

Statistical Signal Processing
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Lecture-4
Random Processing

Hello students, so far we have discussed the basic concepts of probability and random variables. In the last lecture, we discussed the linear algebra of random variables today will give an outline of random processes in fact in statistical signal processing we model the signal as a random process. And therefore, a statistical signal processing is the processing of random processes will start with the definition of random process. Random process that is RP will abbreviate it as RP.

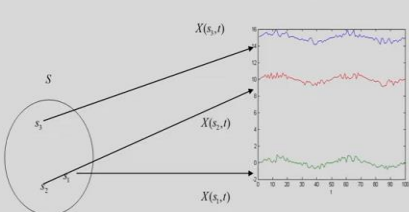
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Random Process (RP)

- ❖ A random process maps each sample point to a waveform

More formally, a *random process* on the sample space S can be defined as an indexed family of RVs $\{X(s, t) \mid s \in S, t \in \Gamma\}$. Γ is an index set usually denoting time.

- ❖ For a fixed $s_0 \in S$, $X(s_0, t)$ is a *single realization of the random process and is a deterministic function*.
- ❖ The random process $\{X(s, t)\}$ is normally denoted by $\{X(t)\}$.



A random process maps each sample point to a waveform unlike a random variable which maps a sample point to a point on a real line. Here it will be mapped to a waveform like this suppose the sample point is mapped to this waveform this is the time axis. So there waveform varying with time and one sample point will be mapped to one of the waveform more formally in a random process on the sample space as can be defined as an index family of random variables.

So a random process comprises of $X(s, t)$ such that s belong to the sample space and t belongs to some index set Γ . So Γ is an index set and Γ usually denotes time. So essentially it is a functions of both s and t , so it varies with s as well as it varies with t if we fix the sample point

for peak sample point s_0 , X_{s_0} realization of the random process and it is a deterministic function because we are considering only one s_0 and our practical realizations whatever signal we observed that is modeled as a single realization of a random process.

This is an important point any observed signal is considered as the realization of a random process like in the case of random variable we omit s therefore X_s is normally denoted by X_t here is the illustration of random process this is the sample space where showing three sample points and each sample point is mapped to a sample function like this wave form like this.

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Discrete-time random process

- ❖ A discrete-time random process $\{X(n)\}$ is defined at discrete points of time. Here $\Gamma \subseteq \mathbb{Z}$.
- ❖ Such a discrete-time random process is more important from the signal processing point of view
- ❖ For example, a sampled speech waveform is modelled as a discrete-time RP

Our concern is discrete time random process. A discrete time random process X_n is defined at discrete points of time. So, here the index set τ is a subset of $\mathbb{Z}$, $\mathbb{Z}$ is the set of integer. So on in integer points the random variables are defined such a discrete time random processes is more important for signal processing point of view. Therefore we will be discussing only discrete time random process as an example a sampled speech waveform is modeled as a discrete time random process. So, sample speech waveform that we can represent as a random process.

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Probability structure of a random process

- Consider a discrete-time random process $\{X(n)\}, n \in \Gamma$
- To describe $\{X(n)\}$ we have to use joint CDF of the RVs at different instants n .
- For any positive integer k , $X(n_1), X(n_2), \dots, X(n_k)$ represent k joint RVs.

Thus a random process can be described by the joint CDF

$$F_{X(n_1), X(n_2), \dots, X(n_k)}(x_1, x_2, \dots, x_k), \forall k \in \mathbb{N} \text{ and } \forall n_k \in \Gamma$$

Probability structure of a random process, we know that random variable is described by CDF similarly random vector it is described by joint CDF and similarly we define also probability density function probability mass functions etcetera. How to have such characterization for a random process to describe X_n we have to use joint CDF because it is an index family of random variables.

So, we have to consider joint CDF only and these joint CDF we have to considered at different instants of n . Suppose you consider any positive integer k then X_{n_1}, X_{n_2} up to X_{n_k} represent k joint random variables because we are considering the random variable at instant n_1, n_2 and so on up to n_k . So, this is there are k joint random variables. Therefore they can be characterized by joint CDF.

