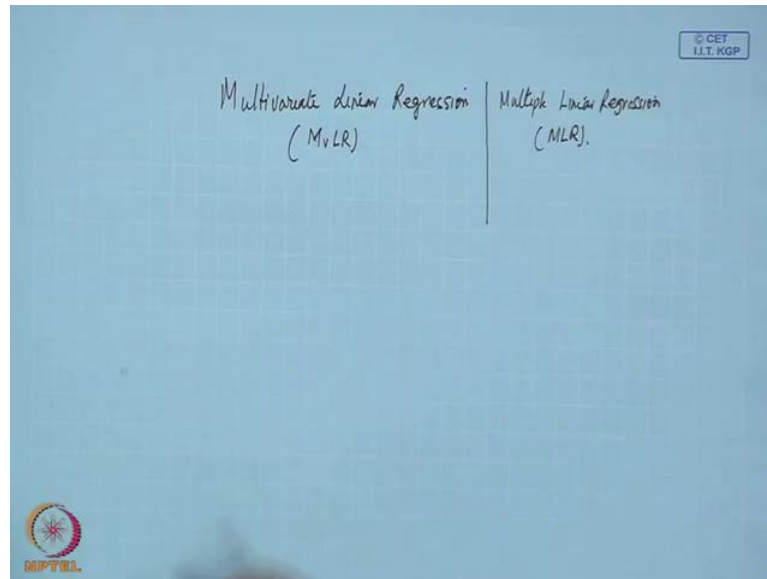


Applied Multivariate Statistical Modeling
Prof. J. Maiti
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Indian Institute of Technology, Kharagpur

Lecture - 27
Multivariate Linear Regression

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Good morning. We will discuss now multivariate linear regression. Multivariate linear regression, we will be denoting it as MvLR, MvLR. There is another regression, which is known as multiple linear regression, which we have already completed that we have denoted like this, MLR. So, the difference between the two will be discussed.

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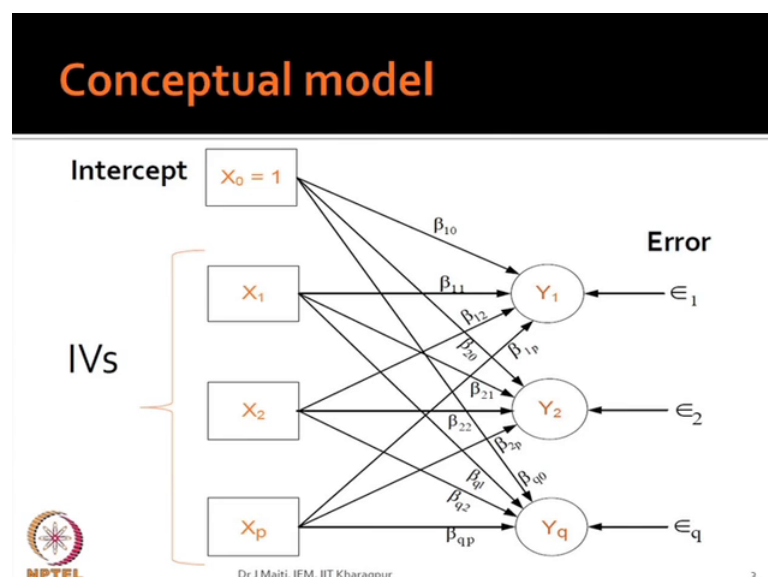
- Conceptual model
- Assumptions
- Estimation of parameters
- Sampling distribution of parameters
- Assessment of overall fit
- Test of individual regression parameters
- Model diagnostics
- Summary



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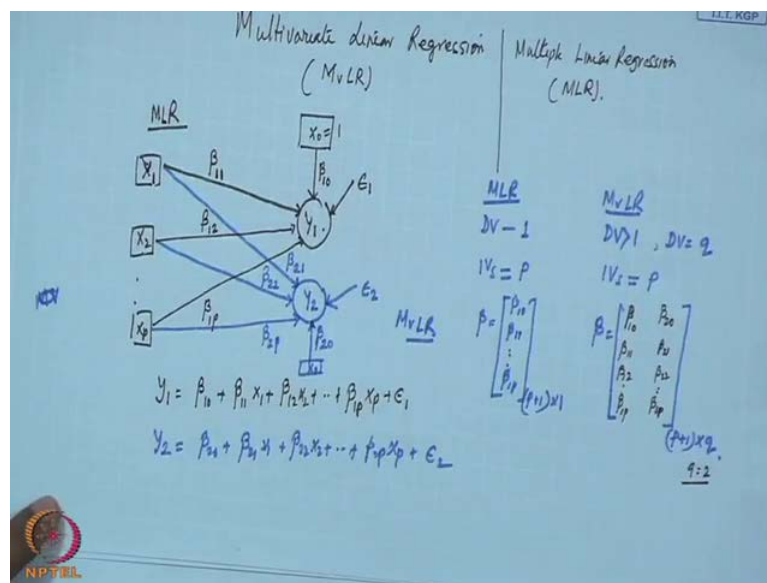
Then the in the same order, just see this slide that we will go by the conceptual model. Then we will describe the assumptions of the model. Then we will estimate the parameters. Then I go for sampling distribution of the parameters. Finally, overall assessment of model fit, that will be for this lecture and then I think test of individual regression parameters, model diagnostics and other things will be will be discussed in the next lecture.

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So, now let us see that what is the difference between MvLR and MLR? So, in this diagram, you see there are two parts. First one is this Y1, Y2, Yq, this side right hand side and left hand side, you see that X0 to Xp. This part, the right hand side part is known as the dependence part, left hand side part is known as the independent, independent part. What what do I mean by this? We are saying that the variables Y1, Y2, Yq are dependant variables and the variable X1 to Xp are independent variables. Now, all the dependent variables starting from Y1 to Yq, they are dependent on the same set of independent variables X1 to Xp. So, for example, let us consider only Y1 variable.

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Y, Y1 and then let us think that there are set of p independent variables and obviously you have seen in multiple linear regression that is if you can recall, if you recall your MLR multiple linear regression, you can find out that we have pictorially depicted the multiple linear regression part like this, where there will be several coefficients. Those are known as regression coefficients that Y1 is affected by X1, so 1 related to dependent side followed by another subscript that 1, that 1 is related to the independent side, so beta 11.

