

Probability Methods in Civil Engineering
Prof. Dr. Rajib Maity
Department of Civil Engineering
Indian Institute of Technology Kharagpur

Module No #01

Lecture No #16

Functions of Random Variables- Different Methods (Contd)

Welcome to this lecture, today we are taking this 3rd lecture of this module. In this module, that means the 4th module, we are discussing about the **function of** functions of random variable. So, basically in the previous modules we have seen the different properties of random variable and their standard distributions, we have discussed. Now, **from the standard** from the standard available random variable, if we generate or if we derive some other function of those random variables, we have discussed that those functions are also a random variable. So, their properties of those functions also we will follow the similar properties of those random variable. So, in this module we are discussing this aspect of this property of this functions of random variables.

Now to start with, what we have done is that, we have discussed the fundamental theory, that is how one particular random variable can be linked to the another one through their functional dependence. So, if I know one random variable **which is** for which the all the properties, that is the distribution is known and we know its functional relation with another random variable, then for that new random variable that is the function of that, how to get different properties.

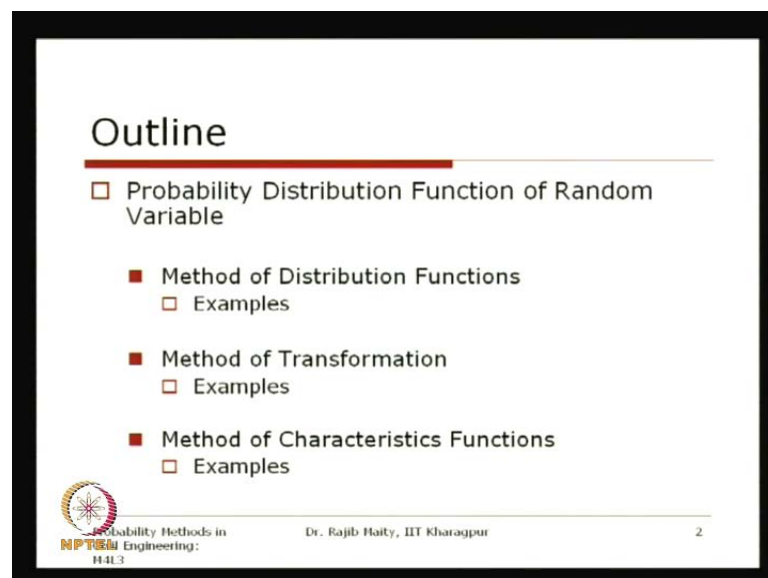
Now, we know that if you want to know the properties of this function, then we have to first of all know what is their probability distribution and the probability distribution to get that in the previous lectures, we have shown some fundamental theorem; and that fundamental theorem, based on the fundamental theorem, we have kept the basic concept same we have derived different methods.

So, in the So, in this lecture also we will continue two different methods, one method is discussed in the earlier lecture and in this lecture also we will continue with two more different methods. So, our basic goal is to know, what the probability density of this new

random variable and once we know this probability density function, then we know that, we can estimate whatever the properties that we need to know, we can estimate that one.

So, **the in the** in the previous two lectures in this module, first we have seen the fundamental theorem and from there we have discuss about the method of distribution function first and in today's lecture, we will discuss about that method of transformation. So, even though we are giving the different names, I repeat that all these methods are based on the same fundamental theorem only with some special assumptions has been taken for different method, that is why the methods look different.

(Refer Slide Time: 03:34)



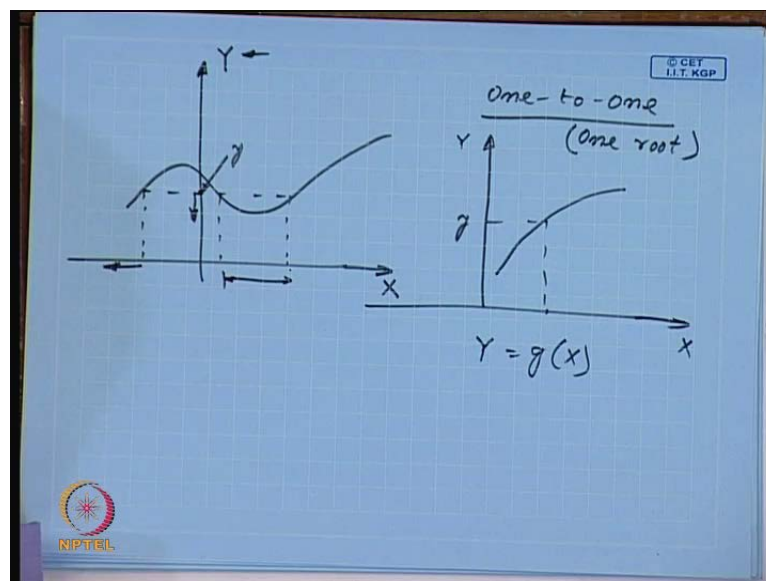
So, **we will** if I just see the different methods that we have discuss earlier, first in **the in** the earlier lecture, we have discuss that method of distribution function with some examples. And today basically we will discuss about this method of transformation with examples. Now, when we are talking about this method of transformation, basically what we are talking that, we are talking about that transformation from one random variable to another one. And here the basic assumption that we are following is that, the transformation is known, that is known in the sense that, the unique transformation is generally assume that is one to one transformation. So, for a particular value of **the of the** original random variable, there is one and only one value for the new random variable, that is derived random variable. So, this transformation is known as one to one.

So, when we are when we discuss that method of transformation, we will see that how this method of transformation is linked, it is, they can be derived from the fundamental theorem for this one to one transformation. A little bit of discussion was there in the previous lecture as well, so will follow that one in details in this **in this** lecture.

And finally, there is another method that is known as this method of characteristics function. So, we know that for the **for the** standard random variables, if we know the PDF then, we can estimate their characteristics function. **So** And we have seen that, usefulness of this characteristics function is that, once we can identify the characteristics function, then using their relationship with the different order of moment, that is first order moment, second order moment, if you can link those things then those if those moments are known, if those initial then the **then the** properties, their location, their central tendency, their dispersion, so these properties also will be known.

So, now if we know the characteristics function, then from that functional relationship for this derived random variable, we can also get the probability density function for the **for the** functions of the random variable, so that we will discuss also under the method of characteristics function. So, before we proceed we will quickly see that, what we have seen in the fundamental theorem as well as for this method of distribution function, we will just recall, so that it will help us to understand that, what is there in the **in the** method of transformation.

(Refer Slide Time: 06:29)



So, what we have seen earlier is that, if this is one random variable, which is the original random variable denoted by x and there is another random variable, which we can derive through some functional relationship with the original random variable x . So, if the if the functions can have any this type of this type of functional relationship, then for a particular value of y , what we have to do, we have to find out the what is the representative set in this in the original random variable that is your x here.

