

Probability Methods in Civil Engineering
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Lecture No # 25

MGF of Multivariate RVs and Multivariate Probability Distributions

Hello, welcome to this lecture. This is seventh lecture of this current module, multivariate random variable. So far, we have discussed their several properties and you know the last lecture that, we are that we are supposed to cover in this lecture is that moment generating function. So, we will complete that part and after that we will discuss something about their different distribution.

So, multiple random variable when we talk it is their joint p d f and joint c d f and if it is discrete then it is joint PMF that we know. So, this thing we will be discussing and after that we will give some standard example of bivariate case means, mostly you know that when we are discussing different properties. We are discussing in case of the bivariate random variable, where there are two random variables are involved and after that we will also take you through the concept, how it can be extended for the multivariate random variable cases, where more than two random variables are there.


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Probability Methods in Civil Engineering

Module 5: Multiple Random Variables

Lecture – 7: MGF of Multivariate RVs and Multivariate Probability Distributions

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So, today's lecture we will start with the discussion of this moment generating function and this moment generating function we have discussed earlier, in case of the single random variable and as this module is on this multiple random variable. We will discuss this one first, in case of this multiple random variable and after that we will take different distributions as well.

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Outline

- ☐ Moment Generating Functions for multiple Random Variables
- ☐ Bivariate Probability Distribution Function
 - ☒ Bivariate Normal Distribution
 - ☒ Bivariate Exponential Distribution
- ☐ Multivariate Random Variables

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So, outline of our today's lecture is, first the moment generating function of multiple random variables. We will know that, how in case of first of all in case of two random variables, how we can generate that moment generating function and after generating after obtaining its MGF that is moment generating function. We will know that their utility that if you take that r th order derivative. Now when we are talking about the r th order derivative that, is in the context of a single random variable that we have discussed earlier, but here as there are more than one random variable.

So, there will be so derivative order also may change for the different random variables and after getting that derivative, if you evaluate that function at the origin. Then we will get that moment that also we will discuss, because that is the main focus of this obtaining this moment generating function. So, we will get different moments and for the multiple random variable when you are talking about the moment, when also there is a sense of joint moment that is for, if it is bivariate case then it is for the both random variable the order for the different order for the different random variable.

So, we will come to know what are their moments, so that is the reason, why will learn about this moment generating function. So, after that we will discuss about this bivariate probability distribution functions and so far as this joint distribution is concerned, there are not many examples. So, there are two things those are important is this are frequently used is this, the bivariate normal distribution and second one is bivariate exponential distribution.

This bivariate exponential distribution we have discussed, while the in this module itself in previous lectures. So, here we will just once again we will mention it just you complete this one. So these are the two bivariate joint p.d.f and that probability function, probability density function is available. So, these are also of much use in different fields including civil engineering.

But here one thing is important that we also mention the earlier that, getting a joint distribution from there from the marginal distribution is not easy and for example, this bivariate normal distribution case also it is not that, if I get some marginal distribution, we cannot create what is their joint distribution opposite side this true.

Opposite side means that if the if both the random variables are jointly a Gaussian distribution then, we can say that their marginals are also Gaussian, but the reverse is not true that also we will discuss.

So, what the major part is will be untouched, so far is that how to get the joint distribution. Because, these are the only two joint distribution forms is available. So, there that the recent technique on copula that will be taking up towards the end of this module, that I mentioned earlier. So, for whatever is the existing effort from the copula we will discuss and this two are the bivariate normal distribution and the bivariate exponential distribution, this we will include in today's lecture.


And after that the property that we are discussing for the case of the bivariate, we will take the same concept for the more than two particularly if I remind that multivariate. So, more than two random variable cases we will extend the same concept.

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Joint Moment Generating Function

- Similar to moment-generating function of a random variable defined in previous module, the moment generating function of two random variable is:
- Discrete case:

$$M_{X,Y}(t,u) = E(e^{tX+uY}) = \sum_{\text{all } x} \sum_{\text{all } y} e^{tx+uy} p_{X,Y}(x,y)$$



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So, start with that joint moment generating function, similar to the moment generating function of a random variable defined in the previous module. The moment generating function of two random variables is again, there are two cases. One is the discrete and other one is that continuous random variable case.

So, in case of the discrete you know, we have discussed for the single random variable. Here, there are two random variables involved then say that x and y . So, for this both this random variable the moment generating function is the expectation of the function of that exponential $t x$ plus $u y$.

So, if we take this expectation of this function now in the earlier lecture, we have discussed for a function how to get that expectation of a function. So, that this nothing but this function multiplied by their probability mass function and we have to sum it over for all x and for all y . So, if we do this on the function that we will get that is the moment generating function of the joint PMF between x and y .

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Joint Moment Generating Function


□ Continuous case

$$M_{X,Y}(t,u) = E(e^{tX+uY}) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{tx+uy} f_{X,Y}(x,y) dx dy$$

□ Moment generating function for marginal:

$$M_X(t) = E(e^{tX}) = M_{X,Y}(t,0)$$

$$M_Y(u) = E(e^{uY}) = M_{X,Y}(0,u)$$



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Similarly, in case of the continuous random variables if this x and y two random variables, which are continuous then it is also that expectation of this of the exponential function, which is the exponential of $t x$ plus $u y$ and to get this exponential. We have to do the integration over the entire range of both the random variables.

Here, this it is integrated from the over the entire real line minus infinity plus infinity. This function exponential of $t x$ plus $u y$ multiplied by their joint pdf $f_{X,Y}(x,y)$ and it is integrated over the entire range. So, this one we will get that moment generating function for the continuous random variable.

Now, just if we want to know their marginals now, what we have done that, it is just the power that we have use the power of this x and power of the y, so which are here both the cases is one. Now, if you want to know for the marginal and we know that how to get that moment generating function of their marginals. So, the marginal means you know that, we discuss earlier is that the distribution of one random variable irrespective of the distribution of the other one.

So, this moment generating function of the x, if I know then we know that this is the will be the expectation of the exponential t x only. So, now expect that exponential of t x means here, then y that term was here just it is the y power 0. This component will not be there not y power 0. So, this component will not be there.

So, that u we can say it to be 0. So, not that y power 0 it is u equals to 0. So, now if you put that u equals to 0, then we can write that the moment generating function of the joint random variable x and y there we have to put that t and for in place of the u, we have put the 0.