So, similarly, second variable from the independent side is X2, which effects Y1, so 1 is affected by 2. So, similarly, 1 is affected by p, pth independent variable. We have also seen that there is a constant term X0, which we are saying that beta1 is affected by X0, X0 always takes value 1. So, this is what is your MLR? In MLR, the equation is Y1

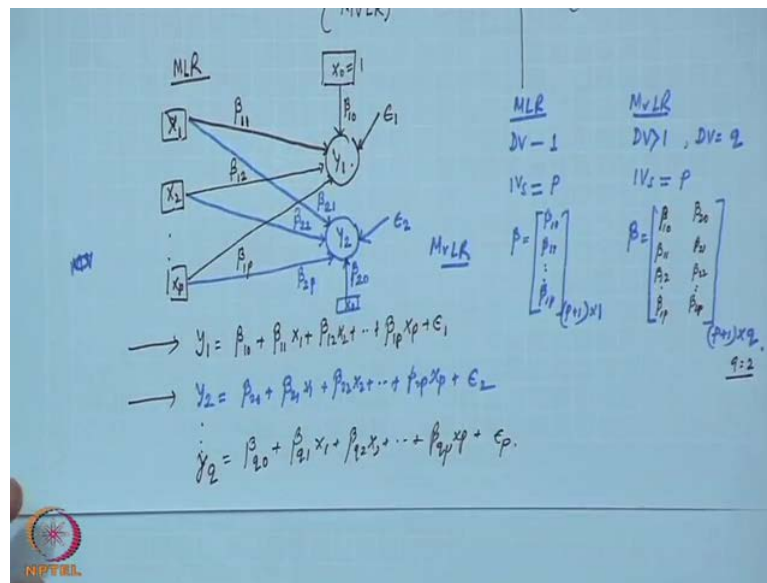
equal to $\beta_{10} + \beta_{11} X_1 + \beta_{12} X_2$ like this $\beta_{1p} X_p + \epsilon_1$. Now, let us add one more variable, which variable? Dependent side. So, we are saying the independent structure remains same X_1 to X_p . You are adding one more dependent variable Y_2 .

Then, if I want to know how they are affecting, then Y_2 is affected by X_1 . So, it will be β_{21} . Then Y_2 is affected by X_2 . It will be β_{22} . Then similarly, at the end, finally, Y_2 is affected by X_p . So, it also has that X_0 component. So, we can say that that I can, we can bring from here. I am writing here again. So, Y_2 is affected by β_{20} .

So, if I write down this $Y_2 = \beta_{20} + \beta_{21} X_1 + \beta_{22} X_2 + \dots + \beta_{2p} X_p + \epsilon_2$, so in this diagram, that means, then there is another component which is ϵ_2 . Similarly, in the in for the Y_1 case, there will be another component which is ϵ_1 . So, this one, when we have added Y_2 , it becomes MvLR, multivariate linear regression. So, if I now think the difference, then if I write MLR, then DV will be 1 and if I write MvLR, your DV will be greater than 1, any value. Your IVs are p . Here, your IVs are also p , p number of IVs.

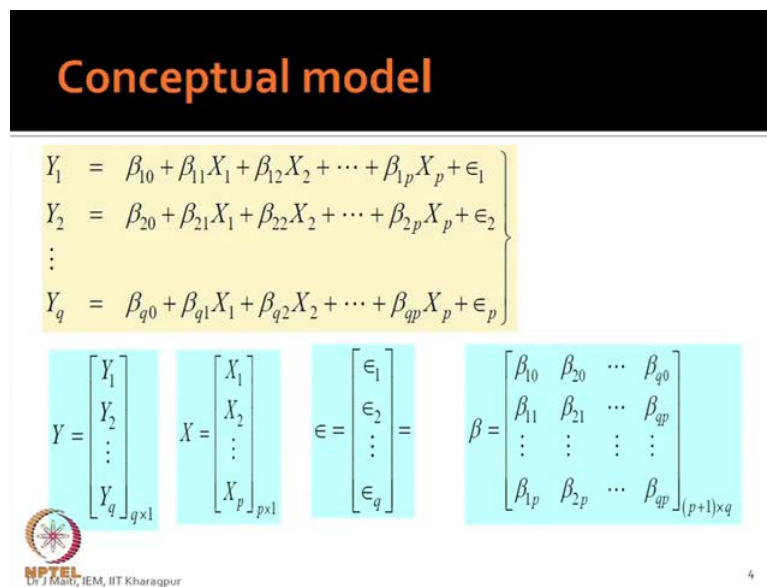
So, here in MLR, you are as dependent variable is 1, so β is the parameter vector which is basically β_{10}, β_{11} like β_{1p} , which will be your $p + 1$ cross 1. But, here what happen? If we assume that total DV equal to q , then your β is no longer $p + 1$ cross 1, it will become $p + 1$ cross q because of q variable, q dependent variables. So, for this case, you can see that here what we have written here in this case. So, if I consider this particular example, then your this one, your β is $\beta_{10}, \beta_{11}, \beta_{12}$, like this β_{1p} , then $\beta_{20}, \beta_{21}, \beta_{22}, \beta_{2p}$. Getting me? So, here that means q equal to 2.

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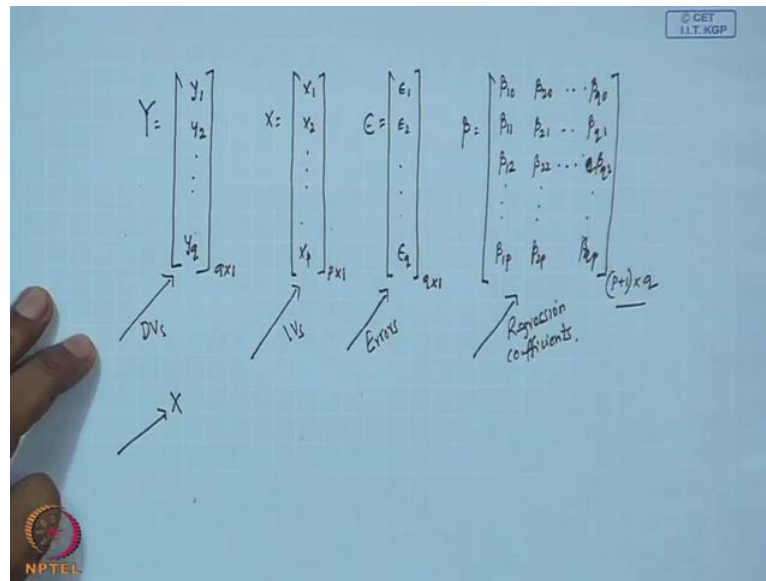
This is what now you may be interested to add many more variables in the Y side; instead of 2q, maybe 10 and then you will be having ten simultaneous regression equations. This is for first variable, second variable, like this, you will be having equation for Yq, which is beta q is affected by the constant term plus beta q1 X1, beta q2 X2, so like this, beta qp Xp plus epsilon p.

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So, then we can write in matrix term. Matrix term, if we want to write, what we will get?

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We will get Y variable, which is Y_1, Y_2 like this Y_q , q cross 1. We have X independent variable side X_1, X_2, X_p , p cross 1. You also have your ϵ_1, ϵ_2 , like this ϵ_q , q cross 1 random that error variables. You have also beta that will be your $\beta_{10}, \beta_{11}, \beta_{12}$, then $\beta_{1p}, \beta_{20}, \beta_{21}, \beta_{22}, \beta_{2p}, \beta_{q0}, \beta_{q1}, \beta_{q2}$, so like this β_{qp} , this will be $p + 1$ cross q . So, I am saying this is my DV. This one my IVs, this is my errors; this is my regression coefficients, regression coefficients. So, if you consider X as design matrix, suppose X is the design matrix, then see this X_1 to X_p , these variables are there, but X_0 will come in to consideration that we will see when we collect data. Getting me?

