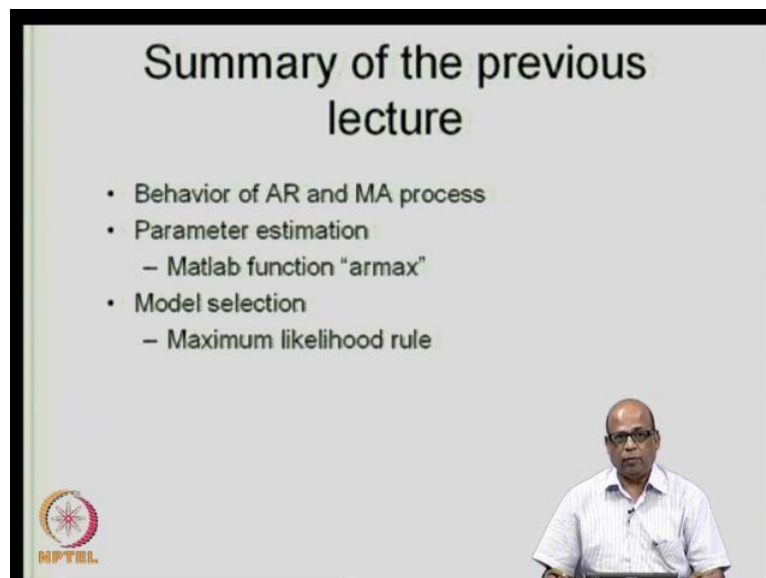


Stochastic Hydrology
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Lecture No. # 17
ARIMA Models – IV

Good morning and welcome to this, the lecture number 17 of the course Stochastic Hydrology.

(Refer Slide Time: 00:25)



If you recall in the last lecture, we discussed the ARIMA models and specifically, how the AR and MA components of the model behave. For example, we considered AR 1 model, formulated a theoretical AR 1 model and saw how the correlogram of the theoretical AR 1 model behaves and how the spectrum behaves and how the PAC function behaves that is the Partial Auto Correlation function behaves and then we also looked at two examples of AR 2 models and similarly, for MA 1 process we saw how the autocorrelation, partial autocorrelation function and the spectral density function behave.

So, we essentially identify the number of AR terms and the number of MA terms in a AR ARIMA model by looking at the PAC function as well as the correlogram. But **I** as I said

in the last lecture, when you have both AR and MA terms in the model and they are slightly of higher order, let us say that you are talking about ARIMA 4, 5 ARIMA 4, 2; 4, 1, these kind of models where the number of AR terms is quite large and the MA terms are also reasonably large in number, then identification becomes quite difficult.

And therefore, what we do is, we formulate candidate models, several of candidate models and then for the candidate models, we examine, we estimate the parameters and then examine whether the model is valid or not, then we also examined in the last lecture, the procedure for parameter estimation. As I mentioned, there are several algorithms available for parameter estimation and we introduced the mat lab function ARMAX, which can be readily used and the **the** parameters for any given ARIMA type of model can be estimated.

We, towards end of the lecture, we examined the maximum likelihood criteria for selection of the models. The specific problem is, that we have a number of candidate models, we have estimated all the parameters for the models based on the data, please do not lose sight of the hydrologic aspects of this, we have observed data, for example the stream flow at a particular location. And then on this observed data, we are doing all of these exercise so that we can fit a model for the time series that has been observed.

So, once we formulate the candidate models and estimate the parameters for each of these models, then we have to choose which among these candidate models is best suited for the observed data, this we do by two methods as I mentioned in the last lecture, one is the maximum likelihood criterion which we use for, when your application is for long term simulation of the data, then we use the maximum likelihood criterion.

What we did in that, we defined a likelihood function, log likelihood function to be precise and estimate the log likelihood function value for each of the models, so each of the model will have one log likelihood value, then pick up that particular model which results in the maximum value of the log likelihood function value.

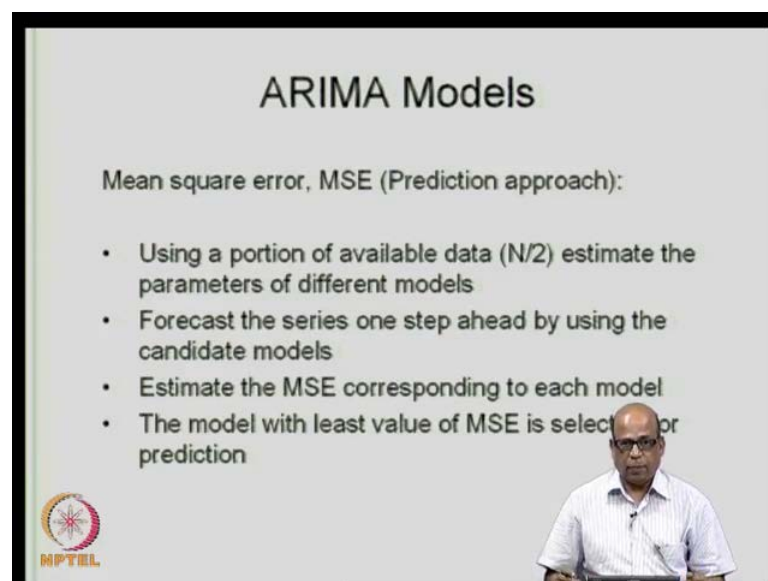
Now, we will go to the second part of the applications of the model where the models are meant not so much for a long term synthetic data **data** generation, but for **one time real time**, one time forecasting models that means, one time step ahead forecasting models, like I mentioned, these applications will be, for example monthly forecasting of the stream flows which may be useful for your real time operation of the reservoirs, where

you are operating the reservoir for irrigation, hydropower, etcetera depending on the storage level and depending on the forecast of the flows, you would like to operate the gates.

This also becomes important when we are looking at smaller time periods like 10 day time periods for real time irrigation scheduling and so on, when we look at the applications, these points will be clear. But essentially we use the time series models to develop forecasting develop forecast for the reservoir in flows, evapotranspiration, rainfall and so on, so forth.

So, when our applications are for short term forecasting, then the maximum likelihood criterion is not the criterion that we use, we must then go for the minimum mean square error criterion. Now, that is what we will discuss in the lecture today, how we develop the minimum mean square error values for the models and then pick the best model that gives the minimum mean square error.

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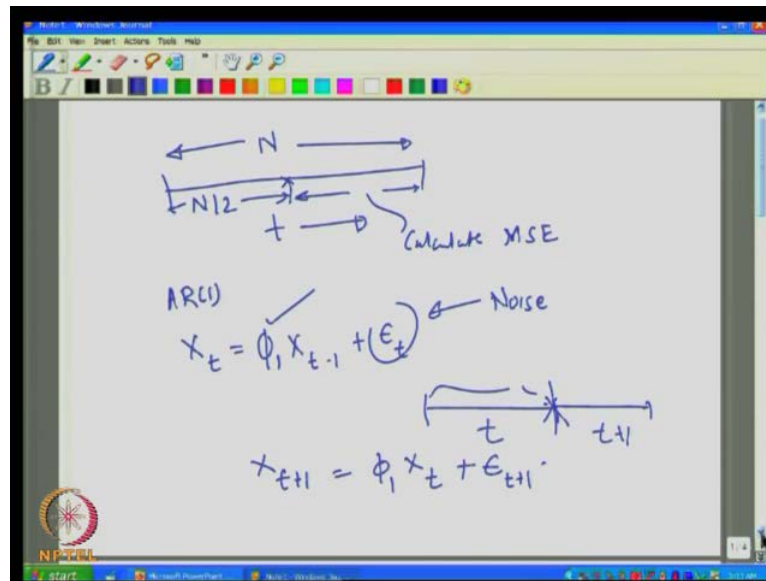
The slide is titled "ARIMA Models" and discusses the Mean square error, MSE (Prediction approach). It lists four steps for model selection:

- Using a portion of available data ($N/2$) estimate the parameters of different models
- Forecast the series one step ahead by using the candidate models
- Estimate the MSE corresponding to each model
- The model with least value of MSE is selected for prediction

The slide also features the NPTEL logo in the bottom left corner and a small inset image of a man in a white shirt and glasses in the bottom right corner.

So, the minimum mean square error criterion or the MSE in short, which is also called as a prediction approach, what we do in this is that, let us say you have 50 years of data, monthly stream flow data, then we use the first 50 first 20, 25 years of data to formulate the mean square error.

