

Total Quality Management - II
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Lecture – 39
Fitting Regression Models – IV

Welcome back my dear friends, a very good morning, good afternoon and good evening to all of you. This is the TQM-II lecture series under the NPTEL MOOC and I am Raghunandan Sengupta from IME department IIT Kanpur. So, we are in the penultimate stage of this TQM-2 and lecture series and we have two lectures left, which is the 39th and the 40th in the 39th I will try to cover more topics, in 40th I will try to basically wrap it up I will discuss what are the things we should lay more stress on in the 40th lecture. So, let us continue the discussion.

So, if you remember we were discussing that giving given a multiple linear regression; a linear one and which way if we check keep saying time and again. Our main hypothesis was basically to have them in a in the very simple linear format find out the beta hats, which and they should be unbiased, then try to find out the generally the plot of the errors with respect to all the x's; x's are the independent variable then have the q-q plots to find out the distribution of for the errors or it can be done for the x's, y's because intrinsically we are considering they are all normal.

Then try to have a look at the covariance matrix or betas and then we went into how in if there are utilizing the concept of the regression with the fractional factorial models combining them; like 2 to the power k 3 to the power k. So, k is the number of variables or it can be attributes also. So, you can have a regression where some of them some of the x's are discrete, sum of the x's are continuous.

Then we considered that if they are orthogonal, how they can be done? Then on then the not orthogonal how the regression models can be done? And you can find out the beta hats. Then we also consider that if the folds can be done folding and trying to find out what is the overall relationship between x's and the y's. Then we went into basically testing with the linearity holds at the fag end of the last lecture which is the 30th lecture we considered that if all the betas are 0; that means, there is no linear relationship

between y and x. Because if beta 1 is 0 it means the rate of change of y with respect to x is not there it does not exist.

So, that means, there is no dependence and then we will consider the further on discussion from now on. But you know all these things which I which I did mention is that at the end of the day your main focus is to find basically have a table of the anova which is analysis of variance. Where in the first column I am again repeating it will be all the variables and the attributes and the second column will be the sum of the squares then in the third column it will degrees of freedom, fourth column it will the mean square; that means, the sum of the squares divided by the degrees of freedom.

Then using the ratios of the or the mean squares or the ratios of the sum of the squares divided by the degrees of freedom will give you their statistics and you will basically have enough value of alpha, level of confidence or considering the p value you will comment whether your hypothesis null is validated or not. And then where we would basically consider as things are in the subsequent lecture which is which we are going to cover to it today.

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If $H_0: \beta_j = 0$ is not rejected, then this indicates that x_j can be deleted from the model. The test statistic for this hypothesis is

$$t_0 = \frac{\hat{\beta}_j}{\sqrt{\hat{\sigma}^2 C_{jj}}} \quad (10.28)$$

where C_{jj} is the diagonal element of $(\mathbf{X}'\mathbf{X})^{-1}$ corresponding to $\hat{\beta}_j$. The null hypothesis $H_0: \beta_j = 0$ is rejected if $|t_0| > t_{\alpha/2}$. Note that this is really a partial or marginal test because the regression coefficient $\hat{\beta}_j$ depends on all the other regressor variables x_i ($i \neq j$) that are in the model.

The denominator of Equation 10.28, $\sqrt{\hat{\sigma}^2 C_{jj}}$, is often called the **standard error** of the regression coefficient $\hat{\beta}_j$. That is,

$$se(\hat{\beta}_j) = \sqrt{\hat{\sigma}^2 C_{jj}} \quad (10.29)$$

Therefore, an equivalent way to write the test statistic in Equation (10.28) is

$$t_0 = \frac{\hat{\beta}_j}{se(\hat{\beta}_j)} \quad (10.30)$$

Now, if H_0 which is the null hypothesis β_j is 0 is not rejected, which means there is some linear relationship. Then this indicate that that x_j can be deleted from the model. So, if it is not rejected we will and consider not consider depending on what the actual hypothesis is. The test statistic for the hypothesis would be the t value. So, this t

value or the t statistics would be $\hat{\beta}_j$. Because $\hat{\beta}_j$ is basically the estimated value of β and divided by $\sqrt{\hat{\sigma}^2 c_{jj}}$.