This is the joined CDF of the random variable X_{n_1}, X_{n_2} up to X_{n_k} at point x_1, x_2 up to x_k . Now this CDF must be defined for all k belonging to $\mathbb{N}$ because this index set is an infinity. So all possible values of k we have to consider, similarly this n_1 and n_2 etcetera. We can place also anywhere on the time axis. So this for all n_k belonging to Γ we have to considered. Therefore all these considerations make the joint CDF characterization of a random process difficult.

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Moments of a Random Process

We can define various moments and central moments.

❖ Mean of the random process

$$E(X(n)) = \mu_X(n), \forall n$$

❖ Autocorrelation function $R_X(n_1, n_2) = E(X(n_1)X(n_2)), \forall n_1, n_2$

❖ Auto-covariance function

$$C_X(n_1, n_2) = E(X(n_1) - \mu_X(n_1))(X(n_2) - \mu_X(n_2)) = R_X(n_1, n_2) - \mu_X(n_1)\mu_X(n_2)$$

and so on

So, we have to have some simpler description of a random process. We can define various moments and central moments of a random process. For example you can define mean of the random process E of X_n . Now this we have to consider at all possible instant of time. So E of X_n is equal to it will be again a function of n for all instant of time n . Similarly other moments also we can consider for example autocorrelation function this is denoted by $R_{X_n 1 n 2}$. What is this E of $X_n 1$ into $X_n 2$.

This is the joint expectation of $X_n 1$ and $X_n 2$. This parameter $R_{X_n 1 n 2}$ also we have to consider for all possible values of $n_1 n_2$ and autocorrelation function it gives the similarity of 2 random variables at different instant of time $X_n 1$ into $X_n 2$ and then they were taking the average below, we can also characterize a random process by auto covariance function C_X of $n_1 n_2 = E$ of $X_n 1 - \mu_{X_n 1}$ into $X_n 2 - \mu_{X_n 2}$ so you subtract the means from the corresponding random variable and then you find joint expectation.

And this we can show that this quantity is equal to R_X of $n_1 n_2 - \mu_{X_n 1}$ into $\mu_{X_n 2}$ and similarly we can define the different moments. For example higher order moments for higher moments etcetera.

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Example Gaussian Random Process

The process $\{X(n)\}$ is called Gaussian if for any $k \in \mathbb{N}$ and any time points $n_1, n_2, \dots, n_k$ the random vector $\mathbf{X} = [X(n_1), X(n_2), \dots, X(n_k)]'$ is jointly Gaussian with the joint PDF

$$f_{X(n_1), X(n_2), \dots, X(n_k)}(x_1, x_2, \dots, x_k) = \frac{e^{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_X)' \mathbf{C}_X^{-1} (\mathbf{x} - \boldsymbol{\mu}_X)}}{(\sqrt{2\pi})^n \sqrt{\det(\mathbf{C}_X)}}$$

where $\mathbf{C}_X = E(\mathbf{X} - \boldsymbol{\mu}_X)(\mathbf{X} - \boldsymbol{\mu}_X)'$

and $\boldsymbol{\mu}_X = E(\mathbf{X}) = [E(X_{n_1}), E(X_{n_2}), \dots, E(X_{n_k})]'$

Will give an example Gaussians random process the process X_n is called Gaussian if for any k belonging to $\mathbb{N}$ and any time points $n_1, n_2, \dots, n_k$ the random vector given by this column vector with elements $X_{n_1}, X_{n_2}, \dots, X_{n_k}$ is jointly Gaussian with joint PDF this is the PDF we have already discussed about this PDF. So PDF is the Gaussian PDF and given by e to the power $-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_X)' \mathbf{C}_X^{-1} (\mathbf{x} - \boldsymbol{\mu}_X)$.