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Let us look at this sketch here, let us say that your time is on this scale and you have N values of the time series, we use the first N by 2 values typically for developing the model and then the remaining N by 2, we use for calculating the MSE. So, we calculate the mean square error with the remaining N by 2 values.

What do I mean by that; let us say that you have fit a AR 1 type of model for using this particular data. When I say you have fit the model, it means that the parameters have been estimated for this particular model. How do I write this? We write this as X_t is equal to in our notation $\phi_1 X_{t-1}$ plus epsilon t . Now, this is a model that we **we** have written and we have estimated the parameter ϕ_1 using one of the algorithms.

Now, when we want to do, use this model for forecasting, what does the forecast mean, remember that this is a noise term here or a residual term, which has a zero mean and they are all uncorrelated, the epsilon t are all uncorrelated. When we want to use this particular model for forecasting purpose, the forecast is the expected value for the next time period, let us say that you are standing in time period t and you want to forecast for time period t plus 1. So, you are here now, and you know the information that is available up to this particular point, and you want to forecast for X_{t+1} . Let us say that I write X_{t+1} is equal to $\phi_1 X_t$ plus epsilon t plus 1, I am writing it for t plus 1 now.

Then for forecasting, whenever we say we want to forecast this using this particular model, what is it that we are doing, we want to get the expected value of X_{t+1} given the entire history up to time period t , that is a problem, so we want to see, what is the expected value of the flow for time period $t+1$ given the value until X_t .

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$$\hat{X}_{t+1} = E[X_{t+1}] = \phi_1 E[X_t] + \underline{E[\epsilon_{t+1}]}$$

$$\hat{X}_{t+1} = \phi_1 E[X_t] + 0 = \phi_1 E[X_t]$$

ARMA(1,1)

$\begin{array}{|c|c|} \hline t & t+1 \\ \hline \end{array}$

$$X_{t+1} = \phi_1 X_t + \theta \epsilon_t + \epsilon_{t+1}$$

Now, to do this, I will take the expected value of X_{t+1} now, let us say I write \hat{X}_{t+1} expected value of X_{t+1} . So, I am taking the expected value of this particular expression, is equal to ϕ_1 into expected value of X_t plus expected value of ϵ_{t+1} . Now, this is the forecast, and we generally denote the forecast by putting a cap on that particular variable. So, \hat{X}_{t+1} which is the forecast for the time period $t+1$ is equal to ϕ_1 expected value of X_t plus expected value of ϵ_{t+1} .

So, I will write expect the forecast \hat{X}_{t+1} is equal to ϕ_1 , expected value of X_t plus, what is this? This has the mean of 0, the ϵ_t remember is a sequence which has a mean of 0. So, when you take the expected value of ϵ_t or ϵ_{t+1} that is 0 here, so this is simply ϕ_1 expected value of X_t .

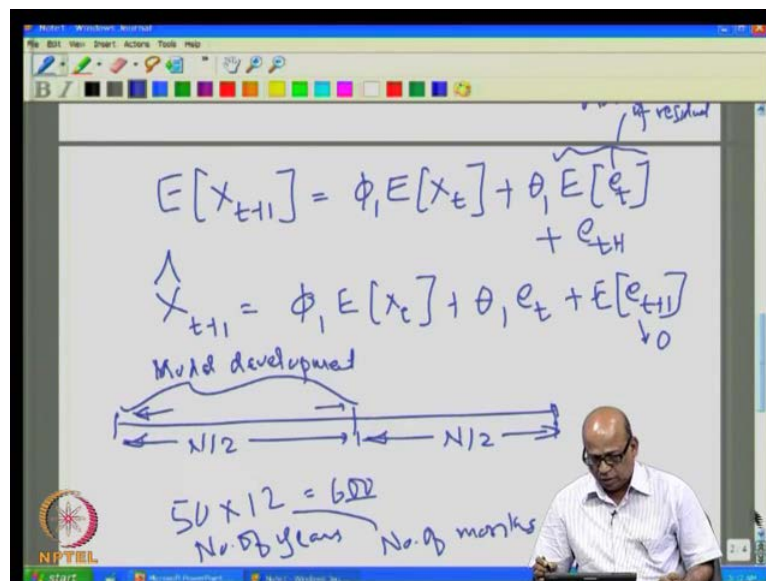
Similarly, you can look at any models, let us say that you want to take ARMA 1, 1 model here, and then use it for forecasting. So, ARMA 1, 1 model we write it as X_{t+1} . Let us say again we look at two time periods, two adjacent time periods, time t , time $t+1$, you are here now and you want to forecast for the next time period. To keep the relevance with hydrology, let us say that you are standing in the month of June and you

have the flow information up to the month of June. And you would like to forecast what is likely to happen to the next month that is the flow during the next month, if you are looking at the forecast for flows.

So, we will write this as X_{t+1} is equal to expected value of **I am sorry** I will may be I will write first the model here. So, ARMA 1, 1 model we will write first, we write ARMA 1, 1 model as for X_{t+1} is equal to I have one AR term. So, I will write this as $\phi_1 X_t$, I have one MA term. So, I will write it this as $\theta_1 e_t + e_{t+1}$, because I am writing for time period $t+1$.

Now, you would like to forecast, use this model for forecasting. As I said, what do we mean by forecast? Forecast is the expected value of the particular variable, stream flow in this particular case for example, for the next time period $t+1$.

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So, we take the expected value of X_{t+1} , I will write that as expected value of X_{t+1} is equal to ϕ_1 expected value of X_t plus θ_1 expected value of e_t plus e_{t+1} .

Remember here, when we are writing a ARMA 1, 1 model, the e_t term here is the residual arising out of the application of this model in the time period t . So, when you write this expected value of e_t in this particular term, the e_t value is a constant. It is not the expected value of $t+1$ that becomes 0. For example, we are may be talking about

expected value of five units and therefore, this becomes e_t itself. So, I will write this as X_{t+1} is equal to ϕ_1 expected value of X_t plus $\theta_1 e_t$ plus expected value of $X_{t+1} e_t$ which is 0.

What is a difference here? What I did is that, this is an actual value of residual that resulted from this, the this is actual value of residual that results from applying this particular model for the time period t (Refer Slide Time: 13:58). So, there was a residual, there was an actual value that was available and then the **the** forecasted value was available. So, the actual value minus forecasted value gave me this e_t value and therefore, when I take the expected value, it becomes e_t itself. So, I will write this as ϕ_1 expected value of X_t plus $\theta_1 e_t$ and this becomes 0, because these are all noise terms and therefore, you write the forecasted value for time period $t+1$.

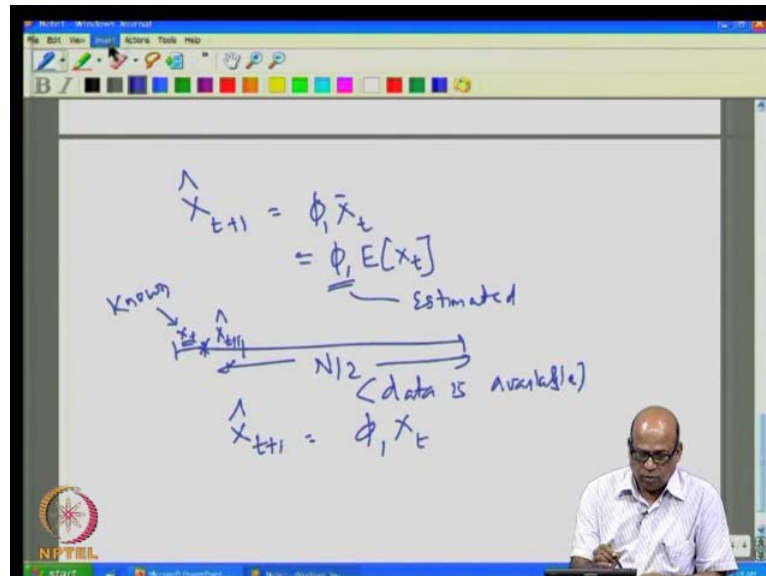
Essentially then, we are getting the forecasted values and applying for the first $N \times 2$ terms that you have, as I mentioned you have $N \times 2$ here, number of values and you have another $N \times 2$ number of values (Refer Slide Time: 15:00). So, what we did, we have developed the model using this $N \times 2$ number of values, this is for model development, by model development I mean we used all of these values to obtain the parameters of the candidate models.

We had let us say for the forecasting, we had the models AR 1, AR 2, ARMA 1, 1, ARMA 1, 2 and so on. As I mentioned in the last class, when we are talking about the forecasting models, typically the lower order models in terms of the number of parameters will suffice. In fact, when we see in the applications most of one, the **AR 1** or AR 1, AR 2 model themselves will be sufficient for one type step away forecasting, especially when we are talking about smoothen processes like monthly stream flow, seasonal flow flows and so on.