So, c_{jj} are the diagonals of the $X^T X$ inverse matrix. So, that was one part when you are trying to find out the beta hats. So, beta hats would basically be found out as that in the long term, the expected value of the beta hats are exactly equal to the β .

So, that is basically the characteristics of unbiasedness. The t statistics are under H_0 would be $\hat{\beta}_j$ and divided by square root of $\hat{\sigma}^2 c_{jj}$. So, they would be the corresponding columns of the principal diagonal of that. Because it is c_{jj} it will be the values accordingly. The null hypothesis is H_0 is basically if β is β_0 ; it will be rejected, if the value of both because null hypothesis can be you are taking both sides it side it.

So, if it both sided obviously, it will be either on to the left of the line depending on which way we are looking at. The word what I am saying is that which way you are looking at is that whether the value of β_0 is equal to 0 or within the band range considering the interval estimation it is plus or minus; that means, plus some delta value with right of 0 and minus some delta value on the left of 0, if you are basically drawing the number line where 0 is the value where the null hypothesis basically is validated.

Now once you remember this concept of β_0 being equal to 0 can be changed as β_0 is equal to say for example, some fixed value where we consider the linearity relationship between the x's or x_j 's and the y's are constant and in another case a null hypothesis we can consider it is changing.

So that means, consider that a under H_0 β_j is say for example, 3. So that means, the rate of change of y with respect to β_j with x that corresponding x is 3 units which is the regression coefficient and on the null hypothesis it is 3 and on the alternative hypothesis is not equal to 3. Then also the interval estimation would be done accordingly where it will be 3 plus some delta 3 minus some delta you will basically consider that hypothesis is satisfied outside this interval is not satisfied.

Now this interval plus delta and minus delta which I mean mentioning would depend on the level of confidence and the distribution you are taking. So, it will be t distribution with corresponding values being there what are those values? I am going to come to that.

So the null hypothesis β_0 is equal to β_j under the null hypothesis being rejected would be true if this is true; that means, it is greater.

So, now check here what I did mention the right hand left hand. So, I will just change the color. So, this $\alpha/2$ basically means the level of confidence. So, if you have I am just drawing the distribution very simply it is not normal it is basic consider is a t distribution. So these are the intervals. These values should be equal to $\alpha/2$, this value equal to $\alpha/2$.

And the region in between this color or this region would basically be equal to α . hence this oh $1 - \alpha$ sorry my mistake, $1 - \alpha$. So, $1 - \alpha/2 + \alpha/2$ the total area is 1; as it should be. And finally, this $n - k - 1$ is basically the degrees of freedom depending on the experiment we are doing.

So, n is the total number of observations, k if you remember is the number of variables depending on which y can be estimated and minus one is basically you are losing the degrees of freedom for the for that was one. Note that this is really a partial marginal test because the regression coefficient β_j depends on all the other regression coefficient which is true. Because β_j would depend on $\beta_0, \beta_1, \beta_2, \dots, \beta_{j-1}, \beta_{j+1}, \beta_{j+2}$ so on and so forth β_k .

On other regression coefficients β_i I should basically mention not depend on β_0 till β_{j-1} and β_{j+1} till β_k , I am saying they should depend on the x values. So, x_1 to x_{j-1} from then x_{j+1} till say for example, x_k . So, so let me continue reading it. So, the regression coefficient β_j depends on all the other regression variables x_i ; i is not equal to j that are in the model considered; the denominator of this equation which is the square root of $\sigma^2_{\epsilon_j}$ is called the standard error of the regression coefficient β_j because it is basically are trying to multiply the variance with the multiplying factor.

