

Introduction To Adaptive Signal Processing
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Lecture No # 05
Multivariate Gaussian Density

Ok. So, today we will be discussing Gaussian density, but not just for one variable for many real continuous random variables together. So, that will be called joint, the jointly Gaussian. Gaussian is sometimes also called normal, jointly Gaussian random variables. This jointly is sometimes also called multivariate, there are different names. So, here we are given capital N, real continuous random variables say x_1, x_2, \dots, x_N .

They are jointly Gaussian. So, every experiment I will observe them together. They may have a relation between them one can be temperature, another can be humidity, another can be anything. They will be called jointly Gaussian if they are joint probability density.

We have discussed what is joint probability density for two variables $X Y$ then for three variables. So, in general the joint probability density if it takes a particular formula form, then it will be called that they will be called jointly Gaussian. Before that let us see, suppose x is the vector that is all the random variables in stack, they are put in a stack in a vector. We have already discussed it yesterday towards the end, same thing I am saying, and a mean vector is a vector of μ_1 is a mean of expected value of x_1 , μ_2 expected value of x_2 μ_N then we know we can take the deviation the increment vector $x - \mu$, that is every random variable, the incremental value it takes around its mean $x_1 - \mu_1$. So, sometimes can be positive, sometimes can be negative and for all other variables.

$$\underline{x} = \begin{bmatrix} x_1 \\ x_2 \\ \cdot \\ \cdot \\ x_N \end{bmatrix}$$

$$\underline{\mu} = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \cdot \\ \cdot \\ \mu_N \end{bmatrix}$$

So, they are in a stack $x_1 - \mu_1, x_2 - \mu_2, x_N - \mu_N$ that is what this is. We have seen that this vector if I take expected value of this is a column vector and its transpose, we have seen in the last class that is the covariance matrix.

$$C_{XX} = E[(\underline{x} - \underline{\mu})(\underline{x} - \underline{\mu})^t]$$

See maybe you can say x, x because $1 \times$ minus μ comes here, another x minus μ row vector comes here, you multiply 1 element here 1 element here. So, this has x , this has x . So, 2 elements of x get multiplied, maybe $x_1 - \mu_1, x_3 - \mu_2$.

So, $x_1 x_3$ terms come that is why we indicate it by $2x$. This is second order thing; 2 random variables are multiplied. It is called the covariance matrix for this random vector. It is a random vector; everybody is a random variable and they are jointly random. So, this we discussed last time μ is the mean alright, then this random variable x_1, \dots, x_N : jointly Gaussian, if they are probably joint probability density that is p capital X_1 capital X_2 dot dot dot capital X_n and the values $x_1 x_2$ dot dot dot dot x_n that takes this form.

$$P_{X_1, X_2, \dots, X_N}(x_1, x_2, \dots, x_N)$$

First $1 \times 2 \pi n \times 2$ and then determinant of this matrix C_{XX} its square root, positive square root into e to the power into. I do not have space here. So, I am writing here. e to the power minus half this vector this difference vector x minus μ . So, you are choosing $x_1 x_2 \dots x_n$.

So, you are putting them in a vector subtracting the mean vector from them that is \underline{x} minus, you are putting those values. So, for those x_1, x_2, \dots, x_N . So, those values of x_1, x_2, \dots, x_N you have to put here in the \underline{x} vector, corresponding μ values $\mu_1, \mu_2, \dots, \mu_N$ in this vector the column vector. So, in the formula they come as row, column transpose is row, then C_{XX} and I am assuming C_{XX} matrix is invertible that is inverse exists. Not all matrices are invertible, but we are assuming assumed when we can assume that that we will discuss later assumed to be invertible.

$$P_{X_1, X_2, \dots, X_N}(x_1, x_2, \dots, x_N) = \frac{1}{(2\pi)^{\frac{N}{2}} \sqrt{\det(C_{XX})}} e^{-\frac{1}{2}(\underline{x} - \underline{\mu})^t C_{XX}^{-1} (\underline{x} - \underline{\mu})}$$

That is C_{XX} inverse exists. So, again C_{XX} was a matrix is inverse is a matrix, matrix then again, the same vector \underline{x} minus $\underline{\mu}$. Now, this is a column vector, this is a matrix, a matrix times a column vector is a column vector. So, this is a column vector, this is a row vector. So, row into column is a scalar.

