

Randomized Algorithms
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Lecture – 10
Chernoff Bound

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Chernoff Bounds.

Toss n coins. Compute the probability that the number of heads $\geq \frac{3n}{4}$. $X = \# \text{ Hc.}$ then $E[X] = \frac{n}{2}$.

$P(X > a)$

Markov Inequality $P(X > \frac{3n}{4}) \leq \frac{E[X]}{(\frac{3n}{4})} = \frac{\frac{n}{2}}{2 \cdot \frac{3n}{4}} = \frac{2}{3}$

Chebyshev's Inequality $P(|X - E[X]| > a) < \frac{\text{Var}[X]}{a^2}$

$P(|X - \frac{n}{2}| > \frac{n}{4}) < \frac{\text{Var}[X]}{(\frac{n}{4})^2} = \frac{\frac{n}{4}}{\frac{n^2}{16}} = \frac{4}{n}$

$\text{Var}(X) = n \cdot p \cdot (1-p) = \frac{n}{4}$

So, in today's lecture we will learn about what is known as Chernoff Bounds ok. So, the question that we will address in this lecture is the following: let us say we toss 1 lakh coins and we will get some number of heads and some number of tails. What is the probability that the number of heads is greater than 75000. So, we look at the following problem toss n coins, compute the probability that the number of heads is greater than say $3n$ by 4 . Clearly we expect; so, if we denote by X the number of heads then expectation of x is equal to n by 2 .

So, now here in this case, the random variable takes a value much larger than its expectation ok. What is the probability of such a thing happening. So, this is a question of the following form; compute the probability that X is greater than a . We already have some tools to address these kind of questions, the first one the Markov inequality. If we use Markov inequality probability that X is greater than $3n$ by 4 will be less than or equal to expectation of x divided by $3n$ by 4 . So, this is equal to n by 2 into $3n$ by 4 . So,

that is equal to 2 by 3 ok. So, we can just conclude that the probability that the number of heads is significantly greater than n is going to be less than 2 by 3.

If you use Chebychev's inequality, so, this is the probability that X minus expectation of x greater than a this is going to be less than variance of x divided by a square. So, here expectation of x is n by 2. So, we have probability that X minus n by 2 to be greater than n by 4 ok. Whenever n is greater than 3 n by 4 X minus n by 2 is greater than the absolute value of that is going to be greater than n by 4. This happens with probability less than variance of x divided by n by 4 the whole square ok. Now variance of this particular random variable is going to be n times p into 1 minus p where, p is the bias of the coin. So, that is going to be equal to n by 4. Therefore, here we will get this as n by 4 into n square by 16.

So, this probability is going to be 4 by n much better than 2 by 3 ok. So, if you toss say 1 million coins the probability that the number of heads is greater than 0.75 million probability that the number of heads is greater than 0.75 million that is going to be something like 4 divided by 4 million ok. So, 1 in a million chance whereas earlier bound would have told us 2 by 3. So, this is a Chebychev's inequality is already a much better bound Markov inequality. We will try to get a significantly better bound in both of these bounds and that is what we will look at in this class.

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Moment Generating Functions.

Let X be a r.v. Clearly tX and e^{tx} are h.v.s.

$$M_X(t) \triangleq \mathbb{E}[e^{tx}]$$

$$= \mathbb{E}\left[1 + \frac{tx}{1!} + \frac{t^2 x^2}{2!} + \frac{t^3 x^3}{3!} + \dots\right]$$

$$M_X(t) = \mathbb{E}\left[1 + t \cdot \frac{\mathbb{E}[X]}{1!} + t^2 \cdot \frac{\mathbb{E}[X^2]}{2!} + \dots\right]$$

$$\left.\frac{d}{dt}(M_X(t))\right|_{t=0} = \mathbb{E}[X]$$

$$\left.\frac{d^2}{dt^2}(M_X(t))\right|_{t=0} = \mathbb{E}[X^2]$$

$$\left.\left(\frac{d}{dt}\right)^n (M_X(t))\right|_{t=0} = \mathbb{E}[X^n]$$

So, first in order to understand what is Chebychev, Chernoff bounds we will look at moment generating functions ok. So, let X be a random variable; so, clearly tx is also a random variable for any value of t and e raised to tx are so, clearly tx and e raised to tx are random variables; now we can compute the expectation of e raised to tx . So, moment generating function of this variable x is defined as the expectation of e raised to tx . So, look at this particular random variable; its expectation is going to be called the moment generating function. Note that the moment generating function is a function of a variable t ; for each value of t we will get a different random variable e raised to tx and those random variables can have their own expectations that expectation is called the so, that expectation is the value of $M_x(t)$.