So, we choose the candidate models which may be different from the candidate models that you would have chosen for long term synthetic generation of the data. We chose that and then developed the model using the first $N \times 2$ value, let us say if you have 50 years of values, you would have 50 into 12 that is 600 values is the data, this is number of years and this is number of months (Refer Slide Time: 16:35). So, you would have 600 values, so use the first 25 years data which is $N \times 2$ to develop the model and then calculate the errors of using this model on one time step ahead forecasting for the

remaining data, remaining part. What do I mean by that; let us say that you chose your AR 1 model.

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So, in the AR 1 model, you would have written your expected value of X_t is equal to ϕ_1 into expected value of X_t that is forecast for time period t plus 1, you would have written as ϕ_1 into X_t bar, or is equal to ϕ_1 expected value of X_t .

So, this ϕ_1 , you would estimate based on the first $N/2$ values of the data and then start applying this, when you start applying this, let us say that you are now applying for the remaining $N/2$ values and this is one time step ahead focus. So, given X_t now, you had this X_t value known, you want to apply this. So, I will get X_{t+1} cap, this is known, X_t is known, typically let us say you are standing at the end of one month, June month. The flow during the month of June is known, you want to forecast the flow during the month **June** July.

Then we apply this, ϕ_1 is known, ϕ_1 is estimated already. So, we apply this as X_{t+1} is equal to ϕ_1 into expected value of X_t that becomes X_t itself, because it is known. So, this you get X_{t+1} bar, but because you have the data already for this $N/2$, this data is available. What we do, this is the forecasted value and X_t is the actual known value of the flow. Let us say that you have X_t available with you, so I will **I will** take the error **error** of forecast.

(Refer Slide Time: 19:27)

The image shows a whiteboard with handwritten mathematical expressions and a diagram. At the top, it defines X_{t+1} as a "Known data value" and \hat{X}_{t+1} as a "Forecasted Value". Below this, the error term is defined as $e_{t+1} = X_{t+1} - \hat{X}_{t+1}$. A diagram shows a horizontal timeline with a vertical line at time $t+1$. To the left of this line, the value X_{t+1} is labeled as "known". To the right, the forecasted value \hat{X}_{t+1} is shown. The error e_{t+1} is indicated as the difference between the actual value and the forecasted value at time $t+1$. Below the diagram, the forecast for time $t+2$ is given as $\hat{X}_{t+2} = \phi_1 X_{t+1}$, where X_{t+1} is the known value from the previous period.

So, what will be the error of forecast, X_{t+1} is your known data, I will say this is known data and \hat{X}_{t+1} is your forecasted value (No audio from 19:40 to 19:46), let us say known data value. So, the error I will write for the time period $t+1$ is equal to X_{t+1} minus \hat{X}_{t+1} , this is what is forecasted from the model, so like this I get error.

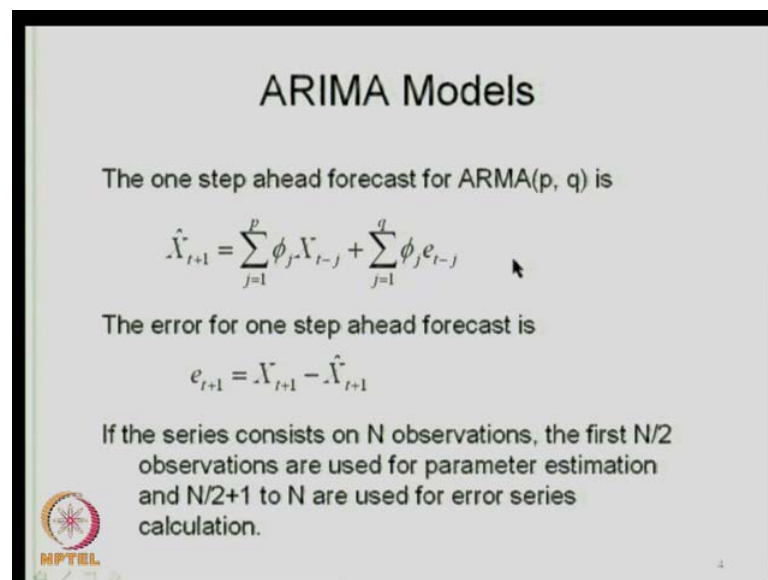
Now, we are writing for the remaining $N-2$ time periods. So, next time when I go, let us say that I want to write for X_{t+2} now. So, I finish my forecasting problem for $t+1$. So, I come to $t+2$, when I am writing for the next time period, we use the actual value that is known, X_{t+1} to get for X_{t+2} the forecast, not the forecasted value, remember. Because this data is already available, and the question that we are asking is standing at this point, knowing the value up to this point, what is my forecast for the next time period?

So, X_{t+2} , I will write this as $\phi_1 X_{t+1}$, the expected value of X_{t+1} which becomes a X_{t+1} itself, because this is known, this is known here plus expected value of the error term which is 0. So, I will simply get X_{t+1} based on this, X_{t+2} based on this (Refer Slide Time: 21:00). Then again calculate the error corresponding to X_{t+2} , because the X_{t+2} forecast, because you know the X_{t+2} value itself. And therefore, you get errors of $t+1$, $t+2$, the errors corresponding to $t+1$ to

plus 2 and so on. So, you formulate the error sequence, using the error sequence, then you can get the mean square error.

So, this is the procedure that we use to obtain the errors of forecast. So, the same thing is summarized here, using a portion of available data which is typically $N/2$, estimate the parameters of different candidate models. Use these forecasted models so developed, all the candidate models one by one to get the series of forecast one time step ahead by using the candidate models. And corresponding to each of these forecasts, **you know** the error term, get the mean square error and then pick up that particular model which results in the minimum mean square error.

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The slide titled "ARIMA Models" contains the following text and equations:

The one step ahead forecast for ARMA(p, q) is

$$\hat{X}_{t+1} = \sum_{j=1}^p \phi_j X_{t-j} + \sum_{j=1}^q \theta_j e_{t-j}$$

The error for one step ahead forecast is

$$e_{t+1} = X_{t+1} - \hat{X}_{t+1}$$

If the series consists on N observations, the first $N/2$ observations are used for parameter estimation and $N/2+1$ to N are used for error series calculation.

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It is written more formally like this, so once time step ahead forecast for ARIMA p q model now can be written as \hat{X}_{t+1} , you have p of AR terms $\sum_{j=1}^p \phi_j X_{t-j}$ plus you have q of moving average terms $\sum_{j=1}^q \theta_j e_{t-j}$, the noise term has mean 0, therefore it when you take the expected value, it vanishes and then the error for one time step ahead focus is e_{t+1} is equal to X_{t+1} which is the known value minus the forecasted value.

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ARIMA Models

The MSE for model is

$$MSE = \frac{\sum_{i=\frac{N}{2}+1}^N e_i^2}{N/2}$$

MPTEL

Once you know e_{t+1} , then you get the mean square error, mean square error is because you have used $N/2$ number of values, you sum over sum over all these errors, the squares of errors and then divide by $N/2$, you get the mean square error. And choose that particular model among the candidate models, among the number of candidate models to pick the particular model which results in the minimum mean square error.

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ARIMA Models

3. Model testing / Validation:

Calibration data Test data

X_1 X_{T-2} X_{T-1} X_T X_{T+1} X_N

First 'T' values are used to build the model (say 50% of the available data) and the rest of data is used to validate the model.

All the tests are carried out on the residual

MPTEL

Now, what did we do, whether it is for long term simulation of the data or it is for short **short** term, one time step ahead forecasting, we use the models ARIMA models and integration we use for differencing, so typically we difference the series first and then apply the ARIMA models. So, whenever I say ARIMA models, the integration is **applied** implied that is the differencing is already done and on that we are applying the ARIMA models.

So, in either case, whether it is being used for long term sequence synthetic generation or for short term duration of the data, we use part of the data for calibration and parameter estimation, the remaining part we are using it for validation type. Let us say you are using it for long term simulation and it is of ARIMA type of model, ARIMA 1, 1 or ARIMA 2, 1 etcetera. So, there is also a MA **MA** model available, MA term that is available.

Then what we do, that we applied for X_{t+1} , get X_{t+1} , you already have a known X_{t+1} . So, the residue term e_{t+1} is got from the available data minus the model data, same principle as we did for forecasting. So, you generate corresponding to this term, the residual series e_t series.