So, e to the power minus half into some scalar. So, it is a scalar, 2π to the power $N/2$ scalars, determined is a scalar. Because probability cannot be a vector, probability is a scalar. So, whole thing is a scalar. So, in the case of jointly Gaussian random variables, this should be the joint probability density.

If capital N is 1 there is only one random variable, then you will see this boils down to or well-known Gaussian density function for one variable. That is if you take N to be 1 only then obviously, 2π to the power $N/2$, N is 1. So, you have got I am writing here N equal to 1, you see $1/2\pi$ to the power $1/2$. So, $1/\sqrt{2\pi}$ then C_{XX} , $C_{XX} \underline{x}$ has only one element x_1 and $\underline{\mu}$ is μ_1 whole vector. So, C_{XX} means what it was a column vector

earlier into row vector, but now column has only one element $x_1 - \mu_1$, corresponding row also has only one element $x_1 - \mu_1$.

So, it is nothing but $x_1 - \mu_1$ whole square, expected value of that and that is called the variance. σ_x^2 square σ_x^2 square and now square root. So, it will be just σ_x , x_1 is a variable, σ_x . Then e to the power minus half again $x - \mu$ transpose has only one value $x_1 - \mu_1$. Matrix is a scalar now that was as we told the variance C_{xx} is a variance σ_x^2 .

So, inverse of that is $1/\sigma_x^2$. So, you have got minus half, So, 2 I bring here, σ_x^2 I bring here, the scalar, and $x_1 - \mu_1$ again $x_1 - \mu_1$ only one element one element.

$$= \frac{1}{\sqrt{2\pi} C_{xx}} e^{\frac{-(x_1 - \mu_1)^2}{2\sigma_x^2}}$$

So, this is what you have for the normal ordinary Gaussian density. we all know general form of the Gaussian density function for one variable, but this is more general. If this happens then they are of course, they are real and continuous in if they are complex then the formula will be more general, I mean more modified that we will not consider in this course.

If you are interested you can consult books. Now, we will see a very interesting thing. Earlier I had said that if a set of random variables they are statistically independent. That is if the joint density is product of individual marginal densities, then they are also uncorrelated. Their covariance is I mean, between any two-element covariance is 0, between any two elements.

Covariance between any two element 0 means what will be this matrix it will be a diagonal matrix because if you take i th element, in the previous class we have seen if you take i th

element and jth element, ith comma jth element that will be x_i minus μ_i times x_j minus μ_j , product and their expectation which is a correlation covariance between x_i and x_j . If i and j are not same then if they are uncorrelated, this correlation is 0 which means if i and j are not same that is I am not considering diagonal entries because for diagonal entries same i same j , 1 1, 2 2, 3 3, 4 4, but if that does not happen then other elements are 0. So, for a if the variables are uncorrelated then the covariance matrix is a diagonal matrix.

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Lecture 5

Jointly Gaussian (Normal) random variables

Multivariate

We are given N real, continuous random variables x_1, \dots, x_N .

$\underline{x} = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_N \end{bmatrix}$; $\underline{\mu} = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_N \end{bmatrix}$

$\underline{C}_{xx} = E[(\underline{x} - \underline{\mu})(\underline{x} - \underline{\mu})^T]$

Then, x_1, \dots, x_N : jointly Gaussian, if $p_{x_1, x_2, \dots, x_N}(x_1, x_2, \dots, x_N) = \frac{1}{(2\pi)^N \sqrt{\det(\underline{C}_{xx})}} \cdot e^{-\frac{1}{2}(\underline{x} - \underline{\mu})^T \underline{C}_{xx}^{-1}(\underline{x} - \underline{\mu})}$

\underline{C}_{xx} : Assumed to be invertible

Ok, that we have seen, let me again elaborate this. We have seen this I did in the last class though. C_{XX} was alright.