So, this expression expectation of e raised to tx that is going to be dependent on t that function is called as a moment generating function. So, e raised to tx we can alternatively alternately write as expectation of $1 + tx + \frac{t^2 x^2}{2!} + \frac{t^3 x^3}{3!} + \dots$ and so on. And, by linearity of expectation, we can write this as expectation of $1 + t \text{ times } x + \frac{t^2}{2!} \text{ times } x^2 + \dots$ ok; so, this is your $M_x(t)$. So, you can verify that if you differentiate $M_x(t)$ with respect to t and evaluate this function at t equals 0 you will get expectation of x .

So the moment generating function has in itself the value of the expectation of the random variable embedded. Similarly, if you look at the second derivative of the function $M_x(t)$ and evaluated at t equals 0 you will get this to be equal to expectation of x^2 . In general, if you differentiate it n times then you will and evaluate this at t equals 0 what you will get is the expectation of x raised to n . So, these expectations are what are called as the moments and that is the reason why $M_x(t)$ is called as a moment generating function.

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If X, Y are two r.v (ind). then $X+Y$ has
 the M.G.P as $M_X(t) \cdot M_Y(t)$.
 $M_{X+Y}(t) = M_X(t) \cdot M_Y(t)$ when X & Y are ind.
 $M_{X+Y}(t) = E[e^{t(X+Y)}] = E[e^{tx} e^{ty}]$
 $= E[e^{tx}] \cdot E[e^{ty}]$ ($E[XY] = E[X]E[Y]$ when X & Y are ind.)
 $= M_X(t) \cdot M_Y(t)$

So, now let us just look at the; let us look at this moment generating functions more closely. If you have two random variables X and Y such that they are independent then, when you look at the random variable X plus Y so, this will have moment generating; if whenever it has a moment generating function ok. So, what will be the moment generating function of X plus Y ok; so, this will be the moment generating function of x times the moment generating function of y ok. In other words the moment generating function of x plus y is equal to $M_x(t)$ times $M_y(t)$ when X and Y are independent not too difficult to see. So, if you look at the random variable X plus Y its moment generating function is nothing, but e raised to t times x plus y and the expectation of this.

So, this is the moment generating function of x plus y . This is going to be equal to expectation of e raised to tx times e raised to ty , e raised to tx and e raised to ty are random variables in their own right. And, those random variables are going to be independent random variables because x and y are independent. When you look at independent random variables and you want to compute the product of these random variables those I mean the expectation of the product that will be just the product of the expectation. So, this will be equal to expectation of e raised to tx times expectation of e raised to ty . So, this is because expectation of xy is equal to expectation of x times expectation of y when x and y are independent ok.

So, this quantity is nothing, but moment generating function of x and this quantity is nothing, but moment generating function of y . So, the moment generating function of x plus y is just the product the individual moment generating functions. Now let us look at the moment generating functions of indicator random variables some particular examples.

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① $X = \begin{cases} 1 & \text{w.p. } \alpha \\ 0 & \text{w.p. } 1-\alpha \end{cases}$ $e^{tx} = \begin{cases} e^t & \text{w.p. } \alpha \\ e^0 & \text{w.p. } (1-\alpha) \end{cases}$

$$M_X(t) = E[e^{tx}] = \alpha \cdot e^t + (1-\alpha) \cdot 1 = 1 + \alpha(e^t - 1) \leq e^{\alpha(e^t - 1)}$$

$X_1, X_2, \dots, X_n$ $\mu_1, \mu_2, \dots, \mu_n$ $E[X_i] = \mu_i$ $E[\bar{X}] = \mu = \mu_1 + \mu_2 + \dots + \mu_n$