So, this is what we are explaining is that so how to get the moment generating function for the marginal that is if we know the moment generating function of the joint random variable then by replacing the suitable attribute, we can get the moment generating function for the marginals. Similarly, for the y also we have to replace the t by 0 and we have to keep this u in from the joint moment generating function.


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Joint Moment Generating Function

- If X and Y are statistically independent:

$$M_{X,Y}(t,u) = M_X(t)M_Y(u)$$
- Multiple Random Variable:
 - The joint moment-generating function of multivariate random variable (X_1, X_2, \dots, X_n) can also be defined similarly

$$M_{X_1, X_2, \dots, X_k}(t_1, t_2, \dots, t_k) = E \left[\exp \left(\sum_{i=1}^k t_i X_i \right) \right]$$



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Now, this is also again we have discussed earlier that, in case that x and y are statistically independent then so far as the moment generating function is concerned, then we can write that this moment generating function of the joint moment generating function should be the product of their moment generating function of the marginal of the random variable involved.

So, this $m_x(t)$ is nothing but the moment generating function of x and $m_y(u)$ is the moment generating function of y . So, both are from the marginal and if we multiply it and if this product is equal to their joint marginal. So, joint moment generating function then we can conclude that x and y are statistically independent.

Now, for this multiple random variable just we have to simply, explain the idea that we have seen is that in case of if there are k random variables, then this for the joint moment generating function between that x_1, x_2 and up to x_k then, what we have to take is that there are different these are all means attribute that we are considering. So, this will be the expectation of a function which should be that exponential of that $t_1 x_1 + t_2 x_2 + t_3 x_3 + \dots + t_k x_k$.


So, basically from this now, if you just see that this is the for the single random variable is an exponential of $t x$, now for that two random variable. We are just adding that that exponential is taken for this $t x + u y$, if there is another random variable that we have to add with some other attribute. Similarly for this k a number of random variables is explained here you have to take the exponential of this function that is $t_1 x_1 + t_2 x_2 + \dots + t_k x_k$ where i is from 1 to k .

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Joint Moment Generating Function

□ The r^{th} moment of X_i can be determined by differentiating the joint moment generating function r times with respect to t_i and then evaluating the derivative with $t_i=0$

$$E[X_i^r] = \left. \frac{\partial^r M_{X_1, X_2, \dots, X_k}(t_1, t_2, \dots, t_k)}{\partial t_i^r} \right|_{(t_i=0)}$$



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Now, after we get this joint moment generating function, what we generally get what is the interest that, we have seen in case of this single random variable also. That we have to take some particular order of their derivative and you have to evaluate its value at the origin to get that order moment.

So, here also for the multiple random variable also, we have to get the suitable derivative with respect to the different random variable involved and we have to evaluate that one at the origin to get their moment for that order.

Now, we will discuss it in general first that, if you are talking about the r^{th} moment of the x_i . This x_i is what you are talking, whatever the random variable involve so there are some random variables are there. So, anyone if I want to know and I want to know, what is this r^{th} moment then, what we have to do is that we have to differentiate the joint moment generating function r times with respect to that t_i and then evaluating that derivative with t_i equals to 0.

So, the expectation of this i^{th} random variable which is involved here power r . So, this moment will be the r^{th} order moment this one is the derivative the r^{th} order derivative with respect to that t_i which one is taken so it is written at the beginning. So, which is the x_1 but it is not necessary that it has to be put here, so any random variable between this x_1 to x_k .

I think the better exponential, you know that it will be $x_1 \times x_2$ then x is somewhere in between where i is I can take any value between 1 and k . So, and with respect so the and it should be derivative it should be taken with respect to the corresponding attribute that is the t_i there and after getting this derivative, we have to evaluate this value of this function at t_i equals to 0.


So, this is the general expression and it is not only for a single random variable. We can do it for even for this any subset of this random variable. So, just I want to know that the expectation of $x_1 \times x_2$. Say for example out of these k random variables I just want to pick up x_1 and x_2 . So, I have to take that double, this moment generating function with respect to that double t_1 double t_2 . So, that two times derivative first time I am doing with respect to the t_1 and then with respect to the t_2 and after getting this derivative, we have to evaluate it at t_1 equals to 0 and t_2 equals to 0.

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Joint Moment Generating Function

□ The mixed moments of X_i and X_j , $E(X_i^r X_j^s)$ can be generated from the joint mgf by differentiating r times with respect to t_i and s times with respect to t_j , and then evaluating the derivative with $t_i = t_j = 0$

$$E[X_i^r X_j^s] = \frac{\partial^{r+s} M_{X_i, X_j}(t_1, t_2)}{\partial t_i^r \partial t_j^s} \bigg|_{(t_i, t_j) = (0, 0)}$$



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So, this is what we are explaining even in this general case, where there are say that two random variables are there the mixed moments. So, you have the mixed moments means we are we even can change the order, so that there are two random variables that x_i and x_j and for the x_i the order is the r and x_j the order is the s can be generated from the joint moment generating function by differentiating r times with respect to the t_i .

Because, it is x raised r times with respect to t_i and s times with respect to t_j , because this is x_j power this says so s times with respect to t_j and then evaluating that derivative with t_i equals to t_j equals to 0.

So, what is explained here in case, that there are two random variable. So, we have to do this derivative with respect to that the r times with respect to the t_i and s times with respect to the t_j and then we are evaluating that function at t_i equals to 0 and t_j equals to 0.

So, where will this be useful is that now once, we get this expression then I can vary any value of this r whether I if I need that expectation of x . Then I have to put that r equals to 1 and s equals to 0, if I want that their expectation of $x y$ then I have to put r equals to 1 s equals to 1 or other way if I want that second moment of y . Then I have to put that r equals to 0 and s equals to 2. So, in so this general expression can be used whatever way we want to know.


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Example

Q. Life of a structure, consist of two subcomponents, depends on their individual life time X and Y , which are exponentially distributed. The joint density function of X and Y is given as:

$$f_{X,Y}(x,y) = 2e^{-x-2y} \quad x, y \geq 0$$

Determine the joint moment generating function of X and Y . Compute the mean of X and Y . Also compute the covariance between them.



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So, where one example you have taken and this example again also taken very carefully, because the properties of this distribution we already know, we have we have seen now what we are looking for though this expression though this problem is that this thing the reason for this taking of this type of problem is that we know what is this p.d.f.

