^2$	$y' = \ln(y)$
$\sigma^2 \propto [E(y)]^3$	$y' = y^{-1/2}$
$\sigma^2 \propto [E(y)]^4$	$y' = y^{-1}$

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And this is what is the transformation kind of things. Suppose you may if constant variance is satisfied there is no transformation required, but if the error variance is proportional to the mean value then find out that take the square root. If error variance is proportional to mean and this quantity then make sin inverse square root transformation. If it is square of mean log transformation, if it is cuba mean then inverse square root transformation, if it is 4th power mean then you do 1 by y that is inverse of this. These are some guidelines you can use while you do some kind of analytics that is experimental data analysis.

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So, you see that although this kind of transformation will give you the desired result and this is available in montgomery linear statistical models.

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Remedy: Linearity

Figure	Linearizable Function	Transformation	Linear form
a & b	$y = \beta_0 x^{\beta_1}$	$y' = \log y, x' = \log x$	$y' = \log \beta_0 + \beta_1 x'$
c & d	$y = \beta_0 e^{\beta_1 x}$	$y' = \ln y,$	$y' = \ln \beta_0 + \beta_1 x$
e & f	$y = \beta_0 + \beta_1 \log x$	$x' = \log x$	$y' = \beta_0 + \beta_1 x'$
g & h	$y = \frac{x}{\beta_0 x - \beta_1}$	$y' = \frac{1}{y}, x' = \frac{1}{x}$	$y' = \beta_0 - \beta_1 x'$

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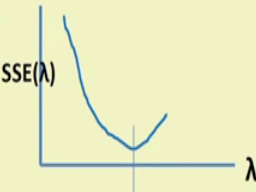
So, you see that if the function is something like this, then you go for log transformation function like this go for this, function like this go for distribute transformation and you do not know which transformation will help you, but that is why what happened you may do all kind of transformation and then find out that whether the assumptions are satisfied particularly in the linearity point of view here.

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Remedy: Normality & heteroskedasticity (Box-Cox method)



$$y^{(\lambda)} = \begin{cases} \frac{y^\lambda - 1}{\lambda y^{\lambda-1}}, & \lambda \neq 0 \\ \ln y, & \lambda = 0 \end{cases}$$

$$\hat{y} = \ln^{-1}[(1/n) \sum_{i=1}^n \ln y_i]$$



$y^{(\lambda)} = X\beta + \varepsilon$

Choose the λ for which $SSE(\lambda)$ is the minimum

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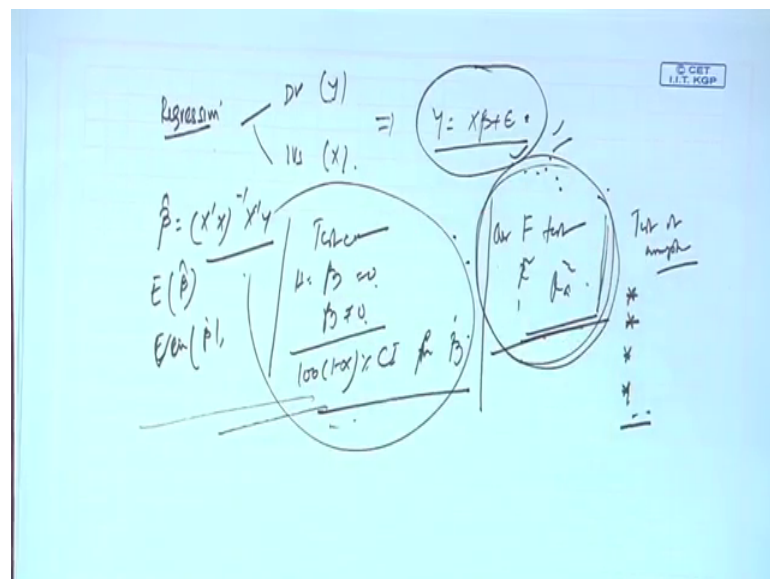
And you do Box-Cox transformation for normality and heteroscedasticity. Here what happened? You basically create a power lambda then this you are creating in the norm

variable y to the power within bracket λ which is y to the power $\lambda - 1$ by λ into $y \cdot \lambda^{-1}$. Now, the λ obviously, not equal to 0 in the first case if $\lambda = 0$, you write $y \cdot \log y$ where $y \cdot$ is inverse log one by n $i = 1$ to $n \log y_i$ that mean you take the log of the original observations then take their sum divide it by the take the average, average of the log transport information and then its inverse log inverse.

So, you choose and this is basically $y \cdot$ and now you choose different λ values feed the regression equation $y = X\beta + \epsilon$ calculate $SS E$ and then plot. So, what happened? You will choose λ different λ value different λ value like this then your y equal to $X\beta + \epsilon$ some model will be there and then you will calculate $SS E$ some value you will get $SS E$ and then you plot λ versus $SS E$; what will happen, you will find out a curve like this where this λ is giving you the minimum $SS E$ minimum $SS E$ value this λ this is a λ step. This is the best transformation for you.

So, choose λ for you is λ $SS E$ λ is the minimum and now again fine that equation is also known you know which λ is minimum an equation is known and that linear equation in huge. So, let me just do one let me summarize the thing.

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We talked about regression and primarily it is linear regression and in linear regression obviously, we have 2 sets of variable dependent variable 1, independent variable several.

We have are interested to find an X equation like this where epsilon is the error term. We have estimated beta using certain equation and then we have estimated the mean value and the variance of beta. Then what we have tested beta tested beta, beta j equal to 0, H 0 beta is equal to 0 beta j not equal to 0. We also found out the con hundred into 1 minus alpha percent confidence interval for beta j, beta j that also you found out.

And then also we found out that whether the model is fit or not using f test overall f test and your R square Ra square and then also we have that a say H 0 a set of variables contributing or not contributing both partial marginal test and subset hypothesis testing we have done. Then we have gone for test of assumptions test of assumptions both linearity, normality, homoscedasticity independence all those things and we have also seen some of the examples particularly here we are using 2 independent variables and one dependent variable and how this model is working or not.

Nutshell what I mean to say you must know that the data fit to this model. You must know that every model has certain assumptions those assumption must be tested. You must be aware that when you are saying a model that model must be fit to the data that is the adequacy test, overall fit test. When the model overall model is fit then you go for individual parameters whether individual they are contributing or not and accordingly you accept or discard those variables which are not contributing, again you rebuild the model with the significant parameters. You may find out some of the estimates are different than the earlier, if that is the case you please be careful if some of the variables become insignificant then there may be the partial correlations and there may be problem in from the beginning data collection to variable selection to maybe the test and test of assumptions and all these things.

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References

- Design and Analysis of Experiments by Douglas Montgomery, Wiley, 8th Edition, 2014
- Applied Multivariate Statistical Modelling by J Maiti NPTEL Video.

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So, thank you very much. Again I must tell you that the Montgomery book and my earlier NPTEL video lecture on multivariate statistical modeling. Thanks a lot.

Thank you very much.