So, if I know this one, then the then the probability of this y less than this specific value y , so towards this one, then this corresponds to this zone as well as the zone between this one. So, this thing if you just add, if we get then in this way we can get what is the what is the distribution of this y , for each and every possible value of this of this new random variable y . Now, when we are talking about this method of transformation, first case that we are considering is that one to one. So, when we are talking about this one to one transformation, that means that, the relationship is such that the for a specific value of y , there will be only one possible only one possible set for this original random variable original random variable x can can have.

So, for this one, for this kind of relationship, from the fundamental theorem we have seen that, this will be just only one root, so what it is reduced to the fact that, it then in the relation to the fundamental theorem is that, only one possible root for a specific value of y . So, with that root, with the fundamental theorem reduced to only one only only one root and using that relationship, we will see how we can get the PDF for this new random variable y . Here, we are assuming that functional relationship, that $g(x)$ is having only only single root for a specific value of y .


(Refer Slide Time: 09:09)

Property of one-to-one transformation

- The probability density function (pdf) of a derived variable by one-to-one transformation.
- Let X be a random variable with continuous pdf, $f_X(x)$. The pdf of the random variable Y defined by the one-to-one transformation $Y=g(X)$ is given by

$$f_Y(y) = \left| \frac{dx}{dy} \right| f_X[x] = |J| f_X[x]$$

where $h(y)$ denotes the inverse function such as $x=h(y)$ if $y=g(x)$. This property requires that the Jacobian of the transformation is nonzero.

 Stability Methods in Engineering: M4L3

Dr. Rajib Hait, IIT Kharagpur

3

So, here when we mention the assumption that we are using that we mention is that, one to one transformation. So, if the x is a random variable with continuous PDF, having the PDF is $f_X(x)$. The PDF of the random variable y , which is defined by this one to one transformation, this one to one transformation is important, that what we are mention here is that y is equals to $g(x)$.

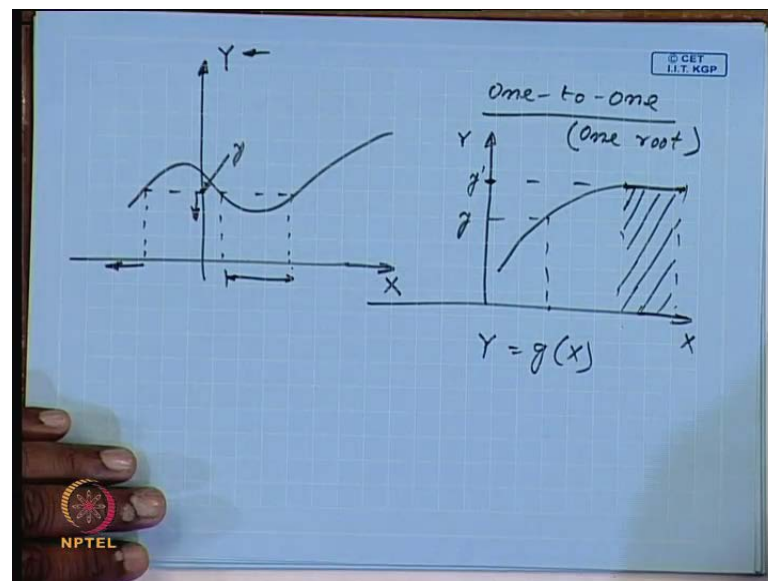
So, this is the functional relationship that is having between two random variable x and y and for this random variable, this distribution is known that is $f_X(x)$, then from the fundamental theorem, we get that this $f_Y(y)$ should be equals to this $f_X(x)$ and multiplying by a factor that $\frac{dx}{dy}$, this we discuss in the last lecture as well, that is this one when we are taking this derivative.

This derivative is the, we are taking the absolute value and this derivative is known as the Jacobin of this transformation, which is denoted as J . So, where this denotes the inverse function, so now, when we are **we are we are** expressing this one, that $f_X(x)$ we should express this one in terms of the new random variable y . So, which is the inverse **inverse** function such as x equals to $h(y)$, if that y is equals to $g(x)$, so **so** we can also write that x equals to as we as discuss earlier lecture that, x equals to $g^{-1}(y)$, that g^{-1} of y is here that $h(y)$.

So, this property requires that, the Jacobin of the transformation is nonzero, now this is for the continuous random variable that, what we have started with and that is why we are

we required that this transformation should be the Jacobin of the transformation should be non zero;and we have also discuss that if it is if Jacobin equals to 0 what does it mean?So, this means that there will be a spike of the **of the** probability distribution that is, there will be a discrete probability for that zone where this, this, this derivative that is a $\frac{dx}{dy}$ is equals to 0. So, $\frac{dx}{dy}$ equals to 0 means the rate of change of x with respect to y is equals to 0 that means.

(Refer Slide Time: 11:52)




So, when So, if the this thing if some functional relationship, it becomes constant for this zone, then we can say that at this point of this value if this specific value y' here, this rate of change of x with respect to y becomes 0. So, if it becomes 0 that means, the this probability, that is probability of this y is equals to y' should be **the**, it should be equal to the probability of this total area, so that means, here **this is be** this is become a discrete probability at this point some probability mass should be concentrated for this specific value of y , this is what is **what is** shown here this what explained in this in this statement.

(Refer Slide Time: 12:47)

One-to-one transformation...Contd.

□ For example, if $Y=mX+c$, the inverse function is $X=(Y-c)/m$, from which $dx/dy=1/m$. Hence,

$$f_Y(y) = \left| \frac{1}{m} \right| f_X\left(\frac{y-c}{m}\right)$$



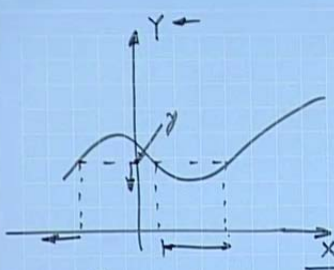
Probability Methods in
Engineering:
M4L3

Dr. Rajib Haity, IIT Kharagpur

4

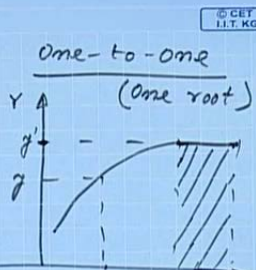
Now, if we take a very a very simple problem if you start with a very simple transformation that is, that y is equals to $m x$ plus c and you know that, this y is equals to $m x$ plus c is a is a transformation of a of a straight line now this.

(Refer Slide Time: 13:07)

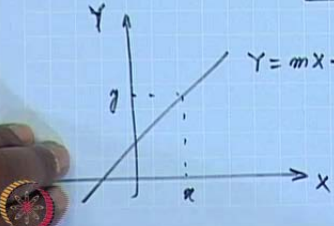


$Y = g(x)$

One-to-One
(One root)



$Y = g(x)$



$Y = mX + c$

$f_Y(y) = \left| \frac{dx}{dy} \right| f_X(x)$
 $x = \frac{y-c}{m}$
 $\frac{dx}{dy} = \frac{1}{m}$

So, the transformation will look will be a will be a straight line like this. So, where, so for this one, if this is your x and if this is your y , then the relationship is that y equals to $m x$ plus c , which is the equation for a straight line for any specific value of this y , there will be a one and only one value of this earlier that original random variable x .