And, in fact just by reading the problem itself, we should know that what the answer is expected. But, whatever the theory we have discussed with respect to this moment generating function or specifically the joint moment generating function, then we will again apply that on to crosscheck that whether what our belief is whether this is being satisfied, through this problem or not that will see.

So, problem states here that this is a life of a structure and this structure the life of that structure depends on the various components, say here it is simplified with respect to the two components only, so just to take this for the discussion problem. So, this life of a structure consists of two subcomponents depends on the individual life time say x and y . So, whatever target is that I have to model the life of a structure, that structure having two subcomponents and the two subcomponents having their life time. Here both the life time of the both the subcomponents are one random variable each.

So, for the first component say the life the random variable denoted for the x for its life time and for the second component the life is denoted for its for a random as a random variable y and both these are exponentially distributed and their joint density function of this x and y is given as this $f(x, y)$ is equal to $2e^{-x-2y}$, where x and y is greater than equal to 0. Determine the joint moment generating function of x and y and then compute the mean of x and y , also compute the covariance between them.

Now before I proceed to solve this problem, why this problem has been taken that I was mention that I mentioned just minutes, before is that this function we know and it can be easily shown that this both this x and y are independent and their joint function is nothing but the product of their individual marginal and marginal for the x is e^{-x} and marginal for the y is $2e^{-2y}$.

This type of problem we have taken up earlier. So, I am not going to how to prove that that onethat I am not taking at this level, because you can just see few lecture previous, where we have solved this type of problem. So, with this onethat what you know is that both this x and y are independent. So, if you know then the covariance between them obviously will be 0, which also we discuss earlier.

Now, you also what we discuss the exponential distribution so $e^{-\lambda x}$ is the 1 exponential distribution for which is the marginal distribution for the x and the λ parameter is here one and for the y the marginal distribution is $2e^{-2y}$ where the λ parameter is 2.

So, that exponential distribution also you know and we also know that the mean of n exponential distribution is $1/\lambda$. So, if you know that 1 then the mean of this x will be $1/1$ that means 1 and mean of the y will be $1/2$ so half.

Now, these things so these things what we will do now for this problem is that, first of all we will determine the joint moment generating function and as we know the answer already, we will just cross-check we take that whether the how we can use this moment generating function to get the respective moments. So, once this answer we can check and then for the any other moment for this joint distribution we can use the same methodology to get that whatever the order of the moment that we need.


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Example...Contd.

Sol.:

We have

$$\begin{aligned}
 M_{X,Y}(t,u) &= E(e^{tx+uy}) = \int_0^\infty \int_0^\infty e^{tx+uy} f_{X,Y}(x,y) dx dy \\
 &= \int_0^\infty \int_0^\infty e^{tx+uy} 2e^{-x-2y} dx dy \\
 &= 2 \int_0^\infty \int_0^\infty e^{(t-1)x} e^{(u-2)y} dx dy
 \end{aligned}$$



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So, we will go for this solution and first what we have to get that we have to get its moment generating function. So, just now what we discuss is this moment generating function is for the two random variables x and y is the expectation of the function which is exponential of $tx + uy$.


Now, this one we can express that the power t^x plus u^y multiplied by its joint density and it should be integrated over the entire support of the random variable, here both the random variables are taking the nonnegative range. So, it is from 0 to infinity and 0 to infinity so both this x and y .

Now, this joint density we know that is 2 exponential minus x minus $2y$ then multiplied with this $e^{tx + uy}$. So, this t and u are the variables to which this moment generating function will be expressed.

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Example...Contd.

$$\begin{aligned}
 M_{X,Y}(t,u) &= 2 \int_0^{\infty} e^{(u-2)y} \frac{e^{(t-1)x}}{t-1} \Big|_0^{\infty} dy \\
 &= \infty - \left(\frac{2}{(t-1)(u-2)} e^{(u-2)y} \Big|_0^{\infty} \right) \\
 &= \frac{2}{(t-1)(u-2)}
 \end{aligned}$$



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Now, after doing some usual steps and doing this integration will get this expression putting that suitable limit up to infinity and will get that it is 2 by t minus 1 multiplied by u minus 2 . So, this is what is shown as to be the moment generating function for the p.d.f. that we have, now our next step is to get their mean of x , mean of y and its covariance.


Now, to get the mean of x what we have to take is that, we have to take that expectation of x and just now we have shown the step, how to get the mean of their marginal distribution that is we have to put the suitable order, so here it is very specifically mentioned that how we are putting the order.

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Example...Contd.

□ $E(X)$

$$E(X) = \frac{\partial^{1+0} M_{XY}(t, u)}{\partial t^1 \partial u^0} = \frac{\partial}{\partial t} \frac{2}{(t-1)(u-2)}$$
$$= \frac{-2}{(u-2)(t-1)^2} \Big|_{t=0, u=0} = 1$$

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So, expectation of x when you are talking about we are putting that the order of this derivative should be for the u for the t , it should be 1 and for the u it should be 0. So, basically we have we are doing a first order derivative with respect to u , this will be t so with respect to t , because t is here for this for this x . So, first order derivative of this moment generating function with respect to t this is wrong, this is t . Here, you can see that $d t$ power 1 and $d u$ power 0, so this is $d t$.

Now, if you do this one and at the end what we have to do, we have to put that t equals to 0 u equals to 0 and if we put this 1 with this expression then we are getting it to be 1. Now, we can see that marginal distribution of the x , then exponential distribution we have seen that 1 and that is e power minus x where that means the λ equals to 1. So, mean is 1 by λ that is 1 by 1 , which is 1 . So, from the moment generating function also we are getting that its mean of the random variable x is 1 .


So, now we will get what is the mean of this x and we know that before I go to the next slide, we know that that mean should be equals to 2.

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Example...Contd.

□ $E(Y)$

$$E(Y) = \frac{\partial^{0+1} M_{XY}(t, u)}{\partial t^0 \partial u^1} = \frac{\partial}{\partial u} \frac{2}{(t-1)(u-2)}$$
$$= \frac{-2}{(u-2)^2(t-1)} \Big|_{t=0, u=0} = \frac{1}{2}$$

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So, how we are putting this once again following the same procedure that is here we are putting that t equals to 0 and u equals to 1. So, this is again the first order derivative with respect to u , so again this is just between this should be u . So, here you can see it from here that should be this should be u , we are doing the first order derivative of the moment generating function with respect to u and after doing so what we are getting is the value at t equals to 0 u equals to 0 and we are getting the value to be half.


