

**Statistical Signal Processing**  
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**Lecture 15**  
**Bayesian Estimator**

Hello students will come to the lecture, in this lecture we will introduce a powerful class of estimators known as Bayesian estimators

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Let us recall:

The method of moments and the MLE method use a probability model of the random data, characterized by unknown constant parameters. The parameters are estimated by exploiting this probability model alone.

❖ The ML estimators are based on elegant mathematical theories and they satisfy the desirable properties like unbiasedness, efficiency and consistency under large sample conditions.

❖ MLE may perform very poorly under limited data conditions.

Let us recall; the method of moments and the maximum likelihood estimation method use a probability model of the random data, this model is characterized by unknown constant parameters. The parameters are estimated by exploiting this probability model alone. The ML estimators are based on the elegant mathematical theories and they satisfy the desirable properties like unbiasedness, efficiency and consistency under large sample conditions.

So to be a good estimators ML estimators required lot of data. MLE may perform very poorly under limited data a condition that is one drawback of MLE.

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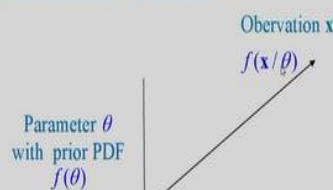
- ❖ In this lecture, we will introduce a class of optimal estimators known as the *Bayesian estimators*.
- ❖ These estimators exploit the probability model of the random data as well as the prior information about the parameters.

In this lecture we will introduce a class of optimal estimators known as the Bayesian estimators. These estimators exploit the probability model of the random data as well as the prior information about the parameters.

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### Bayesian Estimators

- ❖ In the MM and MLE, the parameter  $\theta$  is assumed to be an unknown constant.
- ❖ We may have some prior information about  $\theta$  in a sense that some values or ranges of  $\theta$  are more likely. We can represent this prior information in the form of a prior PDF  $f(\theta)$  or a prior PMF  $p(\theta)$ .
- ❖ Bayesian estimators takes into account these priors along with the random samples while constructing the estimator.



In the MM and MLE, the parameter  $\theta$  is assumed to be an unknown constant. We may have some prior information about  $\theta$  in the sense that some values or ranges of  $\theta$  are more likely. We can represent this prior information in the form of a prior PDF prior density function  $f(\theta)$  or a prior mass function  $p(\theta)$ . So this is important, now whatever prior information about  $\theta$  is available that is expressed in the form of a PDF or a PMF.

A Bayesian estimator takes into account of these priors along with the random samples while constructing the estimator. That is an important aspect. Now we will use the random data as

well as the prior  $f(\theta)$  or  $p(\theta)$ . So that way parameter  $\theta$  is now random and it has a prior PDF  $f(\theta)$  and data is generated according to this model  $f(x|\theta)$  of  $x$  given  $\theta$ . Suppose here  $\theta$  is there which is a random variable and for a given  $\theta$  we will have this  $f(x|\theta)$  given  $\theta$  which will generate the observed data.

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**Prior and Posterior PDFs**

- ❖ The parameter  $\theta$  is now a particular value of a the random variable  $\Theta$ .
- ❖ The likelihood function  $f(x|\theta)$  will now be the conditional PDF denoted by  $f(x|\theta)$ .
- ❖ The joint PDF of the random parameter  $\Theta$  and the random data vector  $\mathbf{X} = [X_1, X_2, \dots, X_N]'$  at point  $\mathbf{x} = [x_1, x_2, \dots, x_N]'$  is given by
$$f(\theta, \mathbf{x}) = f(\theta)f(\mathbf{x}|\theta)$$
- ❖ The posterior PDF  $f(\theta|\mathbf{x})$  can be obtained by applying the Bayes rule
$$f(\theta|\mathbf{x}) = \frac{f(\theta)f(\mathbf{x}|\theta)}{f_{\mathbf{x}}(\mathbf{x})}$$

$$= \frac{f(\theta)f(\mathbf{x}|\theta)}{\int_{\theta \in D_{\theta}} f(\theta)f(\mathbf{x}|\theta)d\theta}, D_{\theta} \text{ is the support where } f(\theta) \neq 0.$$

Let us explain what are prior and posterior PDFs; the parameter  $\theta$  is now a particular value of random variable under this condition that  $\theta$  is random therefore this  $\theta$  whatever parameter we get that is a particular value of a random variable big  $\theta$ . So big  $\theta$  is the random variable and small  $\theta$  is a particular value the likelihood function  $f(x|\theta)$ ,  $\theta$ .