$$C_{XX} = E[(\underline{x} - \underline{\mu})(\underline{x} - \underline{\mu})^T]$$

So, this is a vector the column vector, this is a row vector. So, multiply you get a matrix. i comma j th element of this matrix C_{XX} i comma j , this is a symmetric matrix that we have seen at that time only I showed, i comma j th element will be what? i th i th guy i th element of this vector and j th element of this row vector their product and expectation. So, E with respect to $x_i x_j$ only they matter here in that product.

$$[C_{XX}]_{ij} = E_{X_i X_j} [(X_i - \mu_i)(X_j - \mu_j)]$$

So, x_i minus μ_i and x_j minus μ_j and i and j they are not same, they are two different random variables.

$$= r_{X_i X_j}$$

So, you are taking increment of this around the mean, the increment of this around its mean and multiplying and expecting. So, that is correlation $r_{X_i X_j}$. So, if x_i x_j are uncorrelated this will be 0 that means, as long as then i and j i comma j th element and i and j are not same because they are different random variables then only, I can talk of their correlation.

$$[C_{XX}]_{ij} = 0; i \neq j$$

If i and j are not same and they are uncorrelated then that element is 0 that is i comma j th element is 0 if i and j are not same. If i and j are same it will be x_i minus μ_i , again x_i minus μ_i because i and j are same.

So, x_i minus μ_i whole square expected value which is variance. So, i comma i th element which is i th diagonal element will be $\sigma_{x_i}^2$ variance of the i th element x_i . So, diagonal element will be the variances, but on off diagonal elements will be 0 when the elements are uncorrelated. Now, I had told earlier that if a set of elements are statistically independent, we can show they are uncorrelated that is a correlation between every two elements there, covariance between every two elements there will be 0. But the converse is not in general true that is if a number of variables they are given to be uncorrelated that is covariance between any two is 0.

It does not mean in general that they are statistically independent that is overall joint density will be a product of individual marginal densities, except in the case of jointly Gaussian random variables then one means the other and vice versa both are equivalent

that is what we will see here. Now, suppose that is suppose we are given the fact that x_1 to x_n are uncorrelated, meaning C_{XX} i comma j 0, if as I told here if i not equal to j , and C_{XX} i th diagonal into i comma i would be the variance, $\sigma_{x_i}^2$, variance of x_i fine.

$$[C_{XX}]_{i,i} = \sigma_{x_i}^2$$

So, suppose this is given and of course, they are jointly Gaussian then what happens determinant, number 1, determinant of C_{XX} , now C_{XX} is a diagonal matrix, all other elements are 0 only, diagonal elements are $\sigma_{x_1}^2, \sigma_{x_2}^2, \sigma_{x_2}^2$ dot dot dot their product. So, determinant is $\sigma_{x_1}^2, \sigma_{x_2}^2$ all right.

$$\det[C_{XX}] = \sigma_{x_1}^2, \sigma_{x_2}^2, \dots \sigma_{x_N}^2$$

So, in that Gaussian formula $p(x_1, x_2, \dots, x_n)$ we remember x_1, x_2, \dots, x_n , we had 1 by $\sqrt{2\pi}$ whole to the power N by 2 .

ok, and then in the denominator had positive square root of the determinant. So, if you take positive square root, it will become $\sigma_{x_1}, \sigma_{x_2}, \dots, \sigma_{x_N}$ like that. So, what I do the whole probability density function I write as a product of some terms, one term, similar terms, one is just $\sqrt{2\pi}$, but I had 2π to the power N by 2 . So, I should have 1 by $\sqrt{2\pi}$, again 1 by $\sqrt{2\pi}$, again 1 by $\sqrt{2\pi}$ dot dot dot capital n times. So, I will have maybe 1 by $\sqrt{2\pi}$ here dot dot dot lastly again another 1 by $\sqrt{2\pi}$ capital N number.

Then $\sigma_{x_1}, \sigma_{x_2}, \dots, \sigma_{x_N}$ there is I am taking positive square root this product. So, I take σ_{x_1} here, I take σ_{x_2} here, I take σ_{x_N} here. So, you multiply the denominators you get back what we had earlier then e to the power is something that we have to see. what we had, e to the power, this is where I do the calculation, we had e to the power minus half C_{XX} inverse x minus μ . Now, C_{XX} inverse means what, C_{XX} is a diagonal matrix, all diagonal entries are positive because they are variances.