So, which is again as for our experience that if the marginal density is $2t$ power minus 2 x , that is it is an exponential distribution with lambda parameter 2 and its means should be $1/\lambda$ which is equals to $1/2$, so this is what is there.

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Example...Contd.

□ $E(XY)$

$$E(XY) = \frac{\partial^{1+1} M_{XY}(t, u)}{\partial t^1 \partial u^1} = \frac{\partial^2}{(t-1)(u-2)}$$
$$= \frac{2}{(u-2)^2 (t-1)^2} \Big|_{t=0, u=0} = \frac{1}{2}$$

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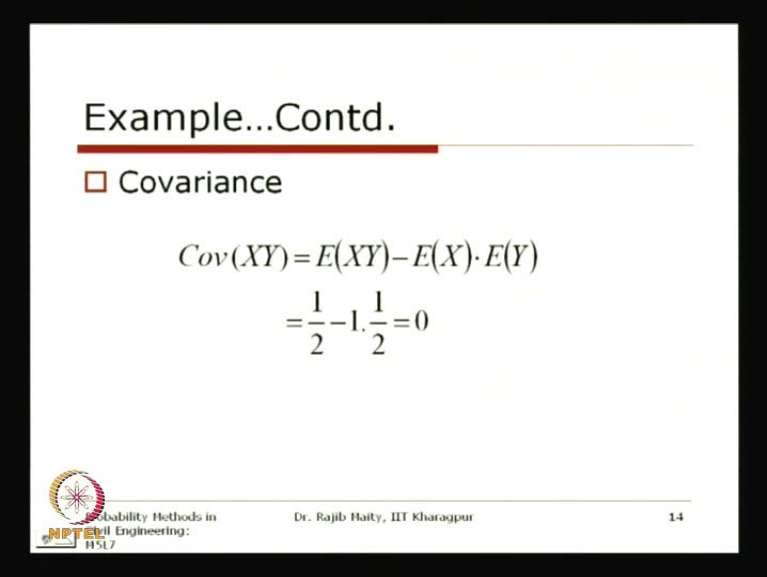
Now, what you have to do we have to calculate its covariance and you know the covariance to get the covariance. First of all, we have to get that expectation of that joint moment that is covariance is covariance between x and y is equals to expectation x y minus expectation of x multiplied by expectation of y, so that expression we know.

So, if we know that expression we have to first get this quantity x y and this one we have done it earlier with respect to the p d f, in fact all this things that is the expectation of x, expectation of y, expectation of x y we did it with the help of the p d f. But here what you are what we have to do now we are showing that how we can use the moment generating function.

So, here that expectation of x y, that means here the both that r that notation that we used minutes before the r equals to 1 and s equals to 1. So, this is that second order derivative with respect to t and d u both. So, this that expression that is the derivative of this moment generating function with respect to t as well as with respect to u. so, this will be d² this will be d² here, so 1 plus 1 we can see as this will be d².

So, if we do this derivative when we will get this expression and again we are we will be evaluate this function at t equals to 0 and u equals to 0, which will yield the results to be half. So, expectation of x y is equals to half. Now, the second one we already know that what is expectation of x and what is expectation of y, the covariance is just this minus product of those.


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Example...Contd.

□ Covariance

$$\begin{aligned} \text{Cov}(XY) &= E(XY) - E(X) \cdot E(Y) \\ &= \frac{1}{2} - 1 \cdot \frac{1}{2} = 0 \end{aligned}$$

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So the covariance is the expectation of $x y$ minus expectation of x into expectation of y , we just simply put whatever we got and we can see that this covariance is 0. Now, covariance is 0 as, I mentioned earlier also again I mentioned, here that the covariance 0 does not mean that random variables are independent but opposite is true.


If the random variables are independent covariance is 0. We started this problem we knew that this to random variables x and y are independent. So the expected are what we should get I cannot use the word expected, because this is a statistical word that we are discussing now. So we knew that covariance will come 0 and that is what we have seen, when we are getting those estimates from the moment generating function directly.

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Bivariate Exponential Distribution

- Bivariate exponential distribution
 - The pdf and cdf of the bivariate exponential distribution is:
$$f_{X,Y}(x,y) = \{(u + wy)(v + wx)\}e^{-ux - vy - wxy}$$
$$F_{X,Y}(x,y) = 1 - e^{-ux} - e^{-vy} + e^{-ux - vy - wxy};$$

for $x, y \geq 0; u, v > 0$ and $0 \leq w \leq uv$
 - Various examples related to Bivariate exponential distribution are discussed.



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Next, we will take up two bivariate standard distributions which are there and in fact that bivariate exponential distribution that we are going to discuss now. We have used this one earlier also and in case of this bivariate joint distribution, we are having only two random variables. If not only two there maybe, but these are the two things that are mostly used that bivariate exponential distribution and next is the bivariate normal distribution. This we will discuss a little and afterwards we will go for this multivariate one and again we will come back to this, how to get the joint distribution from the marginal through copula later.

So, now the bivariate exponential distribution here this is a pdf and cdf of this bivariate exponential distribution is expressed like this that this is the pdf, you know that $f_{X,Y}(x,y)$ is equal to that $u + wy$ multiplied by $v + wx$ multiplied by exponential of minus $ux - vy - wxy$ and if we do their if you want to know their cdf that is cumulative distribution function.

Then this is equal to $1 - e^{-ux} - e^{-vy} + e^{-ux - vy - wxy}$. Now range for this x and y of the random variables for which the range is greater than equal to 0. So it cannot take the negative value the support for this distribution is from 0 to infinity.


The parameters, there are three parameters we can see one is u , v and other one is w . So, this u and v is straightforward which is the nonnegative parameter this is greater than 0. So, far as the w is concern it is having it is limited by two things. One is that it should be greater than or equal to 0 and other 1 is that it should be lower than this u , v . We took this problem earlier, just to know that what the possible range of this third parameter w is. Now we can see here if the w can take from the 0 to u , v so upper limit we have seen the lower limit, if the w is equals to 0 then this component u will not be there.