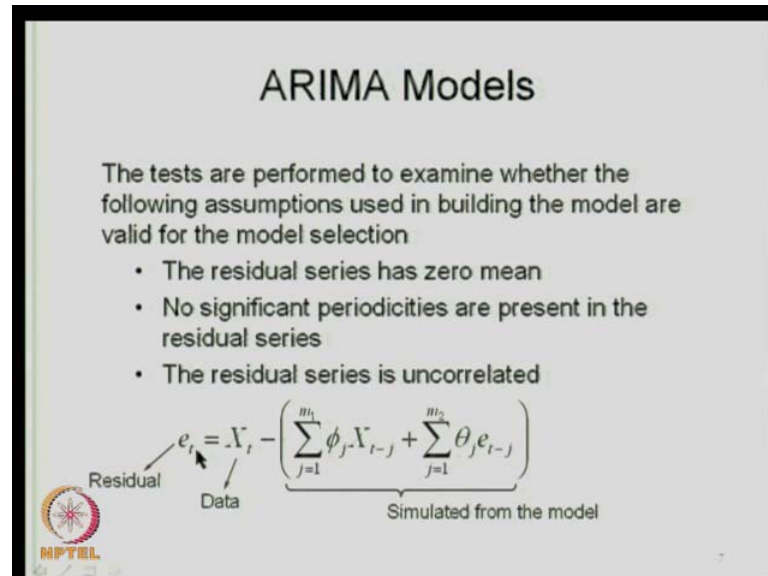
Similarly, if you are using it for the forecast, the error of the forecast that we just obtained there, that error term becomes a residual series. So, for the test data, you have generated e_{t+1} , e_{t+2} , e_{t+3} , etcetera up to all the value, all the remaining values. So, you have essentially the residual series available with you, after you apply the model.

Now, for the validation of the model, we test this residual series that you so obtained. That means, you applied the model for the remaining $N-2$ values or whatever number of values that you have chosen for validation, corresponding to each of the term you obtain either the residual or the forecast error. So, e_t sequence is known, on this e_t sequence, now we carry out all the tests.

What are the tests? The test if you **if you** recall, when we formulated these models, we wrote the residual or the noise term e_t and said that the noise term should have a zero mean that is the first assumption of the model, next that it should be divide of periodicities and it should be uncorrelated. So, we do three primary test, primarily three test we do, one is to test the series that we have generated, namely e_t series. This series

has a zero mean and that the series is divide of any periodicities and also that the series is uncorrelated.

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The tests are performed to examine whether the following assumptions used in building the model are valid for the model selection

- The residual series has zero mean
- No significant periodicities are present in the residual series
- The residual series is uncorrelated

$$e_t = X_t - \left(\sum_{j=1}^{n_1} \phi_j X_{t-j} + \sum_{j=1}^{n_2} \theta_j e_{t-j} \right)$$

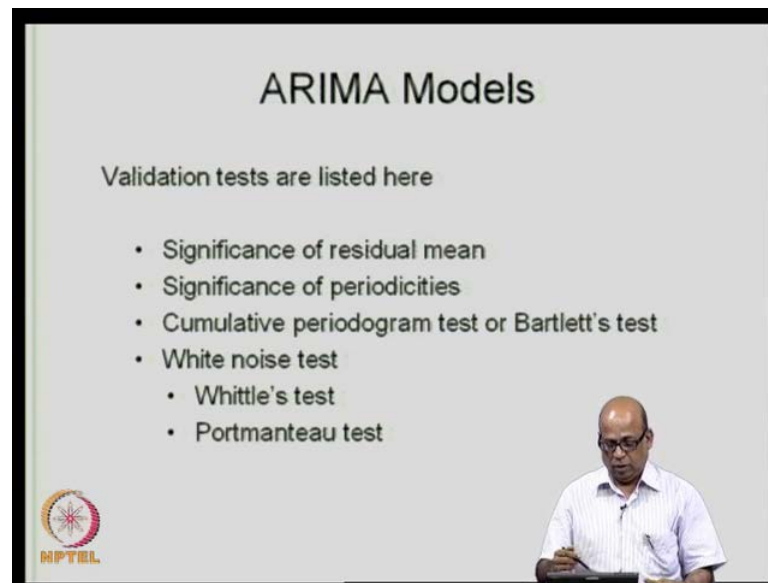
Residual Data Simulated from the model

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So, we perform the test, so examine whether these assumptions that we have made that is the residual series has a zero mean, no significant periodicities are present in the residual series and that the residual series is uncorrelated.

How do we formulate the residual series? As I said, this is e t is the residual is equal to X t which is a known data, minus you have this term corresponding to the simulated data from the model. So, this is phi j X t minus j plus theta j e t minus j. So, this is how you calculate the e t term, so you have the sequence of e t's now.

(Refer Slide Time: 27:51)



The slide is titled "ARIMA Models" and contains the following text:

Validation tests are listed here

- Significance of residual mean
- Significance of periodicities
- Cumulative periodogram test or Bartlett's test
- White noise test
 - Whittle's test
 - Portmanteau test

In the bottom right corner, there is a small inset image of a man in a white shirt and glasses, likely the presenter, looking at a document. In the bottom left corner, there is a logo for "MPTEL" with a circular emblem.

On this sequence of e_t , we do the following validation test, one is significance of the residual mean that, what **what** is it that we want to test here that, the mean of the residuals that we so obtained have a mean of is zero, the mean is zero. But obviously, it will not be exactly equal to 0, therefore the mean should be not far away from 0. So, we say that the mean is not significantly different from 0 that is the test that we make in this.

Similarly, significance of periodicities, you may still have periodicities present in the data when you do the spectral analysis, the spikes may still appear. But the periodicities that you identify on the residual series when you carry out the spectral analysis must be insignificant; all of them must be insignificant.

Now, in this we do the cumulative periodogram test for this and then we also do the white noise test to make sure that the series is uncorrelated. In the white noise test, we will carry out whittle's test and portmanteau test, typically most of this test we formulate in appropriate statistic and then **knowing the** knowing that **that** statistic follows either way F distribution or t distribution corresponding to critical values of F and t, we decide whether the particular series that we have passes the test or not. Exception to that is the cumulative periodogram test where all the periodicities are tested at once in one go, we will see the details of this now.

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Significance of residual mean

Significance of residual mean:

- This test examines the validity of the assumption that the error series $e(t)$ has zero mean
- A statistic $\eta(e)$ is defined as

$$\eta(e) = \frac{N^{-1/2} \bar{e}}{\hat{\rho}^{1/2}}$$

Where:

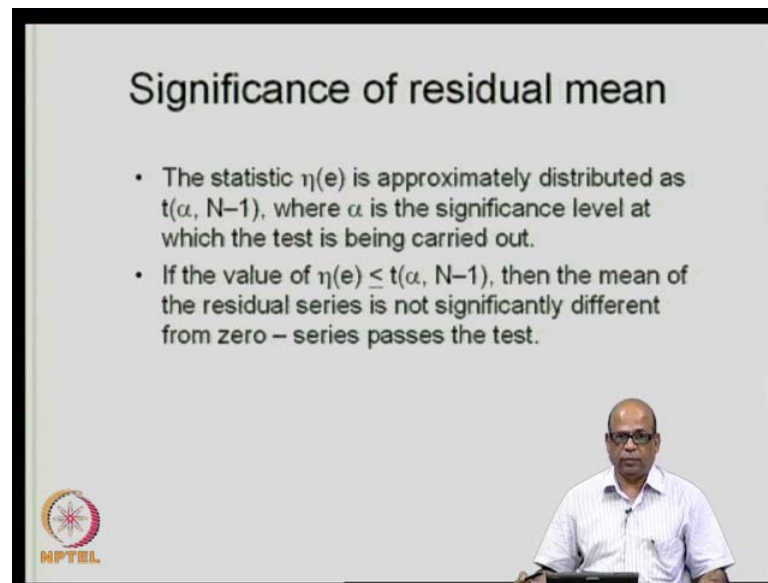
- \bar{e} is the estimate of the residual mean
- $\hat{\rho}$ is the estimate of the residual variance

Ref: Kashyap R. L. and Ramachandra Rao. A. "Dynamic stochastic models from empirical data", Academic press, New York, 1976

So, for the significance of residual mean, it examines the validity of the assumption that the error series $e(t)$ has a zero mean or the mean of the $e(t)$ is not significantly different from zero. So, we define I follow Kashyap and Rao's book here, this is a book, reference is given here.