So, inverse means it will be again diagonal matrix, but it will be a σ_{x_1} by σ_{x_1}

square dot dot dot 1 by sigma x n square and 0 on this side 0 on this side this is cxx inverse this matrix. And then you have x1 minus mu 1 top guy then x is x2 minus mu 2 dot dot xn minus mu n and of course, this term. So, when you multiply this with this, the diagonal matrix, So, just this multiplied by this, this multiplied by the next one, this multiplied by this. So, you get a row vector same, but they are scaled this is multiplied by this this is multiplied by like that with that if I multiply x1 minus mu 1 dot dot dot xn minus mu n. So, this after this product if I do the multiplication, you will have x1 minus mu 1 here and here whole square divided by this plus really plus because a row vector and this matrix times this a column vector.

So, row into column means term by term multiply and add plus. So, x1 minus mu 1 into x1 minus mu 1 by sigma x square x1 square plus x2 minus mu 2 here also. So, x2 minus mu 2 whole square by sigma x2 square plus dot dot dot dot. So, e to the power some terms which are summed that mean, I can separate them out while e to the power minus half is common one term will be just x1 minus mu 1 whole square by this another will be dot dot dot another will be you can just multiply this exponential.

$$P_{X_1, X_2, \dots, X_N}(X_1, X_2, \dots, X_N) \\ = \frac{1}{\sqrt{2\pi}\sigma_{x_1}} e^{\frac{-1}{2} \frac{(x_1 - \mu_1)^2}{\sigma_{x_1}^2}} \frac{1}{\sqrt{2\pi}\sigma_{x_2}} e^{\frac{-1}{2} \frac{(x_2 - \mu_2)^2}{\sigma_{x_2}^2}} \dots \frac{1}{\sqrt{2\pi}\sigma_{x_N}} e^{\frac{-1}{2} \frac{(x_N - \mu_N)^2}{\sigma_{x_N}^2}}$$

So, they will add up the exponent's powers will get add up I will get that, but this is what this is the Gaussian density of the single random variable x1.

This is the Gaussian probability density of a single random variable x2 when they are jointly Gaussian, they are individually Gaussian also ok. There is a meaning I forgot to tell you if they are jointly Gaussian, they are individually Gaussian also.

$$= P_{X_1}(x_1)P_{X_2}(x_2) \dots P_{X_N}(x_N)$$

So, that means, this is nothing, but the probability density of this guy capital X1, this is nothing, but probability density of this guy dot dot dot is nothing, but probability density

of x_n . So, you see now in this case overall joint density is a product of the individual marginal densities.

So, they are statistically independent. So, in the Gaussian case if you are giving the variables to be uncorrelated then they are statistically independent also. converse is always true, but this is not always true, this is happening in the case of Gaussian. I mean in it does not happen in the case of other well-known probability densities, but here all right.

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The image shows a handwritten derivation on a blackboard. The text and equations are as follows:

$$C_{XX} = E[(x - \mu)(x - \mu)^T]$$

$$[C_{XX}]_{ij} = E[(x_i - \mu_i)(x_j - \mu_j)] = \sigma_{x_i x_j}$$

Suppose, we are given the fact that x_1, \dots, x_N are uncorrelated

$$\Rightarrow [C_{XX}]_{ij} = 0, \text{ if } i \neq j \text{ and } [C_{XX}]_{ii} = \sigma_{x_i}^2$$

$$\det[C_{XX}] = \sigma_{x_1}^2 \sigma_{x_2}^2 \dots \sigma_{x_N}^2$$

$$p_{x_1, x_2, \dots, x_N}(x_1, x_2, \dots, x_N) = \left(\frac{1}{\sqrt{2\pi} \sigma_{x_1}} e^{-\frac{1}{2} \frac{(x_1 - \mu_1)^2}{\sigma_{x_1}^2}} \right) \left(\frac{1}{\sqrt{2\pi} \sigma_{x_2}} e^{-\frac{1}{2} \frac{(x_2 - \mu_2)^2}{\sigma_{x_2}^2}} \right) \dots \left(\frac{1}{\sqrt{2\pi} \sigma_{x_N}} e^{-\frac{1}{2} \frac{(x_N - \mu_N)^2}{\sigma_{x_N}^2}} \right)$$