And divided by this normalization factor $(\sqrt{2\pi})^n \sqrt{\det(\mathbf{C}_X)}$ as we know $\mathbf{C}_X$ is the covariance matrix that is given by $E(\mathbf{X} - \boldsymbol{\mu}_X)(\mathbf{X} - \boldsymbol{\mu}_X)'$. Similarly $\boldsymbol{\mu}_X$ vector it is the mean vector and it is the vector comprising of mean of its component random variables. So that this jointly Gaussian random processes important for us and it has comparatively their description because we have a for any $\mathbb{N}$ or any k we can have this joint PDF.

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Stationary Random Process

❖ An RP $\{X(n)\}$ is called **strict-sense stationary (SSS)** if its probability structure is invariant with time. In terms of the joint CDF

$$\begin{aligned} F_{X(n_1), X(n_2), \dots, X(n_k)}(x_1, x_2, \dots, x_k) \\ = F_{X(n_1+h), X(n_2+h), \dots, X(n_k+h)}(x_1, x_2, \dots, x_k), \\ \forall k \in \mathbb{N} \text{ and } \forall h, n_1, \dots, n_k \in \Gamma \end{aligned}$$

❖ Analysing an SSS random process is highly complex. We look for a weak form of stationarity

An RP $\{X(n)\}$ is called **wide sense stationary process (WSS)** if $\forall h, n, n_1, n_2$

1. $EX(n) = EX(n+h) = \text{constant}$ and
2. $R_X(n_1, n_2) = R_X(n_1+h, n_2+h)$

If we put $h = -n_1$, then

$$R_X(n_1, n_2) = R_X(0, n_2 - n_1) \quad \forall n_1, n_2 \text{ is a function of lag } n_2 - n_1 \text{ only.}$$

Because the probabilities structure of the random process is very complex. We have to have some simple assumptions one such assumptions is stationary random process. A random process X_n is called strict sense stationary that is abbreviated SSS. If its probability structure is invariant with time, so whatever probability description is at this point of time suppose the same probability description will hold for another instant of time may be after 1 month in terms of joint CDF.

We can write that is joint CDF of random variable X_{n_1}, X_{n_2} up to X_{n_k} at point x_1, x_2 up to x_k . So joint CDF at this points is same as joint CDF random variable now we will shift the time axis by an amount into the amount is so X of n_1 is X of n_2 is up to X of $n_k + h$. If we consider this set of random variable instead of this. Then also this joint CDF will remain same. So that means is random variable we are shifting by the same amount X_{n_1} has become X of $n_1 + h$ and is X_{n_2} has become $X_{n_2 + h}$.

And under such conditions the CDF at a set of points will remain invariant and this is the strict sense stationarity now to analyses strict sense random process is highly complex will look for a weak form of stationarity. So that is the week sense stationary random process or wide sense stationary random process. A random process X_n is called a wide sense stationary process WSS that is if for all h, n_1 and n_2 number 1 E of $X_n = E$ of X_{n+h} . So to instant of time you have the average value same.

That means average value must be constant, R_X of n_1, n_2 is same as R_X of $n_1 + h, n_2 + h$ this is for all h , now we can put $h = -n_1$ then R_X of n_1, n_2 will be equal to now this h will be, if we will put minus n_1 this 1 will become 0. And here it will become $n_2 - n_1$. So R_X of n_1, n_2 will become R_X of $0, n_2 - n_1$. So if we want this condition then this is equivalent to the statement that R_X of n_1, n_2 is a function of lag $n_2 - n_1$ only.

It does not depend absolutely on n_1 and n_2 it depends only on the time difference $n_2 - n_1$ so such a process therefore for a wide sense stationary process the mean E of X_n is constant and the autocorrelation function R_X of n_1, n_2 is a function of $n_2 - n_1$ only.