So, the standard error so se is the standard error I will repeat I have mentioned it very clearly in the last class and also I faintly remember I did mention in the 37th class also. So, basically it means that standard error is the standard deviation of the sample. So, which is the counterpart of the standard deviation which you have from the population its counterpart in the sample is basically standard error.

So, let me continue reading it standard error for beta j is equal to square root of sigma square hat c j therefore, an equivalent way to write this test statistics would be as given t naught is equal to beta j hat obviously and divided by the standard error of be beta j hat. So obviously, you are trying to normalize it and find out the standard the tests statistics.

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Consider the following model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon} = \mathbf{X}_1\boldsymbol{\beta}_1 + \mathbf{X}_2\boldsymbol{\beta}_2 + \boldsymbol{\epsilon} \quad (10.32)$$

where \mathbf{X}_1 represents the columns of \mathbf{X} associated with $\boldsymbol{\beta}_1$ and \mathbf{X}_2 represents the columns of \mathbf{X} associated with $\boldsymbol{\beta}_2$.

$SS_R(\boldsymbol{\beta})$ is called the regression sum of squares due to $\boldsymbol{\beta}$.

- It can be mathematically expressed as:

$$SS_R(\boldsymbol{\beta}) = \hat{\boldsymbol{\beta}}' \mathbf{X}' \mathbf{y} \quad (p \text{ degrees of freedom})$$

Now consider the model as given. So, you have basically y is equal to so there is no beta naught so is equal to beta 1 X 1 plus beta 2 ma X 2 and plus epsilon. Now the word which I mentioned there is no beta naught technically see the equation they are all bolds; bolds mean they all are matrices. So, let me if I read it will become clear to you where X i represents the columns of X associated with beta i.

So, basically have been able to divide the whole matrix into two parts need not be equal. So, X the corresponding value of X where and those that is X i, X 1 is a matrix is associated with the beta 1 vector and X 2 which is the matrix is associated with beta 2 vector. So, the stand this is the sum of squares are for the residuals which is beta is called the regression sum of squares due to beta and obviously you will have sum of squares coming considering the error part where the suffix will be capital E

It can be mathematically expressed as sum of the squares for the betas being equal to beta hat transpose. So, it was basically a vector it will now become column vector will become row vector multiplied by X transpose into y. So, all these are y is basically of

size n cross 1. Similarly your X value would be of the size of k into n because then in that case X transpose y would basically now become k into cross 1 and so on and so forth.

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The regression sum of squares due to β_1 given that β_2 is already in the model is

$$SS_R(\beta_1|\beta_2) = SS_R(\beta) - SS_R(\beta_2) \quad (10.35)$$

This sum of squares has r degrees of freedom. It is the "extra sum of squares" due to β_1 . Note that $SS_R(\beta_1|\beta_2)$ is the increase in the regression sum of squares due to inclusion of variables x_1, x_2, \dots, x_r in the model.

Now, $SS_R(\beta_1|\beta_2)$ is independent of MS_E , and the null hypothesis $\beta_1 = 0$ may be tested by the statistic

$$F_0 = \frac{SS_R(\beta_1|\beta_2)/r}{MS_E} \quad (10.36)$$

If $F_0 > F_{\alpha, r, n-k-r}$ we reject H_0 , concluding that at least one of the parameters in β_1 is not zero, and, consequently, at least one of the variables x_1, x_2, \dots, x_r in \mathbf{X}_1 contributes significantly to the regression model. Some authors call the test in Equation 10.36 a **partial F test**.

So, if you basically do the multiplication accordingly for the matrices. The regression sum of the squares due to beta 1 and beta 2 k is already given in the model. So, given the sum of the squares for beta 1 given beta 2 or beta 2 given beta 1 would basically be given by this formula; where some of those squares of betas minus sum of the squares are from coming out from the beta 2 would be the sum of squares for beta 1 given beta 2. And if you replace 1 and 2 interchange them it will be exactly the same.