$$= p_{x_1}(x_1) p_{x_2}(x_2) \dots p_{x_N}(x_N)$$

On the right side of the blackboard, there is a matrix representation of the joint density function for correlated Gaussian variables:

$$e^{-\frac{1}{2} (x - \mu)^T C_{XX}^{-1} (x - \mu)}$$

$$\begin{bmatrix} (x_1 - \mu_1) & \dots & (x_N - \mu_N) \end{bmatrix} \begin{bmatrix} \frac{1}{\sigma_{x_1}^2} & 0 & \dots & 0 \\ 0 & \frac{1}{\sigma_{x_2}^2} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \frac{1}{\sigma_{x_N}^2} \end{bmatrix} \begin{bmatrix} (x_1 - \mu_1) \\ (x_2 - \mu_2) \\ \vdots \\ (x_N - \mu_N) \end{bmatrix}$$

This is one important thing; another thing is complex random variables. All right complex random variables, here, I will say Z is a complex random variable if there are two real variables, two real continuous random variables capital X and Y .

$$Z = X + jY$$

So, I conducted experiment, I observed two real valued continuous valued random variables, say capital X and Y and I take their values and write in this form a real variable I call, a complex variable Z , then obviously, Z is also random, but it is complex, then Z

will be called a complex random variable, it has got two real random variables. So, therefore, even though Z is a single random variable it has actually two. So, I should have one X is capital X capital Y . So, Z will be, maybe if you take this to be small x and this to be small y . So, this point is Z small z , as small x plus small $j y$, this point.

$$Z = X + jY$$

And then if I go by dx and this is by dy then this is, you understand one thing dz will be dx plus $j dy$ right. So, dx plus $j dy$. So, this point is nothing, but small z plus dz because x and dz means dx . So, x plus dx and take j common j here and in dz and j here. So, j into y plus dy here, y plus dy is this much and x plus dx is this much.

So, x plus dx plus j times y plus dy this is dz . So, dz is dx plus $j dy$, dx plus $j dy$ and originally z at the two phasors. So, they basically x plus $j y$ and it is dx plus $j dz$. So, x plus dx , z plus dz and j common y plus dy to this point.

$$dz = dx + jdy$$

That means, probability of capital Z lying between small z and z plus dz is equivalent to what will happen that is capital Z will lie between these two this means it will lie within this box.

It will fall anywhere in this box there is a meaning of this. That means, the same as probability of capital X lying between here to here simultaneously with capital Y lying between these which we know is nothing, but p_{xy} joint density small x small y $dx dy$.

$$\equiv \text{prob} (x \leq X \leq x + dx, y \leq Y \leq y + dy) \equiv P_{XY}(x, y) dx dy$$

So, when expressed in terms of $x y$ the real and individual component real and imaginary components then their joint density that will be same as the probability of capital Z falling here. Event is same capital Z lying here means capital X lying between here-to-here capital Y lying between here to here. So, the two things are two events are same capital Z falling in the box is same as capital X lying from here to here and capital Y lying from here to here.

So, this probability and this probability they are same all right. Now, given a function which could be complex or real $f(Z)$ then it takes z , it can give you a real number or it can give you a complex number, does not matter $f(Z)$, every time you find a z there is some x and y , you put that value x plus $j y$ here and get a value. Since that is random this function also takes value randomly. So, what is the expectation $E_z f(Z)$, it is same as if you write working in terms of z , if you write it as x plus $j y$ then what is the meaning of this E_z ? there is $f(Z)$ taking a value from with the I mean when capital Z lies falls anywhere in the box since the widths are infinitely small we assume the value of $f(Z)$ does not change it will have the same value as here at this small z or at this value at this pair ok, but chance of that I mean it will not always happen.