Now with the conditional PDF denoted by  $f(x|\theta)$  given  $\theta$  earlier we used the likelihood function as a function of  $\theta$  but now it will be a conditional PDF  $f(x|\theta)$  of  $x$  given  $\theta$ . the joint PDF of the random variable  $\theta$  and the random data vector  $\mathbf{X}$  this is the random data vector at the point small  $\mathbf{x}$  vector that is  $x_1, x_2$  up to  $x_N$  transpose is given by this expression, this is the expression for a joint PDF  $f(\theta, \mathbf{x})$  of  $\theta, \mathbf{x}$  is equal to  $f(\theta)$  multiplied by  $f(x|\theta)$  given  $\theta$ .

The posterior PDF  $f(\theta|\mathbf{x})$  of  $\theta$  given  $\mathbf{x}$  can be obtained by applying the Bayes rule  $f(\theta|\mathbf{x})$  of  $\theta$  given  $\mathbf{x}$  that will be equal to  $f(\theta)f(\mathbf{x}|\theta)$  into  $f(\mathbf{x})$  given  $\theta$  that is designed PDF of data  $N \times 1$  divided by  $f_{\mathbf{x}}(\mathbf{x})$  of  $\mathbf{X}$ , this is the marginal PDF of  $\mathbf{X}$ . now this marginal PDF  $f_{\mathbf{x}}(\mathbf{x})$  of  $\mathbf{X}$  can be written in terms of the prior and the conditional PDF, that is integration of  $f(\theta)f(\mathbf{x}|\theta)$  over  $\theta$ .

given  $\theta$   $D_\theta$  over  $\theta$  belonging to  $D_\theta$  where  $D_\theta$  is the support of  $f_\theta$  where  $f_\theta$  is not equal to 0.

So that way we have expressed the posterior PDF in terms of the prior PDFs and conditional PDF. So  $f_\theta$  of  $\theta$  given  $X$  is given as  $f_\theta$  into  $f(x|\theta)$  divided by integration  $f_\theta$  of  $x$  given  $\theta$   $d\theta$ ,  $\theta$  belongs to  $D_\theta$  where  $D_\theta$  is the support of  $f_\theta$ .  
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**Example:** Let  $X$  be Gaussian random samples with unknown mean  $\theta$  and variance 1. Given  $\Theta \sim N(0,1)$ , find the a posteriori PDF  $f(\theta/x)$  for a single observation  $x$ .

**Solution:** We have  $f(\theta) = \frac{e^{-\frac{1}{2}\theta^2}}{\sqrt{2\pi}}$  and  $f(x|\theta) = \frac{e^{-\frac{1}{2}(x-\theta)^2}}{\sqrt{2\pi}}$

$$\begin{aligned} \therefore f(x) &= \int_{-\infty}^{\infty} \frac{e^{-\frac{1}{2}\theta^2}}{\sqrt{2\pi}} \frac{e^{-\frac{1}{2}(x-\theta)^2}}{\sqrt{2\pi}} d\theta \\ &= \frac{e^{-\frac{1}{2}x^2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \frac{e^{-\theta^2 + x\theta}}{\sqrt{2\pi}} d\theta \\ &= \frac{e^{-\frac{1}{2}x^2 + \frac{x^2}{4}}}{\sqrt{2\pi} \times 2} \int_{-\infty}^{\infty} \frac{e^{-(\theta^2 - 2\theta \frac{x}{2} + \frac{x^2}{4})}}{\sqrt{2\pi \times \frac{1}{2}}} d\theta = \frac{e^{-\frac{1}{4}x^2}}{\sqrt{4\pi}} \end{aligned}$$

Let us consider one example; let  $X$  be Gaussian random samples with unknown mean  $\theta$  and variance 1. Given  $\theta$  distributed as normal 0, 1 find a posteriori PDF  $f_\theta$  given  $x$  for a single observation  $x$ . In this case only one observation is there and given that  $\theta$  that parameter is distributed as a normal random variable with mean 0 and variance 1. So let us solve this problem we have to find the a posteriori PDF  $f$  of  $\theta$  given  $x$ .

Now we have  $f_\theta$  this is normal with mean 0 and variance 1, so it will be  $f_\theta$ ,  $\theta$  will be  $e$  to the power - half  $\theta^2$  divided by root over 2 pi and the likelihood function or conditional PDF is given by  $f(x|\theta)$  is equal to  $e$  to the power - half of  $(x - \theta)^2$  divided by root over 2 pi. So we have to find out the marginal PDF  $f(x)$ , so that will be a joint PDF and if we integrate with respect to  $\theta$ .

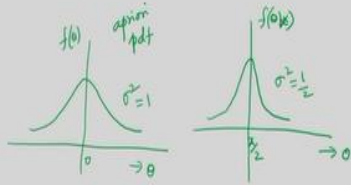
So that way it will be integration from - infinity to infinity  $e$  to the power - half  $\theta^2$  divided by root 2 pi into  $e$  to the power - half of  $(x - \theta)^2$  divided by root 2 pi into  $d\theta$ . So this integration now we will carry out so we are integrating with respect to  $\theta$  so we can take out all the terms involving  $x$  only, so that way this  $e$  to the power - half  $x^2$

square will come out and then we can make it a whole square, so that it may be expressed as a Gaussian.