That means sum of the squares of beta 2 given beta 1 would be sum of the squares of beta minus sum of the squares of beta 1 the sum of the squares as r degrees of freedom; it is an extra sum of the squares due to the beta 1 which is happening. So, k the number of variables which are independent had been divided into r and k minus r .

So, whether r is for beta 1 or r is for beta 2 is basically a nomenclature and just a idea method how you are trying to do it, but remember that this matrix multiplication should be commensurate in the sense there are some and the rows and the columns should be such that the matrix multiplication concept stands true.

Note that the sum of the squares of beta 1 given beta 2 is the increase in the regression sum of the squares due to the inclusion the variables x_1, x_2, x_3 till x_r . So that means, in a one case I am considering there to be added to the models and then what is the increase in the standard errors or sum of the squares in other case I am excluding them. So, what is the increase in the sum of the squares? Now sum of the squares are give beta 1 given beta 2 is independent of the mean squares and thus the null hypothesis where beta one is equal to 0 may be tested by the statistics given of f statistics, given some of the squares of beta 1 given beta 2 divided is by its degrees of freedom divided by the mean square for the errors.

So, the errors also you have the sum of the squares divided by the corresponding degrees of freedom. So, you are taking the ratio of the sum of the squares divided by their degrees of freedom similarly in the denominator you also have the sum of the squares divided by the degrees of freedom you will give the f statistics under the null hypothesis and then you will test depending on the value of alpha and the p value whether you basically reject or accept H_0 .

If F_{naught} which is under null hypothesis is greater than that r f suffix alpha again I am repeating is the level of significance r is the degrees of freedom and minus p is basically the total dig. So, in F values you basically at you try to take the ratios of two samples. So, in this case one sample has degrees of freedom r and another case the degrees of freedom is n minus p . So, p if you remember depending on the problem it has can be either k plus 1 or k only depending on whether you have beta naught yes or no.

So, so that can be changed accordingly. So if F_{naught} is greater than this value what is this value I will just circulate, so this one. So, the here F_{α} again I am repeating is the this value.

So, this is one sided that is why it is not divided alpha by 2. So, the in the in the in the other equation we had basically left hands right hand side and left hand side. Now you are taking only one sided; because either greater than or less than So, it is greater than F_{α} r is the degrees of freedom coming from the addition n minus p is the degrees of freedom which is coming from the total we reject H_0 in concluding then at least one of the parameters in beta 1 is not 0 and consequently at least one the variables x_1 to x_r in X capital X 1 which is basically the matrices contribute significantly to the

regression model. Some authors call the test as a partial F test because you are basically trying to find out the overall effect of x_2 or a overall effect of not coming for our depending on that you basically pass a judgement whether you want to accept H_0 or reject H_0

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EXAMPLE 10.6

Consider the viscosity data in Example 10.1. Suppose that we wish to investigate the contribution of the variable x_2 (feed rate) to the model. That is, the hypotheses we wish to test are

$$H_0: \beta_2 = 0$$

$$H_1: \beta_2 \neq 0$$

This will require the extra sum of squares due to β_2 , or

$$SS_R(\beta_2 | \beta_1, \beta_0) = SS_R(\beta_0, \beta_1, \beta_2) - SS_R(\beta_0, \beta_1)$$

$$= SS_R(\beta_1, \beta_2 | \beta_0) - SS_R(\beta_2 | \beta_0)$$

Now from Table 10.4, where we tested for significance of regression, we have

$$SS_R(\beta_1, \beta_2 | \beta_0) = 44,157.1$$

which was called the model sum of squares in the table. This sum of squares has two degrees of freedom. The reduced model is

$$\hat{y} = \beta_0 + \beta_1 x_1 + \epsilon$$

The least squares fit for this model is

$$\hat{y} = 1652.3955 + 7.6397x_1$$

and the regression sum of squares for this model (with one degree of freedom) is

$$SS_R(\beta_1 | \beta_0) = 40,840.8$$

Note that $SS_R(\beta_1 | \beta_0)$ is shown at the bottom of the Minitab output in Table 10.4 under the heading "Seq SS."