We form a statistic $\eta(e)$ as $N^{-1/2} \bar{e} / \hat{\rho}^{1/2}$, remember $\hat{\rho}$ here is not correlation, but it is a residual variance. So, you had the residual series and the residual series, you have calculated the mean that is the residual series is $e(t)$, this is given now and then \bar{e} is associated with this $e(t)$ and $\hat{\rho}$ is the variance of this $e(t)$ and N is the number of values, N is the data sample length, so you get N of e .

(Refer Slide Time: 31:00)



The slide is titled "Significance of residual mean". It contains two bullet points:

- The statistic $\eta(e)$ is approximately distributed as $t(\alpha, N-1)$, where α is the significance level at which the test is being carried out.
- If the value of $\eta(e) \leq t(\alpha, N-1)$, then the mean of the residual series is not significantly different from zero – series passes the test.

In the bottom right corner of the slide, there is a small image of a man in a white shirt and glasses, presumably the presenter. In the bottom left corner, there is a logo for NPTEL (National Programme on Technology Enhanced Learning) featuring a stylized sun or starburst design.

Once you get eta of e **i am sorry eta of e**, eta of e is known to be approximately distributed as t alpha N minus 1 where alpha is the significant level for example, you pick a 95 percent or 99 percent and then you get the t alpha comma N minus 1, N is the number of data.

If the value of eta e that you so compute using this is less than the critical t value, this is a critical t value corresponding to the level of significance that you choose. If this value is less, then the mean of the residual series is not significantly different from 0 and then we say that the series passes the test.

It is pretty simple as you can see the entire series is considered first, residual series e t and then you formulate for the entire series. You formulate one value of eta e which is your statistic and then look at, compare the eta is so, you so compute with the t distribution of a specified alpha value and then look at whether this eta e value that you have computed is in fact, less than the critical t value and then we say the series passes, so that was the test for significance of the mean.

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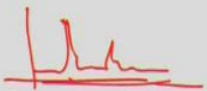
Significance of periodicities

Significance of periodicities:

- This test ensures that no significant periodicities are present in the residual series
- The test is conducted for different periodicities and the significance of each of the periodicities is tested.
- A statistic $\eta(e)$ is defined as

$$\eta(e) = \frac{\gamma_k^2 (N_k - 2)}{4\hat{\rho}_1}$$

γ_k corresponds to the periodicity being tested



Now, the mean may be 0, but still you may have periodicities present in the residuals and that residual series that you obtain from the model should be free of any periodicities, therefore we look for significance of the periodicities. And in fact, the residual series should not have any significant periodicities present in the data. Now, I will discuss two tests, both of the tests are valid, but one is slightly superior to the other as we presently state.

So, the first states again you formulate a statistic $\eta(e)$ is equal to $\gamma_k^2 (N_k - 2)$ divided by $4\hat{\rho}_1$, remember we are doing the test for periodicities here, we pick one periodicity at a time. So, the test is conducted for different periodicities, let us say that the error series that you get has some periodicities present, let us say that this is a residual series and then you see that there are certain periodicities here like this. So, we pick the periodicity corresponding to this particular ω value first and then carry out this test.

So, the k that I am mentioning here corresponds to the particular periodicity that you want to test, one at a time we are testing. So, corresponding to that particular value of k , you calculate γ_k^2 and $\hat{\rho}_1$ again is the variance of the residual series. So, we know $\hat{\rho}_1$ and γ_k , we compute simply $\alpha_k^2 + \beta_k^2$ as we did in our spectral analysis, N is the total number of values that you have.

(Refer Slide Time: 34:17)

Significance of periodicities

Where $\gamma_k^2 = \alpha_k^2 + \beta_k^2$

$$\hat{\rho}_1 = \frac{1}{N} \left[\sum_{t=1}^N \left\{ e_t - \hat{\alpha}_k \cos(\omega_k t) - \hat{\beta}_k \sin(\omega_k t) \right\}^2 \right]$$
$$\hat{\alpha}_k = \frac{2}{N} \sum_{t=1}^N e_t \cos(\omega_k t)$$
$$\hat{\beta}_k = \frac{2}{N} \sum_{t=1}^N e_t \sin(\omega_k t)$$

$2\pi/\omega_k$ is the periodicity for which test is being carried out.

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So, gamma k square is equal to alpha k square plus beta k square for the particular value of k, as I mentioned here gamma k corresponds to the periodicity being tested (Refer Slide Time: 34:26). So, in this case we may be testing for this particular value and the omega k that we do in our spectral analysis, let us say that this is omega k we are writing and this is I k with our notations. We pick up that particular omega k and then for that particular k, we compute gamma k.

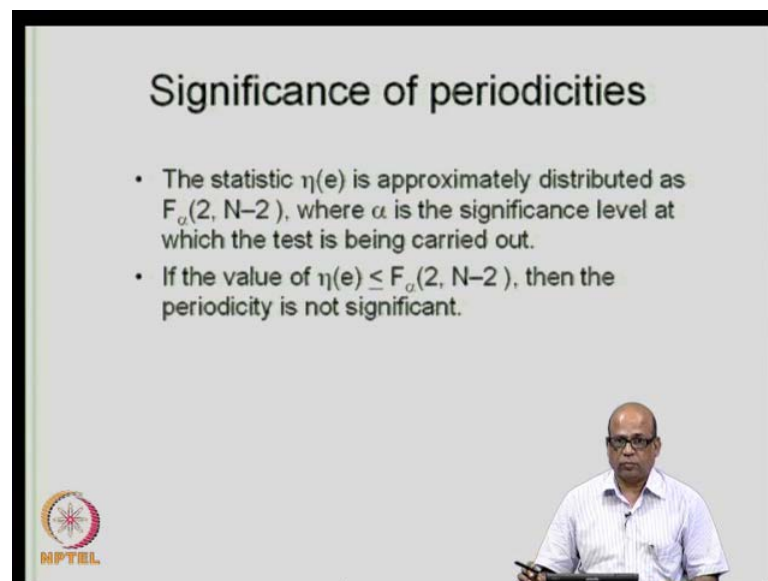
Then rho 1 cap I am sorry I mention this as the variance, rho 1 cap is computed based on your e t, this is actually minus alpha 1 cos omega k t, this is corresponding to that particular k value, omega k minus beta cap sin omega k t whole square. And alpha k we know 2 by N, this is from our spectral analysis, so alpha k and beta k you get directly from the spectral analysis.

So, rho 1 cap is obtained from this for that particular periodicity which we are examining now. I again repeat. This test, test one periodicity at a time. So, we pick that particular omega k and then corresponding to that k, we calculate gamma k and similarly, rho 1 cap is known. See, here your all these values are known, alpha k is known, alpha cap is alpha k actually and then you have the beta k value which is also known, for completeness sake let me make this correction. This is alpha k cap and this is beta k cap, for that particular k and then you calculate rho 1 cap (Refer Slide Time: 36:05). So, rho 1 cap is known in

this statistic γ_k is known, ρ_1 is known and therefore, you can compute the statistics.

Now, the periodicity for which the test is being carried out is 2π by ω_k , let us say you obtain a 12 month periodicity even in the residuals, then the corresponding value of ω_k is 2π by that particular periodicity.

(Refer Slide Time: 36:39)



Significance of periodicities

- The statistic $\eta(e)$ is approximately distributed as $F_{\alpha}(2, N-2)$, where α is the significance level at which the test is being carried out.
- If the value of $\eta(e) \leq F_{\alpha}(2, N-2)$, then the periodicity is not significant.

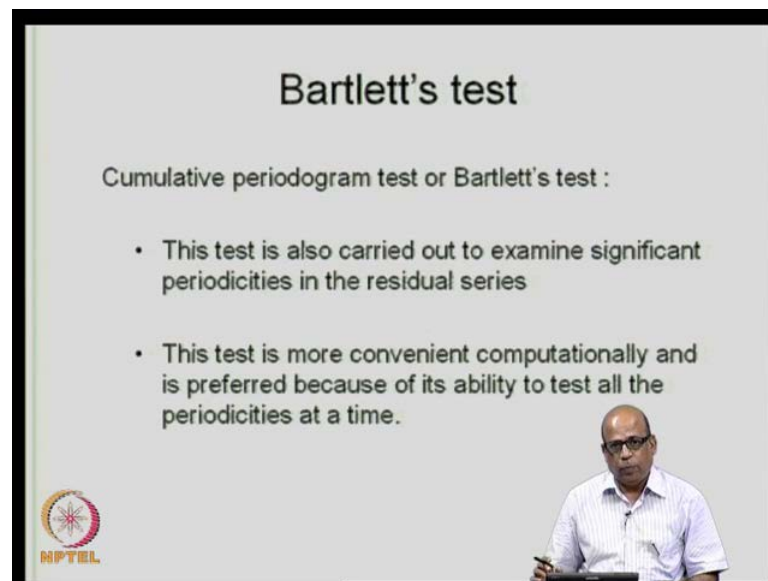
So, the statistic $\eta(e)$ is approximately distributed as $F_{\alpha}(2, N-2)$ where α is the significance level.