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Conceptual model: An example

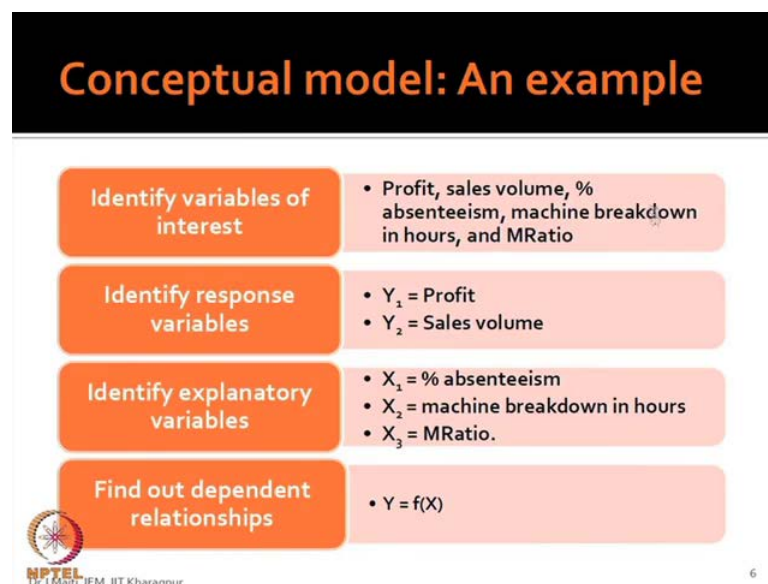
Sl. No.	Months	Profit in Rs million	Sales volume in 1000	Absenteeism in %	Machine breakdown in hours	M-Ratio
1	April	10	100	9	62	1
2	May	12	110	8	58	1.3
3	June	11	105	7	64	1.2
4	July	9	94	14	60	0.8
5	Aug	9	95	12	63	0.8
6	Sep	10	99	10	57	0.9
7	Oct	11	104	7	55	1
8	Nov	12	108	4	56	1.2
9	Dec	11	105	6	59	1.1
10	Jan	10	98	5	61	1.0
11	Feb	11	103	7	57	1.2
	March	12	110	6	60	1.2

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Then, you see this example. I told you that profit, sales, absenteeism, these are the things. Now, here profit and sales, they can be dependent variables and absenteeism, machine, M ratio can be independent variables.

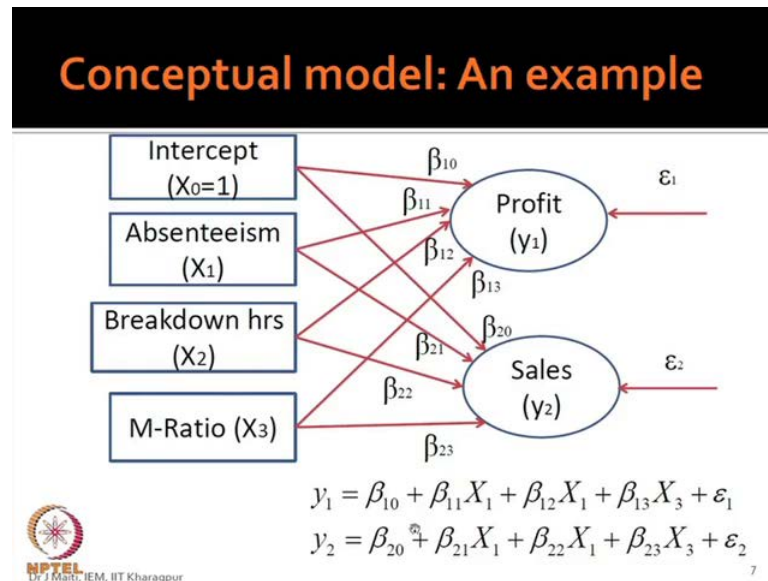
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Then, this is the example from this conceptual model point of view what we are discussing about the cities can. So, identify variables of interest. This is the first step. You have to, you identified that profits, sales, volume, percentage, absenteeism, machine breakdown, M ratio, and these are the variables of interest. Now, you have to identify

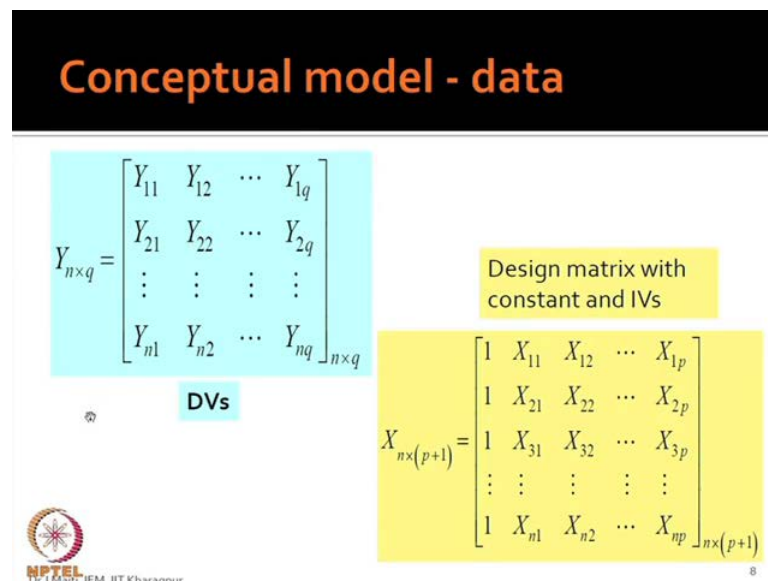
what are the dependent or response variables. So, in this case, Y1 is profit and Y2 is sales volume are the dependent variable. You have to identify the explanatory variables or independent variables of interest your X1, X2, X3, three independent variables. Then you have to find out the relationship, Y equal to function of X. In this case, we are talking about linear relationships.