Consider the viscosity model again, suppose that we wish to investigate the contribution of the variables x_2 and that is the hypothesis we wish to test is accordingly under H_0 β_2 is 0. Under H_1 β_2 is not equal to 0. So, this will require the extra sum of the squares due to coming from β_2 . So, we can find out the sum of the squares of β_2 given β_1 β_0 .

So now you have been able to differentiate and find out the matrices into two parts one coming from β_2 one coming from β_1 and β_0 . So now from this tables we can find out the sum of the squares of β_1 β_2 given β_0 is 44157 which is called this model sum of the squares in the table. This sum of the tables has two degrees of freedom depending on that the deduce model would be like this.

So, you have \hat{y} now it is basically \hat{y} because you are trying to estimate will be equal to β_0 $\hat{\beta}_0$ which is 1, 6, 5, 2 that is 1652.39 plus β_1 $\hat{\beta}_1$ x_1 where β_1 $\hat{\beta}_1$ is equal to 7.6397 plus ϵ would not be there because this is \hat{y} hence we are trying to estimate so errors would come accordingly later on. And the regression sum of the squares of these models are given by sum of the squares of the of the of r given β_0

1 for beta 1 given beta naught it comes out to be 40840. So, note that this value is shown in the bottom of the of the output and depending on this on the way you have basically been able to mention them.

So obviously, it can be sum of the squares of beta 2 given beta naught also depending on how you have been able to the so called partition the effects of r and n minus k.

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Therefore,

$$SS_R(\beta_2 | \beta_0, \beta_1) = 44,157.1 - 40,840.8 = 3316.3$$

with $2 - 1 = 1$ degree of freedom. This is the increase in the regression sum of squares that results from adding x_2 to a model already containing x_1 , and it is shown at the bottom of the Minitab output on Table 10.4. To test $H_0: \beta_2 = 0$, from the test statistic we obtain

$$F_0 = \frac{SS_R(\beta_2 | \beta_0, \beta_1) / 1}{MS_E} = \frac{3316.3 / 1}{267.604} = 12.3926$$

Note that MS_E from the full model (Table 10.4) is used in the denominator of F_0 . Now, because $F_{0.05, 1, 13} = 4.67$, we would reject $H_0: \beta_2 = 0$ and conclude that x_2 (feed rate) contributes significantly to the model.

Because this partial F test involves only a single regressor, it is equivalent to the t -test because the square

of a t random variable with v degrees of freedom is an F random variable with 1 and v degrees of freedom. To see this, note from Table 10.4 that the t -statistic for $H_0: \beta_2 = 0$ resulted in $t_0 = 3.5203$ and that $t_0^2 = (3.5203)^2 = 12.3926 \approx F_0$.

Handwritten green notes on the slide include:

- α (pointing to the significance level in the F-distribution table)
- r (pointing to the degrees of freedom for the numerator, 1)
- $n - p$ (pointing to the degrees of freedom for the denominator, 13)

So therefore, S S R of beta 2 given beta 2 and beta naught is 3316.3. So, in this case you have degrees of freedom as 1 this is an increase in the regression sum of those squares there is results from adding x_2 to the model already occurred. So, you had x_1 you are now basically adding another variable x_2 . So, if you keep adding what is the sum of the squares change.

Now and whether the linearity of the variables holds to test $H_0: \beta_2 = 0$ versus $\beta_2 \neq 0$, you can find out the F statistics that comes out to be 12.39. Known at the mean square of errors which is basically given in the denominators 267.604 is used, once you find out the values the F value F statistic based on which you will cross a judgment where the F naught is greater or smaller is 4.67.

Now, if you should note down these values. So, this is basically alpha this is basically the degrees of freedom which is r and this is basically n minus p; consider p p is basically k plus 1 or k depending on how you have been able to mention in a model. Because this is

partial F's test involves only a single regressor, it is equivalent to the t-test because this square of the t random variable with chi degrees considered so it is written as v consider this is v degrees of freedom is an F random variable with 1 and v degrees of freedom accordingly.