The F distribution tables, if you look up in any standard text book, they will give for each of the significance level that number of degrees of freedom and for the number of values. So, let us say you choose F_{α} as $F_{0.95}$ or α as 0.95, 95 percent significance corresponding to 95 percent significance you compute these (Refer Slide Time: 37:15). In fact, this is just the same test that we did in our spectral analysis to identify the significance of periodicities and then you take the statistic that value of the statistic that you have calculated $\eta(e)$. If this is less than $F_{\alpha}(2, N-2)$, then the periodicity is not significant.

Let us say that your spectral analysis like this shows of two or three different periodicities, first you test for this periodicity. If this is not significant, then naturally these periodicities may not be significant if they are all in decreasing order like this.

However, you may have a case where this may be significant if one of them is significant, then you have to test for the next periodicity also (Refer Slide Time: 38:10). So, you keep testing for the periodicities until you are satisfied that all the periodicities that are thrown up by the spectral analysis of the residual series are all insignificant, then the series passes the test, even if one periodicity is significant, then the particular model does not pass the test.

(Refer Slide Time: 38:45)



The slide is titled "Bartlett's test" and describes the "Cumulative periodogram test or Bartlett's test". It includes two bullet points: "This test is also carried out to examine significant periodicities in the residual series" and "This test is more convenient computationally and is preferred because of its ability to test all the periodicities at a time." The slide also features the NPTEL logo in the bottom left and a presenter in the bottom right.

Bartlett's test

Cumulative periodogram test or Bartlett's test :

- This test is also carried out to examine significant periodicities in the residual series
- This test is more convenient computationally and is preferred because of its ability to test all the periodicities at a time.

Now, for the significance of periodicities, we also have another test which is called as the cumulative periodogram test or the Bartlett's test. The advantage of this Bartlett's test is that unlike the previous test that I just explained, this test examines all the periodicities at one time rather than going by one periodicity to another and therefore, this is computationally convenient. So, the test is more convenient and it is preferred because of its ability to test all the periodicities at a time.

(Refer Slide Time: 39:35)


Bartlett's test

$$\gamma_k^2 = \left\{ \frac{2}{N} \sum_{t=1}^N e_t \cos(\omega_k t) \right\}^2 + \left\{ \frac{2}{N} \sum_{t=1}^N e_t \sin(\omega_k t) \right\}^2$$

$k = 1, 2, \dots, N/2$

$$g_k = \frac{\sum_{j=1}^k \gamma_j^2}{\sum_{k=1}^{N/2} \gamma_k^2} \quad 0 \leq g_k \leq 1$$

The plot of g_k vs k is called as cumulative periodogram

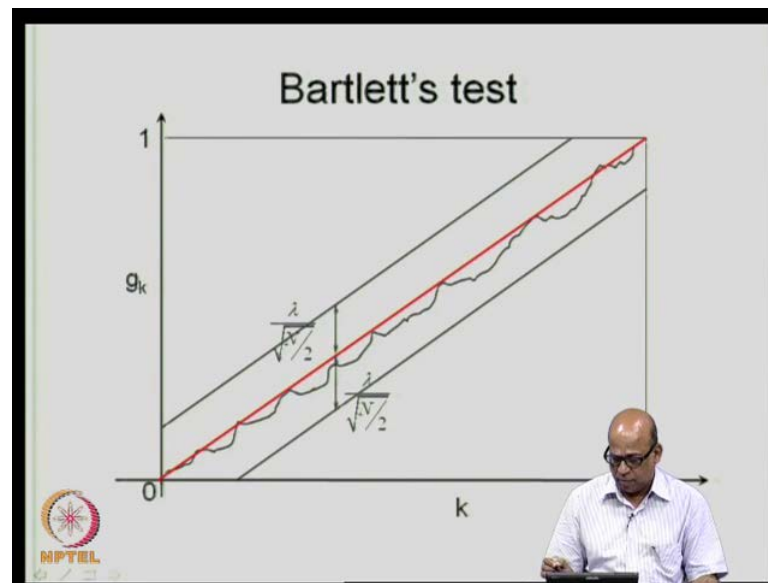
 Ref. Kashyap R.L. and Ramachandra Rao.A. "Dynamic stochastic models from empirical data", Academic press, New York, 1976

So, essentially what we do in this test is that we form a cumulative periodogram as follows; we define gamma k for k is equal to 1 to N by 2, as I mentioned this is a validation test. So, this is validation period, for the N by 2 values you calculate gamma k square, corresponding to each of the k, you know omega k from your spectral density, you **you** would have computed this (Refer Slide time: 40:00) and e t is that particular value of the residual series and you are summing over t is equal to 1 to N, t is equal to 1 to N and therefore, all the terms are known here (Refer Slide Time: 40:12).

So, for k is equal to 1, you get gamma k, k is equal to 2 you get gamma k and so on. So, like this for k is equal to N by 2, you have gamma k square, then you **you** determine g k as j is equal to 1 to k summation gamma j square. So, from 1 to k, you are adding up gamma j square. So, g 1 you have one term, g 2 you have two terms, 3 you have three terms and so on, like this you calculate summation of this divided by over the entire period, k is equal to 1 to N by 2 gamma k square.

So, the plot of g k verses k is called as the cumulative periodogram. So, you know how to compute g k, all values are known here, therefore you can calculate gamma k. Once you know gamma k, you sum up to k and then get g k by normal I think with respect to the entire sum here. And because of this nature, g k varies from 0 to 1; the maximum value it can take is 1.

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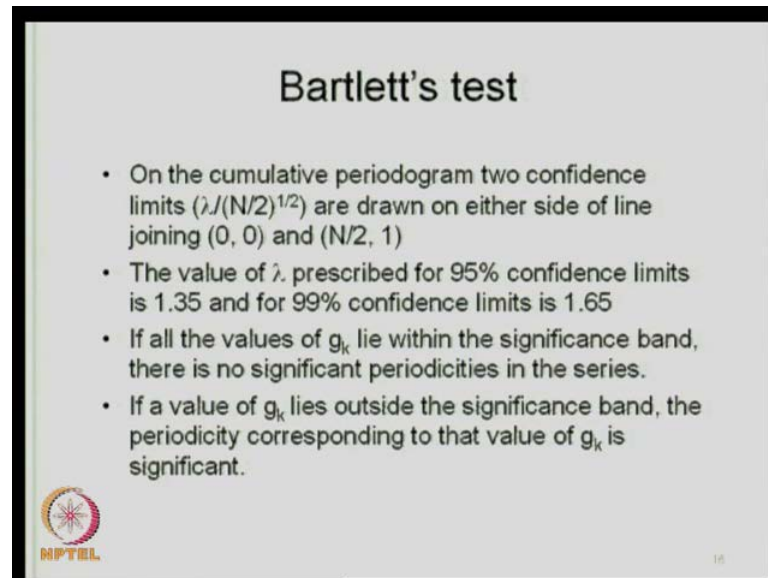
Then we plot k on the x axis and g_k on the y axis, now I will explain this with respect to this figure. So, you computed g_k and you have the k values, so this is the plot, this black line you are seeing is the plot of g_k versus k for the residual series (Refer Slide Time: 41:45).

Now, on this we have used $N/2$ values and the maximum value of g_k is 1. So, draw a line between 0 and $N/2, 1$, this corresponds to a frequency of 0.5, so 0.5, 1 on the frequency diagram. So, draw a line, this is the red line here, on this line on either side of this red line, you draw the confidence bands, the confidence band is γ that is $\lambda / \sqrt{N/2}$. So, this is $\lambda / \sqrt{N/2}$ on either side of the line so drawn. So, like this you get a band **band** now, this is a band.

If your cumulative periodogram that you have drawn, any part of that goes above or below, goes beyond the bound that you have drawn, then the series does not pass the test. In fact, it means that corresponding to that particular, let us say that it was beyond this particular band at a particular k value, the periodicity corresponding to that particular k value in the residual series is significant. If the periodogram completely lies within this band, then it is not significant, so this is how we examined the significance of the periodicities.