$X = X_1 + X_2 + \dots + X_n$

$$M_X(t) = \prod_{i=1}^n M_{X_i}(t) \leq \prod_{i=1}^n e^{\mu_i(e^t - 1)} = e^{\sum_{i=1}^n \mu_i(e^t - 1)} = e^{\mu(e^t - 1)}$$

$M_X(t) = e^{\mu(e^t - 1)}$

Annotations: Poisson Trials, Bernoulli Trials

So, let us say X is a random variable which takes value 1 with probability α and 0 with probability $1 - \alpha$. What will be the expectation of x sorry, what will be the moment generating function of; so, we want to calculate expectation of e^{tx} this is the moment generating function of x ok. So, this is equal to e^{tx} can take 2 values e^{tx} is equal to either e^t this happens with probability α and this is when x is equal to 1 the value of e^{tx} is equal to e^t and this is equal to e^0 this with probability $1 - \alpha$ and that happens when x is equal to 0.

So, the 2 values that it can take are e^t and 1. So, the expectation of e^{tx} is $\alpha e^t + (1 - \alpha) \cdot 1$. So, this can be written as $1 + \alpha(e^t - 1)$ ok. So, this is clearly less than $e^{\alpha(e^t - 1)}$. Now, let us look more closely at the problem that we were interested in; you toss a coin many times what is the probability the number of heads is much larger than the expected

value ok. So, in that the individual toss we will set up as indicator random variables whether a head has appeared or not and we will look at this sum of these ok.

So, that is and we will look at the moment generating function of these independent random variables. So, let us say that we have all these random variables X_1, X_2 up to X_n they are all random variables which takes 0 or 1 value ok. So, this takes value let us say X_1 takes the value 1 with probability μ_1 , X_2 takes the value 1 with probability μ_2 and X_n takes the value n with probability μ_n . So, clearly expectation of X_i is equal to μ_i because this is 0 and random variable their expectation is same as the probability that they take the value 1 and if you look at this random variable X is equal to X_1 plus X_2 plus X_n the moment generating function of X is going to be equal to these are independent random variables. So, your moment generating function is just going to be a product of the individual moment generating functions ok.

So, if $M_{x_i}(t)$ is the individual random variable sorry if $M_i(t)$ is the individual random variable $M_{x_i}(t)$ will refer to the moment generating function of the i th random variable and this is equal to product i going from 1 to n this will be less than or equal to e raised to. So, here $1 + \mu_i t$ is less than $e^{\mu_i t}$ ok. So, here this product is going to be less than $e^{\sum \mu_i t}$ instead of μ_i you have $\mu_i t$ and this is nothing but $e^{\sum \mu_i t}$ will translate into summation in the exponent summation over i μ_i and summation of the individual μ_i s will be equal to the mean of the entire random variable.

So, if you denote expectation of X to be μ which will be equal to $\mu_1 + \mu_2 + \dots + \mu_n$ then this will be equal to $e^{\mu t}$ ok. So, we know that if we look at these random variables X_1 to X_n which takes value 0 or 1 with probability μ_i each when μ_1, μ_2, μ_i then, the moment generating function of the sum of these random variables is going to be less than or equal to $e^{\mu t}$. These random variables have a name they are called as Poisson Trails ok. When all these random variables are identically distributed this is also known as Bernoulli trials ok.

So, in our experiment what we have is, we have these coin tosses or them having the same bias. So, we have identical coins being tossed. So, we have a sum of Bernoulli

trials. So, our random variable is a sum of Bernoulli trials and we know that in case of these kind of trials the moment generating function has to be less than the quantity that we have derived here ok. So, we will use this fact later on, but right now let us just try and see how we to put these things together and obtain what is known as the Chernoff bound ok.