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Example of a WSS process Sinusoid with random phase

Consider the random process $X(n) = A \cos(\omega_0 n + \phi)$, A and ω_0 are constants and $\phi \sim U(0, 2\pi)$

Note $f_\phi(\phi) = \begin{cases} \frac{1}{2\pi} & 0 \leq \phi \leq 2\pi \\ 0 & \text{otherwise} \end{cases}$

$\therefore EX(n) = EA \cos(\omega_0 n + \phi)$

$$= \int_0^{2\pi} A \cos(\omega_0 n + \phi) \frac{1}{2\pi} d\phi = 0$$

$R_X(n_1, n_2) = EX(n_1)X(n_2)$

$$= EA \cos(\omega_0 n_1 + \phi) A \cos(\omega_0 n_2 + \phi)$$

$$= \frac{A^2}{2} E[\cos(\omega_0(n_1 + n_2) + 2\phi) + \cos(\omega_0(n_1 - n_2))]$$

$$= \frac{A^2}{2} \cos(\omega_0(n_1 - n_2))$$

Hence $\{X(n)\}$ is WSS.

Handwritten note: $2 \cos A \cos B = \cos(A+B) + \cos(A-B)$

Will give an example of a WSS process that is a sinusoid with random phase consider this is the sinusoid $X_n = A \cos \omega_0 n + \phi$. In ω_0 are deterministic constant suppose A the amplitude and ω_0 is the frequency these are known only ϕ is unknown and ϕ . If we model ϕ as uniformly distributed random variable. This is the symbol for uniformly distributed random variable and it is uniformly distributed between 0 to 2π .

So, this phase part is a random quantity which is uniformly distributed between 0 to 2π so part we can find out the PDF of this random variable because this is the random part. So PDF will be therefore since it is uniform it is $1 / 2\pi$ for ϕ lying between 0 and 2π and 0 elsewhere. Now how to find out E of X_n because I did a function of this random variable ϕ therefore we can we find a expectation of this function of random variables that is range of random variable is to 2π .

Therefore this is integration 0 to 2π of $A \cos(\omega_0 n + \phi)$ into $1/2\pi d\phi$, and this is a cosine function in a period therefore integration will be equal to 0. So what we see that mean is constant. Now let us find out R_X of $n_1 - n_2$ this is equal to E of $X_{n_1} X_{n_2}$. That is by definition and that is equal to E of $A \cos(\omega_0 n_1 + \phi)$ into $A \cos(\omega_0 n_2 + \phi)$. Now there is a product of 2 cosine terms we can use the relationship that is $\cos$ of twice $\cos A \cos B = \cos$ of $A + B + \cos$ of $A - B$.

So if we use this identity then we will have this is equal to A^2 by 2 into expectation of $\cos$ of $\omega_0 n_1 + n_2 + 2\pi$ this and these are added plus $\cos$ of these and these are subtracted $\cos$ of $\omega_0 n_1 - n_2$, so this part now this is the cosine function if we integrate over 0 to 2π you will get 0. So therefore expectation of this part is 0 and this part does not involve π . So there it is a constant.

Therefore what we will get $A^2/2$ into $\cos$ of $\omega_0 n_1 - n_2$, so what we have seen that this autocorrelation function is a function of the time lag $n_1 - n_2$ and its mean is constant therefore X_n is a WSS process.

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Important Properties of $R_X(m)$

$$\diamond R_X(0) = EX^2(n) = \text{Mean-square value}$$

(Average power)

$$\diamond R_X(-m) = R_X(m)$$

$$\diamond |R_X(m)| \leq R_X(0)$$

$$R_X(m) = E[X(n)X(n+m)] = E[X(n+m)X(n)] = R_X(-m)$$

This follows from the Cauchy Schwartz inequality

$$|<X(n), X(n+m)>| \leq \|X(n)\| \|X(n+m)\|$$

$$\therefore |EX(n)X(n+m)| \leq \sqrt{EX^2(n)} \sqrt{EX^2(n+m)}$$

$$\Rightarrow |R_X(m)| \leq \sqrt{R_X(0)} \sqrt{R_X(0)} = R_X(0)$$

$R_X(m)$ is maximum at $m=0$

Now a WSS process is characterized by the autocorrelation function R_X of m . This is a very important quantity for WSS random process because it is characterized in terms of R_X of m only. So what is R_X of 0 if we put R_X of 0 by definition it will be E of X_n into X_n that is E of

X^2 and this is the mean square value that means this is the square term you are taking the average therefore it is this is the mean square value and this is also interpreted as average power.