Now, we know that, if we just want to know what should be the distribution of this y . So, from the equation we get that this f_y of y , that is PDF of this y is the Jacobian dx/dy of this f_x of x . Now, this x I have to write in terms of this inverse relationship, so this x will be equal to your from this relationship can be written that y minus c by m . So, this, so from here also we can also write that dx/dy is equal to your 1 by m , because c by m is constant. So, this becomes 1 by m . So, we can just replace this part here and this one here to get this f_y .

(Refer Slide Time: 14:50)

The image shows a handwritten derivation on a blue grid background. At the top right, there is a small logo that says "CET I.I.T. KGP". The derivation starts with the formula for the probability density function of y :

$$f_y(y) = \left| \frac{1}{m} \right| f_x\left(\frac{y-c}{m}\right)$$

Below this, the PDF of x is given as:

$$f_x(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

Then, the PDF of y is derived by substituting $x = \sigma y + \mu$ and $\frac{dx}{dy} = \sigma$:

$$f_y(y) = \frac{\sigma}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{\sigma y + \mu - \mu}{\sigma}\right)^2} = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$$

The final result is boxed:

$$f_y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$$

At the bottom, it specifies $\mu = 0$ and $\sigma = 1$. In the bottom left corner, there is a logo for "NPTEL".


So, in this what we will get is that f_y is equal to 1 by m . This should be taken the absolute value depending on what is this, the value of this m the slope of this straight line, this one multiplied by f_x of this y minus c by m . Now, if we know this form, that if you know the PDF of this x will just replace this value and will get the PDF of y , so this expression.

(Refer Slide Time: 15:29)

One-to-one transformation...Contd.

□ For example, if $Y=mX+c$, the inverse function is $X=(Y-c)/m$, from which $dx/dy=1/m$. Hence,

$$f_Y(y) = \left| \frac{1}{m} \right| f_X\left(\frac{y-c}{m}\right)$$

 Stability Methods in
Engineering: M4L3

Dr. Rajib Haity, IIT Kharagpur

4

So, this is what is explained here, so for a relationship, which is a straight line relationship y equals to m x plus c , then the inverse function that we are **that we are** telling in the last slide it is the y minus c by m from which the dx/dy is $1/m$ thus this $f_Y(y)$ equals to $1/m$ absolute value multiplied by f_X of y minus c by m .

(Refer Slide Time: 15:57)

Example


Q. X is a normal variate with parameters μ and σ . Determine the density function of $Y=(X-\mu)/\sigma$

Sol.:
The inverse function is: $x = \sigma y + \mu$
and $dx/dy = \sigma$, thus

$$f_Y(y) = \left| \frac{dx}{dy} \right| f_X(x)$$

$$= \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2} \frac{(\sigma y + \mu - \mu)^2}{\sigma^2}\right] \sigma = \frac{1}{\sqrt{2\pi}} e^{-y^2/2}$$

Therefore Y is a standard normal variate with density function $N(0,1)$

 Stability Methods in
Engineering: M4L3

Dr. Rajib Haity, IIT Kharagpur

5

Another example, we also can see is that standard normal distribution and normal distribution. So, here if where the x is a random variable, which is which the normal distribution is which follow the normal distribution with its parameter μ and σ . So,

if the relationship now determine the density of the function y , which is having a functional relationship with x is x minus μ by σ .

Now, when we have also discuss earlier that why this example is taken here just to show that, how we can get these one, because both this distribution we know that is x as well as y , so we know because this is standard distribution that we discuss in the last module, this is normal distribution or Gaussian distribution with two parameters μ and σ . And for a Gaussian distribution if it transfer that random variable through this relationship that is x minus μ by σ , we know that this the reduced variety which is also a normal distribution having its mean equals to 0 and standardization equals to 1. So, this is nothing but, this will follow a normal this will follow a standard normal distribution, standard normal distribution means its mean is 0 and standardization is 1. So, if this one is and we know the distribution for both the random variable x , which is a normal distribution and the y which is the standard normal distribution, so here now, the inverse relationship we will see first, where this x equals to σy plus μ , this is the inverse relationship; and from here we should calculate that Jacobian, Jacobian is your $\frac{dx}{dy}$, so here the $\frac{dx}{dy}$ is equals to σ . So, again using that same relationship that is $f_y(y)$ is equals to this absolute value of this Jacobian $\frac{dx}{dy}$ multiplied by this $f_x(x)$ of x , express in terms of the new variable y , express in terms of new variable y means we have to place this one this instead of this x , we have to use that σy plus μ .

Now, this $f_x(x)$ of x now, we know that. So, we know that for a normal distribution that $f_x(x)$ is equals to $\frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$. Now so, this has to be replaced by this as, so this y if when we are talking about that is a when this relationship, that we are talking that y is equals to x minus μ by σ and this x equals to σy plus μ just now we have shown. So, here that $\frac{dx}{dy}$ is equals to σ .

So, this $f_y(y)$ is nothing but, now that σ we should multiplied that σ that divided by this original one $\sigma \sqrt{2\pi}$ exponential half in place of this x , what we are write in is that your σy plus μ minus this μ divided by σ whole square. So, this is becoming that one by square root two π exponential minus. So, this σ cancels, so it becomes y^2 by 2. Now, so this one is that $f_y(y)$ of y we got. So, this is also a normal distribution, which is having that μ equals to that mean equals

to 0 and sigma that is standard deviation is equals to here 1, so which is nothing but, standard normal **normal** distribution.

(Refer Slide Time: 21:03)

The image shows a handwritten derivation on a grid background. At the top left, it states $X \rightarrow N(\mu, \sigma^2) \rightarrow -\infty < x < \infty$. Below this, a box contains $Y = e^X$. To the right, another box contains the PDF of Y: $f_Y(y) = \frac{1}{y\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{\log y - \mu}{\sigma}\right)^2}$ with the domain $0 < y < \infty$. The derivation continues with $X = \log Y$, where 'Normal' is written under X and 'Log-Normal' under Y. It then shows $x = \log y \Rightarrow \frac{dx}{dy} = \frac{1}{y}$. The PDF of Y is derived as $f_Y(y) = \left| \frac{dx}{dy} \right| f_X(x) = \left| \frac{1}{y} \right| f_X(\log y)$. Finally, it substitutes the normal PDF for f_X to get $f_Y(y) = \frac{1}{y\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{\log y - \mu}{\sigma}\right)^2}$ for $0 < y < \infty$. A checkmark is at the bottom right. Logos for 'CET I.I.T. KGP' and 'NPTEL' are visible.