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$$P(X > a) = P(e^{tx} \geq e^{ta}) \quad (\text{where } t > 0)$$

$$P(X > a) \leq \frac{E[e^{tx}]}{e^{ta}}$$

$$P(X > a) \leq \min_{t > 0} \frac{E[e^{tx}]}{e^{ta}} \quad \leftarrow \text{Chernoff Bound}$$

$$P(X < a) = P(tx \geq ta) \quad (\text{where } t < 0)$$

$$= P(e^{tx} \geq e^{ta}) \leq \frac{E[e^{tx}]}{e^{ta}} \min_t$$

So, we were interested in questions of this kind probability that X is greater than a ok. Now this is same as this probability is same as probability that e raised to tx is greater than e raised to ta where t is greater than 0 ok. So, we had an arbitrary random variable x , we have converted that into a positive valued random variable this applies only when t is greater than 0 . And therefore, once we have converted this into a positive valued random variable we can apply Markov inequality. And therefore, we will get this to be less than or equal to expectation of the random variable e raised to tx divided by a , a here being e raised to ta . Now this statement is true for any value of t in particular we can choose that t which minimizes this expression.

So, probability that X is greater than a is going to be less than minimum chosen over all t greater than 0 expectation of e raised to tx divided by e raised to ta . Note that the numerator and denominator are both functions of the variable t ok. So, you can try and optimize depending upon the random variable X and how its distribution is e raised to tx its expectation is going to have some functional form dependent on t and so, numerator

and denominator both has t in them you can try and figure out what is the best value of t , but is the smallest value that you can choose, so as to minimize this expression and even for that value this probability will be less than that ok. So, these kind of bounds are in general known as Chernoff bounds.

Now, this is not a very useful form of Chernoff bound, but we will convert this into a nice form when X is when we know something more about the distribution of x . We could also look at inequalities of the form I mean we want to compute the probabilities that X is less than a particular a now this is going to be equal to the probability that minus let us say we choose a negative number t . So, then this probability is going to be equal to the probability that tx is greater than ta here we will choose where t is less than 0 and this probability is going to be equal to the probability that e raised to tx is greater than or equal to $t a$. Now e raised to tx . So, this is e raised to ta .

Now e raised to tx is going to be always a positive valued random variable and therefore, we can say that this probability is less than or equal to expectation of e raised to tx by e raised to ta . Again we can minimize. So, this is less than minimum over all values of t of this expression ok. So, these bounds are what is called as Chernoff bounds, it can apply both in the case where X is greater than a and X is less than a and for any random variable X ok. So, now, what we will do is we will convert Chernoff bound into a more useful form ok.

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Handwritten notes on a grid background showing Chernoff bound derivations. The notes include the following formulas and steps:

- Top row: $P(X > (1+\delta)\mu)$ and $P(X < (1-\delta)\mu)$. Below the first formula is $\mu + \delta\mu$.
- Middle row: $P(|X - \mu| \geq \delta\mu)$. To the right, a diagonal line separates this from the text "Tail Bounds".
- Below the middle row, the first formula is repeated: $P(X > (1+\delta)\mu) \leq \min_{t \geq 0} \frac{E[e^{tx}]}{e^{(1+\delta)\mu}}$. An arrow points from $E[e^{tx}]$ to the right.
- To the right of the arrow, the formula $P(X > (1+\delta)\mu) \leq \frac{E[e^{tx}]}{e^{(1+\delta)\mu}}$ is written.
- Below this, the formula $X = X_1 + X_2 + \dots + X_n$ is written.
- Below that, the formula $E[e^{tx}] \leq e^{\mu(e^t - 1)}$ is written.
- At the bottom, the final simplified bound is shown: $P(X > (1+\delta)\mu) \leq \left(\frac{e^{\delta}}{1+\delta} \right)^\mu$. To the right of this is the text "RHS".

So, we will try and obtain the Chernoff bound for these kind of events X greater than say $1 + \delta \mu$. So, δ is a parameter that is given to us and μ let us imagine that it is the expectation of the random variable we want to compute the probability that X is greater than $1 + \delta \mu$. So, you expect it to be around μ but, what is the probability that it is greater than μ by an amount $\delta \mu$ ok. So, and we will also look at probability that X is less than $1 - \delta \mu$ and together we will also look at we will combine these two things together and compute the probability that $X - \mu$ is greater than $\delta \mu$ ok.