Because this average power $E\{X^2}$ is the power delivered on 1 ohm resistance unit resistance, so that way $E\{X^2}$ denote average power. So to interpretation this is A^2 square mean value and this is also the average power of the random process. And $R_X(m)$ is an even function because what we can write is that $E\{X^n X^{n+m}$ this is same as if I interchanged the order $E\{X^{n+m} X^n$.

Now this quantity by definition this is my $R_X(m)$. And by definition this quantity will be this minus this is the argument therefore this will be $R_X(-m)$. So what do we get is that autocorrelation function is an even function of m . Then magnitude of $R_X(m)$ at any m is less than equal to $R_X(0)$. So this we can prove using the Cauchy Schwarz inequality that we know. So inner product of X^n and X^{n+m} if we take the inner product and then magnitude of that will be less than normal X^n into normal X^{n+m} .

So inner product is expectation magnitude of $E\{X^n X^{n+m}$ that will be less than equal to normal of X^n is square root of $E\{X^2}$ normal X^{n+m} is square root of $E\{X^{2+n+m}$ therefore this quantity is autocorrelation function magnitude of $R_X(m)$ is less than equal to square root of $R_X(0)$ this is equal to $R_X(0)$. And similarly this 1 is also $R_X(0)$ because this is the product of X^{n+m} into X^{n+m} .

So there is lag will be 0 so that way both square root together is $R_X(0)$. So what we have seen that magnitude of $R_X(m)$ is less than equal to $R_X(0)$ that means the autocorrelation function has a maximum at the origin at m and $m = 0$ there is a maximum. So $R_X(m)$ is maximum at $m = 0$. There are other properties also I will not discuss them.

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Cross-correlation Function

- ❖ The cross-correlation function two discrete-time random processes $X(n)$ and $Y(n)$ is defined as

$$R_{XY}(n, m) = E X(n) Y(n + m)$$

- ❖ $X(n)$ and $Y(n)$ are called jointly WSS if they are individually WSS and

$R_{XY}(n, m)$ is a function of time lag m . Thus we can write

$$R_{XY}(n, m) = R_{XY}(m)$$

- ❖ Symmetry property of $R_{XY}(m)$

$$R_{XY}(m) = E X(n) Y(n + m)$$

$$= E Y(n + m) X(n)$$

$$= R_{YX}(-m)$$

$$\therefore R_{XY}(m) = R_{YX}(-m)$$

We can consider also multiple random processes for example X_n and Y_n . So if X_n and Y_n are 2 random processes then their correlation structure will be given in terms of cross correlation function and this is by definition $R_{XY}(n, m)$ that is cross correlation argument n and m it is the expectation of E of X_n that is expectation of X n into Y $n + m$. So cross correlation R_{XY} at point n, m is defined as E of X_n into Y $n + m$, so that we can define cross correlation function. Here also I want to make some simplistic assumption about the correlation between X_n and Y_n .

And we define what is known as the jointly wide sense stationary random process. Therefore jointly WSS process X_n and Y_n are characterized by their individual autocorrelation functions and the cross correlation function $R_{XY}(n, m)$ that is equal to $R_{XY}(m)$. Now we know that autocorrelation function is even symmetric similar to that there is a symmetry property of $R_{XY}(m)$. $R_{XY}(m) = E$ of X n into Y $n + m$ this is by definition.