So, here one is r and n minus k is v. To see note that that in the table we can find out H naught as beta naught is equal to 0 and you can test the t take t values comes out to be 3.55 and you can test using the t statistics and the values of t statistics are such that is almost equal to F naught. So, depending on whether you want to add 1 or take about 1 they would be equivalent to the t statistics, but if you can keep adding more than 1 then obviously in the F statistics value would change accordingly.

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Confidence Intervals on the Individual Regression Coefficients

Because the least squares estimator $\hat{\beta}$ is a linear combination of the observations, it follows that $\hat{\beta}$ is normally distributed with mean vector β and covariance matrix $\sigma^2(X'X)^{-1}$. Then each of the statistics

$$\frac{\hat{\beta}_j - \beta_j}{\sqrt{\hat{\sigma}^2 C_{jj}}} \quad j = 0, 1, \dots, k \quad (10.37)$$

is distributed as t with $n - p$ degrees of freedom, where C_{jj} is the (jj) th element of the $(X'X)^{-1}$ matrix, and $\hat{\sigma}^2$ is the estimate of the error variance, obtained from Equation 10.17. Therefore, a $100(1 - \alpha)$ percent confidence interval for the regression coefficient β_j , $j = 0, 1, \dots, k$, is

$$\hat{\beta}_j - t_{\alpha/2, n-p} \sqrt{\hat{\sigma}^2 C_{jj}} \leq \beta_j \leq \hat{\beta}_j + t_{\alpha/2, n-p} \sqrt{\hat{\sigma}^2 C_{jj}} \quad (10.38)$$

$$\hat{\sigma}^2 = \frac{SS_E}{n - p} \quad (10.17)$$

Now dif so this initially was about the hypothesis testing. Now if you want to basically consider the interval estimation of the individual regression coefficients because the least square estimates beta hat is a linear combination of the observation it follows that b hat is normally b hat beta hat is basically normally distributed with mean value of beta as mentioned according the concept of unbiasedness.

And covariance matrix beam if you remember for the for covariance of betas so it will be a k plus 1 into k plus 1 or k cross k matrix with the principal diagonals are all the standard errors or the square of the standard square of the standard errors or the betas and the off the diagonal element are the so called covariance is existing between beta j and

beta k or beta i into beta j where basically i and j change from 1 to k or 1 to $k + 1$ depending on how many such variables you have.

So, then each of the statistics would basically be that we want to find out the considering the normality holds. So it obviously it would become a very simple test of z value. So, in the z value what you have when you are converting an x which is a random variable which is normal into z statistics or the standard normal distribution. So, it will be x in the new denominator you will have basically the square root of the variance or the standard deviation and in the numerator you will have basically that that random variable minus its expected value.

So, if you basically want to find a similar to that what I mentioned just look at this equation. So, it will be $\hat{\beta}_j - \beta_j$. So, $\hat{\beta}_j$ is the random variable β_j is basically is the expected value divided by square root of $\sigma^2 c_{jj}$ which is if use mean remember the standard at standardized standard error.

So, it is distributed as t distribution with $n - p$ degrees of freedom, where c_{jj} is the j th element of the on the matrix $X^T X$ inverse of that and σ^2 is the estimate of the error variance obtained earlier. So obviously, this $100 - 1 - \alpha$ is basically the levels of significance and based on that if you basically find out the interval; the interval would come like this.

So, this is basically the central value and plus minus a further interval in which you want to basically find out. So, the left value would be $\hat{\beta}_j - t_{\alpha/2, n-p} \sqrt{\sigma^2 c_{jj}}$ and on the right hand side it will be plus. Now note down one thing it is we are trying to find our interval. So, if it is interval obviously that value which you have so it will be $\alpha/2$ here $\alpha/2$ here. So, the whole area inside is basically $1 - \alpha$. And $n - p$ is basically a degrees of freedom $n - k + 1$ or n in the bracket $k + 1$ or $n - k$ would depend on number of variables which you have.