So that we have we write  $x$  square by 4 and subtract a square by 4 here, so that way this part will be a Gaussian now and integration of the Gaussian is equal to 1, therefore what will be left to it,  $e$  to the power  $-\frac{1}{4}$  of  $x$  square divided by root 4 pi.

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Posterior PDF

$$\begin{aligned} \therefore f(\theta/x) &= \frac{f(\theta, x)}{f(x)} \\ &= \frac{e^{-\frac{1}{2}\theta^2} e^{-\frac{1}{2}(x-\theta)^2}}{\sqrt{2\pi} \sqrt{2\pi}} \\ &= \frac{e^{-\frac{1}{4}x^2}}{\sqrt{4\pi}} \\ &= \frac{e^{-\frac{(\theta - \frac{x}{2})^2}{2}}}{\sqrt{\pi}} \sim N\left(\frac{x}{2}, \frac{1}{2}\right) \end{aligned}$$


We have to find out the posterior PDF  $f$  of  $\theta$  given  $x$ , it is given by  $f$  of  $\theta$   $x$  divided by  $f$  of  $x$  and if we substitute  $f$  of  $\theta$   $x$  already we have seen that  $e$  to the power  $-1/2$  of  $\theta$  square into  $e$  to the power  $-\frac{1}{2}$  of  $x - \theta$  square divided by  $2\pi$  into root  $2\pi$  and  $f$  of  $x$  will be equal to this quantity  $e$  to the power  $-x$  square by 4 divided by root over  $4\pi$

And this if we simplify, we will get  $e$  to the power  $-\frac{(\theta - \frac{x}{2})^2}{2}$  whole square divided by root  $\pi$ , this is a normal distribution with mean  $x$  by 2 and variance  $\frac{1}{2}$ . So what we have got that this posterior distribution is a normal distribution with mean  $x$  by 2 and variance half, earlier we have  $f$   $\theta$  that is a priori  $\theta$ ,  $f$   $\theta$  and this is 0 and earlier it was it like this normal distribution, this is the a priori PDF it has variance 1.

Now if we consider  $f$  of  $\theta$  given  $x$  a posteriori PDF, so here  $\theta$  now it will be a distribution this point is  $X$  by 2 and variance is now earlier variance was 1 now variance is half. So, that way what we see that given the information about the data the a posteriori PDF is now governed by the data the data will determine the a posteriori PDF and also we see that earlier whatever uncertainty was there that variance was 1.

Now it has billions has reduced, so this once we have the data the uncertainty about the parameter is reduced earlier this PDF was how its standard deviation Sigma squared it is equal to 1 and here Sigma square is equal to half. So uncertainty is reduced, as well as the mean is now shifted towards the data so that way from 0 we have x by 2 that way this distribution is now biased towards the actual data. So that way we have the idea of prior PDF and posteriori PDF.

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**Cost functions.**

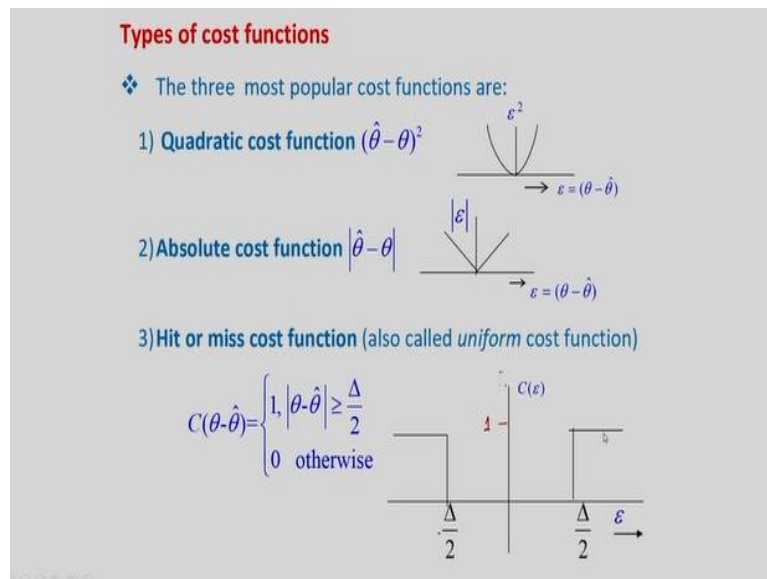
- ❖ The parameter  $\Theta$  is a random variable and the estimator  $\hat{\theta}(X)$  is another random variable.
- ❖ Estimation error  $\varepsilon = \theta - \hat{\theta}$ .
- ❖ We have to minimize some error measures, termed as **cost functions**.
- ❖ We associate a **cost function** or a **loss function**  $C(\theta - \hat{\theta})$  with the estimator  $\hat{\theta}$ .
- ❖ It represents the positive penalty with each wrong estimation.
- ❖ Thus  $C(\theta - \hat{\theta})$  is a non-negative function of the error.