$$\begin{aligned} E_Z[f(Z)] &= \iint f(X + jY) P_{XY}(x, y) dx dy \\ &= E_{XY}[f(X + jY)] \end{aligned}$$

So, I have to multiply it by the chance. Chance means you have to multiply this f of z value that is f of this value by the probability of capital Z occurring here or equivalently this capital X occurring capital $X Y$ occurring in this box. So, that is I have to when I express z in terms of $x y$, I have to multiply this by the joint density. And then average ok, multiply everywhere and take average then as we have done many times in the past weighted average. So, this to be divided by summation of the weights. weights are this part, but their average their integral is 1, certainty probability of certainty is 1.

So, you get this which is nothing, but E_{xy} of this function f now you write in terms of $x y$ x plus $j y$ all right. Then one thing E_z suppose this is given and I suddenly and f of z is a complex function. So, it takes z as argument and gives you a complex number f of z . Now instead of f of z I take the conjugate that is whatever value f of z takes I take a complex conjugate of that. So, then it becomes a function $f^* z$ $f^* z$ what will happen to that.

$$E_Z[f^*(Z)] = \iint f^*(X + jY)P_{XY}(x, y)dx dy$$

So, again just same thing I have to do, but instead of f it will be f star small x plus these are small within the integral that the variables of integral. So, now you see probability density is real is never complex, d x d y real and this is star. So, I can as well take the star out I mean this product I can take put the star outside, the star of a product means star will be put on everybody, but they are real. So, no change. So, this product and star and then integral is a summation.

$$\begin{aligned} &= [\iint f(x + jy)P_{XY}(x, y)dx dy]^* \\ &= [E_Z[f(Z)]]^* \end{aligned}$$

So, summation of complex quantities with the conjugate you will get the same thing if you first sum and then conjugate that is z 1 star plus z 2 star that is first conjugate z 1 star z 2 star then add that is integral you will have the same thing if you do first addition z 1 plus z 2 then star that is z 1 star plus z 2 star we all know is z 1 plus z 2 star first add. So, first you integrate and then star that is first you integrate and this product but this is what I have here. So, that means, this will be nothing, but E z of the original function f z and then star. So, if you are taking f star z and then expectation value with respect to z what you have to do you have to first take the expectation value of the original f z no star and then put star on that and implication of that is suppose f z is z. So, that means, E z f z star is z star and z is suppose mu plus sorry 1 minute z is suppose x plus j y, E z z star will be what first you have to take E z of z and then star E z of z means expected value of this plus j times expected value of this.

$$f(Z) = Z = X + jY$$

$$E_Z[Z^*] = [E_Z[Z]]^*$$

$$= [\mu_X + j\mu_Y]^*$$

$$= \mu_x - j \mu_y$$

So, that was μ_x it was μ_y expected value. So, μ_x plus this is μ_x plus $j \mu_y$ star which is same as μ_x minus $j \mu_y$. So, this is very important if you have z and you are taking expectation of z star which is suppose z it will be original expected value and then conjugate.

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Complex random variables

$Z = X + jy$ μ_x μ_y

$dz = dx + jdy$

Prob of Z lying between z and $z+dz$
 \equiv Prob of $(x \leq X \leq x+dx, y \leq Y \leq y+dy) \equiv p_{xy}(x, y) dxdy$

Given a function $f(z)$

$$E_z[f(z)] = \iint f(x+jy) p_{xy}(x, y) dxdy = E_{xy}[f(x+jy)]$$

$$E_z[f^*(z)] = \iint f^*(x+jy) p_{xy}(x, y) dxdy$$

$$= \left[\iint f(x+jy) p_{xy}(x, y) dxdy \right]^* = [E_z[f(z)]]^*$$

$f(z) = Z = x + jy$
 $E_z[Z^*] = [E_z[Z]]^*$
 $= [\mu_x + j\mu_y]^*$
 $= \mu_x - j\mu_y$

So, μ_x minus $j \mu_y$. So, I stop here and we will build up on these complex random variables and in particular we will consider complex covariance and complex covariance matrices in the next class. Thank you very much.