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EXAMPLE 10.7

We will construct a 95 percent confidence interval for the parameter β_1 in Example 10.1. Now $\hat{\beta}_1 = 7.62129$, and because $\hat{\sigma}^2 = 267.604$ and $C_{11} = 1.429184 \times 10^{-3}$, we find that

$$\hat{\beta}_1 - t_{0.025,13} \sqrt{\hat{\sigma}^2 C_{11}} \leq \beta_1 \leq \hat{\beta}_1 + t_{0.025,13} \sqrt{\hat{\sigma}^2 C_{11}}$$
$$7.62129 - 2.16 \sqrt{(267.604)(1.429184 \times 10^{-3})} \leq \beta_1$$
$$\leq 7.62129 + 2.16 \sqrt{(267.604)(1.429184 \times 10^{-3})}$$
$$7.62129 - 2.16(0.6184) \leq \beta_1 \leq 7.62129 + 2.16(0.6184)$$

and the 95 percent confidence interval on β_1 is

$$6.2855 \leq \beta_1 \leq 8.9570$$

So, we can construct a 95 percent, 90 percent, 99 percent confidence in interval. Now beta 1 hat is equal to 7.62 and because sigma square hat is given as 267 and c 1 is given as 1.42 into 10 to the power minus 3. We can find out the confidence interval where the confidence interval for beta 1 comes out to be in between at a 95 confidence level it will come out to be 6.2855 is the left limit and right limit is basically 8.9570.

So, change 95 and the values of corresponding of the left limit and the right limit will change depending on that you can find out the very new limits for the confidence interval for betas depending on the level of significance.

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Prediction of New Response Observations

A regression model can be used to predict future observations on the response y corresponding to particular values of the regressor variables, say $x_{01}, x_{02}, \dots, x_{0k}$. If $\mathbf{x}_0' = [1, x_{01}, x_{02}, \dots, x_{0k}]$, then a point estimate for the future observation y_0 at the point $x_{01}, x_{02}, \dots, x_{0k}$ is computed from Equation 10.39:

$$\hat{y}(\mathbf{x}_0) = \mathbf{x}_0' \hat{\boldsymbol{\beta}}$$

A $100(1 - \alpha)$ percent prediction interval for this future observation is

$$\begin{aligned} \hat{y}(\mathbf{x}_0) \pm t_{\alpha/2, n-p} \sqrt{\hat{\sigma}^2 (1 + \mathbf{x}_0' (\mathbf{X}'\mathbf{X})^{-1} \mathbf{x}_0)} &\leq y_0 \\ &\leq \hat{y}(\mathbf{x}_0) + t_{\alpha/2, n-p} \sqrt{\hat{\sigma}^2 (1 + \mathbf{x}_0' (\mathbf{X}'\mathbf{X})^{-1} \mathbf{x}_0)} \end{aligned} \quad (10.42)$$

In predicting new observations and in estimating the mean response at a given point $x_{01}, x_{02}, \dots, x_{0k}$, we must be careful about extrapolating beyond the region containing the original observations. It is very possible that a model that fits well in the region of the original data will no longer fit well outside of that region.

Prediction on the new response observations; now consider a regression model it can also be used to predict future observation so obviously we have been mentioning time and again you want to basically find out \hat{y} for the n th plus 1 \hat{y} for the n th plus 2 so on and so forth find out the error difference for n plus 1 or n plus 2 or n plus 3 and if you want to double checks find out the sum of the errors they should actually be equal to 0 ok.

So, without repeating let me come to the problem a regression model can be used to predict future observations on the response y corresponding to particular values of the regression variables say x_1, x_2 so on and so forth in x_k , so that means you are taking the first second third fourth the k th variable and their responses say for example, for the n th plus 1. Now in this case you will basically write \mathbf{X} transpose as very simply mentioned is now in basically the scalar vector or a sorry is basically a row vector or a or a column vector where the first element remains as 1.