So, these will be special cases of what known as tail bounds ok. We will do this for the case where, X is a sum of Bernoulli trials or even Poisson trials. So, we know that probability that X is greater than $1 + \delta \mu$ is going to be less than expectation of e^{tX} raised to t divided by $e^{t(1 + \delta \mu)}$ minimized over all values greater than or equal to 0. Now if X is the sum of Poisson trials of n Poisson trials then this quantity on the numerator expectation of e^{tX} we know is surely less than some quantity we had derived this. So, this is the moment generating function of X we had shown that that is going to be less than $e^{\mu(e^t - 1)}$.

So, we will use this result. So, we have this to be less than $e^{\mu(e^t - 1)}$ divided by $e^{t(1 + \delta \mu)}$ ok. So, this is going to be this is equal to $e^{\mu(e^t - 1 - t(1 + \delta \mu))}$ ok. So, what is the value of t that we will choose in order to make this quantity on the LHS sorry the RHS we need to make as small as possible what is the value of t that we can choose.

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$$P(X > (1+\delta)\mu) \leq \left(\frac{e^{(e^t-1)}}{e^{t(1+\delta)}} \right)^\mu$$

$e^t = 1+\delta$ (Choose t satisfying $e^t = 1+\delta$)

$$P(X > (1+\delta)\mu) \leq \left(\frac{e^{1+\delta-1}}{(1+\delta)^{1+\delta}} \right)^\mu$$

$$P(X > (1+\delta)\mu) \leq \left(\frac{e^\delta}{(1+\delta)^{1+\delta}} \right)^\mu$$

$E[X] = \mu/2$

$$P\left(X > \frac{3\mu}{4}\right) = P\left(X > \left(1+\frac{1}{2}\right)^{\frac{\mu}{2}}\right) \leq \left(\frac{e^{\frac{1}{2}}}{1.5^{1.5}} \right)^{\frac{\mu}{2}} = \left(\frac{\sqrt{e}}{1.5^{1.5}} \right)^{\frac{\mu}{2}}$$

So, probability that X is greater than 1 plus δ times μ shown this is equal this is less than or equal to e raised to e raised to t minus 1 divided by e raised to 1 plus δ the whole raised to sorry there is a t here e raised to the expectation is this quantity and this probability that the probability that X is greater than 1 plus δ times μ is going to be less than this quantity ok

This is because this probability that X is greater than 1 plus δ times μ is going to be less than expectation of e raised to tx divided by e raised to t times 1 plus δ times μ . So, that is how the t comes in the denominator ok. So, this is our expression we need to choose a value of t ok. So, any positive value of t would be fine. So, $1/t$ which can simplify things is choose t equals. So, let us say we choose e raised to t to be equal to 1 plus δ ok. So, this quantity e raised to t we will choose it to be 1 plus δ clearly. So, choose the t satisfying this expression and clearly that t is going to be positive t is 0 you will get 1 . So, little greater than 0 you will get 1 plus δ as δ is δ some positive number ok.

So, now this so, if we had chosen this as our value of t we know that probability that X is greater than or greater than 1 plus δ times μ is going to be less than or equal to e raised to 1 plus δ minus 1 divided by 1 plus δ the whole raised to 1 plus δ the whole raised to μ and this is equal to e raised to δ by 1 plus δ the whole raised

to $1 + \delta$ the whole raised to the power μ . So, this is the probability that X is greater than $1 + \delta$ times μ .

Now, let us look at this expression closely in our scenario, we had X to be the sum of n i.i.d random variables and the expectation of X was $n/2$. We wanted X to be greater than $3n/4$ and we wanted to determine the probability of this probability that X is greater than $3n/4$. Now in this case, δ would be $1/2$. So, probability so, we were interested in probability that X is greater than $1 + 1/2$ times the expectation $n/2$ and this by Chernoff bound is going to be less than or equal to $e^{-\delta^2 \mu}$ here δ is $1/2$ divided by 1.5 the whole raised to 1.5 the whole raised to μ and μ is going to be $n/2$ ok. So, this is equal to so, so, this is $\sqrt{e}$ divided by $1.5^{n/2}$ sorry 1.5 raised to 1.5 you can verify that this quantity is less than 1.