So, let me summarize that, so you draw first the cumulative periodogram by considering g_k by constructing g_k for each of the k values, you have $N/2$ values and therefore, you construct g_k versus k for $N/2$ values.

(Refer Slide Time: 43:52)



Bartlett's test

- On the cumulative periodogram two confidence limits $(\lambda/(N/2)^{1/2})$ are drawn on either side of line joining $(0, 0)$ and $(N/2, 1)$
- The value of λ , prescribed for 95% confidence limits is 1.35 and for 99% confidence limits is 1.65
- If all the values of g_k lie within the significance band, there is no significant periodicities in the series.
- If a value of g_k lies outside the significance band, the periodicity corresponding to that value of g_k is significant.

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And then on the cumulative periodogram two confidence limits given by $\lambda/\sqrt{N/2}$ are drawn on either side of the line joining $0, 0$ and $N/2, 1$. What is 1 ? 1 is the maximum value of this and $N/2$ corresponds to your k_{\max} that is the maximum lag, in fact this corresponds to a frequency of 0.5 (Refer Slide Time: 44:06).

Then what is this λ value? The value of λ for 95 percent confidence limit is 1.35 and for 99 percent confidence limit is 1.65. In fact, this test also I have taken it from Kashyap and Rao. So, you can refer to Kashyap and Rao, these are the values prescribed for this.

So, you know how to draw the bounds now, **you know** you have drawn the cumulative periodogram, you also know how to draw the bounds. Now, if all the g_k values lie within the significance band, there is no significant periodicity present in the series, which means that the series passes the test. If a particular value of g_k lies outside the significance band, the periodicity corresponding to that value of g_k is significant and therefore, the series does not pass the test. So, this is how we **we** carry out the Bartlett's test or the cumulative periodogram test.

In most of the cases if the models are acceptable, the residual series typically gives, typically lies well within the bounds that we have drawn. It hardly comes very close to these bounds, in fact when I discuss the case studies it will be clear. However, just for your curiosity, what you can do is, you draw the cumulative periodogram also for the original data, and then you will see that there are several periodicities which are lying way above and below this particular band, either on this side or on this side, so indicating that the periodicities are significant.

(Refer Slide Time: 46:10)


Whittle's test for white noise

White noise test (Whittle's test):

- This test is carried out to test the absence of correlation in the series.
- The covariance r_k at lag k of the error series $e(t)$

$$r_k = \frac{1}{N-k} \sum_{j=k+1}^N e_j e_{j-k} \quad k = 0, 1, 2, \dots, k_{\max}$$

- The value of k_{\max} is normally chosen as $0.15N$

 Ref: Kashyap R. L. and Ramachandra Rao. A. "Dynamic stochastic models from empirical data", Academic press, New York

Now, we will see the white noise test. So, what we did is, for the periodicities we have now two test and typically we prefer the Bartlett's test, because of its ability to test all the periodicities in one go and also computationally it is very simple. That means, you simply calculate g_k much like your correlogram, you formulate the periodogram with respect to lag k , you calculate g_k values and plot this and immediately this significance comes out. Whereas, in the first two tests that we had, we formulated a statistic corresponding to each of the periodicities that we suspect to be significant, we have to make that particular test and therefore, the Bartlett's test is preferred for test of periodicities.


Now, we will test for white noise or the assumption that we made that the series is uncorrelated, the residual series that we formulate is uncorrelated. So, in this we have the whittle's test for white noise, whittles test we formulate again the covariance matrix,

remember we are dealing with the e_t series that is the residual series. So, the covariance r_k at lag k of the error series e_t is calculated, this is simply r_k is equal to $1 - 2k/N$ for k from 0 to k_{max} , k_{max} can be typically 0.5 \sqrt{N} here, e_j into e_{j-k} .

(Refer Slide Time: 48:06)

Whittle's test for white noise

- The covariance matrix is

$$\Gamma_{n1} = \begin{bmatrix} r_0 & r_1 & r_2 & \cdot & \cdot & r_{k_{max}} \\ r_1 & r_0 & r_1 & \cdot & \cdot & r_{k_{max}-1} \\ r_2 & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ r_{k_{max}} & r_{k_{max}-1} & \cdot & \cdot & \cdot & r_0 \end{bmatrix}_{k_{max} \times k_{max}}$$


Then once you know the **the** covariance that is r_k is given, you formulate the covariance matrix. So, covariance matrix will be also size k_{max} by k_{max} , the symmetrical $r_0, r_1, r_2, \dots, r_{k_{max}}$, like this you formulate; Γ_{n1} , this is a covariance matrix and we denote it by Γ_{n1} , then we formulate a statistic.

(Refer Slide Time: 48:38)

Whittle's test for white noise


- A statistic $\eta(e)$ is defined as

$$\eta(e) = \frac{N}{n1-1} \left(\frac{\hat{\rho}_0}{\hat{\rho}_1} - 1 \right) \quad n1 = k_{\max}$$

Where $\hat{\rho}_0$ is the lag zero correlation =1, and

$$\hat{\rho}_1 = \frac{\det \Gamma_{n1}}{\det \Gamma_{n1-1}}$$

The matrix Γ_{n1-1} is constructed by eliminating the last row and the last column from the Γ_{n1} matrix.



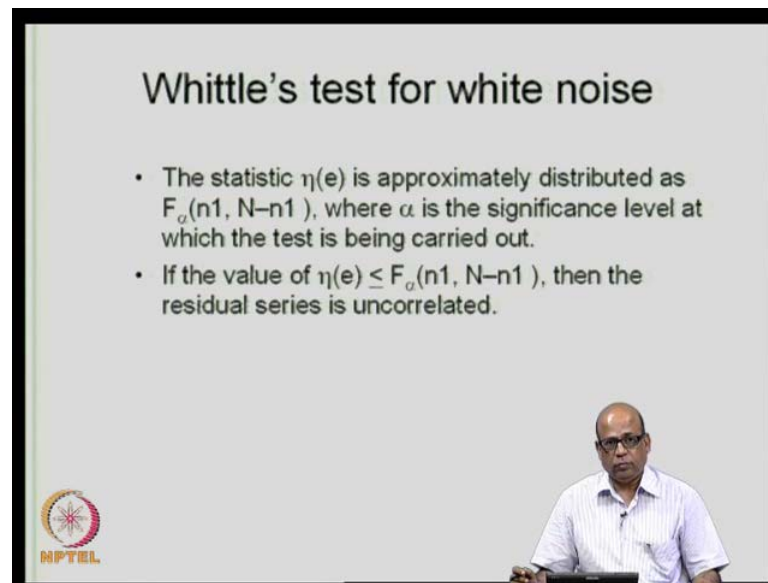
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So, essentially we determined the covariance and formulated the covariance matrix. Using the covariance matrix, we define a statistic $\eta(e)$ is equal to N is given $n1 = k_{\max} - 1$ $\hat{\rho}_0$ is lag zero correlation which is 1 and $\hat{\rho}_1$ cap are defined now minus 1.

$\hat{\rho}_1$ cap here is determinant τ_{n1} by determinant τ_{n1-1} , τ_{n1} is given here (Refer Slide Time: 49:06). So, this is a matrix, the determinant of this is used here, determinant τ_{n1} , now τ_{n1-1} is constructed by eliminating the last row and the last column from the **τ_{n1}** the τ_{n1} matrix that is, you have this matrix, you delete this row **I am sorry** the last row and the last column, you delete this row and this column (Refer Slide Time: 49:15). So, you will have τ_{n1-1} and the determinant of that, that is what is used here, so you can formulate $\hat{\rho}_1$ cap.

So, you know N here, $n1$ is known, k_{\max} $\hat{\rho}_0$ is 1 and $\hat{\rho}_1$ cap is formed by $\hat{\rho}_1$ cap is calculated by taking the determinant τ_{n1} , τ_{n1} is defined here, then take out the last row and last column, take the determinant and that is what defines determinant of τ_{n1-1} . Therefore, $\hat{\rho}_1$ cap is known and therefore, you calculate this particular statistic.