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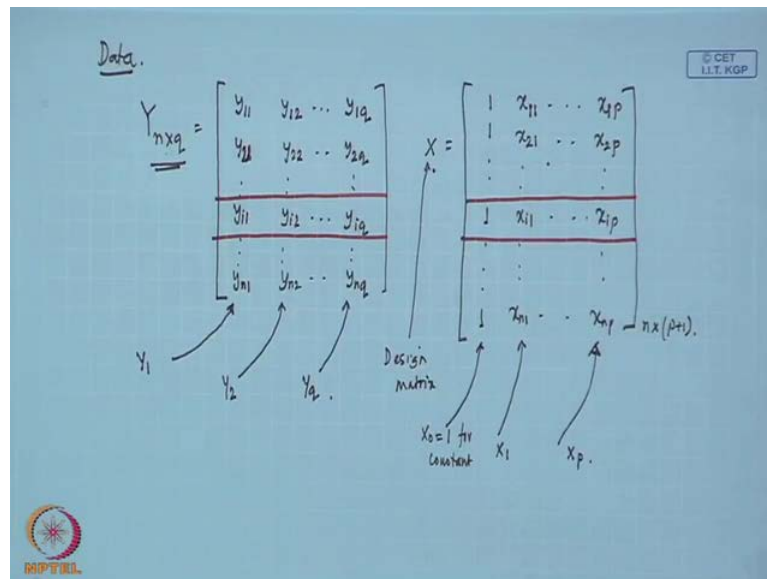
Then, this is your diagram for multivariate linear regression model. So, in general, what will happen? There will be p and there will be as I said, there will be q. That is the case.

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Now, come to the data point, come to data. So, we require data for Y.

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Suppose we will collect n data points on q Y variables that n data points on q Y variables will give you Y n cross q matrix. So, let us write down this y_{11}, y_{12}, y_{21} you write y_{11}, y_{21} , like this y_{i1} , like this y_{n1} . What is this? This data set, this data vector is on variable Y_1 , so n data points on variable Y_1 , so 1, 1 to n . So, I will go for variable 2, so first observation on variable 2, second observation on variable 2, so like this i th observation on variable 2, so like this y_{n2} , this is my second observation on variable 2.

So, in the same manner, you will be finding out first observation on variable 1, second observation, first observation on variable q , second observation on variable q , so like this i th observation on variable q . When I am saying variable 1 or variable, variable 2 or variable q , these are all related to Y variable, so same manner n observation on Y variable q . So, this is my Y_q . Now, this is what is data, this is related to the data. Now, so long, you are not collecting this data. These all data points are random and we, we, we will describe later that we assume multivariate normality. Then your another data is point is X , which we talk about as design matrix in this case.

So, if I say design matrix, then how many variables will be there, p variables plus constant, so all 1. So, this is the X_0 part equal to 1, this for constant. Then I will come to the second one, x variable observation 1 on variable 1, observation 2 on variable 1, observation i on variable 1, then observation n on variable 1, this is my X_1 . So,

similarly, you will be finding out x variable observation 1 on x variable p, variable observation 2 on p, so like this i on p, so like this n on p. So, this is your X_p independent side.

So, then this matrix is n cross q Y matrix and the X matrix will be n cross p plus 1, p plus 1. So, if you take with this, if you take the ith observation, now, now, we will take the ith observation for all the cases, we will be taking the ith observation, so ith observation is this here. Here, it is ith observation is this. So, let us write the ith observation. What is the status?

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The image shows a whiteboard with handwritten mathematical equations. At the top right, there is a small box containing the text "CET IIT KGP". The main content consists of several equations for different observations:

$$y_{i1} = \beta_{10} + \beta_{11}x_{i1} + \dots + \beta_{1p}x_{ip} + \epsilon_{i1}$$

$$y_{i2} = \beta_{20} + \beta_{21}x_{i1} + \dots + \beta_{2p}x_{ip} + \epsilon_{i2}$$

$$y_{ik} = \beta_{k0} + \beta_{k1}x_{i1} + \dots + \beta_{kp}x_{ip} + \epsilon_{ik}$$

$$y_{iq} = \beta_{q0} + \beta_{q1}x_{i1} + \dots + \beta_{qp}x_{ip} + \epsilon_{iq}$$

Below these equations, it is noted that $i = 1, 2, \dots, n$ observations. At the bottom, the general matrix form is written as:

$$Y_{n \times q} = X_{n \times (p+1)} \cdot B_{(p+1) \times q} + \epsilon_{n \times q} \quad \leftarrow \text{MLR}$$

The NPTEL logo is visible in the bottom left corner of the whiteboard image.

So, if you write the ith observation, you the general equation is y. So, we will write ith observation for which one for the first variable, so ith observation first y. So, this can be written like this beta1 is affected by 0 plus, plus beta1 is affected 1, then x ith observation. So, in this manner, you will be getting beta1 is affected by p, x ip. So, similarly, you will be getting for y i2, beta20, beta 2 is affected 1 x i1, like this beta2 is affected by p, x ip. Same manner, you can if you write the kth variable, y ik equal to beta I think k 0 plus beta k is affected 1 x i1, like this beta k is affected by p, x ip. Then you go for y ip k iq, this will be beta q0, beta q1, x i1, like this beta qp, x ip.

So, this is from the x point of view, y and x, but x cannot capture y in totality. So, the error term, we also require to add everywhere that we have not added here. So, this will be your related to first variable related to second y variable related to third k y variable

related to q y variable, but these are all related to ith observation, so x_{i1} , ϵ_{i1} , ϵ_{i2} , ϵ_{ik} , ϵ_{iq} . So, if you take this in general that for i equal to 1 to n observations because you have total n observations, so this set will be repeating n times. Getting me?

So, that is why, what happened? You are getting Y n cross q ; that Y_1 , this i is changed from 1 to n , this one, that variable change from q . Then this one, this side, you see that X , that X will be your, it also will be n cross p plus 1. Then there will be beta. This beta will be p plus 1 cross q plus, you have epsilon, which is n cross q . This is a general equation in terms of data, the regression equation multiple, multivariate multiple regression. Here, we are, we have included the data points.


So, this is our model. This is our model. Now, see that what is required? Similar to MLR, you require to estimate beta, you require to estimate the epsilon also, but here it is not a one dependent variable case, q dependent variable case. Your estimation is now more in number. Getting me? The totality is more.

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Data - example

$$Y = \begin{pmatrix} 10 & 100 \\ 12 & 110 \\ 11 & 105 \\ 9 & 94 \\ 9 & 95 \\ 10 & 99 \\ 11 & 104 \\ 12 & 108 \\ 11 & 105 \\ 10 & 98 \\ 11 & 103 \\ 12 & 110 \end{pmatrix}$$

$$X = \begin{pmatrix} 1 & 9 & 62 & 1 \\ 1 & 8 & 58 & 1.3 \\ 1 & 7 & 64 & 1.2 \\ 1 & 14 & 60 & 0.8 \\ 1 & 12 & 63 & 0.8 \\ 1 & 10 & 57 & 0.9 \\ 1 & 7 & 55 & 1 \\ 1 & 4 & 56 & 1.2 \\ 1 & 6 & 59 & 1.1 \\ 1 & 5 & 61 & 1 \\ 1 & 7 & 57 & 1.2 \\ 1 & 6 & 60 & 1.2 \end{pmatrix}$$



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So, we will proceed with estimation and this is one the same example what we have discussed. So, we have considering profit and sales, Y_1 , Y_2 and our absenteeism, breakdown hours, M ratio, so this is X . As I told you for X , there will be 1 and these are the X_1 , X_2 , X_3 , so this is 12 cross 1, 12 cross 1, 2, 3, 4, 4, 12 cross 4. This one is your 12 cross 2 matrix. So, we I will show you the result of this matrix basically this data set.