So, that case basically when we take that two random variables x and y are independent, then this component cancels out, so that just now you have seen. So, this w parameter is basically taken care of if there is any means if there is any dependence, between x and y that is being taken care through this w . So we have discussed this problem just now also and earlier also and we have seen their parameters earlier. So, this is the one example of this bivariate distribution which is exponential and other one is this bivariate normal distribution.

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Bivariate Normal Distribution

- ☐ Bivariate Normal Distribution is an example of joint density function of two continuous RVs, say X and Y .
- ☐ In such cases, marginal density functions of both the RVs are also Normal distribution.
- ☐ However, reverse is not true, i.e., if both X and Y are Normally distributed their joint distribution not necessarily be joint Normal distribution.



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This is having many fields we use this whenever even earlier that, whenever we are talking about this joint distribution. We generally go for this bivariate normal distribution, because this it is grounded on the form theoretical basis and to get for the other distribution, other joint

distribution is not that easy and so we have found earlier there is a tremendous application of this bivariate normal distribution in many research fields including civil engineering.

So, this bivariate normal distribution, when we are talking this is then an example of the joint density function of two continuous random variables say that x and y and in such cases the marginal density function of both the random variables are also normal distribution. So, you remember that first we are stating this and after that we are stating this statement, that is first it is if their joint density function is Gaussian means Gaussian or normal distribution, then we can say that their marginals are also normal distribution.

However, the reverse is not true that is if both x and y are normally distributed their joint distribution not necessarily be joint normal distribution, so this not necessarily be joint normal distribution means that the reverse is not true, what we are we have also mention this point earlier as well.


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Bivariate Normal Distribution

□ The joint pdf is expressed as:

$$f_{X,Y}(x,y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}} \exp \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_X}{\sigma_X} \right)^2 - 2\rho \frac{x-\mu_X}{\sigma_X} \frac{y-\mu_Y}{\sigma_Y} + \left(\frac{y-\mu_Y}{\sigma_Y} \right)^2 \right] \right\} \quad -\infty < x, y < \infty$$

where, μ_X is the mean of X ; μ_Y is the mean of Y ; σ_X^2 is the variance of X ; σ_Y^2 variance of Y and ρ is the correlation coefficient between X and Y



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How this p d f is it can be express as like this? So, there are basically five parameters for this one, you know for a normal distribution there are there are two parameters, one is that its mean and another one is that variance.

So, here there are two random variables are there so that for one random variable, this will be the mean of 1 random variable variance of that random variable and mean and variance of the other random variable and their joint dependence. So, there are total 5 parameters will be there and this 5 parameters are here you can see that this that sigma x is the standard deviation for the x, sigma y is the standard deviation of the y, this rho is the correlation coefficient that we discuss earlier. So, this rho is the correlation coefficient between x and y, then this mu x is the mean of this random variable x and this mu y is the mean of the random variable y.

Now, just to read out this one is that $\frac{1}{2} \pi \sigma_x \sigma_y \sqrt{1 - \rho^2}$ exponential of there are three parameters, that there are three parts of this what is there inside this exponential. So, first one this is taken common that is minus half by this $1 - \rho^2$ after this this is multiplied with the three different parts.

The first one is this $(x - \mu_x)^2 / \sigma_x^2$. We can see here this is written in such a way that you can see now this is also can be taken to the reduced random variable, that is transformed that is $x - \mu_x$ by sigma x. So, this is why how it is written just in a minute we will take up that one. So, this is $(x - \mu_x)^2 / \sigma_x^2$ and third component is the $(y - \mu_y)^2 / \sigma_y^2$ and second component is their joint again that is $2 \rho (x - \mu_x) / \sigma_x (y - \mu_y) / \sigma_y$.

Now, you can see here it that if they are independent that is if the x and y are independent then we know that this will be the product of their marginals. So, that is the check for this independence, so if they are independent then you know the covariance will be 0. Then if the covariance is 0 then correlation coefficient is also will be 0. You put that correlation coefficient that is rho equals to 0, here then we can easily separate out two distributions, which are the first one will be the normal distribution for x and other one will be the normal distribution for y.

So, this we can say so this rho is basically taken care of, if there is any dependence between x and y and also you know that x and y are the support is from the minus infinity to the plus infinity. So, whatever I told it is written here that is mu x is the mean of x mu y is the mean of y sigma x square is the variance of x sigma y square is the variance of y rho is the correlation coefficient between x and y.


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Bivariate Normal Distribution

□ Now taking $Z_1 = (X - \mu_X) / \sigma_X$ and $Z_2 = (Y - \mu_Y) / \sigma_Y$;
The pdf for bivariate standard normal distribution is expressed as:

$$f_{Z_1, Z_2}(z_1, z_2) = \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left[-\frac{(z_1^2 - 2\rho z_1 z_2 + z_2^2)}{2(1-\rho^2)}\right]$$

for $-\infty < z_1, z_2 < \infty$



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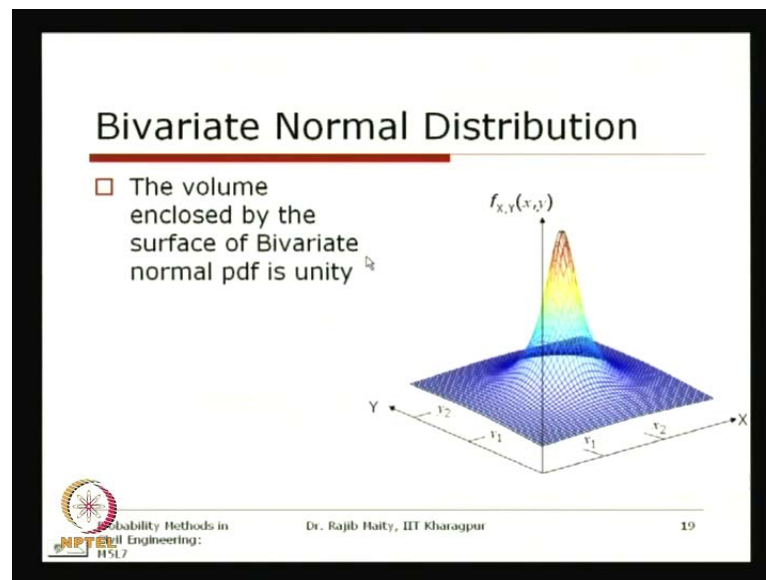
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Now, just what I was talking that if you just replace this on that replace, means that if I put that another random variable z_1 which is the $x - \mu_x$ by σ_x and z_2 is the $y - \mu_y$ by σ_y . The pdf of the bivariate standard normal distribution, so earlier we have seen that after this transformation, we are getting the standard normal distribution here also we are getting that is again bivariate.