Let us define another term what is known as cost function. The parameter theta is a random variable and the estimator theta hat X is another random variable. So we have two random variable and estimation error at a particular point is given by theta hat - theta. In Bayesian framework, we have to minimize some error measures termed as cost function. These error measures are called cost functions.

We associate a cost function or a loss function of these forms C of theta - theta hat with the estimator theta hat. So if theta hat is the estimator this cost function is a function of the error theta is estimated value. So it is a function of error. It represents the positive penalty with each wrong estimation. So we have a this function will be a positive function if we commit more error than this function should be large.

Thus C of theta - theta hat is a non-negative function of the error theta - theta hat is the error, therefore C of theta - theta hat is a non-negative function of the error. So the cost function is a non-negative function of the error.

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The three most popular cost functions are; number one quadratic cost function or it is also called square error cost function because this is the square of the error, squared error cost function. So if we plot it suppose this x is epsilon is equal to theta - theta hat error and this side is the cost function. So if we put epsilon squared that will be a parabola, this is the quadratic or square error cost function.

Absolute cost function, instead of considering the square here we can take the absolute value of theta hat - theta, so then we get this cost function is plotted here actually it is a plot of mode of e. So that way there will be two lines like this. The third cost function is hit-or-miss cost function and it is also called the uniform cost function. it is given by C of theta - theta hat is equal to 1, that is uniformly 1 for absolute value of theta - theta hat to get our than equal to delta by 2 and 0 otherwise.

So this is the plot of the hit-or-miss cost function here, this delta is a small quantity. So it is a very small quantity within this error is and that is cost to the error is 0 and if we cross delta by 2 this side already said that error will be uniformly 1. So that is a very strict condition on the cost function

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### Risk function and Bayesian estimation problem

- ❖ The parameter  $\Theta$  and the estimator  $\hat{\theta}$  are joint RVs.
- ❖ The Bayesian risk function or the average cost is defined as

$$\bar{C}(\hat{\theta}) = EC(\Theta - \hat{\theta}) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} C(\theta - \hat{\theta}) f(\mathbf{x}, \theta) d\mathbf{x} d\theta$$

The estimator  $\hat{\theta}$  seeks to minimize the Bayesian Risk.

- ❖ Thus, the Bayesian estimation problem is

$$\text{Minimize}_{\text{over } \hat{\theta}} \bar{C}(\hat{\theta}) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} C(\theta - \hat{\theta}) f(\mathbf{x}, \theta) d\mathbf{x} d\theta$$

- ❖ The Bayesian estimator is given by

$$\hat{\theta}_{\text{Bayes}} = \arg \min_{\hat{\theta}} \bar{C}(\hat{\theta})$$

And now, we will define this function and the Bayesian estimation problem the parameter  $\theta$  and the estimator  $\hat{\theta}$  are joint random variables. The Bayesian risk function or the average cost function is defined as this, actually we have defined the cost function expected value of the cost function is the least function that we denote by  $\bar{C}(\hat{\theta})$  it is a function of  $\hat{\theta}$  and it is given by expected value of  $C$  of  $\theta - \hat{\theta}$ .

And if we expand it, it will be integration from  $-\infty$  to  $\infty$  -  $\infty$  to  $\infty$   $C$  of  $\theta - \hat{\theta}$  into  $f(\mathbf{x}, \theta) d\mathbf{x} d\theta$ . The estimator  $\hat{\theta}$  seeks to minimize the Bayesian risk. So a Bayesian estimation problem is to minimize this risk function and thus the Bayesian estimator problem is given by minimize over  $\hat{\theta}$   $\bar{C}(\hat{\theta})$  that is equal to integration from  $-\infty$  to  $\infty$  -  $\infty$  to  $\infty$  of  $C$  of  $\theta - \hat{\theta}$  into  $f$  of  $\mathbf{x}, \theta d\mathbf{x} d\theta$ .

The Bayesian estimator is now given by,  $\hat{\theta}_{\text{Bayes}}$  is equal to  $\arg \min_{\hat{\theta}} \bar{C}(\hat{\theta})$ . So the value of  $\hat{\theta}$  corresponding to the minimum is the base estimator.