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Assumptions

- Errors ($\epsilon_{n \times q}$) are multivariate normal
- Error variances are equal (homogenous) across observations conditional on predictors
- Errors have common covariance structure across observations, and
- Independent observations.

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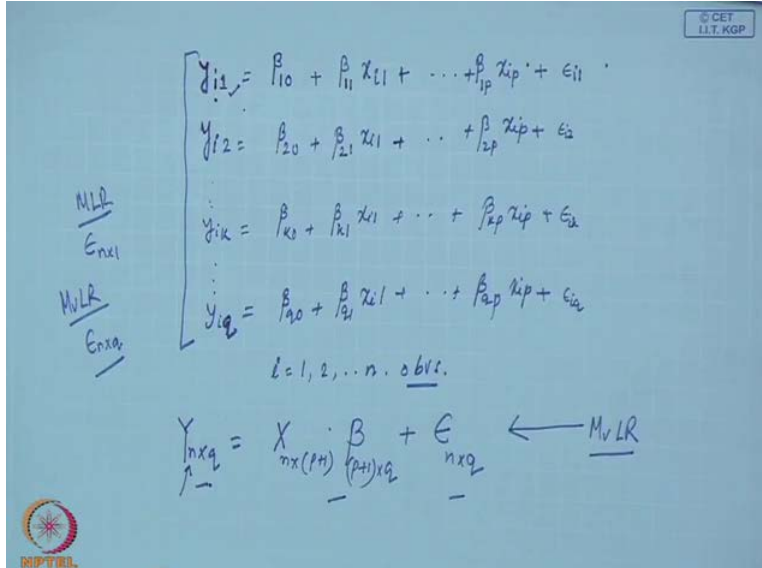


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Now, what are the assumptions? This model, MvLR is also applicable under certain assumptions like MLR. The assumptions are; you see errors are multivariate normal. Error is now epsilon n cross q. In MLR, what is the error? It is n cross 1.

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MLR
 $\epsilon_{n \times 1}$

MvLR
 $\epsilon_{n \times q}$


$$y_{i1} = \beta_{10} + \beta_{11}x_{i1} + \dots + \beta_{1p}x_{ip} + \epsilon_{i1}$$

$$y_{i2} = \beta_{20} + \beta_{21}x_{i1} + \dots + \beta_{2p}x_{ip} + \epsilon_{i2}$$

$$y_{ik} = \beta_{k0} + \beta_{k1}x_{i1} + \dots + \beta_{kp}x_{ip} + \epsilon_{ik}$$

$$y_{iq} = \beta_{q0} + \beta_{q1}x_{i1} + \dots + \beta_{qp}x_{ip} + \epsilon_{iq}$$

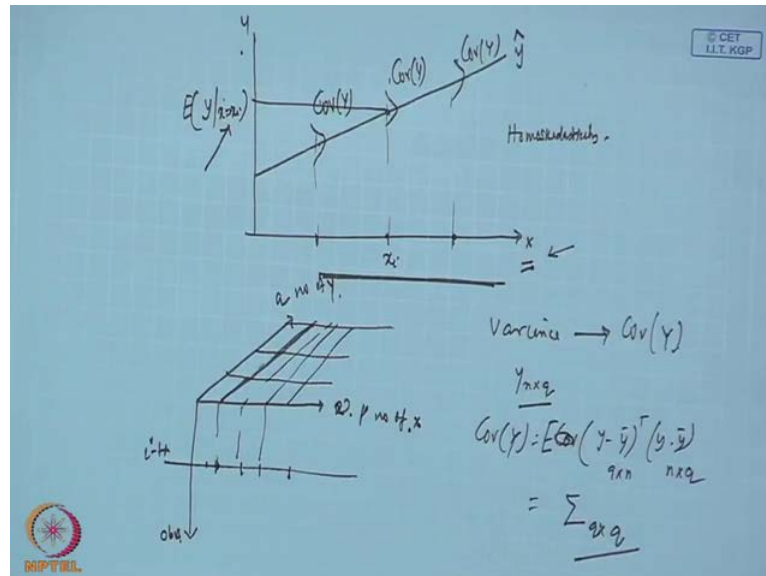
$i = 1, 2, \dots, n$. obv.

$$Y_{n \times q} = X_{n \times (p+1)} \cdot \beta_{(p+1) \times q} + \epsilon_{n \times q} \quad \leftarrow \text{MvLR}$$


In MLR, error is n cross 1. In MvLR, error is n cross q. Keep in mind this difference. So, in MLR, what we have discussed that the error variance, error variance across all x observations, they will be same, which is known as homogeneity condition, so

homogenous, sorry, that we can say homoskedasticity. Homoskedasticity, homoskedasticity problem we have discussed.

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Homoskedasticity problem with respect to simple regression, we have discussed x and y . Then we say this is my regression line, which is \hat{y} . Then when there are different observations, different observations, we say that the variance of y across x observations depends on values, this value, this value, this value, they are same that is homoskedasticity. So, here also, this remains same that error variances are equal across observations conditional on predictors. It is true, true because if this one is x_i , this particular value is expected value of y given x equal to x_i . This is conditional on predictor.

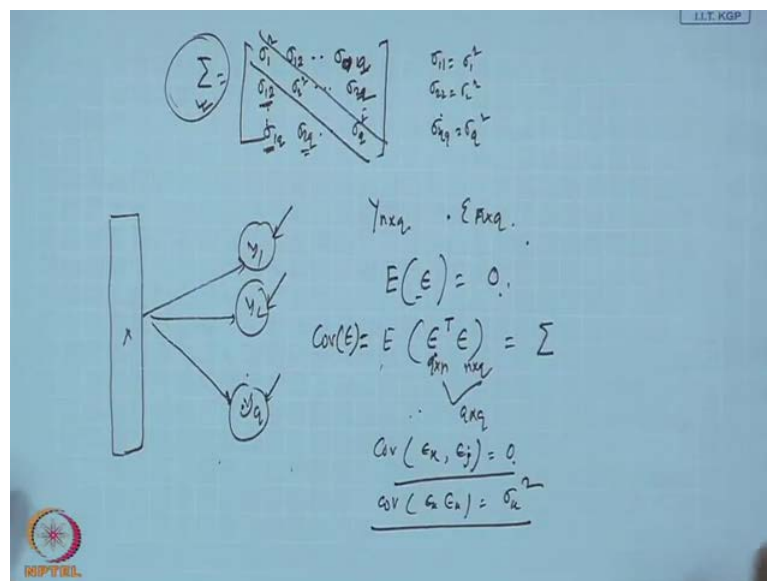
Now, errors have common covariance structure across observations. This is another assumption. Please understand that when we talk about MLR, we talk about across observations of X , we talk about error variances are equal, but here what is happening? Your x_i remains same, but y side is now multiple means what I mean to say? This is my X . Suppose this is my observations, this is the X variables, but you have q number of this is p number of X , q number of Y .