So, this bivariate standard normal distribution if we want to know then for this expression you have to put that z_1 . So, this then they are pdf joint will be $1 / (2\pi\sqrt{1-\rho^2})$ exponential of minus $z_1^2 - 2\rho z_1 z_2 + z_2^2$ divided by $2(1-\rho^2)$. For which again this z_1, z_2 are also not from minus infinity to plus infinity this is from 0 to 1 this is from 0 to 1.

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This is in an example, just to view that how the joint distribution is the joint distribution looks like; here we have taken one example of this joint normal distribution. We can see that now this is kind of surface and that is a three-dimensional surface and the volume below this surface that is a three-dimensional volume and that volume if we just extend the idea of this that, what we have discuss for this single random variable that area below the curve for a single random variable is equals to one.

Herein case of the joint, if I take the bivariate case, because this pictorial representation is possible only up to the bivariate case this is x and y . So, this is the kind of surface and there is a volume is there between the $x-y$ plane and that surface.


So, what we are talking if you refer to this figure is this is this is your x axis, this is your y axis, so this is your $x-y$ plane. So, the volume in between this surface and that $x-y$ plane should be equals to 1, just to satisfy what the requirement for a p d f. That is the total volume that is between that is covered by this surface is equals to 1. So, volume of enclosed by the surface of bivariate normal p d f is unity is 1 is shown here.

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Bivariate Normal Distribution

□ The Bivariate normal density function equation in the previous slide can be expanded as:

$$f_{X,Y}(x,y) = \frac{1}{\sqrt{2\pi}\sigma_X} \exp\left[-\frac{1}{2}\left(\frac{x-\mu_X}{\sigma_X}\right)^2\right] \cdot \frac{1}{\sqrt{2\pi}\sigma_Y\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2}\left\{\frac{y-\mu_Y - \rho(\sigma_Y/\sigma_X)(x-\mu_X)}{\sigma_Y\sqrt{1-\rho^2}}\right\}^2\right]$$



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Now, this one we will just express this in a similar way just after doing some algebraic manipulation just we will see that why we need this one. So, the same expression so this expression that I am going to tell you now and the joint distribution that is between the x and y are same. What we are doing is that, we are first we are trying to separate out what is the marginal distribution of x. So, marginal distribution of x which is having the mean μ_X and variance σ_X^2 , then we can we know that what is this distribution by this time.

So, if we just want to take out that component out what is this remaining, that is what can be done in few steps and this is what is done. So, we first of all what we are writing is this $1/\sqrt{2\pi}\sigma_X \exp[-\frac{1}{2}(\frac{x-\mu_X}{\sigma_X})^2]$. So, this is nothing but the marginal distribution of x and this quantity is taken out and what is remaining we are just expressing it here that $1/(\sqrt{2\pi}\sigma_Y\sqrt{1-\rho^2}) \exp[-\frac{1}{2}(\frac{y-\mu_Y - \rho(\sigma_Y/\sigma_X)(x-\mu_X)}{\sigma_Y\sqrt{1-\rho^2}})^2]$. So, we consider this is the outside constant and this is another constant, if we multiply this 2 it is coming that $2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}$.

So, similarly in that way so whatever was the inside, we have just taking we have taken this one out so what is remaining is noted here. Why this is important, why we are explaining this, why we are expressing this same expression in a different way.

We have just to know their what is the conditional distribution and you know that we have discuss for the bivariate case the conditional distribution is expressed and this has to be just divided by their marginal densities. So, that is why the marginal density is taken out.

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Bivariate Normal Distribution

- Now the conditional density function X given Y=y is as:

$$f_{Y|X}(y|x) = \frac{1}{\sqrt{2\pi}\sigma_Y\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2}\left\{\frac{y-\mu_Y-\rho(\sigma_Y/\sigma_X)(x-\mu_X)}{\sigma_Y\sqrt{1-\rho^2}}\right\}^2\right]$$

- Marginal density function of X is:

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma_X} \exp\left[-\frac{1}{2}\left(\frac{x-\mu_X}{\sigma_X}\right)^2\right]$$



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And if I just want to know, what is the conditional density function for this x given y equals to this. So, this f y given x is equals to so this will be the joint density divided by the marginal density of x.

So, now you can see that this is just said this is wrong so this will be the conditional density function of y given x equals to x. So, this is what is expressed here also so this conditional distribution of y given x is that the joint density divided by the marginal density of x now what we will get. So, this is the marginal density of the x, so this has to be taken out. So, what is remaining here is nothing but the conditional density of y given x.

So, this is what is explained here now what you can see here is that this is also taking the form of a form of a normal distribution as well. How it is taking the form, because you can see that this is square root of 2 pi sigma y this 1 so this pull component, that is the standard deviation that is sigma y multiplied by square root of 1 minus rho square is the standard deviation of the conditional distribution of y given x.

Similarly, exponential half y minus this full quantity is the mean of the conditional distribution of y given x. So, the unconditional mean of y is μ_y , now on the conditional mean is μ_y minus ρ multiplied by σ_y by σ_x multiplied by x minus μ_x so we can see that the conditional mean is changing with respect to that it is dependent on the value of the x it is here and similarly $\sigma_y^2 (1 - \rho^2)$ which is the standard deviation of that conditional distribution.

Marginal density function of x here is so that this is the marginal density function of the x which we have divided it by to get this conditional density. How to get the conditional density, that we discuss earlier you can recall that $f_{Y|X}(y|x)$ is equal to $f_{X,Y}(x,y)$ divided by $f_X(x)$, that is that joint divided by the marginal. So this marginal was divided by to get this conditional.