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### Minimization of the Bayesian risk function

We have to solve the minimization problem,

$$\underset{\text{over } \hat{\theta}}{\text{Minimize}} \bar{C}(\hat{\theta}) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} C(\theta - \hat{\theta}) f(\mathbf{x}, \theta) d\mathbf{x} d\theta$$

Applying Bayes rule, the problem can be rewritten as

$$\begin{aligned} \underset{\text{over } \hat{\theta}}{\text{Minimize}} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} C(\theta - \hat{\theta}) f(\theta | \mathbf{x}) f(\mathbf{x}) d\theta d\mathbf{x} \\ = \int_{-\infty}^{\infty} \left( \int_{-\infty}^{\infty} C(\theta - \hat{\theta}) f(\theta | \mathbf{x}) d\theta \right) f(\mathbf{x}) d\mathbf{x} \end{aligned}$$

❖ Since  $f(\mathbf{x})$  is always +ve, the above integral will be minimum if the inner integral is minimum.

Now let us see, how to minimize the Bayesian risk function; we have to solve the minimization problem, that is minimize over  $\hat{\theta}$   $\bar{C}(\hat{\theta})$  is equal to integration - infinity to infinity, - infinity to infinity  $C(\theta - \hat{\theta}) f(\mathbf{x}, \theta) d\theta$ . Now we will apply the Bayes rule and express this joint PDF as a product of the prior PDF and the conditional PDF.

So that way now minimizes some problem will become minimize over  $\hat{\theta}$  integration - infinity to infinity, - infinity to infinity of  $C(\theta - \hat{\theta}) f(\theta | \mathbf{x})$  that is the conditional PDF into  $f(\mathbf{x}) d\mathbf{x}$ . Now we can express this integration as two integration like this one is involving  $\theta$  only and other part involving  $\mathbf{x}$ . So that way the inner integral involves  $\theta$ ,  $\hat{\theta}$  and  $\mathbf{x}$  but the outer integral involves only  $\mathbf{x}$ .

So that way this minimization problem we have written like this. Now, we know that this  $f(\mathbf{x})$  is always a positive quantity, therefore this entire double integral will be minimum whenever the inner integral is minimum. So since  $f(\mathbf{x})$  is always positive the above integral will be minimum, if the inner integral is minimum.

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### Minimization of the Bayesian risk function...

❖ The minimization problem can be simplified as

$$\underset{\text{over } \hat{\theta}}{\text{Minimize}} \int_{-\infty}^{\infty} C(\theta - \hat{\theta}) f(\theta | \mathbf{x}) d\theta$$

❖ Three different estimators result corresponding to three types of cost functions.

The minimization problem is now simplified and it can be written as minimize that inner integral basically over theta hat integration - infinity to infinity C of theta - theta hat into f of theta given x, d theta. Now three different estimators result corresponding to three types of cost function. We know that we have three cost function; squared error cost function, absolute error cost function and uniform error cost function.

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### Quadratic Cost Function and the Minimum Mean-square Error Estimator

❖ Quadratic cost function  $C(\theta - \hat{\theta}) = (\theta - \hat{\theta})^2$

❖ The estimation problem is

$$\underset{\text{over } \hat{\theta}}{\text{Minimize}} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (\theta - \hat{\theta})^2 f(\mathbf{x}, \theta) d\mathbf{x} d\theta$$

❖ This is equivalent to minimizing

$$\begin{aligned} \underset{\text{over } \hat{\theta}}{\text{Minimize}} & \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (\theta - \hat{\theta})^2 f(\theta | \mathbf{x}) f(\mathbf{x}) d\theta d\mathbf{x} \\ & = \int_{-\infty}^{\infty} \left( \int_{-\infty}^{\infty} (\theta - \hat{\theta})^2 f(\theta | \mathbf{x}) d\theta \right) f(\mathbf{x}) d\mathbf{x} \end{aligned}$$

So first we will be considering quadratic cost function or square error cost function and the the corresponding minimum mean square error estimator. Quadratic cost function or square error cost function that is given by C of theta - theta hat is equal to theta - theta hat all square. The estimation problem is this minimize over theta hat integration - infinity to infinity - infinity to infinity of theta - theta hat whole square f of x theta dx, d theta

And then as I told this we will write in terms of the prior and conditional PDF, so this is the expression and as I did earlier so I will have an inner integral like this and outer integral involving  $f(x)$  like this.

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**Minimum mean-square error (MMSE) estimator**

❖ As argued earlier, the simplified problem is

$$\text{Minimize}_{\hat{\theta}} \int_{-\infty}^{\infty} (\theta - \hat{\theta})^2 f(\theta | \mathbf{x}) d\theta$$

❖ At the minimum,

$$\frac{\partial}{\partial \hat{\theta}} \int_{-\infty}^{\infty} (\theta - \hat{\theta})^2 f(\theta | \mathbf{x}) d\theta = 0$$

$$\Rightarrow -2 \int_{-\infty}^{\infty} (\hat{\theta} - \theta) f(\theta | \mathbf{x}) d\theta = 0$$

$$\Rightarrow \hat{\theta} \int_{-\infty}^{\infty} f(\theta | \mathbf{x}) d\theta = \int_{-\infty}^{\infty} \theta f(\theta | \mathbf{x}) d\theta$$

$$\therefore \hat{\theta} = \int_{-\infty}^{\infty} \theta f(\theta | \mathbf{x}) d\theta$$