We can also before you go to a specific civil engineering problem, we can also check one more **one more** relationship this we have discussed earlier is that, **is that** if say that x is a **is a** normal distribution with the two parameters sigma and say mu and sigma with the sigma square is the variance, for which that you know that support is from the entire real axis minus infinity to plus infinity. Now, if we have a new relationship say that y is equals to e power x, if this relationship holds then, we can also express that x equals to your log of y.

Now, if we recall **from our** from our previous module lecture that, when this x is **x is** normal distribution x is your normal distribution, that is for a random variable if I take the log it becomes normal distribution that means, this y is nothing but, your log normal distribution.

Now, if this one, so we know the distribution of x and we also know the distribution of y, so distribution of y means the log normal **log normal** distribution from our earlier modules lecture, we know that for the log normal distribution, the distribution looks like this that is $\frac{1}{y\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{\log y - \mu}{\sigma}\right)^2}$, sigma is out this root exponential minus half of this log y minus mu by sigma and this one is that 0 to y infinity, so this is a nonnegative number, so this distribution we have seen earlier.

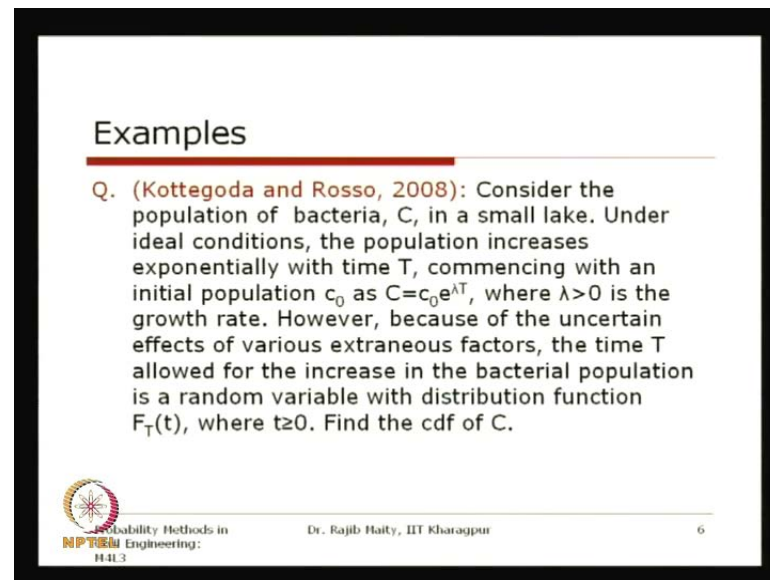
Now, if we know this one, this distribution the normal distribution if you know that, this is the functional relationship that holds between this x and y can we can we obtain this log normal distribution from this normal distribution or not that we will see from this one, because this transformation is also one to one transformation.

So, now to get this one, we have to find out that this the inverse relationship, that is x equals to $\log y$ that we have got. So, that $\frac{dx}{dy}$ equals to your $\frac{1}{y}$ now this x is a . So if I just want to know from here that, what is this. $f(y)$ is your Jacobian $\frac{dx}{dy}$ and $f(x)$ of x . Now, from this one, so this will be your $\frac{1}{y}$ $\frac{1}{y}$ and multiplied by $\frac{1}{y}$ this one we know that this is normal distribution square root 2π .

So, this x now will be replaced by your this $\log y$, so I can just write that once more one more step, that is $f(x)$ is of $\log y$; so, which is nothing but, $\frac{1}{y}$ by this y comes square root 2π sigma from this normal distribution e power minus half $\log y$ minus μ , so this $\log y$ we are writing in place of this x minus sigma whole square this zone as we know that this we are taking this relationship this one is valid for this 0 to infinity as we have taken this relationship. So, this can never come below 0 , so this is now this one is nothing but is equals to this one.


So, if this kind of relationship holds, we can see that if x is the normal distribution y is your y is your log normal log normal distribution, which we have got from this one to one transformation method and using this Jacobian of this transformation. So, now. So, this is the standard, so all this distribution just now what we discuss is that normal then standard normal distribution and after that we have discuss that from the normal to this log normal distribution. So, all this distribution that we now that, these are some some standard distribution. Now, will take up one example from this civil engineering, where the distribution of the original random variable is known, so we have to find out that what is the distribution of a $of a$ derived random variable.

(Refer Slide Time: 26:09)



Examples

Q. (Kottegoda and Rosso, 2008): Consider the population of bacteria, C , in a small lake. Under ideal conditions, the population increases exponentially with time T , commencing with an initial population c_0 as $C = c_0 e^{\lambda T}$, where $\lambda > 0$ is the growth rate. However, because of the uncertain effects of various extraneous factors, the time T allowed for the increase in the bacterial population is a random variable with distribution function $F_T(t)$, where $t \geq 0$. Find the cdf of C .

 Stability Methods in
Engineering:
M4L3

Dr. Rajib Halty, IIT Kharagpur

6

So, this problem is taken from the book Kottegoda and Rosso. So, **there are** this is on the growth of the population of bacteria in a lake. So, let us consider that the population of the bacteria that is C that is in terms of concentration in a small lake. So, this C the population denoted by C under ideal condition the population increases exponentially with time T , so as time increases this population of bacteria will also increase, commencing with an initial population c_0 . So, c_0 is the, at T equals to 0 the concentration is c_0 .

Now, at any time T the C the new concentration is equals to it is increasing exponentially. So, the relationship is $c_0 e^{\lambda T}$, where λ is greater than 0. So, λ greater than 0 here means that this one will increase as the time pass that is as T increases this C will also increase, so this growth rate.

However, because of uncertain effect of various extraneous factor, the time T allowed for the increase in bacterial population is a random variable with the distribution function F_T . Now, this T that is the time that is taken for this increase in the population, this is we are **this is this is we are** for from the real condition we have seen this is also a **also a** random variable, that random variable this F_T that is the cumulative distribution function of that one is known, some standard distribution is known. So, if this distribution is known then, that we have to find out the CDF of C . So, basically there is a

functional relationship is there, this T distribution of this T is known we have to find out, what is the distribution for the population of bacteria C **at a time** at time T?

(Refer Slide Time: 28:26)

Examples...Contd.

Sol.:

The bacteria immediately prior to flushing has a population with distribution function $F_C(c)$ given by


$$F_C(c) = \Pr[c_0 e^{\lambda T} \leq c] = \Pr[\lambda T \leq \ln \frac{c}{c_0}] = F_T\left(\frac{1}{\lambda} \ln \frac{c}{c_0}\right)$$

where $c \geq c_0$. Thus,

$$t = \frac{1}{\lambda} \ln \frac{c}{c_0}$$

and the first derivative of the inverse function is

$$\frac{dt}{dc} = \frac{1}{\lambda c}$$



Probability Methods in
 Engineering: H4L3

Dr. Rajib Haity, IIT Kharagpur

7

So, to solve this problem, first of all we have to find out that what is this distribution for the concentration? The bacteria immediately prior to the flushing have a population with the distribution function $F_C(c)$. So, this population which is we are talking about this should be means this should be express in terms of the this is **a this is from the** from the standard properties of the CDF that we have discuss in earlier lectures, that is the $F_C(c)$ the c is the variable here random variable is c is nothing but, the probability of that this concentration at any time, that is concentration at any time now here is that c naught e power λT . So, this one should be less than as specific value which is nothing but, the expression of this cumulative distribution function.