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Bivariate Normal Distribution

□ Also the conditional density function has the mean as:

$$E(Y | X = x) = \mu_y + \rho \left(\frac{\sigma_y}{\sigma_x} \right) (x - \mu_x)$$

□ Variance as:

$$Var(Y | X = x) = \sigma_y^2 (1 - \rho^2)$$



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So, once again we will just see what is their conditional mean and conditional variance. So this conditional mean is this that is conditional mean of this y given x equals to x is equals to μ_y plus ρ multiplied by σ_y by σ_x multiplied by x minus μ_x . Just what we have seen it from this expression, this expression is also the Gaussian distribution, that is normal distribution and from this normal distribution form, we can say that this component is the mean of that conditional distribution, which is basically we are taking this minus common.

That is why this sign becomes plus so this $\mu_y + \rho \sigma_y \text{ by } \sigma_x$ multiplied by x minus μ_x , this is the conditional mean, which is shown here and the conditional variance is the $\sigma_y^2 (1 - \rho^2)$ variance we are talking about, because this is the standard deviation so this conditional variance is $\sigma_y^2 (1 - \rho^2)$.

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Multivariate Random Variables

- So far we have discussed with two random variables and their distribution functions.
- In many real case we are interested in more than two random variable representing different quantities of interest from the same experiment and the same sample space
- Concept can be extended for multivariate RVs having more than two RVs



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Now, as when we started this lecture that we will take some of this bivariate concept to this multivariate one, the concept can be easily extended so we will discuss about that one with the context of this multivariate random variable having more than two random variables. Basically, we will try to generalize it in terms of that n random variable involved.

So, far we have discussed with the two random variables and their distribution function their different properties and etcetera. We have discussed we have taken of some problems also, but in many real cases we are interested in more than two random variables, representing different quantities of interest from the same experiment and the same sample space.

So, what is the volume, what is the total traffic volume at the junction, if you want to know then this kind of analysis is there were the number of random variables involved will be more. Similarly the different source of the fresh water is joining to a reservoir; it is draining to a reservoir.


So, if we just want to know their joint behavior there also this kind of application can be found out. So, now to do this on the concept that we have discussed for this bivariate case can be extended for the multivariate random variables as well having more than two random variables.

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Joint Probability Distribution

□ **n-dimensional discrete random variable**

■ If X_1, X_2, \dots, X_n are random variables defined on the same probability space, then (X_1, X_2, \dots, X_n) is defined as a n-dimensional discrete random variable if it can take values only at a countable number of points (x_1, x_2, \dots, x_n) . These variables (X_1, X_2, \dots, X_n) are said as joint discrete random variables.



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So, now discuss this things in the context of the general framework that is the n dimensional discrete random variables, if we want to discuss first so this is n dimensional means there are n random variables n triple cases are we have to consider jointly and these are the discrete random variable, in case of discrete random variable we are discussing.

So, if x_1, x_2 up to x_n of the random variable defined on the same probability space, then x_1, x_2 up to x_n is defined as an n dimensional discrete random variable, if it can take values only at a countable number of points x_1, x_2 up to x_n . These variables x_1, x_2, x_n are said to be the joint discrete random variables.

Basically, means what we discuss earlier is that for a two dimensional case, just take x_1 and x_2 there are two axes orthogonal to each other and take some points. So, it can take only few points on that plane made by two axes x_1 and x_2 . Now just extend this on to this n dimensional space, where there are n axes which are x_1, x_2 up to x_n and take some point which point will be should be denoted by the n coordinates. Now, n coordinates means x_1 equals to this x_2 equals to this x_3 equals to this like this up to x_n equals to this.

Now, those few those specific points are consist that probability space. So, this if I take the full I should not say then volume, because volume is again the three dimensional. It can go even more than three dimensional, so for those specific cases, if the probability mass is concentrated then that is known as to be that multivariate joint discrete random variable.


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Joint Probability Mass Function (Joint PMF)

□ **Definition:**

- The joint pmf of a n-dimensional random variable (X_1, X_2, \dots, X_n) is defined as the intersection probability of the n sequence of events $(X_1=x_1), (X_2=x_2), \dots, (X_n=x_n)$ if (x_1, x_2, \dots, x_n) are points in the n-dimensional sample space of this variable, and 0 otherwise

$$P_{X_1, \dots, X_n}(x_1, \dots, x_n) = P(X_1 = x_1, \dots, X_n = x_n) \quad \forall (x_1, \dots, x_n) \in S$$



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Now, know that what is there probability involve is this one, so that the joint PMF, because you know this is discrete. So, we have to discuss this one with respect to the PMF probability mass function of that n dimensional random variable that is x_1, x_2 up to x_n is defined as the intersection probabilities of the n sequence of the event.

So, this n sequence of the event is this that is if x_1 equals to x_1 that is now, you can know that these are the random variable this is specific value of the random variable. So, this x_1 takes the specific value x_1 x_2 take specific value x_2 and up to like this up to the n takes this specific value x_n . So, for this joint case means all this things are occurring together. So, this is that if that x_1, x_2, x_n are the points in the n dimensional sample space of this variable. So, that probability is taken only for those few those specific points otherwise it is 0.

Then, that is the joint PMF between the random variables x_1, x_2, x_n and this is what is shown here that is the probability mass function is defined as that for this x_1, x_2, x_n is the probability

of the joint occurrence of all these events and this belongs to that probability space. This outcome, the specific outcome, joint outcome, belongs to that sample space.

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Joint pmf...Contd.


□ Properties

$$p_{X_1, \dots, X_n}(x_1, \dots, x_n) \geq 0$$

$$\sum_{x_1, \dots, x_n} p_{X_1, \dots, X_n}(x_1, \dots, x_n) = 1$$

□ Cumulative distribution function

$$F_{X_1, \dots, X_n}(x, \dots, x) = \sum_{x_1, \dots, x_n \leq x} p_{X_1, \dots, X_n}(x_1, \dots, x_n)$$



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Now properties again extending the same idea that, it should be the total if I take this each and every point. If I take what are the probabilities so that probabilities should be greater than or equal to 0, so it should not be then a negative number. Again if I just sum up for all the probability masses for all those feasible points then it should be equal to 1.

These are two basic criteria to satisfy the condition for a valid PMF that we have discussed earlier for the single random variable. For the bivariate random variable, now here it is for this multivariate for the n number of cases.

Keeping the basic conception, these are the two properties to satisfy the condition of a PMF and the cumulative distribution function is that you know that cumulative distribution function here, just I will just take a minute that all these attributes now I have taken for this x_1, x_2, \dots, x_n . Why we have taken I will just explain in a minute, this could have again been also that that different activity that is for x_1 it is x_1 , for x_2 it is x_2 that also it can be done.