Thus, the MMSE estimator is

$$\hat{\theta}_{\text{MMSE}} = \int_{-\infty}^{\infty} \theta f(\theta | \mathbf{x}) d\theta, \text{ the conditional mean or mean of the a posteriori PDF.}$$

So this simplified problem is now minimize integral from - infinity to infinity of  $(\theta - \hat{\theta})^2$  into  $f(\theta | x)$   $d\theta$ . So this is the mean square error we are trying to minimize. At the minimum, now we will take the partial derivative with respect to  $\hat{\theta}$  of this expression must be equal to 0 and we can take this differentiation inside because these limits do not involve in any  $\hat{\theta}$  time.

Therefore we can take this partial differentiation inside and carrying out the partial differentiation we will get that  $-2$  times integral from - infinity to infinity of  $(\hat{\theta} - \theta)$  into  $f(\theta | x)$   $d\theta$  is equal to 0 and after simplifying we will get that corresponding to this term  $\hat{\theta}$  integration - infinity to infinity  $f(\theta | x)$   $d\theta$  is equal to integral - infinity to infinity  $\theta f(\theta | x)$   $d\theta$ .

Now this part I know because this is a PDF, if you integrate from - infinity to infinity you will get 1, so left side is simply  $\hat{\theta}$  and right side this quantity is the conditional expectation  $\theta f(\theta | x)$   $d\theta$  integral of from - infinity to infinity. So this is the conditional expectation, therefore  $\hat{\theta}$  will be equal to integration from - infinity to infinity of  $\theta f(\theta | x)$   $d\theta$ , this is the conditional expectation.

So therefore what we get is that the MMSE, minimum mean square error estimator is given by  $\hat{\theta}_{MMSE}$  is equal to integral - infinity to infinity of  $\theta f(\theta | \mathbf{x}) d\theta$ , this is the conditional mean or mean of the a posteriori a PDF. Therefore we have to find out the a posteriori PDF  $f(\theta | \mathbf{x})$  and mean of this will be the  $\hat{\theta}_{MMSE}$ .

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**MMSE steps**

- ❖ Given a *priori* density function  $f(\theta)$  and the conditional PDF  $f(\mathbf{x} | \theta)$ .
- ❖ Determine a posteriori density  $f(\theta | \mathbf{x})$ . This is determined from the Bayes rule:
 
$$f(\theta | \mathbf{x}) = \frac{f(\theta)f(\mathbf{x} | \theta)}{\int_{\theta \in D_\theta} f(\theta)f(\mathbf{x} | \theta) d\theta}$$
- ❖ Find  $\hat{\theta}_{MMSE} = E(\Theta | \mathbf{X} = \mathbf{x})$

So therefore we can find out the MMSE, minimum mean square error estimate or in the following steps given a *priori* density function  $f(\theta)$  and the conditional PDF of  $\mathbf{x}$  given  $\theta$  determined a posteriori PDF  $f(\theta | \mathbf{x})$ , this is determined using this Bayes rule;  $f(\theta | \mathbf{x})$  equal to  $f(\theta)f(\mathbf{x} | \theta)$  divided by integral  $\theta$  belonging to  $D_\theta$  of  $f(\theta)f(\mathbf{x} | \theta) d\theta$ ,  $D_\theta$  is the support of  $f(\theta)$ ,  $f(\theta)$  into  $f(\mathbf{x} | \theta)$  divided by integral  $\theta$  belonging to  $D_\theta$  of  $f(\theta)f(\mathbf{x} | \theta) d\theta$ .

So this integral is the expression for  $f(\mathbf{x})$  marginal density of  $\mathbf{x}$  and once we have this conditional PDF, now we can find out the conditional mean that is  $\hat{\theta}_{MMSE}$  will be equal to  $E(\Theta | \mathbf{X} = \mathbf{x})$  that is the parameter vector given that data vector  $\mathbf{X}$  is equal to  $\mathbf{x}$ , that is the observed data. So that way we can find out the conditional mean and that is  $\hat{\theta}_{MMSE}$ .

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### Example:

Let  $X_1, X_2, \dots, X_N$  be samples of  $X \sim N(\theta, 1)$  with  $\theta \sim N(0, 1)$ . Find

$\hat{\theta}_{MMSE}$ .