So, now this one after some algebraic transformation that is λT should be is less than equals to \log natural c by c naught, which is also further written that is the $f(t)$ of this. So, this is T less than equals to 1 by $\lambda \log$ natural c by c naught. So, now, when we are writing that probability of t less than 1 by $\lambda \log$ natural c by c naught that means, that is the probability distribution of the cumulative distribution of the random variable T and here that 1 by $\lambda \log$ natural c by c naught has come.

So, now where this c is greater than c naught, so as this we have seen that it from the initial population, this c at any time it is always increase as this λ is greater than


0. Thus this T is equal to $1/\lambda \ln c$ by c naught. So, this is our **that that** inverse relationship that, we have got that is given that c equals to c naught $e^{\lambda T}$. So, this is our inverse relationship which states that T is equal to $1/\lambda \ln c$ by c naught. So, if you know this one then we can calculate its derivative that is dT/dc which is needed for the Jacobian which is equal to, so $1/\lambda c$.

(Refer Slide Time: 30:57)

Examples...Contd.

- we have

$$f_Y(y) = \left| \frac{dx}{dy} \right| f_X[x] = |J| f_X[x]$$
- i.e. $f_C(c) = \frac{1}{\lambda c} f_T\left(\frac{1}{\lambda} \ln \frac{c}{c_0}\right)$
 where $c \geq c_0$. Now if $f_T(t)$ is known, then $f_C(c)$ can be found from the foregoing equation
- For example, if T is exponentially distributed with pdf



Probability Methods in
Engineering: H4L3

Dr. Rajib Halty, IIT Kharagpur

8

So, once we know this Jacobian then we have to just get that, f_T is that is also we know that this is the functional relationship that is $f_Y(y)$, the new variable is the Jacobian multiplied by $f_X(x)$ express in terms of y . So, this $f_C(c)$ is nothing but, then $1/\lambda c$ multiplied by $f_T(1/\lambda \ln c)$ by c naught. So, this T is express in terms of its inverse **inverse** relationship, where c is greater than c naught now if this f_T is known then f_C can also be found from this equation.

Now, what was given in this problem that, if from this cumulative distribution of this t is known, **so** as this is known then you can easily get **what should be the** what should be expression for this part now here we can take one specific example that is if this T is exponentially distributed. Now, if this T is exponentially distributed then we have to assume some we know that how this exponential distribution form is that is $a e^{-at}$ minus T .

(Refer Slide Time: 32:14)

Examples...Contd.

$$f_T(t) = ae^{-at}$$


with $a > 0$, one gets

$$f_C(c) = \frac{1}{\lambda c} ae^{-[(a/\lambda)\ln(c/c_0)]} = \frac{a}{\lambda c} \left(\frac{c_0}{c}\right)^{a/\lambda} = \frac{\theta c_0^\theta}{c^{\theta+1}}$$

where $\theta = a/\lambda > 0$, and $c \geq c_0$. The cdf of C is given by

$$F_C(c) = 1 - \left(\frac{c_0}{c}\right)^\theta$$

with $\theta > 0$ and $c \geq c_0$.



Probability Methods in
Engineering: M4L3

Dr. Rajib Hait, IIT Kharagpur

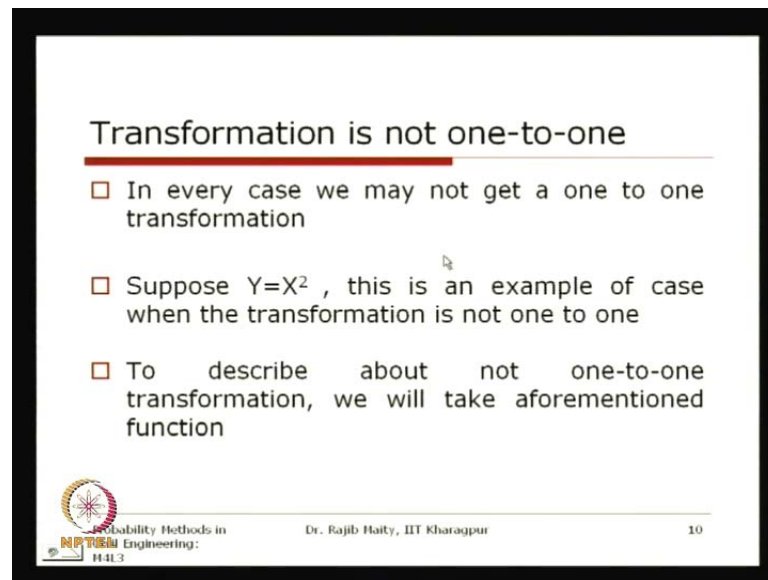
9

You recall that from our earlier lectures we generally use this lambda as parameter for the exponential distribution, but here the lambda is used for some other growth rate. So, we are using some other variable that is a, so this is the form of this exponential distribution. This t is obviously greater than 0 that is a e power minus a t with this a greater than 0.

Now, So, once you know this one, then this thing can be shown if it is given like this then we are just replacing this value with that expression, that is a by lambda log natural c by c naught, so which can be rearranged and express that a by lambda c naught by c power a by lambda, which is equal to theta c naught power theta by c power theta plus 1. This theta is now a new variable, which is equal to a by lambda which is greater than 0 and c is greater than c naught.


So, the CDF of c is given by f c c is equal to 1 minus c naught by c power theta, so this is your PDF that is probability density function from here if we integrate, we will get that cumulative distribution function for c, which is one by one minus c naught by c power theta and here this theta is greater than 0 and c is greater than c naught. So, similarly if this f t follows some other distribution, so this expression we will **will** change accordingly.

(Refer Slide Time: 34:11)



Transformation is not one-to-one

- In every case we may not get a one to one transformation
- Suppose $Y=X^2$, this is an example of case when the transformation is not one to one
- To describe about not one-to-one transformation, we will take aforementioned function

 Stability Methods in
NPTU Engineering:
M4L3

Dr. Rajib Haity, IIT Kharagpur

10

Now, So, far what we discuss is that method of transformation is for the **for the** transformation is unique that is one to one now we will see one more case that where the transformation is not one to one then what should we do? So, again we will recall that same the fundamental theorem if the roots are basically roots are more than one then **then** how this transformation changes.

So, here we are taking that the transformation is not one to one, so in every case we may not get a one to one transformation suppose that y is equals to x square if this is the relationship, this is an example of the of the case, when the transformation is not one to one.