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The slide contains the following text:

Whittle's test for white noise

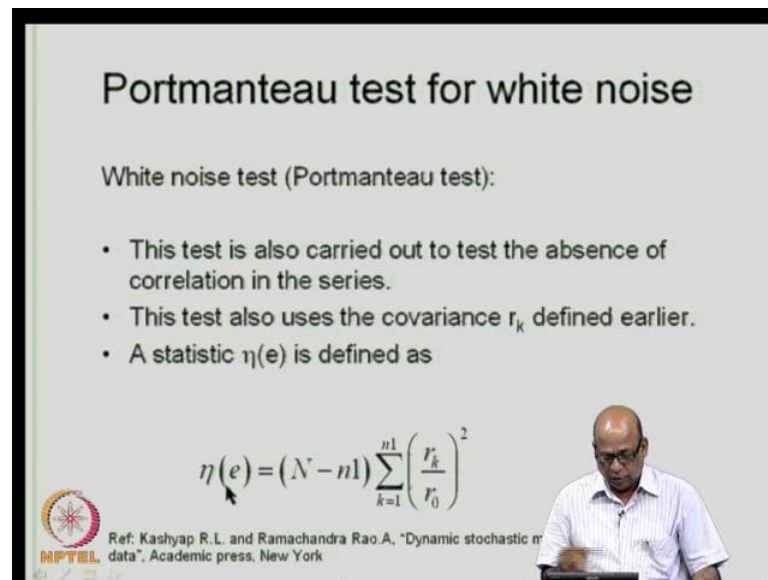
- The statistic $\eta(e)$ is approximately distributed as $F_{\alpha}(n1, N-n1)$, where α is the significance level at which the test is being carried out.
- If the value of $\eta(e) \leq F_{\alpha}(n1, N-n1)$, then the residual series is uncorrelated.

In the bottom right corner of the slide, there is a small image of a man in a white shirt and glasses, and a logo in the bottom left corner that says 'MPTEL'.

Once that statistic is known, this statistic is approximately distributed as $F_{\alpha} n1 N$ minus $n1$, $n1$ is your k max, typically taken as 15 percent of your N . So, as we did for earlier test using the F distribution, you fix your confidence level, typically 95 percent or 99 percent, $n1$ is known, capital N is known which is a number of values and therefore, you know the critical value of F_{α} .

If the value of $\eta(e)$ that you calculate, this value if the value that you calculate is less than F_{α} , then the residual series is uncorrelated. So, this is what we do for white noise test, white noise is the series is uncorrelated, this is called as the whittle's test.

(Refer Slide Time: 51:17)



Portmanteau test for white noise

White noise test (Portmanteau test):

- This test is also carried out to test the absence of correlation in the series.
- This test also uses the covariance r_k defined earlier.
- A statistic $\eta(e)$ is defined as

$$\eta(e) = (N - n1) \sum_{k=1}^{n1} \left(\frac{r_k}{r_0} \right)^2$$

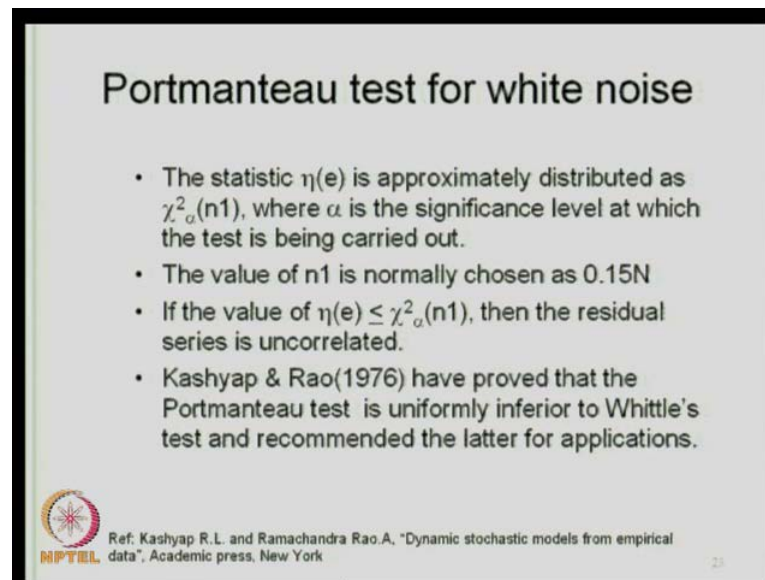
Ref: Kashyap R. L. and Ramachandra Rao. A. "Dynamic stochastic data". Academic press, New York

In fact, we have another test for examining whether the series e_t is in fact, white noise or not, it is called as the portmanteau test.

Now, this test is also carried out to test the absence of correlation in the series and this also uses the r_k as defined earlier that is the covariance matrix as we have defined earlier, this covariance matrix is yours (Refer Slide Time: 51:50). So, you know how to formulate the covariance matrix, from the covariance matrix we start defining another statistic in the portmanteau test.


So, using the covariance matrix r_k here, your statistic $\eta(e)$ is defined, N is the number of values, n_1 is your maximum lag and r_k is corresponding to the covariance of order k . And then r_0 is the covariance of order 0, so k is equal to 1 to n_1 , so you know how to calculate the statistic value.

(Refer Slide Time: 52:27)



Portmanteau test for white noise

- The statistic $\eta(e)$ is approximately distributed as $\chi^2_{\alpha}(n_1)$, where α is the significance level at which the test is being carried out.
- The value of n_1 is normally chosen as $0.15N$
- If the value of $\eta(e) \leq \chi^2_{\alpha}(n_1)$, then the residual series is uncorrelated.
- Kashyap & Rao(1976) have proved that the Portmanteau test is uniformly inferior to Whittle's test and recommended the latter for applications.

 Ref. Kashyap R.L. and Ramachandra Rao.A. "Dynamic stochastic models from empirical data", Academic press, New York

Now, this statistic is approximated as, approximately distributed as chi square distribution with argument n_1 . So, you fixed α and n_1 , n_1 is your k_{\max} which is a maximum lags up to which you have gone. And from the chi square tables, you can get this value, the value n_1 is chosen again as $0.15N$. So, get the chi square, α, n_1 value from the tables, if your $\eta(e)$ that you have calculated is less than the critical value of α , then the residual series is uncorrelated.

So, for the white noise test, that means to examine whether the series is uncorrelated or not, we have two tests, namely Whittle's test and the portmanteau test. In both of them we use the covariance matrix, covariance matrix of the residual series and then formulate a statistic.

Now, between these two Kashyap and Rao have proved, the same text book that I have been mentioning, they have proved that the portmanteau test is uniformly inferior to Whittle's test and we prefer the latter for applications. So, while there are two tests that I have mentioned, if you **if you** are in a fix which one to use, you always go for the Whittle's test as **as** shown by Kashyap and Rao.

Now, before we go to the applications, essentially what we have done is that we have formulated the models, ARIMA type of models and then we have made the test and then we now start applying to different case studies. So, let us summarize what we did on the

choice of the ARIMA models, this is not just what I have covered in this lecture, but also on the previous two lectures.

We formulate the ARIMA models after differencing the series, ARIMA that is autoregressive integrated moving average models that order of differencing, whether it is a first order or second order differencing depends on the non-stationarity present in the data. So, you do the differencing essentially to convert the data into a stationary data, once you have the data as the stationary time series, then you apply this ARIMA type of models.

Now, in the ARIMA models, you know how to estimate the number of AR parameters and the number of MA parameters. In the event that you are not very clear about, how many AR terms to use, how many MA terms to use etcetera, you form the candidate models, a large number of candidate models and then apply this candidate models to the observed time series, estimate the parameters and calibrate the models using part of the series, typically we use half the length of the data to calibrate the series and use these models for validation for the remaining half, remaining part of the data. When you apply these models for the remaining part of the data, you get the residual series that is e_t series.

We do the test of validation on the residual series, we do essentially three different tests, one is to test whether the residual series has a zero mean or the mean of the residual series is not significantly from significantly different from zero that is the first test that we do. Then we test whether the residual series, the that we have obtained by applying the model or is divide of any significant periodicities. So, you may come up with some periodicities, you test for the periodicities, we sort two tests for periodicity, for significance of periodicities and we normally prefer the cumulative periodogram test for significance of periodicities.

Then we also examine whether the series e_t that we have obtained is, in fact white noise or it is uncorrelated. For that again we had two tests, the portmanteau test and the whittle's test, it is shown in the standard text book, Kashyap and Rao that the whittle's test is better than the portmanteau test.

So, we apply the models, make all these tests and then make sure that the model that you have chosen, either based on the maximum likelihood criteria or the minimum mean

square criteria passes all these tests and then choose that particular model for application. We will continue the discussion and specifically, we will see how to apply these models, whatever procedures that I have explained so far, we will see the applications of these in the coming lectures. Thank you for your attention.