Computation will tell you that this quantity is less than 1 and therefore, the probability of this happening is some quantity less than 1 the whole raised to $n/2$. So, when n was something like 4 million this is some quantity less than 1 the whole to the power 2 million. So, it is a much smaller number than whatever we had computed before $1/1$ million ok. Chebychev bound gave us the probability to be less than $1/1$ million, but here were getting a probability some number which is less than 1 the whole raised to 2 million ok. Now this number is it really a small number is it very close to 1 ok. So, in some sense this formula is not really easy to work with does not really tell us how far is this away from 1, if this was let us say half and we readily know if this is less than half we readily know that this is a this gives a very good bound ok. So, let us try and convert this into a more workable form ok.

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Handwritten mathematical derivation on a grid background:

$$P(X < (1+\delta)n) < (q_n)^n$$

$$\frac{e^\delta}{(1+\delta)^{1+\delta}} = \frac{e^\delta}{e^{\ln(1+\delta) \cdot (1+\delta)}} = \frac{1}{e^{(1+\delta)\ln(1+\delta) - \delta}}$$

Let $T = (1+\delta)\ln(1+\delta) - \delta$

Define functions f_1 and f_2 :

$$f_1 = (1+\delta)\ln(1+\delta)$$

$$f_2 = \delta$$

At $\delta = 0$, $f_1 = 0$ and $f_2 = 0$.

Derivatives:

$$f_1' = 1 + \ln(1+\delta)$$

$$f_2' = 1$$

Since $f_1' > f_2'$ for $\delta > 0$, f_1 grows faster than f_2 .

Graph showing f_1 (upper curve) and f_2 (lower curve) starting from the origin.

So, we have this probability to be e raised to δ divided by $1 + \delta$ the whole raised to $1 + \delta$. This is the same as e raised to δ divided by e raised to $\log(1 + \delta)$ times $1 + \delta$. And that we can write it as 1 by e raised to $1 + \delta \log(1 + \delta) - \delta$. So, if we can show that the exponent here, let us call that as T . If T is greater than 0 for any value of δ , then we have some number which is strictly less than 1 . So, this quantity, this entire quantity, will be less than 1 . So, this will be less than 1 ok; so, will try to show that. So, if you look at this expression $1 + \delta \log(1 + \delta) - \delta$, you can plot the curves of these.

So, this is a curve δ and $1 + \delta \log(1 + \delta) - \delta$. If that curve lies strictly above this curve, then we know that $1 + \delta \log(1 + \delta) - \delta$ is going to be always positive. How do we show that? At 0 , the function f_1 and the function f_2 are both 0 because this is $1 + 0 \log(1)$ and this is 0 . So, $\log(1)$ is 0 . So, both of them are 0 . So, f_1 and f_2 agree at 0 and if you look at the derivative of f_2 , that is so f_2' is 1 . f_1' is going to be $1 + \log(1 + \delta)$. f_1' is going to be $1 + \log(1 + \delta)$ which is greater than 1 . So, f_1 as a function which grows faster than f_2 and they agree at 0 . So, clearly, this is going to lie above the curve of f_2 . So, this is a curve of f_1 , f_2 , and this is the curve of $f_1 - f_2$.

So, this difference is always going to be positive. Since it is always going to be a positive number, we can write this is $1 - f(\delta)$ where $f(\delta)$ is going to be

always greater than 1 ok. So, this entire this tells us that this is going to be some f let us say. So, the probability of success or the probability that X is less than 1 plus delta times μ is going to be always less than some let us say g delta raised to μ and g delta is a number less than 1. Now we will determine the form of g delta.

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$$P = \frac{1}{e^{(1+\delta)\ln(1+\delta)} - \delta} = \frac{e^\delta}{(1+\delta)^{1+\delta}} = \frac{e}{2^2} = \frac{e}{4}$$

$$(1+\delta)\ln(1+\delta) = (1+\delta) \left[\delta - \frac{\delta^2}{2} + \frac{\delta^3}{3} - \frac{\delta^4}{4} + \dots \right]$$

$$= \delta - \frac{\delta^2}{2} + \frac{\delta^3}{3} - \frac{\delta^4}{4} + \dots + \delta^2 - \frac{\delta^3}{2} + \frac{\delta^4}{3} - \dots$$

$$= \delta + \delta^2 \left(\frac{1}{1} - \frac{1}{2} \right) - \delta^3 \left(\frac{1}{2} - \frac{1}{3} \right) + \delta^4 \left(\frac{1}{3} - \frac{1}{4} \right) - \dots$$

$$= \delta + \delta^2 \left(\frac{1}{2} \right) - \frac{\delta^3}{6} + \frac{\delta^4}{12} + \dots$$

$$\tau = \boxed{(1+\delta)\ln(1+\delta) - \delta} = \frac{\delta^2}{1 \cdot 2} - \frac{\delta^3}{2 \cdot 3} + \frac{\delta^4}{3 \cdot 4} - \frac{\delta^5}{4 \cdot 5} + \dots$$