So, to describe about that not one to one transformation, we will take this as a aforementioned **aforementioned** function. So, we will take this function that y equals to x square x square, we will define some PDF for this x and we will try to find out what is the PDF of this **of this** y .

(Refer Slide Time: 35:26)

Example


Q. Suppose X is a continuous random variable with probability density function: $f(x) = x^2/3$ for $-1 < x < 2$. What is the pdf of $Y = X^2$?

Sol.:

The transformation $Y = X^2$ is not one-to-one over the interval $-1 < x < 2$

Taking the inverse function of $Y = X^2$, we have

$$\begin{aligned} X_1 &= -\sqrt{Y} & \text{for } -1 < x < 0 \\ X_2 &= +\sqrt{Y} & \text{for } 0 < x < 2 \end{aligned}$$

 Probability Methods in
H4L3

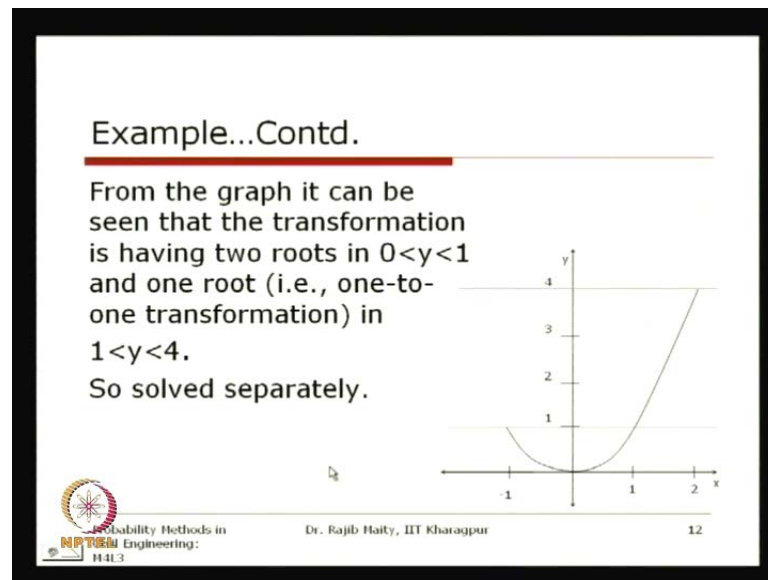
Dr. Rajib Haity, IIT Kharagpur

11

So, here we are taking this example, that is suppose that x is a continuous random variable with the probability density function $f(x)$ is equals to $x^2/3$, for this zone that is minus 1 to 2, so minus 1 to 2 in this zone this density is defined that is $x^2/3$. Now, we can check that whether this integration of this total of this PDF is equals to 1, that is why this by 3 is coming as a normalizing factor.

So, this is a PDF and the functional relationship that we have taken is that y is equals to x^2 , so we know what the PDF of x is and we also know that, this relationship is not one to one then, what should be the distribution of this y . So, to solve this one that this transformation is not one to one over the interval of this minus 1 to this 2. So, taking the inverse function, that is y equals to x^2 that is x equals to plus minus square root y , so this x has also two roots we are getting now the x equals to minus square root y , so this minus square root y we will get for this x when it is minus 1 to 0; and for the x 2 is the plus square root y for the region 0 to 2. Now, this can be, so if we just plot this function this will be clearer to express this 1.

(Refer Slide Time: 37:14)



So, here the PDF is shown here from this minus 1 to this two and we have just given two lines here, one is that y equals to 1 and one is that y equals to 4. So, for this zone that is from this x equals to minus 1 to 1 the root, if I take the square root of this one this square root of y that minus will give the this root that is from this minus one to 0; and **and** for this one from 0 to 2 that this zone that there we can take that root of this plus square root of y that is from this 0 to up to 2.

Now, what we can see here that from for this y this variable for this y , there is if we start from 0 to 1, so in this zone there are basically two roots for this relationship. So, one is on the negative side another one is on the positive side. So, we have to calculate this zone separately that is 0 to 1; and for this 1 to 4 in this zone of for the random variable y , for this zone that is 1 to 4 then this transformation is one to one.

So, from. So, we have. So, we have divided this random variable for this **two zone** two zones one is from the 0 to 1, where there are two roots exist and another zone is from 1 to 4 where there one root exists. So, from the graph it can be seen that the transformation is having two roots in the zone 0 to 1 for y and one root that is one to one transformation in the zone 1 to 4, so these two zones are solved separately.

(Refer Slide Time: 39:07)

Transformation is not one-to-one
...Contd.

□ Considering the first case $0 < y < 1$

$$F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y})$$

$$= F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

□ The probability density function $f_Y(y)$

$$f_Y(y) = f_X(\sqrt{y}) \frac{1}{2} y^{-1/2} - f_X(-\sqrt{y}) \left(-\frac{1}{2} y^{-1/2}\right)$$

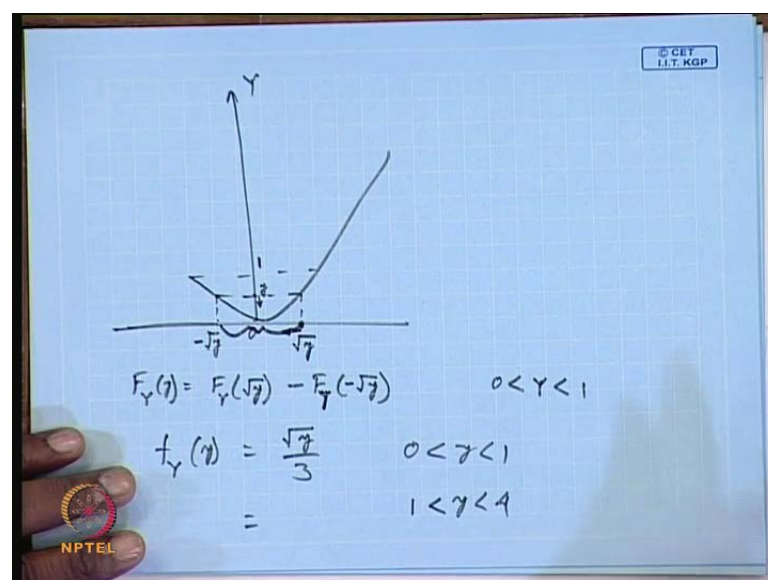
$$= \frac{(\sqrt{y})^2}{3} \frac{1}{2} y^{-1/2} + \frac{(-\sqrt{y})^2}{3} \frac{1}{2} y^{-1/2}$$

$$= \frac{1}{6} y^{1/2} + \frac{1}{6} y^{1/2} = \frac{\sqrt{y}}{3}$$

NPTEL Probability Methods in Engineering: M4L3 Dr. Rajib Maity, IIT Kharagpur 13

So, the first zone that is when we are taking that y is between 0 and 1 then $f_Y(y)$ is equals to probability of y less than y , this specific value should consist of the two zones, that will show that is probability of x square is less than equals to y , which is equals to probability of minus square root y to the plus square root of y . So, this f_X of square root of y minus f_X of square root of y .