So, if I say that p this is my i th observation is this i th observation on x , so fine, this is 1, 2, 3 like this, but for this i th observations as you have q number of y variables, so what will happen ultimately here? So, everywhere, you take any observation, you are getting

this q Y combinations. So, that is why, the third assumption is when we have q variable, so this variance will converted to covariance, variance will be leading to covariance. so covariance of y. So, that means here I will be talking about covariance of y. Here, I will be talking about covariance of Y. Here, I will be talking about covariance of Y. Understood?

So, covariance of Y as Y is n cross q. What will be covariance of Y that is covariance of we can you can write that y minus this mean and all those things. So, finally, you can write y minus y bar that transpose, this is the expected value, expected value of this y minus y bar this will be q cross n, n cross q. So, you will be getting something like this, which is q cross q. Go back to the first class first or second lecture that is what we have discussed. So, we say that there are p variables. Now, I can write down this covariance structure like this.

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Now, I can write down this covariance structure like this, covariance sigma 1 square, sigma 12, sigma 1q here because we are taking q y variables sigma 12, sigma 2square, sigma 2q, now sigma 1q, sigma 2q, sigma q square. Actually, sigma 11 is sigma 1 square. So, sigma 22 is sigma 2 square, like this, sigma qq is sigma q square. This is the variance part.

So, in third assumption, we are saying error have covariance equal or common covariance structure, basically, common covariance structure means for this observation,

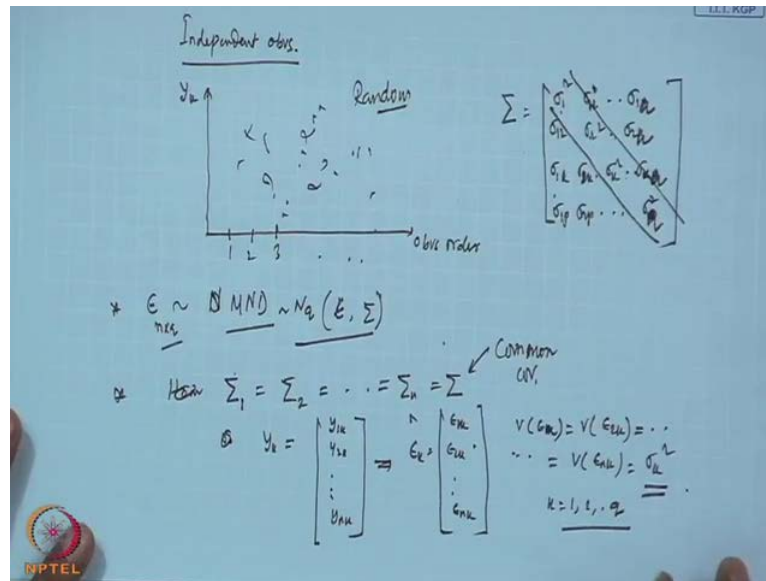
this observation, this observation, this observation, will be related to x , related to x . This is the i th observation and here I think, observation this side I think, we are saying n , basically these are all x variable for every observation, so i th observation. Similarly, for every observation what will happen ultimately? You will be finding σ , capital σ . This is the y s variability.

We are assuming here that see y , if you see the diagram, we are saying that y_1, y_2 like this, then y_q and this side is a set of x and we are linking to everywhere here. Then we are saying this is my multivariate, but we are not talking about anything related to y_1, y_2, y to y_q . We are assuming that are their correlation structure is insignificant or negligible basically. So, if this is the case, what will happen? These off diagonal elements become 0.

So, as all of you know that in the regression side that this variability part ultimately will be captured by error, so we can say this is the error covariance also. Your Y is n cross q . Similarly, error is also n cross q . So, if I want to find out what is the expected value of error, this will become 0. This is the assumption and expected value of error, transpose error. The reason is we want to find out the covariance of error. So, covariance of error is expected value of this minus mean transpose and this so as mean is 0, so this is the case. So, this one will be q cross n , this is n cross q , so ultimately this will be q cross q . So, this we are saying this.

So, that means the error, the regression error covariance matrix is this and we are saying that covariance from one dependent variable to other dependent variable that variability is like this the covariance between ϵ_k and ϵ_j , two, two different variables this 0 and covariance ϵ_k, ϵ_k that is σ_k^2 . Then so you have to, you require to test this.

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Then, the last one is independent observations, independent observations. How do we test independent observation? So, that means no auto correlation, independent observation means one observation Y , we are talking about related to Y . One observation is independent of others. It is not influenced by other observation. So, if you see the observation order vis a vis the plot for every variable, you can do this observation order. Then suppose this is Y_k , then you will be getting a plot random, a random plot observation, first observation, second observation, third observation like this.

So, nutshell, what is there? The errors are very important component in regression and that too true in multiple, multivariate multiple regression also. So, error will ultimately tell you all the assumptions. First one is it will be your multivariate normal distribution. So, it will be MND. So, we say error is n cross q . So, definitely it will be q . That means, it will be $n \times q$ that the mean value epsilon and you have to find out the variance, variance component, what will be the value? So, this you have to find out. Now, this is multivariate one. Second assumption is our homoskedasticity, equal covariance.

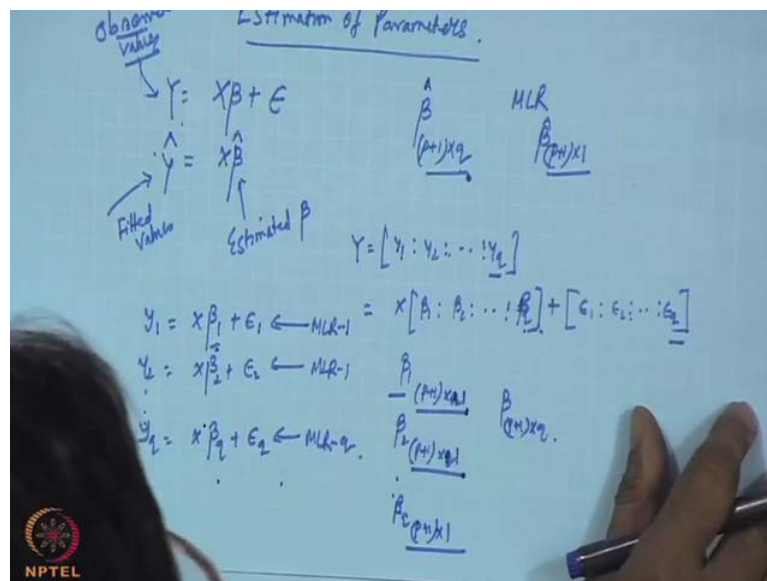
So, if you have observations, different observations, suppose n observations you have collected, then all will be equal to this. This is the common covariance matrix, your error variances are equal we say across observations, common covariance structure. That means another one is error variance means for every Y variable, if I say Y_k , then if I take the observation related to Y_k , then Y_{1k} , Y_{2k} like this Y_{nk} observations will be there and

corresponding error will be epsilon k. This one will be epsilon1k, epsilon2k like this epsilon nk.