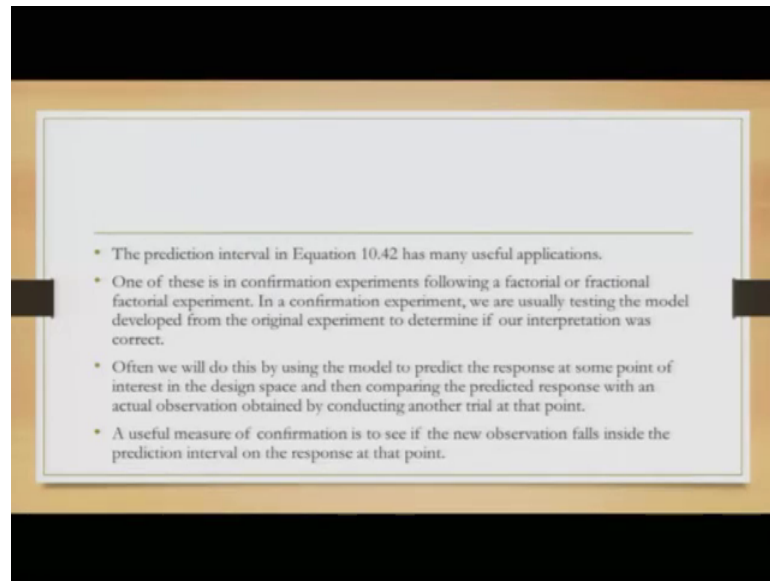
Because we are considering β is there. So, hence point estimate or the future prediction model would be found out when we multiply if you remember correctly we will multiply what $\hat{\beta}$ plus $\hat{\beta}_1$ into x_{n+1} which is given as x_0 1 here plus $\hat{\beta}_2$ into x_2 which is the random variable its n th plus 1 value which is given as x_2 plus dot dot till the last value which will basically be $\hat{\beta}_k$ multiplied by x_k values and it is n plus 1 reading; which here is given as x_k

So, if you I want to basically predict person 100 minus 1 minus alpha which is the level of confidence which you have person prediction interval for this feature observations you can find out that it will basically be y_{n+1} ; y_{n+1} is basically the without the hat is basically y suffix $n+1$ reading. So it is left limit and right limit would basically be comprised of the values which you have for the using the t statistics and you can utilize that to find out that how the interval and depending on the level of confidence you have how the interval will change.

So, if we change alpha obviously, the intervals will increase or decrease depending on how you have considered that. In predicting new observations and in estimating the mean response to a given point set of values which is x_1 , x_2 , x_3 till x_k we must be careful about extrapolating beyond the region containing the original observation because if you remember and they would definitely be increase or decrease of the variance and increase and decrease on the trends so on and so forth.

So, how you are going to consider that that would be important we always be careful about extrapolating beyond the region containing the original observations it is very possible that our model that fits well in the region on the original data will no longer fit felt outside the region because as I mentioned for those points depending on change in the trend, change in the seasonality or change in say for example, pressure for of this example change in the viscosity, change in the pressure, change in the flow rate, change in the temperature and humidity and so on and so forth.

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So, the prediction interval has would have many useful applications one of those is the confirmation experiment following a factorial or fractional factorial experiment. In a confirmation experiment we are easily testing the model develop for original experiment to determine if our interpretations were correct. So, given an in sample and out sample we will first test our model for the in sample and then try to utilize that model and that model for the out sample and find out the robustness of model.

Often we will do this by using the model to predict the response at some point of time of interest and the design space and then comparing the predicted response with an actual observations or it can be a useful measure for the confirmation. So, as to see if the new absorption falls within that in confidence interval such that we are quite confident that the model or the predicted values or the x 's, y 's and whatever model we have proposed is right. So, with this I will end the 39th lecture and continue some discussion the 40th one and then wrap up this course in the end of the 40th lecture.

Thank you very much for your attention. Have a nice day.