### Solution:

$$f(\theta) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}\theta^2}$$

$$\begin{aligned} f(\mathbf{x} / \theta) &= \prod_{i=1}^N \frac{1}{\sqrt{2\pi}} e^{-\frac{(x_i - \theta)^2}{2}} \\ &= \frac{1}{(\sqrt{2\pi})^N} e^{-\frac{\sum_{i=1}^N (x_i - \theta)^2}{2}} \end{aligned}$$

$$\text{Also, } f(\theta / \mathbf{x}) = \frac{f(\theta)f(\mathbf{x} / \theta)}{f(\mathbf{x})}$$

Let us consider one example; let  $x_1, x_2$  up to  $x_N$  be samples of normal distribution  $X$  is normally distributed with mean  $\theta$  and variance 1 and  $\theta$  is also a random quantity with distribution normal 0, 1. That is the PDF of  $\theta$ , find  $\hat{\theta}_{MMSE}$ . So let us solve this problem  $f(\theta)$  is given by this is the normal distribution with mean 0 and variance 1 and conditional PDF  $f(\mathbf{x} / \theta)$  of  $\mathbf{x}$  given  $\theta$  there are  $N$  observations are there, therefore they are iid.

So it will be product  $i$  going from 1 to  $N$  of  $1 / \sqrt{2\pi} e^{-\frac{1}{2}(x_i - \theta)^2}$ , okay? And this, if I carry out the product and then I will get in terms of summation like this, so it will be  $1 / (\sqrt{2\pi})^N e^{-\frac{1}{2} \sum_{i=1}^N (x_i - \theta)^2}$ . Now the a posteriori PDF  $f(\theta / \mathbf{x})$  of  $\theta$  given  $\mathbf{x}$  is given by  $f(\theta)$  into  $f(\mathbf{x} / \theta)$  divided by  $f(\mathbf{x})$ ,  $f(\mathbf{x})$  is the marginal PDF.

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### Example ...

where

$$\begin{aligned} f(\mathbf{x}) &= \int_{-\infty}^{\infty} f(\theta)f(\mathbf{x} / \theta) d\theta \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}\theta^2} \frac{1}{(\sqrt{2\pi})^N} e^{-\frac{\sum_{i=1}^N (x_i - \theta)^2}{2}} d\theta \\ &= \frac{1}{(\sqrt{2\pi})^{N+1}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}\theta^2 - \frac{\sum_{i=1}^N (x_i - \theta)^2}{2}} d\theta \\ &= \frac{1}{(\sqrt{2\pi})^{N+1}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}\theta^2 - \frac{\sum_{i=1}^N x_i^2 - 2\theta \sum_{i=1}^N x_i + N\theta^2}{2}} d\theta \\ &= \frac{1}{(\sqrt{2\pi})^{N+1}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(N+1)\theta^2 + \theta \sum_{i=1}^N x_i - \frac{\sum_{i=1}^N x_i^2}{2}} d\theta \end{aligned}$$

So we can calculate  $f(x)$  like this that is joint density integrated over  $\theta$ , so we can put the expression for  $f(\theta)$  and  $f(x|\theta)$  and then we all carry out the integration and this will give us like this,  $e$  to the power - summation  $i$  going from 1 to  $N$  of  $x_i$  squared divided by  $2 + N$  into  $\bar{x}$  square divided by 2 into  $N+1$ . So this we can simplify it like this,

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**Example ...**

$$\begin{aligned} \therefore f(\theta | \mathbf{x}) &= \frac{f(\theta)f(\mathbf{x}|\theta)}{f(\mathbf{x})} \\ &= \frac{e^{-\frac{1}{2}\theta^2} \sum_{i=1}^N \frac{(x_i - \theta)^2}{2}}{\frac{(\sqrt{2\pi})^{N+1}}{e^{-\frac{1}{2} \sum_{i=1}^N \frac{x_i^2}{2} - \frac{N^2 \bar{x}^2}{2(N+1)}}}} \quad \bar{x} = \frac{1}{N} \sum_{i=1}^N x_i \\ &= \frac{(\sqrt{2\pi})^N}{e^{-\frac{N+1}{2} \left( \theta - \frac{N}{N+1} \bar{x} \right)^2}} \\ &= \frac{1}{\sqrt{2\pi \times \frac{1}{N+1}}} \\ \therefore \hat{\theta}_{MMSE} &= E(\theta | \mathbf{X} = \mathbf{x}) \\ &= \frac{N}{N+1} \bar{x} \end{aligned}$$

We see that  $\hat{\theta}_{MMSE}$  is a scaled version of the MLE.

This  $\bar{x}$  is and this quantity  $\bar{x}$  is  $1$  by  $N$  into summation  $x_i$ ,  $i$  going from  $1$  to  $N$  this is the sample mean. Therefore the posterior PDF is given by  $f(\theta | \mathbf{x})$  is equal to  $f(\theta)$  into  $f(\mathbf{x}|\theta)$  divided by  $f(\mathbf{x})$  and we will substitute all this quantity now, we have the expression for  $f(\mathbf{x})$  also this is the expression for  $f(\mathbf{x})$ . So if we substitute and simplify we will get this expression,  $e$  to the power - half into  $N+1$  into  $\theta - \frac{N}{N+1} \bar{x}$  whole square.