Then, what we mean by that errors are equal variance? That means variance of epsilon1k equal to variance of epsilon2k like this equal to variance of epsilon nk that will be sigma k square. Keep in mind that will be sigma k square. So, k stands from 1 to. Now, where from you will get the k? You will get the k from the common covariance matrix sigma1 square, sigma12, sigma1p, sigma12, sigma2 square, sigma2p, sigma1k, sigma2k, sigma k square, sigma kp, so like this, sigma1p, sigma2p, sigma p will give q. Why q, q variable? P, we are using for independent side.

So, this diagonal matrix, I think you understood fully. So, these are the assumptions, these assumptions. Now, actually what will happen? Later on, you will see that you cannot, you do not know this value, this matrix. You have to estimate, you have to estimate from data. So, our next issue is estimation of parameters.

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I think this estimation of parameters, we will complete this session by using, describing this estimation of parameters. So, our equation is Y is X beta plus epsilon. If we estimate, we say Y cap equal to X beta cap, this is our fitted value. This is the estimated parameter, estimated beta, these are observed values and this is known the error term. So, I have something to say here. I mean something more to say here. What I mean to say that please remember that Y is, there are q Y variables. Getting me?

So, similarly, and p X variables, so that means this beta, this beta, what you estimate that will be that will be how many parameters will be there. It will be matrix. What will be that matrix? That will be p plus 1 cross q. Now, in MLR, what will happen? MLR beta is p plus 1 cross 1. So, additionally, it is coming like this. So, that means if I write like this Y equal to something like this Y1, Y2, then it is Yq, this Y, then this equal to we say X beta. Can I not write like this beta 1, beta 2, beta q, so plus you will get what epsilon? This epsilon is also having epsilon1, epsilon2 like epsilon q, no, no q, this is q, this is q, q will be there.

So, it is this equation is nothing but as we have seen basically q MLR, q number of MLR means Y1 equal to X beta 1 plus epsilon1, Y2 equal to X beta 2 plus epsilon2, similarly Yq equal to x beta q plus epsilon q. So, this is MLR 1, MLR 2, MLR q. So, q MLR is there. Then what we will have to estimate them? What we have to estimate?

We have to estimate beta 1 to beta q and again beta 1 is p plus q into p plus 1 into q and everywhere beta 1 is also p plus 1 cross q vector beta 2 also p plus 1 cross q, sorry beta 1 p plus 1 vector. This is also one vector. Please remember beta in a nutshell, beta is p plus 1 cross q, but beta 1 p plus 1 cross 1 beta 2 p plus cross into 1, like this, beta q that also p plus 1 cross 1 vector. So, if you go, what is the procedure in MLR, what way? We suppose one variable, Y variable and we estimate, we estimate first the error term.

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MLR $\rightarrow y_1 = x\beta_1 + \epsilon_1$
 $\epsilon_1 = y_1 - x\beta_1$
 $\epsilon = y - x\beta$
 $SSE = \epsilon^T \epsilon = (y - x\beta)^T (y - x\beta)$
 Scalar quantity $\frac{\partial SSE}{\partial \beta} = 0$
 $\hat{\beta}_1 = (X^T X)^{-1} X^T y_1$

MVL
 $Y = X\beta + \epsilon$
 $\epsilon = Y - X\beta$
 $SSCP_\epsilon = \epsilon^T \epsilon = \begin{bmatrix} \sum_{i=1}^n \epsilon_{i1}^2 & \sum_{i=1}^n \epsilon_{i1}\epsilon_{i2} & \dots & \sum_{i=1}^n \epsilon_{i1}\epsilon_{iq} \\ \sum_{i=1}^n \epsilon_{i2}\epsilon_{i1} & \sum_{i=1}^n \epsilon_{i2}^2 & \dots & \sum_{i=1}^n \epsilon_{i2}\epsilon_{iq} \\ \vdots & \vdots & \ddots & \vdots \\ \sum_{i=1}^n \epsilon_{iq}\epsilon_{i1} & \sum_{i=1}^n \epsilon_{iq}\epsilon_{i2} & \dots & \sum_{i=1}^n \epsilon_{iq}^2 \end{bmatrix}$

For example, if I recollect Y_1 equal to $X \beta_1$ plus ϵ_1 that one MLR, what you have done? You have done, first find out ϵ_1 that is Y_1 minus $X \beta_1$. So, if I forget about 1 because we are talking this subscript 1, we are talking about Y_1 variable in this case. So, general term is Y minus $X \beta$. Then you have found out $Y Y^T$ transpose that is $\epsilon^T \epsilon$ that is $(Y - X \beta)^T (Y - X \beta)$. Then you have this is what we say sum square errors, sum square errors.

Then, you have regress sum square errors. You have found out the derivatives of sum square SSE with respect to β and then you put it to 0. Finally, you got the equation like this $\hat{\beta} = (X^T X)^{-1} X^T Y$, this was the estimation procedure there. This is I am talking in terms of MLR multiple linear regression. Now, if we consider Y_1 in that, we want to keep that notation, then this will be β_1 and this will be Y_1 . X part remains same because of no multivariate regression, X part is unchanged. Y part is having additional Y , that is the issue.

So, in the same manner, so what happen ultimately here? In this case, when you collected n into 1 data points and ultimately, this one is 1 into n , n into 1. Ultimately, SSE is 1 into 1, so it is a scalar quantity. SSE is a scalar quantity in case of MLR. Now, what will happen in this case? Suppose we are coming to MvLR. MvLR same thing, Y is $X \beta$ plus ϵ , but this one we are saying n cross q , this one n cross p plus 1, this one is n p plus 1 cross q , this is your n cross q . So, let us see what will happen. We want to see a errors terms. So, error ϵ is now $Y - X \beta$. So, error is n cross q ; in sense also, $Y - X \beta$ is n cross q .

Now, we want to find out the sum square errors like MLR that is to be minimized. So, if I want to do this, then if I will make like this $\epsilon^T \epsilon$, this will be q cross n into ϵ , which is n cross q . So, what is happening here? You are getting a q cross q matrix. So, I cannot say it is a scalar quantity is a matrix. As a result, this will give you a structure like this and your this, this elements will be sum of i equal to 1 to n ϵ_i^2 . The second diagonal element will be i equal to 1 to n ϵ_i^2 . Similarly, the last diagonal will be i equal to 1 to n ϵ_i^2 and off diagonal element will be the covariance structure $\epsilon_i \epsilon_j$, $\epsilon_i \epsilon_j$ like this, so like this. Getting me?