And this now because these are 2 real number I will exchange the order, So E of Y $n + m$ into X n now definition of cross correlation function we take the difference of second argument and first argument. So that way this will be now R_{YX} of $-m$. Therefore $R_{XY}(m) = R_{YX}(-m)$. This is this symmetry property of cross correlation function. So that way if these signals are jointly WSS then we can characterize them with the help of autocorrelation function and cross correlation function.

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Frequency-domain Analysis of a WSS signal

- ❖ Recall that a discrete-time deterministic signal $\{g(n)\}$ has the frequency-domain representation in terms of the discrete-time Fourier transform (DTFT)

$$G(\omega) = \text{DTFT}(g(n)) = \sum_{n=-\infty}^{\infty} g(n)e^{-j\omega n}$$

- ❖ $G(\omega)$ is periodic with a period 2π rad/sample and the inverse is given by

$$g(n) = \frac{1}{2\pi} \int_{-\pi}^{\pi} G(\omega)e^{j\omega n} d\omega$$

- ❖ $G(\omega)$ exists if $\sum_{n=-\infty}^{\infty} |g(n)| < \infty$

$$\sum_{n=-\infty}^{\infty} g^2(n) < \infty$$

- ❖ Such a frequency-domain representation is not possible for a random process.

Frequency domain analyses of a WSS signal. We know that in deterministic signal processing the frequency domain analysis plays an important role. We will see if similar representation is possible for a WSS signal. Recall that a discrete time deterministic signals g_n has the frequency domain representation in terms of discrete time Fourier transform DTFT. So for example if g_n is the signal then its DTFT is given by summation $g_n e^{-j\omega n}$ and n going from minus infinity to plus infinity.

So this is the definition of discrete time Fourier transform and it is denoted by $G(\omega)$ and it is easy to see that $G(\omega)$ is periodic with period 2π radian per sample. This unit of ω is radian per sample. So $G(\omega)$ is periodic. Therefore we can find the inverse by this relationship $g_n = \frac{1}{2\pi} \int_{-\pi}^{\pi} G(\omega) e^{j\omega n} d\omega$. So will integrate over $d\omega$ will get g_n . g_n and $G(\omega)$ form a Fourier transform pair. Now $G(\omega)$ the Fourier transforms exists if this sum exists.

That is sum from n is equal to minus infinity to infinity of absolute value of g_n is less than infinity. In other words g_n is absolutely summable in that case the $G(\omega)$ exist and we can also write a simpler relationship that is instead of $\sum |g_n|$ we can write $\sum g_n^2$. In that case we will have the condition summation g_n^2 n going from minus infinity to infinity that is less than infinity this expression is the energy of the signal.

So if energy of the signal is finite then its DTFT exists. So this is the condition for existence of Fourier transform discrete time Fourier transform of a discrete time signal. But such a frequency domain representation is not possible for a random process because this condition that the energy of the signal is finite is not valid for a random process. The definition of stationarity implies that the signal never decays and therefore each energy will be infinite.

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Power spectral density (PSD)

❖ Average power of a discrete-time RP $\{X(n)\}$

$$R_X[0] = E[X^2[n]]$$

❖ PSD $S_X(\omega)$ indicates the contribution of each component frequency ω to the average power.

❖ For a WSS RP $\{X(n)\}$, $S_X(\omega)$ exists and is given by

$$S_X(\omega) = \lim_{N \rightarrow \infty} \frac{1}{2N+1} E \left| \sum_{n=-N}^N X[n] e^{-j\omega n} \right|^2$$



But the power spectral density PSD of a WSS process exists. So let us see what is PSD first average power of a discrete time random process $X[n]$ is that $R_X[0] = E[X^2[n]]$. PSD $S_X(\omega)$ indicates the contribution of each component frequency ω to the average power this power spectral density is average power per frequency. So this is the contribution of each component of frequency ω to the average power.