So this is again a Gaussian mean is given by this therefore this conditional mean or  $\hat{\theta}_{MMSE}$  will be equal to  $E(\theta | \mathbf{x})$  is equal to small  $\mathbf{x}$  vector and that will come out to be this  $N$  divided by  $N+1$  into  $\bar{x}$ , where  $\bar{x}$  is this sample mean this is the expression for sample mean. We see that  $\hat{\theta}_{MMSE}$  is a scaled version of the MLE. So with the prior information about  $\theta$  we are able to find out  $\hat{\theta}_{MMSE}$  given by this  $N$  divided by  $N+1$  into  $\bar{x}$ , whereas what is this sample mean and we hope it would do better than MLE.

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### Summary

- ❖ In Bayesian estimators, the parameter  $\theta$  is assumed to be random variable with the prior PDF  $f(\theta)$  or a prior PMF  $p(\theta)$ .
- ❖ The priors along with the random samples are taken into account for constructing the estimator.
- ❖ The estimator uses the posterior PDF obtained using the Bayes rule:
$$f(\theta / \mathbf{x}) = \frac{f(\theta)f(\mathbf{x} / \theta)}{f(\mathbf{x})}$$
- ❖ We associate a *cost function* or a *loss function*  $C(\theta - \hat{\theta})$  with the estimator  $\hat{\theta}$ .
- ❖ A Bayesian estimator solves the optimization problem

$$\underset{\text{over } \hat{\theta}}{\text{Minimize}} \bar{C}(\hat{\theta}) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} C(\theta - \hat{\theta}) f(\mathbf{x}, \theta) d\mathbf{x} d\theta$$

Let us summarize the lecture; in Bayesian estimators, the parameter  $\theta$  is assumed to be random variable with the prior PDF  $f(\theta)$  or a prior PMF  $P(\theta)$ . The priors along with the random samples are taken into account for constructing the estimator. So we have to consider both the prior and model of this random samples. The estimator uses the posterior PDF obtained using the Bayes rule; that is  $f(\theta \text{ given } \mathbf{x})$  equal to  $f(\theta)$  into  $f(\mathbf{x} \text{ given } \theta)$  divided by  $f(\mathbf{x})$ .

So this is the marginal PDF, this is the conditional PDF of  $\mathbf{x}$  given  $\theta$  and this is the prior PDF. In Bayesian estimator, we associate a cost function or a loss function  $C$  of  $\theta - \hat{\theta}$  with the estimator  $\hat{\theta}$ . So this is a positive error measure and this cost function we want to minimize. A Bayesian estimator solve the optimization problem, this is the optimization problem here, minimize over  $\hat{\theta}$   $\bar{C}(\hat{\theta})$  that is the expected cost function or average cost function or least function

And that is equal to integral from  $-\infty$  to  $\infty$   $C(\theta - \hat{\theta})$  into  $f(\mathbf{x}, \theta) d\mathbf{x} d\theta$ , this is the Bayesian estimation problem as an optimization problem.

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### Summary...

❖ The three most popular cost functions are:

1) Quadratic or square-error cost function  $C(\hat{\theta} - \theta) = (\hat{\theta} - \theta)^2$ ,

2) Absolute cost function  $C(\hat{\theta} - \theta) = |\hat{\theta} - \theta|$  and

3) Hit or miss cost function (also called *uniform* cost function)

$$C(\theta - \hat{\theta}) = \begin{cases} 1, & |\theta - \hat{\theta}| \geq \frac{\Delta}{2} \\ 0 & \text{otherwise} \end{cases}$$

❖ The MMSE estimator minimizes the mean square error and is given by

$$\hat{\theta}_{MMSE} = \int_{-\infty}^{\infty} \theta f(\theta | \mathbf{x}) d\theta = E(\Theta | \mathbf{x})$$

The three most popular cost functions are; quadratic or square error cost function that is  $C$  of  $\theta$  hat -  $\theta$  is equal to  $\theta$  hat -  $\theta$  whole square. So this is this square error. Absolute cost function that is  $C$  of  $\theta$  hat -  $\theta$  will be equal to mode of  $\theta$  hat -  $\theta$ ; we take the absolute value of error. Third cost function is hit or miss cost function also called uniform cost function.

And it is given by  $C$  of  $\theta$  -  $\theta$  hat is equal to 1 if mode of  $\theta$  hat -  $\theta$  that is greater than equal to  $\Delta$  by 2 and equal to 0 otherwise, where  $\Delta$  is a small positive number the MMSE estimator minimizes the mean square error. This mean square error is minimized and is given by  $\theta$  hat MMSE is equal to this expression integral from - infinity to infinity  $\theta$  of  $\theta$  given  $\mathbf{x}$ ,  $d\theta$  and that is the conditional expectation of  $\theta$  given  $\mathbf{x}$ , this is the MMSE. Thank you.