(Refer Slide Time: 39:51)



So, what is there actually how we are getting this one is this one this. So, for this zone that is from 0 to 1 for any specific value if I take y , then what is this? So, for this less than

y. So, this is your random variable y, so for this any specific value that y is less than this y then what we are getting is that this zone 1 is which is explain that this is for this side, so this zone that we are we are getting, so now, to get that this zone that is what we need that the square roots of, so this is nothing but, your square root of y and this is nothing but, that minus square root of y. So, from this 1 this total minus this 1, so we will get this part, so f y of square root y minus f y of minus square root of y to get this zone which is nothing but, equals to your f y for the zone 0 to 1.

So, this one we are getting, which is shown here in this **in this** relationship, now so this probability density function f y is so, f y we are getting that is in terms of its now we will **now** do that this one this is express in terms of the root and this is you are the Jacobin of the transformation, so this is half square root of y power **y power** minus y power minus half that is 1 by square root of y that is that Jacobin. So, this Jacobin is for this zone this is the positive one and this is the negative one, **sorry** there will be that absolute value symbol here. So, this minus half square root y once you take this absolute value of this Jacobin it becomes plus that half y power minus half.

So, now this f x square root y is express in terms of this, that is this cumulative distribution this f x is given that is PDF is given to us. So, x square by 3, this x square by 3 this x is now replaced by this square root of this y. So, square root of y square by 3 similarly here it is the minus square root of y square by 3.

So, if we just take this one add this 2, then we get that square root of y by 3, so for this zone that is f y of y is becoming square root of y by 3 for the zone y this 1, similarly we have to find one more expression for the zone 1 to 4 that **that** is straight forward because that relationship is one to one as we have discuss.

(Refer Slide Time: 43:00)


Transformation is not one-to-one
...Contd.

□ For the second case $1 < y < 4$
 $F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = P(\sqrt{y}) = F_X(\sqrt{y})$
 and so the probability density function is

$$f_Y(y) = F'_Y(y) = f_X(\sqrt{y}) \frac{1}{2} y^{-1/2}$$

$$= \frac{(\sqrt{y})^3}{3} \frac{1}{2} y^{-1/2} = \frac{\sqrt{y}}{6}$$

Thus, $f_Y(y) = \begin{cases} \frac{\sqrt{y}}{3} & \text{for } 0 \leq y < 1 \\ \frac{\sqrt{y}}{6} & \text{for } 1 \leq y < 4 \end{cases}$

 Probability Methods in
Engineering:
M4L3

Dr. Rajib Hait, IIT Kharagpur

14

So, for this zone that is for 1 to 4 that $f_Y(y)$ is equals to this less than y that is x square is less than y , which is P of this square root y that is $f_X(\sqrt{y})$. So, the probability function that $f_Y(y)$ is equals to that, this f_X of square root y single root only the positive root, we are considering multiplied by this absolute value of this Jacobian, which is half y power minus half. So, this Jacobian multiplied now we are expressing this one in terms of the PDF of x , that is x square by three in place of x we are writing square root of y . So, square root of y square by 3 multiplied by half y power minus half which is equals to square root y by 6.

So, this expression is for the **for the** range 1 to 4, so for thus 1 to 4 what we got that this is square root of y by 6. So, this is the complete statement for the **for the** distribution of the of y that is $f_Y(y)$ of y square root of y by 3 for this zone 0 to 1 and square root of y by 6 for the **for the** zone 1 to 4. So, this is **this is** shown here. So, here also we can once check that if this **if this** transformation or if this density is correct then we can if we just integrate it for the entire zone that is 0 to 4; and we take this one then also this will this should equals to the unity that you can check to the this expression a valid PDF.


(Refer Slide Time: 44:56)

Determination of Density of Function using Characteristic functions

- Consider a function of random variable $Y=g(X)$.
- From earlier lectures we learned that, the characteristic function of a random variable X

$$\phi_X(s) = E[e^{isX}] = \int_{-\infty}^{\infty} e^{isx} f_X(x) dx$$
- Therefore the characteristic function of the $Y=g(X)$ is

$$\phi_Y(s) = E[e^{isg(X)}] = \int_{-\infty}^{\infty} e^{isg(x)} f_X(x) dx$$
- If this can be written as $\int_{-\infty}^{\infty} e^{isy} h(y) dy$
 Then it follows $f_Y(y) = h(y)$ if the transformation of unique or one-to-one



Probability Methods in
Engineering:
M4L3

Dr. Rajib Haity, IIT Kharagpur

15

Now, we will discuss about the another method, which is the method call that, this method of characteristic function. So, we will determine the relationship we determine **we determine** the relationship of the PDF of the new variable, that is function of the random variable through the characteristic function.

Now, this characteristic function we have discuss for **a for** a variable **earlier** in the earlier lecture now in this lecture easing that property we will see that how this that property can be utilize, can be used to find out the density of the function of random variable.

Now, we have seen that, if there is a function of random variable like that y is equals to $g(x)$, now for these random variable x , that the characteristic function of that random variable x can be shown that, this is $E[e^{isx}]$ is equals to that expectation of $e^{isg(x)}$ this I is your square root of minus one that is a complex number s is the variable, that we are using here; and this capital x is your random variable. So, expectation of this quantity is nothing but, the characteristic function that we have explain earlier.

Now, this expectation of **a of a** function also we have we have discuss earlier that is if. So, this can be express, that is for any function that should be multiplied by the PDF of that random variable and integrate over the entire support to get the form of that form of that characteristic function. So, this is integrated from this minus infinity to plus infinity multiplied by $e^{isg(x)}$ multiplied by $f_X(x) dx$, so this is the expression for

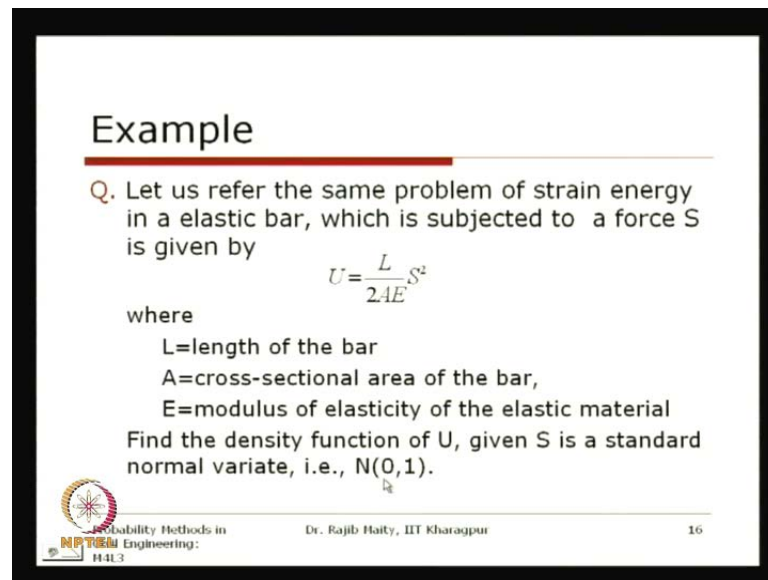
the characteristic function; and this is obviously, for the continuous one and you can see that if it is a **if it is a** discrete random variable then this integration will be replaced by the summation and this f_X that is the PDF we will replace by its pmf that, we generally in standard we call that for pmf is for the discrete random **random** variable. So, this is the **this is the** characteristic function of a random variable x .