But just to show herein this case if we are talking about only a single point in that probability space, then it is the summation of all the feasible points, those are below this specific number x

that is x_1 less than equals to x_1 , x_2 less than equals to x_2 up to this 1 that x_n less than equals to x_n . If I sum up this 1 then I will get the cumulative distribution function for that joint PMF in case of the n random variables involve in that multivariate random variable case.

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Joint probability density function (Joint pdf)

□ Similar to the discrete variables, if X_1, X_2, \dots, X_n are n random variables defined on the same probability space, then (X_1, X_2, \dots, X_n) is a n -dimensional continuous random variable if and only if there exists a function such that:

$$f_{X_1, \dots, X_n}(x_1, \dots, x_n) \geq 0$$

and

$$\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_{X_1, \dots, X_n}(x_1, \dots, x_n) dx_1 \dots dx_n = 1$$



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Now for the same is, if I take for this continuous random variables also if this x_1, x_2, x_n are the continuous random variable that is similar to this discrete random variable. If these are the n random variables defined on the same probability space then x_1, x_2, x_n is an n -dimensional continuous random variable.


If and only if there exist a function such that these are the two conditions again, you know that for the function should be a for all for the entire range of this random variable should be a non-negative number, should be greater than equal to 0 and their integration over the entire range of each random variable should be equals to 1.

And you know that this should be n , so dx_1 up to in this way up to dx_n . So, n number of this integration in this n different direction of this random variable, so dx_1, dx_2 up to this it is dx_n so these are the two conditions that should be satisfied.

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Joint cdf

□ The joint distribution of n continuous random variable can be completely described with their joint cdf as:

$$F_{X_1, \dots, X_n}(x_1, \dots, x_n) = P[(-\infty \leq X_1 \leq x_1) \cap \dots \cap (-\infty \leq X_n \leq x_n)]$$
$$= \int_{-\infty}^{x_1} \dots \int_{-\infty}^{x_n} f_{X_1, \dots, X_n}(\alpha_1, \dots, \alpha_n) d\alpha_1 \dots d\alpha_n$$


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Then joint c d f again, if we are talking of about then here we have just taken as that is a even more general case, for the discrete also that can be the general case, but we have discuss it thein a specific condition.

So, here where we are taking that the x_1 is the specific value up to this x_1 and like this, if we want to know what is their cumulative. What is joint cumulative distribution function is that probability of the x_1 from minus infinity to x_1 , intersection x_2 from minus infinity to x_2 like this up to x_n from the minus infinity to that x_n .

So, if we take this joint occurrence this intersection means, we know that is the joint occurrence of all this events, then this one you know that. We will get it from this integration that is from minus infinity to x_1 for this $d x_1$, this here that is that non variable is represented as α_1 . So, this $d \alpha_1$ like this n times integration up to this $d x_n$. This can go in this way to get this joint c d f for this joint multivariate case.


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Example

Q. The joint density of random variables is given by:

$$f_{I,J,K,L}(i,j,k,l) = \begin{cases} \frac{1}{2}(i+j+k+l) & 0 < i,j,k,l < 1 \\ 0 & \text{otherwise} \end{cases}$$

Calculate the probability $P\left(I < \frac{1}{2}, J < \frac{3}{4}, L > \frac{1}{2}\right)$

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One quick example, we will take this is theirs involve the four random variables i, j, k and l . This is the p d f is defined by and we have to calculate the probability of less than half j less than $\frac{3}{4}$ then l greater than half. So there is no so for k it is not defined anything.

So, just this is you can so here now on nothing that we have not that we have not shown here you can check yourself that whether this is a valid p d f or not. Now to check whether this is a valid p d f or not you have to do that integration for the range between given 1 for all the random variables four times integration, you have to do and you have to put their integration limit and check whether the final is coming to be 1 or not. So, here what you are trying to look at that for this specific space what is the probability.

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Example...Contd.

Sol.:

The joint pdf can be used to compute probabilities as:

$$\begin{aligned} P\left(I < \frac{1}{2}, J < \frac{3}{2}, K < 1, L > \frac{1}{2}\right) \\ &= \int_{\frac{1}{2}}^1 \int_0^1 \int_0^{\frac{1}{2}} \int_0^{\frac{1}{2}} \frac{1}{2} (i + j + k + l) di dj dk dl \\ &= \int_{\frac{1}{2}}^1 \int_0^1 \int_0^{\frac{1}{2}} \frac{1}{2} \left(\frac{1}{8} + \frac{j}{2} + \frac{k}{2} + \frac{l}{2}\right) dj dk dl \end{aligned}$$




Simply, we have to do the integration for this specific range for this 1, you know that for this half for the j, it is the range is the half to 1 and for the i it is the half to 1 for j this is 0 to 1 for this k. For this i it is 0 to half, for this j this is 0 to 2, third for k it is the entire range 0 to 1 and for the l it is half to one.

Because, we can see that then it is satisfying this limit. So, the inner 1 the first integration is for this limit for i which is less than half and the lower range always is 0, so it is 0 to half for i. Similarly, we have defined this integration limits this is the pdf and we are doing this integration.

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Example...Contd.

$$\begin{aligned} &= \int_{1/2}^1 \int_0^1 \frac{1}{2} \left(\frac{3}{16} + \frac{9}{16} + \frac{3k}{4} + \frac{3l}{4} \right) dk dl \\ &= \int_{1/2}^1 \frac{1}{2} \left(\frac{3}{4} + \frac{3}{8} + \frac{3l}{4} \right) dl \\ &= \frac{1}{2} \left(\frac{3l}{4} + \frac{3l}{8} + \frac{3l^2}{8} \right) \bigg|_{1/2}^1 = \frac{27}{64} \end{aligned}$$



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To occur doing this integration we will gradually get, what is their joint now what is the joint probability is 27 by 64. So, in this lecture we have discussed about their joint moment generating function for multiple random variable. Then we have taken that two standard bivariate joint p d f and c d f bivariate exponential distribution and bivariate normal distribution.

After that we have taken the multivariate cases and we have discussed one problem as well. So, in the next lecture we will discuss some more of these properties and we will also discuss about the functions of the multivariate random variable case and after that we will again discuss about this multivariate case to get their joint p d f from the marginal one using copula thank you.