So, your error now this is your sum squares error is no longer a scalar, it is a matrix. It has, it is symmetric, and it has diagonal elements, the variance component of the diagonal elements, the covariance component of this error terms although we assume that the covariance component will be 0 that is our assumption. So, that means this epsilon transpose epsilon is now sum square cross product. Getting me? Earlier, you have, we have described this sum square cross product epsilon E. We can write or you can say sum square epsilon, let it be E. E stands for error.

Now, we what we want? We want this matrix should be as minimum as possible because in MLR, we have seen ML that SSE the minimum. Then we derive, make derivatives, we get derivatives of SSE with respect to beta values, we put it to 0 and we got this equation. We have done the same thing here. We want this matrix to be as minimum as minimum as possible. Now, if I assume that the off diagonal elements are 0, then what we want? We do not want only that individual diagonal elements will be as minimum as possible. We want the total as minimum as possible, the total sum square across all the Y variables as minimum as possible.

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$$\text{trace}(SSCP_e) = \text{tr}(E'E)$$

$$= \sum \epsilon_{i1}^2 + \sum \epsilon_{i2}^2 + \dots + \sum \epsilon_{iq}^2$$

Minimize

$$\beta = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_p \end{bmatrix}$$

$$\frac{\partial \text{trace}(E'E)}{\partial \beta_j} = 0$$

$$j = 0, 1, \dots, p$$

$$E = Y - X\beta$$

$$\text{tr}(E'E) = (Y - X\beta)'(Y - X\beta)$$

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So, what we do that is we take the trace of SS CP e or trace of SS CP e or you can say trace value of epsilon transpose epsilon. This is the diagonal values. So, this will be sum total of diagonal values epsilon i 1 square, epsilon i 2 square, like this epsilon i q square. So, we will minimize the trace. Your aim is minimize this. Getting me? You have to

minimize the trace. So, now how to minimize? What is your control here? Your control is you want the estimate in such a manner that this will minimise this.

So, your control is beta. Now, beta is not single here. Beta is basically; what is the dimension of beta, p plus 1 into q. So, this first one is for beta 1 that is p plus 1 into 1. Second one will be for beta 2, so like this. Getting me? So, what do we require? We require that we will find out derivative of trace, derivative of trace of this with respect to particular beta. Getting me? Now, by saying this, if I say beta k, then it is coming to the univariate case. So, beta k actually again, you have j equal to 1 to 0 to 1, 0 to p number of independent variables. So, ultimately beta k j will come. So, these things you will put into 0, put into 0.

So, actually what will happen? You will, you can write like this epsilon this one, this quantity cannot be, we cannot write like this that epsilon is Y minus X beta. So, epsilon transpose epsilon is Y minus X beta transpose Y minus X beta. So, when we are taking derivative trace that means you are taking trace of this means trace of this. That means only the diagonal elements are coming, so this.

Then, when you are taking the derivative of this with respect to this, when you are talking about the first beta 1 and that too beta 10, then ultimately all other things become 0. Only that quantity will be used for derivation. So, ultimately it will end into that. That means with it end into the individual Y variables, individual Y variables.

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$$\hat{\beta} = (X^T X)^{-1} X^T [y_1, y_2, \dots, y_q]^T$$

$$= (X^T X)^{-1} X^T Y$$

MVR

$$\hat{\beta} = [\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_q]^T$$

$$= \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \\ \vdots \\ \hat{\beta}_q \end{bmatrix} = (X^T X)^{-1} X^T \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_q \end{bmatrix}$$

$$\hat{\beta}_1 = (X^T X)^{-1} X^T y_1$$

$$\hat{\beta}_2 = (X^T X)^{-1} X^T y_2$$

$$\hat{\beta}_q = (X^T X)^{-1} X^T y_q$$

As a result, when you do this type of manipulation, ultimately you will be getting beta cap, which is nothing but $X^T X^{-1} X^T$. I can say y_1, y_2 like this y_q . Getting me? Actually, this quantity will be multiplied to everyone. So, this totality when I write Y , then this one is $X^T X^{-1} X^T$ capital Y . What is the difference these between these MLR beta? Here, Y is q number of variables, Y variables in MLR that will be one. So, this is our estimated case. So, essentially what we can say here? We can say that in this is the least square estimate and in least square estimate for MvLR, you will get beta cap like this, which is your beta 1 cap, beta 2 cap beta q is there q cap, beta q variables are there. This is nothing but here beta 1 cap means beta 10, beta 11, so like this beta 1 is affected by your p . This is your beta 20, and then beta 2 is affected by 1 like this. Beta 2 is affected by p . Similarly, beta q 0, beta q 1, then beta q is affected by p . This is the case.

This one is as we are saying that $X^T X^{-1} X^T$ and we are saying that Y is there. So, can we not write like this? This one is similar to y_1, y_2 like this. This much is not required like this. Then y_1 is again, $y_1, 1$ to that n ; all those things are there, so nutshell. What I mean to say that I mean to say that you, if you want to do this one is $X^T X^{-1} X^T y_1$. Similarly, beta 2 cap is $X^T X^{-1} X^T y_2$. So, like this, beta q cap is $X^T X^{-1} X^T y_q$. Got it?

So, that means you may go for individual MLR and then whatever beta value you get and in multiple MvLR also, you will get the same thing. From beta point of view, there is absolutely no problem because here you are doing q number of calculations, computations. Here, in one computation, you are getting q beta vectors. So, both are same. Then where is the difference if I get the same beta value while going for q MLR? Why should I go for one MvLR? The error comes in, sorry not error, the difference comes in error terms because you have n errors. These errors are multivariate errors. Every observation is multivariate.

Earlier in MLR, your n errors are there, but from every univariate error, one observation per one error, but it is a multivariate observation now on q variables. That is the difference. In other way, there is no difference. So, that means as error plays the significant role in regression, so what do you want to do then? That means your all

treatment will be in the multivariate domain here that q , q every error observation is q data points, q number for y_1 to y_q .

So, test statistics differs. Test statistics will differ. It is similar to ANOVA and MANOVA. If you remember in ANOVA, ANOVA one response variable and several population in one way and two way that two factors and one response variable. Then you have used the ANOVA table. You have gone for f distribution, but in MANOVA, what happened? MANOVA, you have not got the straight cut distribution that is your SS, what I can say factor by your SS error. That means MS, MS factor that you did not get in MANOVA. Instead you got SSCP there and then you have gone for different test like Wilk's lambda test. It happened. Similar thing happened here. Getting me?

So, we will not go for like simple one coefficient of determination r square. We will go for; instead we will go for likelihood ratio test like Wilk's lambda. Getting me? That is the issue. So, ultimately your estimation part no problem, but model adequacy test will differ. Your test statistics will differ, but again the diagonal diagnostics case, you will find out the similarity with MLR. But, the beauty of this MLR is that may be that for the multivariate situation also, it can capture. So, many a times, we will not go for MvLR, but if MvLR, when there are more number of y variables, better to test through MvLR and because similar, it is not tough. Thank you very much. Next class, I will go for in detail the sampling distribution of beta, then model adequacy test and other things.