So, that way if we integrate $S_X(\omega)$ then we will get the average power. So, $E[X^2[n]]$ is equal to integration of $S_X(\omega) d\omega$ this way we can define $S_X(\omega)$. And now how to define $S_X(\omega)$, So $S_X(\omega)$ is equal to limit and tends to infinity $1 / (2N + 1)$ into expected value this quantity. This is the DTFT of $X[n]$ n varying from $-N$ to $+N$. This is the DTFT of the signal in a window from $-N$ to $+N$.

So we are considering a window from $-N$ to $+N$ and we are taking the DTFT of the signal within this window. Now for a WSS process this quantity exists actually though Fourier transform does not like this quantity exist expected value of this quantity divided by $2N + 1$. If we take

limit tends to infinity that will be called that the power spectral density. So we have defined power spectral density.

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Wiener-Khinchin theorem

❖ $R_X(m)$ and $S_X(w)$ form a DTFT pair

$$S_X(w) = \sum_{m=-\infty}^{\infty} R_X(m) e^{-j\omega m} \quad -\pi \leq w \leq \pi$$

Taking the inverse DTFT

$$R_X(m) = \frac{1}{2\pi} \int_{-\pi}^{\pi} S_X(w) e^{j\omega m} dw$$

❖ $S_X(w)$ is a real and even function of w

❖ The generalized PSD in the z-domain is defined as follows

$$S_X(z) = \sum_{m=-\infty}^{\infty} R_X(m) z^{-m}$$

Now we defined what is known as Wiener Khinchin theorem that is the relationship between autocorrelation function and power spectral density $R_X(m)$ and $S_X(\omega)$ form DTFT pair. So they are Fourier transform of each other. So $S_X(\omega)$ is equal to that this is this discrete time Fourier transform $R_X(m)$ into $e^{-j\omega m}$ summation of m is equal to infinity minus to plus infinity.

So this is the definition of power spectral density. Now this power spectral density is periodic function of ω therefore we can take the inverse DTFT and we get the autocorrelation function by this integration. So it is in the $1/2\pi$ into integral minus π to π of $S_X(\omega) e^{j\omega m} d\omega$. So that way we can find out the power spectral density and from power spectral density we can find out the autocorrelation function from this definition autocorrelation function is an even function and.

Since from this definition power spectrum is a real function, it is a real function because the DTFT is a complex function power spectrum is only real function. And since autocorrelation is even function therefore this part is even and since autocorrelation function is also real and therefore $S_X(\omega)$ also must be even so, with this argument autocorrelation $S_X(\omega)$ is real

and even function of omega. Now it is helpful in signal processing if we have the generalized power spectral density in terms of z transform.

So, is the $S_X(z)$ transforms of the autocorrelation sequence as $S_X(z)$ is summation m going from minus infinity to infinity of $R_X(m)$ into z to the power $-m$ and we can get back $R_X(m)$ in using this inverse formula? So that way that is a complex number now instead of e to the power $j\omega$ only we have used z .

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Example

Suppose

$$R_X(m) = a^{|m|} \quad |a| < 1$$

Then the PSD is given by

$$S_X(\omega) = \frac{1-a^2}{1-2a\cos\omega+a^2} \quad -\pi \leq \omega \leq \pi$$

The generalised PSD is

$$S_X(z) = \frac{1-a^2}{1-a(z+z^{-1})+a^2}$$

Handwritten notes on the right side of the slide:

$$e^{j\omega} = z$$

$$e^{-j\omega} = z^{-1}$$

$$\cos\omega = \frac{z+z^{-1}}{2}$$

That will give 1 example suppose $R_X(m)$ is equal to $a^{|m|}$ where $|a| < 1$ then only it will become auto correlation function then PSD we can apply the formula of power spectral density that power series expansion. So we will get $S_X(\omega) = \frac{1-a^2}{1-2a\cos\omega+a^2}$. This is the power spectral density and this power spectral density.