So, we have this expression 1 by e raised to 1 plus delta log 1 plus delta minus delta. So, this we can write it is. So, if you look at 1 plus delta into log 1 plus delta if you just expand this is going to be 1 plus delta into log 1 plus 1 plus delta we can write it as delta by 1 minus delta square by 2 plus delta cube by 3 minus delta raised to 4 by 4 and so on ok. So, this if we just multiply the terms delta minus delta square by 2 plus delta cube by 3 minus delta raised to 4 by 4 and so on, when you multiply the delta you will get plus delta square minus delta cube by 2 plus delta raised to 4 by 3 and so on. So, this is going to be equal to delta plus this is going to be delta square into 1 by 1 minus 1 by 2 minus delta cube into 1 by 2 minus 1 by 3 plus delta raised to 4 into 1 by 3 minus 1 by 4 and so on ok.

Clearly this is going to be equal to delta plus delta square into 1 by 1 into 2 minus delta cube by 2 into 3 plus delta raised to 4 by 4 and so on and what we were interested in this 1 plus delta into log 1 plus delta minus delta. So, this term goes away. So, this is going to be equal to delta square by 1 into 2 minus delta cube by 2 into 3 plus delta raised to 4 by

3 into 4 minus delta raised to 5 into 4 into 5 and so on ok. How large is this is the exponent of this term ok. So, if we call this as T how large is T ok.

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$$\frac{\delta^2}{3} \leq T \leq \frac{\delta^2}{2} \quad \text{when } 0 \leq \delta \leq 1$$

$$T = \frac{\delta^2}{1 \cdot 2} - \frac{\delta^3}{2 \cdot 3} + \frac{\delta^4}{3 \cdot 4} - \frac{\delta^5}{4 \cdot 5} + \dots$$

$$T > \frac{\delta^2}{1 \cdot 2} - \frac{\delta^3}{2 \cdot 3}$$

$$= \frac{\delta^2}{3} \left(\frac{3}{2} - \frac{\delta}{2} \right)$$

$$= \frac{\delta^2}{3} \left(\frac{3 - \delta}{2} \right)$$

$$\frac{-\delta^n}{(n-1)n} + \frac{\delta^{n+1}}{n(n+1)} = \frac{\delta^n}{n} \left[\frac{-1}{n-1} + \frac{\delta}{n+1} \right]$$

$$\frac{-1}{n-1} + \frac{\delta}{n+1} < 0$$

$$\frac{\delta^2}{3} < T < \frac{\delta^2}{2}$$

We will show that T lies between delta square by 2 and delta square by 3 [FL]. So, we will show that T lies between delta square by 2 and delta cube by 3 ok. So, T if we know is equal to delta square by 2 by 1 into 2 minus delta cube by 2 into 3 plus delta raised to 4 by 3 into 4 minus delta raised to 5 by 4 into 5 and so on. Now, if you look at, if you combine any 2 of 2 terms if you look at delta r raised to n the first term is going to be negative term delta raised to n divided by n minus 1 into n, the next term is going to be a positive term delta raised to n plus 1 divided by n into n plus 1.

We can take what is common delta raised to n divided by n and then what is remaining is minus 1 by n minus 1 plus delta by n plus 1. The denominator is larger and the numerator is smaller because delta we may assume is between 0 and 1, if delta is between 0 and 1 when delta is between 0 and 1 these sum of terms taken 2 at a time this is going to be a negative number because, this is smaller than this quantity if you call it as b and this as a minus a plus b. So, this if you call this entire thing as a then a plus b is going to be less than 0 because 1 is greater than delta n minus 1 is smaller than n plus 1.