Now, now the functional relationship as you have decided that y is equal to $g(x)$. Now, for this $g(x)$ now my goal is to get the characteristic function for this function $g(x)$, which is nothing but, equals to the another random variable y . So, if I want to get that characteristic function then, that f_Y should equal to that $e^{i t y}$. So, this y is your random variable here now this y is equal to $g(x)$, we are replacing that y in terms of $g(x)$, so that exponential of is $g(x)$ now this exponential that exponential of $i t g(x)$ this function if I take the expectation then we will get the characteristic function of that function of the random variable x , which is y , so from again following the same that definition of this characteristic function.

We can take this function and this is multiplied by this the probability density function of this random variable x and if we just integrate it from minus infinity to plus infinity, then will get the characteristic function. Similarly, again if it is a discrete random variable then this can be replaced by its summation sign.

Now, again if I just one to know if we do not want to express is in terms of x , if we entirely express is in terms of y then **then** this expression can also be expressed as that minus infinity to plus infinity $e^{i t y} h(y) dy$ now from this two, if you just compare then it follows that f_Y is your $h(y)$ if the transformation is unique one to one. So, if we can get this relationship then, this $h(y)$ that is the inverse that is that **that** is express in terms of here in this expression can we can get that PDF of this y .

(Refer Slide Time: 50:00)



Example


Q. Let us refer the same problem of strain energy in a elastic bar, which is subjected to a force S is given by

$$U = \frac{L}{2AE} S^2$$

where

- L =length of the bar
- A =cross-sectional area of the bar,
- E =modulus of elasticity of the elastic material

Find the density function of U , given S is a standard normal variate, i.e., $N(0,1)$.

 Probability Methods in Engineering: M4L3 Dr. Rajib Haitry, IIT Kharagpur 16

Now, to explain this one we will take one example and this example, we have solved earlier also but using the method of distribution I guess and now we have taken this same **same** example and we will solve this example with the method of characteristic function.

So, this one that let us refer the same problem of this strain and energy the problem was on the strain energy of the elastic bar; and that strain energy is generally related to it is force that is under, which the elastic bar is subject to. So, the relationship between this U and this S is like this that U equals to $\frac{1}{2} \frac{L}{AE} S^2$ now, where this L is length of the bar A is the cross sectional area of the bar and E is the modulus of elasticity of the material.

Now, find the density of the U , given the S is a standard normal **standard normal** distribution standard normal variety that is the n equals to a **normal de** normal distributions, which is having mean 0 and standard deviation equals to one now.

(Refer Slide Time: 51:31)

Example...Contd


Sol.:

Rewriting the equation we have $U=cS^2$
 where $c=L/2AE$
 Using characteristic function of $U=g(S)$, we have:

$$\phi_U(\omega) = \int_{-\infty}^{\infty} e^{i\omega g(s)} f_S(s) ds = \frac{2}{\sqrt{2\pi}} \int_0^{\infty} e^{i\omega c s^2} e^{-s^2/2} ds$$

As s increases from 0 to ∞ and transformation is one-to-one, we have

$$du = 2cs ds = 2\sqrt{cu} du$$



Stability Methods in
NPTI Engineering:
M4L3

Dr. Rajib Haitry, IIT Kharagpur

17

So, this one again we are just clubbing all those constant herethrough this another constant see. So, the relationship is U is equal to $C S^2$ using the characteristic function that is U equals to $g(S)$, we have that $\phi_U(\omega)$ is equal to minus infinity to plus infinity $\int_{-\infty}^{\infty} e^{i\omega g(s)} f_S(s) ds$ and $f_S(s)$ is equal to $f_S(s) ds$. Now, this is the characteristic function if you just replace this one this density we know, which is a standard normal distribution, if you replace this one then we will get the 2 by square root of 2π integrate from 0 to infinity. This is an even function that is why multiplied by 2 and taken from this 0 to infinity $e^{-\omega^2 c s^2}$ multiplied by $e^{-s^2/2} ds$ as s increases from 0 to infinity the transformation is one to one. So, we have that du equals to $2cs ds$ that is $2\sqrt{cu} du$.

(Refer Slide Time: 52:45)

Example...contd.

□ The characteristic equation yields

$$\phi_U(\omega) = \frac{2}{\sqrt{2\pi}} \int_0^\infty e^{j\omega u} e^{-u/2c} \frac{du}{2\sqrt{cu}} = \int_0^\infty e^{j\omega u} h(u) du$$

□ Using $f_U(u) = h(u)$

$$f_U(u) = \frac{e^{-u/2c}}{\sqrt{2\pi cu}} \quad \text{for } u \geq 0$$

Probability Methods in Engineering: M4L3 Dr. Rajib Hait, IIT Kharagpur 18

So, the characteristic equation yields that $\phi_U(\omega)$ is equal to $\frac{2}{\sqrt{2\pi}} \int_0^\infty e^{j\omega u} e^{-u/2c} \frac{du}{2\sqrt{cu}}$. So, we are expressing in terms of u here, so which is equal to $\int_0^\infty e^{j\omega u} h(u) du$ now **now** if we do u .

So, now if we call if we compare then this $f_U(u)$ that is the density of the random variable u is equal to $h(u)$ here. So, if you just compare this one it comes that $f_U(u) = \frac{e^{-u/2c}}{\sqrt{2\pi cu}}$ for this u is greater than equal to 0. So, we got the same distribution that same PDF that you got in the earlier using the method of distribution. So, here also we have seen that using the method of characteristic function we got the same **same** expression as well.

So, starting from the fundamental theorem we have discussed that method of distribution method of transformation one to one as well as not one to one as well as not one to one then we discuss that method of characteristics function with examples.

Now, one thing we have to keep in mind that even though this transformation, we have explain particularly in some of the non-linear transformation, which is very common in the civil engineering problems here; that kind of transformation may or may not always in this kind of close form salutation and they are what is more what will be useful in such cases that if we come to know some of the properties of those derived variables that is its expectation that means, we get the mean variance and it is obviously, if we get the

moment generating function. Then we can know the required information of the function that is the derived random variable. So, these things we will discuss in the next class how we can without knowing the PDF how we can know those properties of the derived random variable **thank you**.