We can convert into the generalized PSD by putting suppose instead of e to the power $-j\omega$ will put that is equal z to that it is minus $j\omega$ is equal to z^{-1} . So that if I delete both then $\cos\omega$ will be $\frac{z+z^{-1}}{2}$, so that way we get this relationship so that we introduced power spectrum and generalize power spectrum power spectral density is also called power spectrum.

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Summary

❖ A random process $\{X(t)\}$ is a mapping from the sample space to a waveform

❖ We defined a discrete-time random process $\{X(n)\}$ as an indexed family of random variables, where the index set $\Gamma \subseteq \mathbb{Z}$

❖ Stationarity plays an important role in analysing a random process.

For a strict-sense stationary process $\{X(n)\}$,

$$F_{X(n_1), X(n_2), \dots, X(n_k)}(x_1, x_2, \dots, x_k) = F_{X(n_1+h), X(n_2+h), \dots, X(n_k+h)}(x_1, x_2, \dots, x_k), \\ \forall k \in \mathbb{N} \text{ and } \forall h, n_1, \dots, n_k \in \Gamma$$

❖ An RP $\{X(n)\}$ is a wide-sense stationary if $EX(n) = \text{constant}$ and

$R_X(m) = EX(n)X(n+m)$ is a function of lag m only

And let us summarize a random process X_t is a mapping from the sample space to a waveform. So this is one important point we define discrete time random process X_n as an index family of random variables where the index set τ is a subset of $\mathbb{Z}$ we can take $\mathbb{Z}$ as the index set that this setup integers. Stationarity plays an important role in analyzing a random process for a strict sense stationary random process this CDF at instant $n_1, n_2, \dots, n_k$ is same as the CDF at instant $n_1 + h, n_2 + h, \dots, n_k + h$.

So, CDF is invariant under (())(34:04) it time exists. And this is true for all k and this is also all placement of $h, n_1, n_2, \dots, n_k$. So that way this is a very complicated expression which we have to apply to test whether a random processes is strict strictly stationary or not. So to realize this condition we defined it a random process X_n is it is a wide sense stationary if E of X_n is constant, and R_X of m that is the autocorrelation function that is by E of X_n into X_{n+m} .

And it is a function of lag m only. So it does not depend on and it depends on only the time difference and then we call it a WSS random process we gave one example of WSS random process.

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Summary contd..

- ❖ A discrete-time WSS random process $\{X(n)\}$ is characterised in terms of the autocorrelation function $R_X(m)$
- ❖ $R_X(m)$ is an even function of m with a maximum at $m=0$.
- ❖ $R_X(m)$ and $S_X(w)$ form a DTFT pair

$$S_X(w) = \sum_{m=-\infty}^{\infty} R_X(m) e^{-j\omega m} \quad -\pi \leq w \leq \pi$$
$$R_X(m) = \frac{1}{2\pi} \int_{-\pi}^{\pi} S_X(w) e^{j\omega m} dw$$

- ❖ The generalized PSD in the z-domain is defined as

$$S_X(z) = \sum_{m=-\infty}^{\infty} R_X(m) z^{-m}$$

THANK YOU

Also a discrete time WSS random process X_n is characterized in terms of autocorrelation function R_X of m because mean is constant anyway. Therefore this autocorrelation function R_X of m characterizes a discrete time random process. R_X of m is an even function of m with maximum at $m = 0$. So that we have proved that R_X of m at the maximum at $m = 0$ that we proved in Cauchy Schwarz inequality. Then this is the Wiener-Khinchin theorem for discrete time random processes that autocorrelation function and power spectral density form a DTFT pair.

That is S_X of ω is the discrete time Fourier transform of the R_X m sequence. Similarly R_X m can be formed from the PSD by this expression and in signal processing the generalized PSD is useful. So it is given by z transform that z transform of R_X n sequence and when we put $z = e$ to the power R_X ω we get the power spectrum that is S_X ω we can get. We also gave one example to illustrate how S_X of z that can be obtained from S_X of ω . In the next lecture we will discuss the output response of a linear system when the stationary WSS signal is applied as an input. Thank you