So, the sum of all these terms all these are all these terms are going to be negative terms and their sum is going to be negative. So, T is going to be at most delta square divided by 1 into 2. Now if you take the summation in a slightly different way that is if you take

these two together and the next two together and the next two together you can see that each of these terms are going to be positive ok. So, T so, therefore, T is going to be greater than delta square by 1 into 2 minus delta cube by 2 into 3 and this is equal to delta square by 3 into 3 by 2 minus delta by 2. So, this is equal to delta square by delta square by 3 into 3 minus delta by 2 the maximum value that delta can take. So, this delta the maximum value that it can take is 1 and even in that case this expression is going to be 1. So, this is no less than 1. So, the entire value of T will lie between. So, that T should surely lie between delta square by 3 and delta raised to 2 by 2.

Now, let us look at this expression slightly more carefully. We did this analysis for delta lying between 0 and 1. But $1 + \delta \log 1 + \delta - \delta^2$ this is of course, always positive. Since it is always positive, the value of this expression at 1 is going to be I mean whatever is the value of the expression at 1 that is going to be smaller than the values at all higher delta ok. So, if delta is greater than 1 greater than or equal to 1 then this expression p, p is going to be less than the value at 1 which is going to be $1 + \delta \log 2 - 1$ ok.

We could have just simply written this as e raised to delta divided by $1 + \delta$ the whole raised to $1 + \delta$ and delta equals 1 this is e divided by 2 square. So, that is e by 4 it is less than 3 by 4 ok. So, when delta is greater than 1 we can readily substitute this by 3 by 4.

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$$\text{Thm: For any } 0 < \delta < 1, P(X > (1+\delta)\mu) \leq \left(e^{-\frac{\delta^2}{3}}\right)^\mu$$

$$= e^{-\frac{\mu \delta^2}{3}}$$

$$\text{If } \delta > 1, \text{ then } P(X > (1+\delta)\mu) \leq \left(\frac{3}{4}\right)^\mu$$

$$\text{Thm: For any } 0 < \delta < 1, P(X < (1-\delta)\mu) \leq e^{-\frac{\delta^2 \mu}{2}}$$

$$\text{Thm: } P(|X - \mu| > \mu \delta) \leq 2 \cdot e^{-\frac{\mu \delta^2}{3}}$$

$$\ln(1+\delta)$$

$$\ln(1-\delta)$$

So, $3/4$ raised to μ and when δ is less than 1 we have so, when δ is less than 1 we have the following theorem. So, for any δ between 0 and 1 probability that X is greater than $1 + \delta \mu$ is going to be less than $e^{-\delta^2 \mu / 3}$ the whole raised to μ ok.

So, this is $e^{-\delta^2 \mu / 3}$ ok. So, this is a form in which we are going to use Chernoff bound. When δ is greater than 1 of course, so, if δ is greater than 1 then probability that X is greater than $1 + \delta \mu$ is less than $3/4$ raised to μ ok. We can also derive a similar bound for any δ between 0 and 1 probability that X is less than $1 - \delta \mu$ this we can show is less than or equal to $e^{-\delta^2 \mu / 2}$ times μ . Now here, if you assume that X is a positive valued random variable δ cannot I mean there is no sense in choosing a δ less than 1 I mean let us say greater than 1 if δ is greater than 1 then $1 - \delta$ is going to be let us say a negative number ok.

So, we can combine these two theorems and say probability that $X - \mu$ is greater than $\mu \delta$ is going to be less than $2 e^{-\delta^2 \mu / 3}$ ok, we take the smaller of these quantities or smaller or larger. So, this probability is going to be the sum of these two probabilities and $\mu \delta^2 / 3$ is going to be a smaller quantity than this. So, we can say that this probability is less than twice; this twice $e^{-\delta^2 \mu / 3}$ ok. So, when we try to prove this statement, we just have to retrace the steps of the previous theorem. But, instead of $\log(1 + \delta)$ has expansion we will have to work with $\log(1 - \delta)$ expansion and then everything else is the same and we will get this as the bound ok.

We will stop here for the time being. This concludes our discussion on Chernoff bounds.